

EXPLORATORY ANALYSIS OF ACCIDENT/DRINKING TIME SERIES

Prepared by A.I. McLeod, E. Vingilis, R. Mann and J. Seeley

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Exploratory Analysis of Motor Vehicle, Trauma, Assault and Alcohol

Introduction

This report is also available in Adobe PDF format on the companion CD under the root directory as ExploratoryAnalysis.pdf.

Datasets compiled in SPSS by Jane Seeley, UWO, Faculty of Medicine are analyzed. These datasets provide information of monthly numbers of fatalities, trauma and arrests for impaired driving and assault. Data on traffic counts and annual alcohol consumption is also briefly examined. The datasets examined are listed below.

Code	Name	#records	Description
А	ALCOHOL	11	Annual consumption of alcohol, Canada, 1989 to 1999
В	TRAFFIC	1,680	MTO traffic counts, Ontario, 1992.1 – 1996.12
С	TIRF	434	Traffic Injury Research Foundation, Alcohol Fatalities, Manitoba & Ontario, 1992.1 – 1998.12
D	MTO	26026	MTO highway driver injury and deaths, Ontario, 1992.1 – 1998.12
Е	MVA	84	Ontario Trauma Registry, motor vehicle accidents, May 1992 – April 1999
F	OTR	84	Ontario Trauma Registry, all cases, May 1992 – April 1999
G	LPS	8,920	London Police, Arrests for assault & impaired driving, 1992.5 – 1999.5
Н	WPS	2,943	Windsor Police, Arrests for assault & impaired driving, 1994.1 – 1998.12
Ι	FARS	3,956	Fatal Accident Reporting System, Michigan & New York, 1992.5 – 1999.12
J	MISTPOL	164,931	Police Accident Reports, 3 counties, Michigan, 1992.5 – 1999.12

These datasets were supplied in SPSS format. They were then converted into tabdelimited ascii files using Perl. These data files where input to S-Plus (Version 6 for Windows was used). The location of these files on the companion CD is shown in the table below:

File or Directory Location	Description
EXPLORATORYREPORT.PDF	this report
SPSS	SPSS data files
SPSS\DAT	ASCII tab-delimited data files
REPORT	MS Word 2000 files used to create this report
REPORT\RTF	Above files converted to RTF format for compatibility
MISC	S-Plus Chapter
ALCOHOL	S-Plus Chapter
TRAFFIC	S-Plus Chapter
TIRF	S-Plus Chapter
MTO	S-Plus Chapter
MVA	S-Plus Chapter
OTR	S-Plus Chapter
LPS	S-Plus Chapter
WPS	S-Plus Chapter
FARS	S-Plus Chapter
MISTPOL	S-Plus Chapter

In datasets B through J, the wkgrp variable is defined as a factor with two levels "SunWed" and "ThuSat". The "SunWed" level corresponds to all counts which occur from 11PM on Sunday until Thursday 3AM (4AM for G,H). The complement of this forms the "ThuSat" factor values.

In our analysis tables and figures are usually labelled with numbers. Because the numerous iterations involved in the analysis some figures and/or tables are not shown. For example, you find Figure 1, 2, 4, 5 with no Figure 3. I hope this causes no confusion.

Methods of Exploratory Analysis

Crosstabs Analysis

With crosstabs and frequency analysis we compare the relative proportions of two factors as well as their possible association. For example in MISTPOL we find that 24% of all accidents for which drinking status is reported have drink=yes but when we look at drinking status and injury class we find the following percentages for drink=yes 74.8%, 49.6%, 48.7%, 21.1% and 20.3% for fatal, incapacitating, non-capacitating, possible-injury and no-injury accidents respectively. Clearly drinking status and accident severity are strongly associated. The Pearson chi-square goodness-of-fit test is used to test statistically the null hypothesis of no association between the factors. In some of the datasets the degrees of freedom is very large so even very unimportant differences could be significant. On the other hand, some of the categories in the tables have sometimes have expectations less than 5 and this means that the chi-square test which is an approximation may be inaccurate. In this case we could resort to bootstrapping could get a more exact answer (this has not yet been done in this report).

With each crosstabs analysis we also examine two additional tables:

- (i) Contribution to the Chi-Square, (observedexpected) /√expected. This table helps to pinpoint the cells that contribute to the Pearson chi-square goodness of fit test.
- (ii) Percentage Error, (observed-expected)/expected. This table helps to pinpoint that may be of interest too.

Time Series Plots, Loess Smoothing, Mann-Kendall Trend Tests

Trellis graphics were introduced by Cleveland (1993). The theory of trellis graphics is further developed in Becker et. al (1996). Trellis graphics, available in S-Plus, represent one of the biggest advances in multivariable data visualization in the last ten years.

Trellis plots are coordinated multi-panel displays of data. Three possible choices for the scales used for the panel axes are:

- (i) same scale
- (ii) sliced scale
- (iii) independent scale

Each method is useful but in particular situations. If the data in each panel vary over roughly the same region then using (i) small scale is a good idea. In this case accurate visual comparisons can quickly be made between panels. However if the data vary across quite different scales, as in the case of fatality rates by drinking status, then using (i) may spoil the resolution on either or both panels. If the resolution is degraded like this it is better to use (ii) or (iii). If the variability in the data is similar in different panels then sliced scaling (ii) is a good idea. Sliced scaling means that the number of data units per centimeter on the graph is kept the same in all panels. For example one panel may have data varying between from 50 to 150 and another panel 450 to 550. Sliced scaling allows for good data resolution and allows one to compare the variability across panels. However with sliced scaling you must pay closer attention to the indicated axes tick lables to interpret which panel is larger whereas with (i) it is visually obvious. If the variability is very different between panels, as for example is the case when the MISTPOL accidents are compared across injury class then in order to achieve good data resolution it may be necessary to use method (iii) independent scaling. With independent scaling, the scaling is chosen separately for each panel in order that the data region fills up the space available in each panel. With independent scaling, it is crucial that the indicated axes tick marks be carefully examined when interpreting the display.

Data resolution can also be degraded by outliers. It is very important to know about outliers since they may be of great interest in their own right but outliers can also squish the bulk of the data into a small region. To overcome this difficulty we can use a monotonic transformation such as logarithms. This is much preferable to allowing the scale to change as is sometimes seen in scientific journals.

The importance of the choice of aspect-ratio for data visualization was extensively researched by Cleveland and is discussed in his books Cleveland (1993, 1994) and also in our **Appendix 1**. In most of the examples we have chosen to use an aspect-ratio of one since this seems to provide a reasonable comprise to visualizing the trend (which would often require a more extreme aspect-ratio much larger than one if banking to 45° is used) and visualizing the time series itself (which typically implies an aspect-ratio less than one if banking to 45° is used).

The robust loess locally linear smoother, introduced by Cleveland and discussed in his books (Cleveland, 1993, 1994) is used for trend assessment. The application of robust loess to trend assessment of environmental time series was discussed by McLeod et al. (1990). A brief overview of loess is given in **Appendix 2**.

The Mann-Kendall trend test (Mann, 1945), described in **Appendix 3** in more detail, is used to assess the statistical significance of monotonic trends. This test consists of computing the Kendall rank correlation of the time series with the observation number. When seasonality is present we can use the Seasonal-Mann-Kendall test of Hirsch et al. (1982) to test for long-term monotonic trend. We can also apply the regular Mann-Kendall trend test to the STL deseasonalized data. The exact algorithm of Panneton & Robillard (1972a, 1972b) is used to evaluate the significance level of the Mann-Kendall test. The significance level of the Seasonal Mann-Kendall test is evaluated using a normal approximation. It should be noted that this trend test is of limited relevance since the major focus of this investigation is not monotonic trend but rather on the possible changes in trend that have occurred over the period 1992 to 1999. For this reason, the loess analysis, which is the visualization approach, is expected to be more useful.

STL Analysis

The STL or seasonal-trend loess decomposition (Cleveland et al., 1991; Cleveland, 1993) is used to examine the trend and seasonal pattern in various time series. STL is based on loess smoothers and provides an excellent tool for the estimation of a dynamic and changing seasonal component. STL provides the following decomposition of the data,

DATA = TREND + SEASONAL + REMAINDER

Additional components may also be added to this model to remove other sources of variation.

Given an STL decomposition the deseasonalized series is defined by,

DESEASONALIZED SERIES = DATA - SEASONAL

The deasonalized series may be tested for trend using the Mann-Kendall trend test.

For the STL analysis eight graphical displays were examined:

- (i) SLT decomposition
- (ii) Monthplot display of the SEASONAL component.
- (iii) Time series plot of deseasonalized series
- (iv) Autocorrelation plot of deseasonalized series
- (v) Autocorrelation plots of the REMAINDER term
- (vi) Normal probability plot of the REMAINDER term
- (vii) Residual-Fit or RF-Spread Plot
- (viii) Spread-Location or SL-Plot

The basic STL graphical display is comprised of four panels showing a time series trace plot of the original data along with its decomposition into three or more components. The scales in each panel are independent to permit maximum graphical resolution but a bar at the right side of each panel provides an indication of the scaling in each panel.

The monthplot (Cleveland & Terpenning, 1984; Cleveland, 1993) of the SEASONAL component is used for visualizing the seasonal effects and how they evolve over time. The monthplot shows the mean level of the seasonal component for each month as a horizontal bar. A time series trace shows the fluctuations over time around this horizontal bar. In this way we can see how the seasonal component changes over time.

The REMAINDER term is like a residual component and should not contain any significant trend or seasonality. In practice the autocorrelation present in the REMAINDER term is close to zero for lower order lags. Also in practice the REMAINDER term is often normally distributed but again this is not required. However if the REMAINDER term is normally distributed and is not autocorrelated then the STL decomposition extracts all information in the data in the sense that what is left over is informationless white noise. STL like most other seasonal adjustments methods over-

corrects for seasonality in the sense that the spectral density of the remainder process has a deficiency at the seasonal frequency and this implies that one can expect a small negative autocorrelation at lag 12 in the REMAINDER term. It is thus information to look at autocorrelation plots of the REMAINDER term. Likewise it is also information to examine a normal probability plot of the residuals. Outliers are always of interest. They can point to possible interesting new phenomenon. The normal probability plot display also includes 0.5% significance limits given by Royston (1993), the W-test for normality (Royston, 1982) and the skewness test for normality.

The RF-Spread Plot (Cleveland, 1993) enables one to visualize how much of the variation in the data is explained by the model. This has traditionally been done using the coefficient of determination, R^2 , but the RF-Spread Plot is more informative. The RF-Spread Plot is a trellis display comprised of two Quantile Plots. Given some data X_I , ..., X_n , let $X_{(1)}$, ..., $X_{(n)}$ denote the data re-arranged in ascending order so that $X_{(1)} \le X_{(2)} \le ... \le X_{(n)}$, then the Quantile Plot is simply the scatterplot of $X_{(1)}$ vs. f_i where $f_I = (i-1/2)/n$. This is quite similar to the empirical distribution function plot but it is quite different in both construction and spirit. The first panel of the RF-Spread Plot shows the Quantile Plot for the fitted values with the mean subtracted out and the second panel shows on the same scale the Quantile Plot for the remainder or residual. By comparing the panels, we can see at a glance the amount of variation accounted for by our model.

The SL-Plot (Cleveland, 1993) has been further developed by McLeod (1996, 1999). The purpose of the SL-Plot is to detect montone spread. Monotone spread occurs when the variability of the data increases systematically with the location or level of the data. Montone spread can usually be corrected by a suitable power transformation such as a log or square root or reciprocal. In modelling situations montone spread sometimes indicates an important inadequacy. Also if montone spread is present, data visualization can often be improved by making a suitable power transformation.

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A. Alcohol Dataset

Overview and Summary 1989 to 1999 (annual)

Source: *Volume of sales of alcoholic beverages and per capita 15 years and over* The numbers are the volume of sales in thousands of litres. Population is not included.

	year	sprtlitr	sprtperc	sprtabs	winelitr	wineperc	wineabs	beerlitr	beerperc	beerabs
1	1989	67227	40.0	26890.80	84741	13.0	11016.330	836569	5.0	41828.45
2	1990	63866	36.9	23566.55	80051	11.7	9365.967	829824	4.9	40661.38
3	1991	59700	36.9	22029.30	79166	11.7	9262.422	813304	4.9	39851.90
4	1992	55389	36.9	20438.54	78442	11.7	9177.714	788218	4.9	38622.68
5	1993	50565	36.9	18658.48	77417	11.7	9057.789	738902	4.9	36206.20
6	1994	48955	36.9	18064.40	77174	11.7	9029.358	733285	4.9	35930.96
7	1995	48551	36.9	17915.32	79861	11.7	9343.737	747182	4.9	36611.92
8	1996	50131	36.9	18498.34	83946	11.7	9821.682	756023	4.9	37045.13
9	1997	51270	36.9	18918.63	88591	11.7	10365.147	733674	4.9	36596.19
10	1998	57223	36.9	21115.87	91019	11.7	10649.220	740017	4.9	36260.83
11	1999	63173	36.9	23310.84	93168	11.7	10900.660	750435	4.9	36771.32

Variables:

-sprtlitr, winelitr, and beerlitr

are just the total volumes

-sprtperc, wineperc, and beerperc

are the estimated % of alcohol for each category

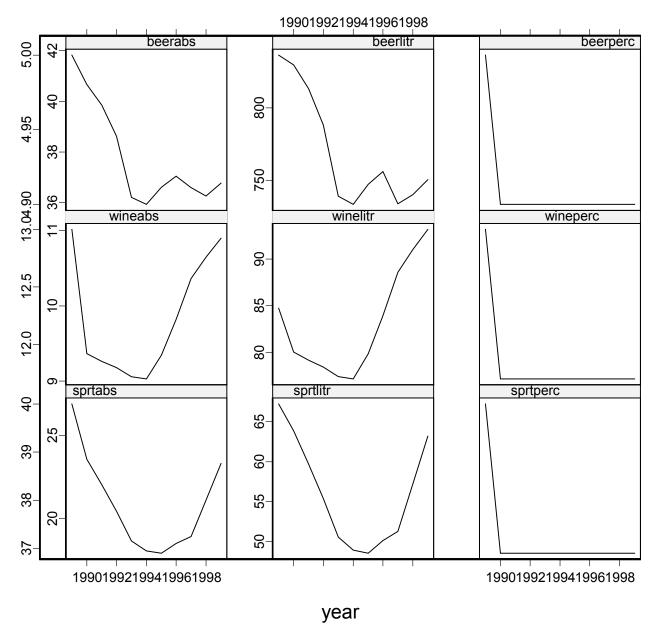
-sprtabs, wineabs, and beerabs

are the volumes multiplied by the % alcohol to calculate the volume of sales in thousands of litres of absolute alcohol (the standard without the population calculation)

Conclusions:

consumption of beer has decreased consumption of spirit alcohol decreased and then increased consumption of wine decreased and increased back to 1989 levels total consumption of alcohol has decreased and is now increasing again

Consumption of Spirits, Wine and Beer



litr and abs are in millions of litres, perc are percentage

B. Traffic Dataset

Overview and Summary

Traffic counts for Ontario, January 1992 to December 1996. TRAFFIC.SAV has 420 records corresponding to each weekday for 5 years ($5 \times 12 \times 7=420$). The weekday data was aggregated to the level of wkgrp using the definition that SunWed is comprised of all counts starting on Sun at 11PM and running through to Thursday 3AM and the rest is in the ThuSat group. Note that there is no data for 1997 and 1998.

A regression analysis of the log counts on the variables wkgrp, hour, month and year accounts for 96.4% of the variation. However year is still very significant even when it is entered last into the regression. This suggests that it would be very desirable to obtain data for 1997 and 1998.

The STL visualization indicates that the effect of the missing values in the years 1992 and 1996 is quite large (Figure <u>STL</u>).

The traffic counts are split very nearly evenly between the ThuSat and SunWed weekgroups. However there is significant variation in traffic counts over the other variables hour, month and year.

With respect to hour and month we show that the proportions remain relatively stable over the years and so we create a proxy for the traffic counts using the fitted values from the regression for the year 1994.

Instead of this proxy it would be very desirable to get the complete data from 1992 to 1998.

If it can be assumed that the annual total traffic counts are relatively constant from year to year then the proxy variable we have defined is adequate. Furthermore since the regression of log counts on the variables wkgrp, hour, month and year accounts for 96.4% of the variation it would be expected that the proxy variable will not be needed in developing statistical models for traffic fatalities since it is so closely related to the other variables.

Analysis: Regression

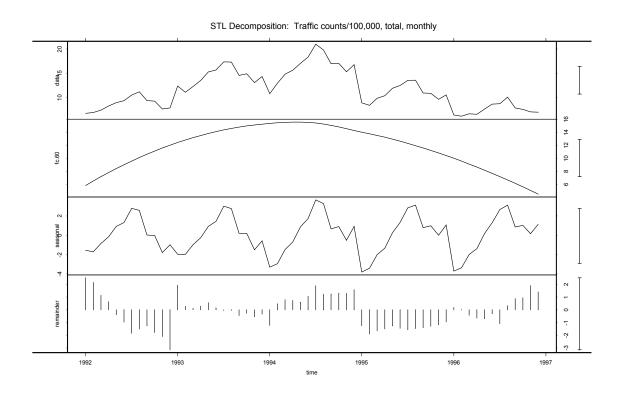
We investigated the possible use of a regression model to fill in the missing data. It was found that a log transformation of the counts was helpful in reducing skewness of the residuals and reducing monotone spread. The regression was fit using the year variable last and is summarized below:

Notice the fit is quite good in terms of the R-squared but that the year variable is quite important. The variation over years is summarized below in the STL analysis and in the crosstab analysis.

Seasonal-Trend Loess Analysis

The seasonal-trend analysis shows the strong seasonal period and remarkable variation in the traffic counts. I now think the low values for 1992 and 1996 are due to the missing stations.

Figure STL

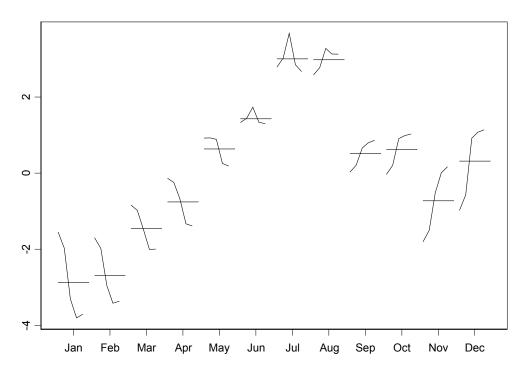


ss.window = 5 , fc.window = 60 , fc.degree = 2

TRAFFIC

Monthplot Analysis

There are no surprises here. We would expect July and August to have the largest values.



Traffic counts/100,000, total, monthly

Crosstab Analysis: year

> crosst Call:	abs(count	s~year,	data=tı	raffic.df)						
crosstab	os(counts 5 cases ir		data =	<pre>traffic.df)</pre>						
N	++ N N/Total ++									
year										
1992	10371415 0.15									
1993	17176609 0.25									
1994	19633001 0.28									
1995	13124602 0.19									
1996	9067319 0.13	г -								

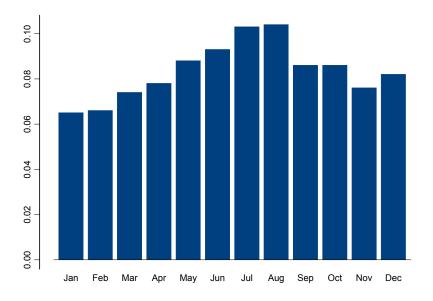
Notice that that there was approximately twice as much traffic in 1994 than in 1992 and 1996. This indicates traffic intensities can vary significantly from year to year. However it may be that relationship between traffic intensities and the covariates wkgrp, hour and month remains linear. In this case a proxy variable for traffic intensity may be useful. This is investigated in the crosstabs analysis in the next section.

Crosstab Analysis: wkgrp, hour and month

The analysis below suggests that the split between SunWed and ThuSat is very close to 50% over the years. Because of the large sample size the factors are not independent but the percentage error between observed and expected under the independence assumption is very small.

A similar story holds for the crosstabs analysis of year with the hour groups. In this case the split over the hour groups "11PM", "12AM", "1AM", "2AM" is very close to 0.42, 0.27, 0.18, 0.13 over all the years.

The crosstabs analysis of year and month does indicate some association between the factors but again it is quite small. The traffic intensity as a percentage of the annual totals are shown in the plot below.



We conclude from the analysis that a proxy variable for the traffic intensity developed from the regression model may be useful.

TRAFFIC

N N/RowTo N/ColTo N/Total	tal									
year	hour 11PM	12AM	1AM	2AM	RowTot					
1992	0.145	0.147	0.155	1401088 0.135 0.160 0.020	+ 103714: 0.15 					
1993	0.418			2104242 0.123 0.241 0.030	171766 0.25 					
1994	0.285	0.285	0.281	2393232 0.122 0.274 0.034	196330 0.28 					
1995	-			1666881 0.127 0.191 0.024	131246 0.19 					
1996	0.131	0.273	0.127	1178185 0.130 0.135 0.017	906731 0.13 					
ColTotal	1				+ 6937294 					
<pre> get.cr [,1 [1,] - [2,] [3,]</pre>	Test for independence of all factors Chi^2 = 21954.57 d.f. = 12 (p=0) Yates' correction not used > get.crosstabs.percenterror(counts~year+hour) [,1] [,2] [,3] [,4] [1,] -3 -2 4 7 [2,] 1 1 0 -3 [3,] 1 1 -1 -3									

> crosstabs(counts~year+wkgrp, data=traffic.df) Call: crosstabs(counts ~ year + wkgrp, data = traffic.df) 69372946 cases in table +---+ N N/RowTotal N/ColTotal N/Total . +---+ year |wkgrp SunWed |ThuSat |RowTotal| 1992 |5242486 |5128929 |10371415| 0.505 0.495 0.15 0.152 0.147 0.076 0.074 1993 |8317516 |8859093 |17176609| 0.484 0.516 0.25 0.241 0.254 0.120 0.128 1994 9683588 9949413 19633001 0.493 0.507 0.28 0.280 0.286 0.140 0.143 6618902 6505700 13124602 1995
 0.504
 0.496
 0.19

 0.192
 0.187
 1

 0.095
 0.094
 1
 1996 4680669 4386650 9067319 0.516 0.484 0.13 0.136 0.126 0.067 0.063 ColTotal | 34543161 | 34829785 | 69372946 | 0.5 0.5

Test for independence of all factors

Chi² = 31245.03 d.f.= 4 (p=0) Yates' correction not used

> get.crosstabs.percenterror(counts~year+wkgrp)

 $\begin{bmatrix} 1, 1 \end{bmatrix} \begin{bmatrix} 2 \\ 2 \end{bmatrix}$ $\begin{bmatrix} 1, 1 \\ 2 \end{bmatrix} = \begin{bmatrix} 2 \\ -3 \end{bmatrix} = \begin{bmatrix} 2 \\ 3 \end{bmatrix} = \begin{bmatrix} 2 \\ -3 \end{bmatrix} = \begin{bmatrix} 2 \\ -4 \end{bmatrix} = \begin{bmatrix} 2$

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TRAFFIC

> crosstabs(counts~year+month, data=traffic.df)

6937294	<pre>> crosstabs(counts~year+month, data=traffic.df) Call: crosstabs(counts ~ year + month, data = traffic.df) 69372946 cases in table ++</pre>												
N N/RowT N/ColT N/Tota	otal otal 1												
year	month	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	RowTotl
1992	0.0654 0.1495	0.0669 0.1522	0.0717 0.1456	0.0799 0.1522	0.0864 0.1473	0.0903 0.1452	0.1015 0.1480	0.1078 0.1554	0.0907 0.1573	0.0895 0.1555		0.0760 0.1393	10371415 0.15
1993	0.0720 0.2727	0.0646	0.0715 0.2403	0.0788 0.2486	0.0889 0.2511 0.0220	0.0913 0.2431	0.1009 0.2435 0.0250	0.1008 0.2407	0.0848 0.2436 0.0210	0.0868 0.2498		0.0835	17176609 0.25
1994	0.0551 0.2385	0.0660 0.2842	0.0754 0.2899	0.0796 0.2869	1704066 0.0868 0.2800	1830829 0.0933 0.2838	2096297 0.1068	1975855 0.1006 0.2747	1695305 0.0863 0.2837	0.0867		0.0854 0.2966	19633001 0.28
1995	0.0681 0.1970	0.0643 0.1850	0.0752 0.1932	0.0792 0.1910	0.0908 0.1960	0.0952 0.1936	0.1033 0.1907	0.1034 0.1887	0.0835 0.1834	0.0830 0.1825	1	0.0803 0.1865	13124602 0.19
1996	0.0712	0.0681 0.1354	0.0738	0.0728 0.1212	0.0843 0.1256	0.0955	0.0966	0.1115	0.0870 0.1319	0.0836	1		
	0.065	0.066	0.074	0.078	0.088	0.093	0.103				5280252 0.076	5654517 0.082	69372946
Test fo: get.cros ata = t: > get.c:	10 -2 16 0 4 -2	ndence o 103146.1 correctic ercenter f) .percenter [,3] [,4 -3 -3 2 2	f all fac d.f.= 4 n not us ror(count error(cou	tors 4 (p=0) ed s ~ year (6] [,7] -3 -1 -2 -2 0 4 2 1	r + mont] r+month) [,8] [,9 4 -3 -3 0	h, d 9] [,10] 5 4 -2 1 0 1 -3 -4	[,11] [-3 0 2 -3	+ -7 2 5 -1 1 -3	3	-5	+	+	++

9

TRAFFIC

Proxy Variable

Since no data is available for 1997 and 1998 we have decided to use a proxy variable for the traffic counts. Alternatively if traffic could be predicted for 1997 and 1998 we could use the predicted traffic intensities to supplement the original data. However since there is a great deal of variation and no simple trend present, we have decided to replace the original data with a proxy.

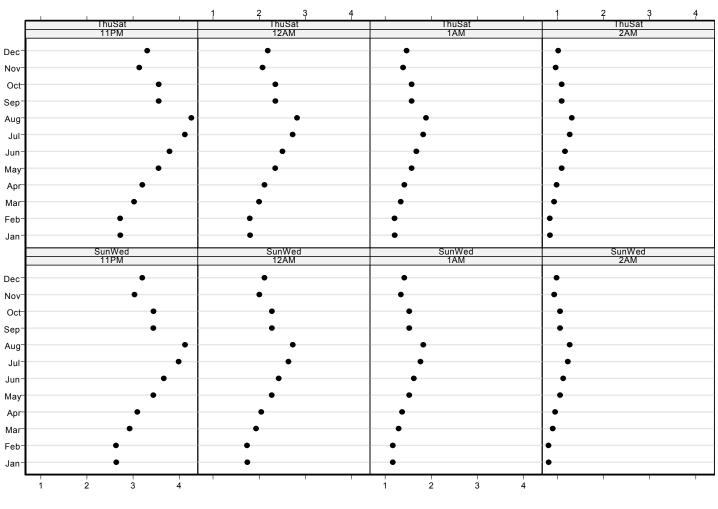
The proxy variable we use is developed from a regression of the log of the traffic counts regressed on wkgrp, hour, month and year. The expected traffic count for a given set of covariates is then obtained by the transformation,

 $Y = \exp(y + \sigma^2/2),$

where y is the predicted log count, σ^2 is the residual variance and Y is estimated expected traffic count. For the proxy we take year=1994 and we use $Y/10^5$. The proxy values are visualized using a dotchart below.

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TRAFFIC



proxy

C. TIRF Dataset - Introduction

Introduction

TIRF, Traffic Injury Research Foundation, provided this data on fatal car accidents in Manitoba and Ontario from January 1,1992 to December 31, 1998. This data is for automobile driver deaths only. The data is analyzed in two parts. First the Late Night Subset and second TIRF – All Times

Late Night Subset

The Late Night Subset of the TIRF data set contains 434 records. The variables of interest in original data set are shown in the table below.

Variables of interest:

drink	alcohol consumption, factor: yes, no, unknown
year	ordered factor, 7 levels:1992 to 1998
month	ordered factor, 12 levels: January to December
wkgrp	ordered factors of week group, 2 levels: SunWed, ThuSat
hour	ordered factors, 5 levels: 11PM, 12AM, 1AM, 2AM, 3AM
prov	factor: Manitoba, Ontario

In the SPSS file the variable is bacpos and the variable drink in our analysis is definied by: drink = yes if bacpos = 7777 or 8888 or 1 drink = no if bacpos = -7 or 0 drink = unknown, otherise

Missing Data:

There are 434 cases in the original SPSS file but case 343 is missing date and hour so it was deleted. Also hour was missing in additional 33 cases so these records were also deleted. This leaves 400 cases for our analysis. In the remaining 400 cases, there are 44 cases with drink=unknown in Ontario and 2 cases in Manitoba. Of the 400 cases, there are 347 in Ontario and 53 in Manitoba.

TIRF – All Hours (TIRFALL)

There were 2234 cases in the original SPSS datafile. Record 1765 was incomplete with respect to moyr so this record was omitted. Also record 1990 missing crsw (day of week).

An ordered factor hour was created defined by the hour in which the accident occurred. There are 24 levels to this ordered factor. All accidents occurring from 0000 to 0059 were put in 12PM, 0100 to 0159 in 1AM, etc.

Variables of interest:

drink	alcohol consumption, factor: yes, no, unknown
year	ordered factor, 7 levels:1992 to 1998
month	ordered factor, 12 levels: January to December
wkgrp	ordered factors of week group, 2 levels: SunWed, ThuSat
hour	ordered factors, 24 levels: 12PM, 1AM,, 11PM.
prov	factor: Manitoba, Ontario

Missing Variables

There are 2233 cases in the S-Plus dataframe. We see there are 33 cases were hour is missing and 342 cases where drink is unknown.

```
> dim(tirfall.df)
[1] 2233 10
> sum(tirfall.df$drink=="unknown")
[1] 342
> sum(is.na(tirfall.df$hour))
[1] 33
```

C. TIRF Late-Night Subset - Crosstabs

Summary

TIRF data set contains 434 observations of car fatality accidents from January 1992 to December 1998 in Ontario (376) and Manitoba (58). Each case represents one driver fatality. After deleting all cases where the hour variable was missing we were left with 400 cases. Of these 400 cases, 347 were in Ontario and 53 were in Manitoba.

DRINK-HOUR. In Ontario, about 70% of fatalities involve alcohol. Although the association between alcohol fatalities and hour is not statistically significant (p=18%) we see that there is a slight increase in alcohol usage associated with deaths at 2AM (11% extra) and 3AM (13% extra). In Manitoba about 77% of fatalities are alcohol associated. Again the association with hour does not appear to be statistically significant but there is a 15% excess of alcohol fatalities in the 3AM window.

<u>WKGRP</u>. The car fatality accidents in SunWed group count for 62% of total accidents while 38% in ThuSat group. If accidents were uniformly distributed over the weekdays then the expected number of accidents in the SunWed group would be $(233+142)\times4/7 = 214.28$ with a binomial standard deviation of 9.583. So the departure from uniformity is just significant at the 5% level on a two-sided test.

<u>DRINK-WKGRP</u> are associated in Ontario but probably not in Manitoba. Among the accidents, in Ontario, in SunWed group 75% of them were owing to driving after drinking while 63% of the accidents in ThuSat group were owing to drinking. It implies that drunken driving more likely happened in Sunday through Wednesday.

<u>DRINK-MONTH</u> are not associated. There are more car fatality accidents in warm season than in cold season. This is also true in alcohol level.

<u>DRINK-YEAR</u> might be associated. The annual car fatality accidents fluctuated over years. There are more accidents in 1994 (17.5%) and fewest in 1998 (7.6%). The percentage of accidents with drink=yes for 1998 is relatively high in both Ontario and Manitoba. Figure 0 shows there is considerable variation here. One explanation is that with the increased use of seatbelts, daytime running lights and airbags only very serious crashes with often involve alcohol result in fatalities.

<u>HOUR-WKGRP</u> are likely unassociated. There are 25% more fatalities in the ThuSat-11PM slot and 28% fewer in the ThuSat-3AM slot than expected under the assumption of independence but the difference is not statistically significant.

<u>HOUR-MONTH</u> are not likely associated. However, in Ontario, there are 231% more fatalities than expected in the 2AM-April slot than would be expected under independence.

<u>HOUR-YEAR</u> are associated in Ontario but probably not in Manitoba. Hour and year are associated. There seems to be an excess of fatalities in the 2AM-1998 and 3AM-1997 windows (97% and 110% respectively).

WKGRP-MONTH are associated in Ontario. In Ontario, there are increased fatalities in ThuSat-Feb (65%) and ThuSat-Dec (58%) and decreased fatalities in ThuSat-May (42%) and ThuSat-Nov (56%).

WKGRP-YEAR are not associated.

Frequency Analysis of Drink

In Ontario out of a total of 375 cases, there are 44 cases where drink=unknown and only 2 cases in Manitoba. Thus in Ontario there remain 331 cases.

```
>crosstabs(~drink, data=tirf.df, subset=province=="Ontario", na.action=na.exclude)
Call:
crosstabs(formula = ~ drink, data = tirf.df, subset = province == "Ontario",
     na.action = na.exclude)
375 cases in table
+---+
| N |
|N/Total|
+----+
drink |
----+
unknown| 44
           |0.12 |
----+
no |107 |
|0.29 |
----+
yes |224
|0.6
            ----+
> crosstabs(~drink, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude)
Call:
crosstabs(formula = ~ drink, data = tirf.df, subset = province == "Manitoba", na.action
= na.exclude)
58 cases in table
+----+
| N
      |N/Total|
+----+
drink |
----+
unknown| 2 |
 0.034
----+
no |12 |
    0.21
-----+
yes |44 |
    |0.76 |
----+
```

Drink and hour are independent

Ontario

About 70% of fatalities involve alcohol. Although the association between alcohol fatalities and hour is not statistically significant (p=18%) we see that there is a slight increase in alcohol usage associated with deaths at 2AM (11% extra) and 3AM (13% extra).

> crosstabs(~drink+hour, data=tirf.df, subset=province=="Ontario", na.action=na.exclude)
Call:
crosstabs(~ drink + hour, data = tirf.df, subset = province == "Ontario", na.action = na.exclude)
316 cases in table

++ N N/RowTotal N/ColTotal N/Total ++										
drink		12AM	1AM	2AM	3AM	RowTotl				
no	0.281 0.370	18 0.188 0.305 0.057	0.281 0.360	0.146 0.226	0.104 0.213	96 0.3 				
-	0.209 0.630	41 0.186 0.695 0.130	0.218 0.640	0.218 0.774	0.168 0.787	220 0.7 				
ColTotl		59 0.19				316 				
Chi Yate > get.c [, [1,]	^2 = 6.2 es' corre rosstabs 1] [,2] 22 0	ndence of 51495 d.f ection no .percente [,3] [,4] 18 -20 -8 12	E.= 4 (p= ot used error(~d:] [,5] 6 -30	=0.181137	,	-				

Manitoba

In Manitoba about 77% of fatalities are alcohol associated. Again the association with hour does not appear to be statistically significant but there is a 15% excess of alcohol fatalities in the 3AM window.

```
> crosstabs(~drink+hour, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude)
Call:
crosstabs( ~ drink + hour, data = tirf.df, subset = province == "Manitoba", na.action = na.exclude)
53 cases in table
```

+ N N/RowT N/ColT N/Tota + drink	otal 1 +	12AM	1AM	2AM	3AM	RowTotl	I
no	0.235	2 0.167 0.222 0.038	0.250 0.273	0.286	0.083 0.111	+ 12 0.23 	+
-	0.765	7 0.171 0.778 0.132	0.195 0.727	0.714	0.195 0.889	+ 41 0.77 	+
ColTotl		9 0.17		0.13	0.17	Ì	+
Chi Yat Som > get.c [, [1,]	^2 = 0.9 es' corr e expect rosstabs 1] [,2] 4 -2	endence o 0669428 d cection no ced values s.percent [,3] [,4 20 2 -6 -	.f.= 4 (] ot used s are le: error.ma] [,5] 6 -51	ctors p=0.9147 ss than s	5, don't		+ ated p-value

Drink and wkgrp are associated

Ontario

```
> crosstabs(~wkgrp, data=tirf.df, subset=province=="Ontario")
Call:
crosstabs( ~ wkgrp, data = tirf.df, subset = province == "Ontario")
375 cases in table
+---+
| N
|N/Total|
+---+
wkgrp |
----+
SunWed |233
            1
 |0.62
            1
-----+------
ThuSat |142 |
    |0.38
-----
```

The car fatality accidents in SunWed group count for 62% of total accidents while 38% in ThuSat group. If accidents were uniformly distributed over the weekdays then the expected number of accidents in the SunWed group would be $(233+142)\times4/7 = 214.28$ with a binomial standard deviation of 9.583. So the departure from uniformity is just significant at the 5% level on a two-sided test.

Among the accidents in SunWed group 75% of them were owing to driving after drinking while 63% of the accidents in ThuSat group were owing to drinking. It implies that drunken driving more likely happened in Sunday through Wednesday.

```
> crosstabs(~drink+wkgrp, data=tirf.df, subset=province=="Ontario", na.action=na.exclude)
Call:
crosstabs ( ~ drink + wkgrp, data = tirf.df, subset = province == "Ontario", na.action = na.exclude)
331 cases in table
+---+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
drink |wkgrp
     |SunWed |ThuSat |RowTotl|
no | 56 | 51 |107 |
     |0.52 |0.48 |0.32 |
     |0.27 |0.42 |
|0.17 |0.15 |
                          ----+
yes |154 | 70 |224 |
|0.69 |0.31 |0.68 |
|0.73 |0.58 | |
     0.47 0.21
----+
ColTotl|210 |121 |331
|0.63 |0.37 |
                          1
----+
Test for independence of all factors
  Chi^2 = 8.411244 d.f.= 1 (p=0.003729074)
  Yates' correction not used
> get.crosstabs.percenterror.ont(~drink+wkgrp)
   [,1] [,2]
[1,] -18 30
[2,] 8 -15
```

Manitoba

In Manitoba there is no association and the pattern is different from Ontario. In Manitoba 79% of fatalities involve drinking as compared with 68% in Ontario and 73% of accidents occur in the SunWed wkgrp as compared with 63% in Ontario.

```
> get.crosstabs.percenterror.ont(~drink+wkgrp)
   [,1] [,2]
[1,] -18 30
[2,] 8 -15
> crosstabs(~drink+wkgrp, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude)
Call:
crosstabs( ~ drink + wkgrp, data = tirf.df, subset = province == "Manitoba", na.action = na.exclude)
56 cases in table
+----+
IN
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
drink |wkgrp
    |SunWed |ThuSat |RowTotl|
----+
no | 9 | 3 |12
                          1
     |0.750 |0.250 |0.21
                          ____
                    |0.220 |0.200
                           |0.161 |0.054 |
----+
    |32 |12 |44
|0.727 |0.273 |0.79
yes
                           1
                          |0.780 |0.800 |
     |0.571 |0.214 |
-----+----+----+
ColTotl|41 |15 |56 |
     |0.73 |0.27 |
-----+
Test for independence of all factors
   Chi^2 = 0.0248337 d.f.= 1 (p=0.8747821)
   Yates' correction not used
   Some expected values are less than 5, don't trust stated p-value
> get.crosstabs.percenterror.man(~drink+wkgrp)
   [,1] [,2]
[1,] 2 -7
[2,] -1 2
```

Drink and month are not associated

Ontario

+----+

There are more car fatality accidents in warm season than in cold season. This is also true in alcohol level.

> crosstabs(~drink+month, data=tirf.df, subset=province=="Ontario", na.action=na.exclude)
Call:
crosstabs(~ drink + month, data = tirf.df, subset = province == "Ontario", na.action = na.exclude)
331 cases in table

N N/RowTotal N/ColTotal N/Total +													
drink	+ month Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	RowTotl
no	6 0.056 0.429 0.018	6 0.056 0.300 0.018	9 0.084 0.391 0.027	2 0.019 0.100 0.006	11 0.103 0.367 0.033	11 0.103 0.314 0.033	13 0.121 0.361 0.039	7 0.065 0.200 0.021	9 0.084 0.281 0.027	14 0.131 0.359 0.042	8 0.075 0.308 0.024	11 0.103 0.524 0.033	107 0.32
yes	8 0.036 0.571 0.024	14 0.062 0.700 0.042	14 0.062 0.609 0.042	18 0.080 0.900 0.054	19 0.085 0.633 0.057	24 0.107 0.686 0.073	23 0.103 0.639 0.069	28 0.125 0.800 0.085	23 0.103 0.719 0.069	25 0.112 0.641 0.076	18 0.080 0.692 0.054	10 0.045 0.476 0.030	224 0.68
ColTotl	L 14 0.042	20 0.060	23 0.069	20 0.060	30 0.091	35 0.106	36 0.109	35 0.106	32 0.097	39 0.118	26 0.079	21 0.063	331

Test for independence of all factors

Chi^2 = 13.11596 d.f.= 11 (p=0.2858142)

Yates' correction not used

Some expected values are less than 5, don't trust stated p-value

> get.crosstabs.percenterror.ont(~drink+month)

[1,] 33 -7 21 -69 13 -3 12 -38 -13 11 -5 62		[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]	[,9]	[,10]	[,11]	[,12]
	[1,]	33	-7	21	-69	13	-3	12	-38	-13	11	-5	62

[1,]	33	-7	21	-69	13	-3	12	-38	-13	11	-5	62
[2,]	-16	3	-10	33	-6	1	-6	18	6	-5	2	-30

1

Manitoba

Less data but otherwise similar to Ontario.

> crosstabs(~drink+month, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude) Call: crosstabs(~ drink + month, data = tirf.df, subset = province == "Manitoba", na.action = na.exclude) 56 cases in table +----+ | N |N/RowTotal| |N/ColTotal| |N/Total | +----+ drink |month |Jan |Feb |Mar |Apr |May |Jun |Jul |Aug |Sep |Oct |Nov |Dec |RowTotl| _____+ |12 no

 |0.167
 |0.083
 |0.000
 |0.083
 |0.000
 |0.167
 |0.167
 |0.000
 |0.167
 |0.000
 |0.167
 |0.000
 |0.167
 |0.000
 |0.167
 |0.000
 |0.167
 |0.000
 |0.167
 |0.000
 |0.167
 |0.000
 |0.21

 |0.500
 |0.500
 |0.000
 |0.167
 |0.000
 |0.286
 |0.222
 |0.400
 |0.000
 |0.222
 |0.000
 |

 |0.036
 |0.018
 |0.000
 |0.036
 |0.036
 |0.036
 |0.036
 |0.000
 |0.036
 |0.000
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 |0.000
 |0. yes |2 |1 |1 |1 |5 |2 |5 |7 |3 |8 |7 |2 |44 |0.045 |0.023 |0.023 |0.023 |0.114 |0.045 |0.114 |0.159 |0.068 |0.182 |0.159 |0.045 |0.79 0.500 |0.500 |1.000 |1.000 |0.833 |1.000 |0.714 |0.778 |0.600 |1.000 |0.778 |1.000 | 0.036 0.018 0.018 0.018 0.089 0.036 0.089 0.125 0.054 0.143 0.125 0.036 ColTotl|4 |2 |1 |1 |6 |2 |7 |9 |5 |8 |9 |2 |56 0.071 |0.036 |0.018 |0.018 |0.107 |0.036 |0.125 |0.161 |0.089 |0.143 |0.161 |0.036 | Test for independence of all factors Chi^2 = 8.051178 d.f.= 11 (p=0.7087121) Yates' correction not used

Some expected values are less than 5, don't trust stated p-value > get.crosstabs.percenterror.man(~drink+month)
> get.crosstabs.percenterror.man(~drink+month)

	[,⊥]	[,4]	[,3]	[,4]	[,]]	[,6]	[,/]	[,8]	[,9]	[, 10]	[, ⊥⊥]	[,12]
[1,]	133	133	-100	-100	-22	-100	33	4	87	-100	4	-100
[2,]	-36	-36	27	27	6	27	-9	-1	-24	27	-1	27
	>											

Drink and year are perhaps associated (p=10%).

Ontario

The annual car fatality accidents fluctuated over years. There are more accidents in 1994 (17.5%) and fewest in 1998 (7.6%).

Although there are fewest accidents in 1998, 76% are associated with drinking alcohol. This suggests that the reduction in deaths is more due to improved car design, day time running lights and air bags than is due to the effectiveness of anti-drinking campaigns.

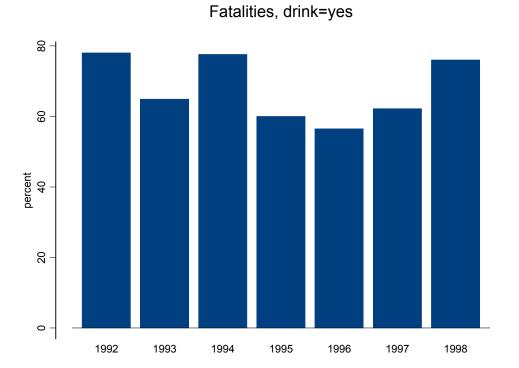
```
> get.crosstabs.percenterror.man(~drink+month)
    [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11] [,12]
[1,] 133 133 -100 -100 -22 -100 33 4 87 -100 4 -100
[2,] -36 -36 27 27 6 27 -9 -1 -24 27 -1 27
> crosstabs (~drink+year, data=tirf.df, subset=province=="Ontario", na.action=na.exclude)
Call:
crosstabs( ~ drink + year, data = tirf.df, subset = province == "Ontario", na.action = na.exclude)
331 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
drink |year
      |1992 |1993 |1994 |1995 |1996 |1997 |1998 |RowTotl|
|11 |20 |13 |20 |20 |17 | 6
|0.103 |0.187 |0.121 |0.187 |0.187 |0.159 |0.056
                                                            |107
no
                                                                    1
                                                            10.32
      0.220 0.351 0.224 0.400 0.435 0.378 0.240
                                                            0.033 0.060 0.039 0.060 0.060 0.051 0.018
_____+
yes |39 |37 |45 |30 |26 |28 |19 |224
                                                                    |0.174 |0.165 |0.201 |0.134 |0.116 |0.125 |0.085 |0.68

        |0.780
        |0.649
        |0.776
        |0.600
        |0.565
        |0.622
        |0.760

        |0.118
        |0.112
        |0.136
        |0.091
        |0.079
        |0.085
        |0.057

                                                            ColTotl|50 |57 |58 |50 |46 |45 |25 |331 |
|0.151 |0.172 |0.175 |0.151 |0.139 |0.136 |0.076 |
Test for independence of all factors
   Chi^2 = 10.60544 d.f.= 6 (p=0.1013636)
   Yates' correction not used
> get.crosstabs.percenterror.ont(~drink+year)
   [,1] [,2] [,3] [,4] [,5] [,6] [,7]
[1,] -32 9 -31 24 34 17 -26
[2,] 15 -4 15 -11 -16 -8 12
```

Figure 0. Percent of fatalities with drink=yes in Ontario



Manitoba

The pattern observed in Ontario appears again. Although 1998 had the fewest number of fatalities, 100% of the fatalities were associated with drinking but there were only a total of 5 fatalities recorded in Manitoba for 1998. There is not enough data for the standard test to obtain an accurate p-value.

> crosstabs(~drink+year, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude)
Call:
crosstabs(~ drink + year, data = tirf.df, subset = province == "Manitoba", na.action = na.exclude)
56 cases in table

+ N N/RowT N/ColT N/Tota	otal							
	+ year 1992	1993	1994	1995	1996	1997	1998	RowTotl
	0.167	0.111	0.182	0.083 0.200	0.375	0.333	0 0.000 0.000 0.000	12 0.21
1	0.833	0.889	0.205	0.800	0.625	0.182		+++ 44 0.79
	+ 6 0.107		11 0.196	1 -		+ 12 0.214	+ 5 0.089	-++ 56
Chi Yate Some > get.c [,	^2 = 4.3 es' corr e expect rosstabs 1] [,2] 22 -48	25773 d. ection n ed value .percent [,3] [,4 -15 -	f all fa f.= 6 (p ot used s are le cerror.ma [] [,5] [7 75 2 -20	=0.63268 ss than n(~drink ,6] [,7] 56 -100	5, don't +year)	trust s	tated p-	-++

Hour and wkgrp are independent (p-value=17%)

Ontario

There are 25% more fatalities in the ThuSat-11PM slot and 28% fewer in the ThuSat-3AM slot than expected under the assumption of independence but the difference is not statistically significant.

> crosstabs(~wkgrp+hour, data=tirf.df, subset=province=="Ontario", na.action=na.exclude) Call: crosstabs(~ wkgrp + hour, data = tirf.df, subset = province == "Ontario", na.action = na.exclude) 347 cases in table +----+ | N |N/RowTotal| |N/ColTotal| |N/Total | +----+ wkgrp |hour |11PM |12AM |1AM |2AM |3AM |RowTotl| -+---+ SunWed |44 |46 |50 |45 |37 |222 |0.198 |0.207 |0.225 |0.203 |0.167 |0.64 | |0.550 |0.676 |0.602 |0.682 |0.740 | |0.127 |0.133 |0.144 |0.130 |0.107 | 1 ThuSat |36 |22 |33 |21 |13 |125
 |0.288
 |0.176
 |0.264
 |0.168
 |0.104
 |0.36
 |

 |0.450
 |0.324
 |0.398
 |0.318
 |0.260
 |

 |0.104
 |0.063
 |0.095
 |0.061
 |0.037
 |
 ColTotl|80 |68 |83 |66 |50 |347 |0.23 |0.20 |0.24 |0.19 |0.14 | 1 Test for independence of all factors Chi^2 = 6.38332 d.f.= 4 (p=0.1722922) Yates' correction not used > get.crosstabs.percenterror.ont(~wkgrp+hour) [,1] [,2] [,3] [,4] [,5] [1,] -14 6 -6 7 16 [2,] 25 -10 10 -12 -28

Manitoba

As in Ontario there are much fewer (55% here) fatalities in the ThuSat-3AM slot but the numbers are small and the difference does not appear to be statistically significant.

```
> crosstabs (~wkqrp+hour, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude)
Call:
crosstabs ( ~ wkgrp + hour, data = tirf.df, subset = province == "Manitoba", na.action = na.exclude)
53 cases in table
```

```
+----+
| N
|N/RowTotal|
[N/ColTotal]
|N/Total |
+----+
wkgrp |hour
    |11PM |12AM |1AM |2AM |3AM |RowTotl|
SunWed |13 | 7 | 7 | 5 | 8 |40
                                             - I
   |0.325 |0.175 |0.175 |0.125 |0.200 |0.75 |
|0.765 |0.778 |0.636 |0.714 |0.889 |
|0.245 |0.132 |0.132 |0.094 |0.151 |

      ThuSat | 4 | 2 | 4 | 2 | 1 |13

      |0.308 |0.154 |0.308 |0.154 |0.077 |0.25

      |0.235 |0.222 |0.364 |0.286 |0.111 |

                                              1
                                             - I
     0.075 0.038 0.075 0.038 0.019
ColTotl|17 |9 |11 |7 |9 |53 |
 |0.32 |0.17 |0.21 |0.13 |0.17 |
                                              1
Test for independence of all factors
  Chi^2 = 1.804387 d.f.= 4 (p=0.7716796)
  Yates' correction not used
  Some expected values are less than 5, don't trust stated p-value
> get.crosstabs.percenterror.man(~wkgrp+hour)
   [,1] [,2] [,3] [,4] [,5]
[1,] 1 3 -16 -5 18
[2,] -4 -9 48 16 -55
```

Hour and Month

Ontario

Overall there is no strong indication of associated. However there are 231% more fatalities than expected in the 2AM-April slot than would be expected under independence.

> crosstabs (~hour+month, data=tirf.df, subset=province=="Ontario", na.action=na.exclude) Call: crosstabs(~ hour + month, data = tirf.df, subset = province == "Ontario", na.action = na.exclude) 347 cases in table +----+ | N IN/RowTotall |N/ColTotal| |N/Total | +---hour |month |Jan |Feb |Mar |Apr |May |Jun |Jul |Aug |Sep |Oct |Nov |Dec |RowTotl| 11PM | 2 | 4 | 6 | 5 | 6 |12 | 9 | 9 | 7 |10 | 6 | 4 |80 | 0.0250 |0.0500 |0.0750 |0.0625 |0.0750 |0.1500 |0.1125 |0.1125 |0.0875 |0.1250 |0.0750 |0.0500 |0.23 | 0.1538 |0.1905 |0.2727 |0.2174 |0.1935 |0.3077 |0.2432 |0.2195 |0.2258 |0.2439 |0.2400 |0.1739 | I 0.0058 0.0115 0.0173 0.0144 0.0173 0.0346 0.0259 0.0259 0.0202 0.0288 0.0173 0.0115 12AM | 4 | 4 | 3 | 1 | 6 | 9 |10 | 8 | 5 | 8 | 5 | 5 | 68 |

 |0.0588 |0.0588 |0.0441 |0.0147 |0.0882 |0.1324 |0.14/1 |0.11/0 |0.0735 |0.1170 |0.0735 |0.1170 |0.00141 |0.0141 |0.0141 |0.0141 |0.0144 |0.02174 |

 |0.3077 |0.1905 |0.1364 |0.0435 |0.1935 |0.2308 |0.2703 |0.1951 |0.1613 |0.1951 |0.2000 |0.2174 |

 1AM | 4 | 8 | 8 | 1 | 9 | 4 |11 |11 | 9 | 8 | 5 | 5 |83 | 0.0482 0.0964 0.0964 0.0120 0.1084 0.0482 0.1325 0.1325 0.1384 0.0964 0.0602 0.0602 0.24 0.3077 |0.3810 |0.3636 |0.0435 |0.2903 |0.1026 |0.2973 |0.2683 |0.2903 |0.1951 |0.2000 |0.2174 | 0.0115 |0.0231 |0.0231 |0.0029 |0.0259 |0.0115 |0.0317 |0.0317 |0.0259 |0.0231 |0.0144 |0.0144 | 1 2AM | 1 | 4 | 1 | 13 | 7 | 6 | 4 | 8 | 7 | 7 | 2 | 6 | 66 | 0.0152 |0.0606 |0.0152 |0.1970 |0.1061 |0.0909 |0.0606 |0.1212 |0.1061 |0.1061 |0.0303 |0.0909 |0.19 |0.0769 |0.1905 |0.0455 |0.5652 |0.2258 |0.1538 |0.1081 |0.1951 |0.2258 |0.1707 |0.0800 |0.2609 | |0.0029 |0.0115 |0.0029 |0.0375 |0.0202 |0.0173 |0.0115 |0.0231 |0.0202 |0.0202 |0.0058 |0.0173 | |0.0202 |0.0202 |0.0202 |0.0173 | |0.0202 |0.0202 |0.0202 |0.0058 |0.0173 | |0.0202 |0.0202 |0.0202 |0.0202 |0.0202 |0.0058 |0.0173 | |0.0202 |0.0202 |0.0202 |0.0202 |0.0058 |0.0173 | |0.0202 |0.0202 |0.0202 |0.0202 |0.0202 |0.0173 |0.0202 |0.0202 |0.0202 |0.0202 |0.0058 |0.0173 |0.0202 |0.0202 |0.0202 |0.0202 |0.0058 |0.0173 |0.0202 |0.0202 |0.0202 |0.0202 |0.0058 |0.0173 |0.0202 |0.0202 |0.0202 |0.0202 |0.0173 |0.0202 |0.0202 |0.0202 |0.0058 |0.0173 |0.0202 |0.0202 |0.0202 |0.0058 |0.0173 |0.0202 |0.0202 |0.0058 |0.0173 |0.0202 |0.0202 |0.0058 |0.0173 |0.0202 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0202 |0.0058 |0.0173 |0.0058 |0.0173 |0.0058 |0.0173 |0.0058 |0.0173 |0.0058 |0.0173 |0.0058 |0.0173 |0.0058 |0.0173 |0.0058 |0.0173 |0.0058 |0.0173 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0.0058 |0

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 |0.1538 |0.0476 |0.1818 |0.1304 |0.0968 |0.2051 |0.0811 |0.1220 |0.0968 |0.1951 |0.2800 |0.1304 | | I 0.0058 |0.0029 |0.0115 |0.0086 |0.0086 |0.0231 |0.0086 |0.0144 |0.0086 |0.0231 |0.0202 |0.0086 |

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Test for independence of all factors

Chi^2 = 53.12007 d.f.= 44 (p=0.1630167)

Yates' correction not used

Some expected values are less than 5, don't trust stated p-value

Manitoba

> get.crosstabs.percenterror.ont(~hour+month) [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11] [,12] [1,] -33 -17 18 -6 -16 33 6 -5 -2 6 4 -25 [2,] 57 -3 -30 -78 -1 18 38 0 -18 0 2 11 [2,]

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 > crosstabs(~hour+month, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude) Call: crosstabs (~ hour + month, data = tirf.df, subset = province == "Manitoba", na.action = na.exclude) 53 cases in table +----+ IN I |N/RowTotal| |N/ColTotal| |N/Total | +---+ hour |month |Jan |Feb |Mar |Apr |May |Jun |Jul |Aug |Sep |Oct |Nov |Dec |RowTotl| 0.118 |0.000 |0.000 |0.000 |0.118 |0.059 |0.118 |0.235 |0.118 |0.176 |0.059 |0.000 |0.32

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 ColTotl|4 |2 |1 |1 |5 |2 |7 |9 |3 |8 |9 |2 |53 | |0.075 |0.038 |0.019 |0.019 |0.094 |0.038 |0.132 |0.170 |0.057 |0.151 |0.170 |0.038 | | _____+ Test for independence of all factors Chi^2 = 48.19941 d.f.= 44 (p=0.3068645) Yates' correction not used Some expected values are less than 5, don't trust stated p-value > get.crosstabs.percenterror.man(~hour+month) [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11] [,12] [1,] 56 -100 -100 -100 25 56 -11 39 108 17 -65 -100 [2,] -100 194 -100 -100 -100 194 152 -35 -100 [3,] -100 -100 382 -100 -4 -100 38 61 -100 -26 -35 7 194 -40 141 [4,] 279 -100 -100 -100 -100 -100 -100 -100 -5 237 -100 [5,] -100 194 -100 489 136 -100 -100 -35 96 47 -35 -100

Hour and year are associated in Ontario

Ontario

Hour and year are associated. There seems to be an excess of fatalities in the 2AM-1998 and 3AM-1997 windows (97% and 110% respectively).

> crosstabs(~hour+year, data=tirf.df, subset=province=="Ontario", na.action=na.exclude) Call: crosstabs(~ hour + year, data = tirf.df, subset = province == "Ontario", na.action = na.exclude) 347 cases in table +----+ IN |N/RowTotal| |N/ColTotal| |N/Total | +----+ hour |year |1992 |1993 |1994 |1995 |1996 |1997 |1998 |RowTotl| 11PM |12 |10 |14 |14 |15 |11 | 4 |80 0.1500 0.1250 0.1750 0.1750 0.1875 0.1375 0.0500 0.23 |0.2105 |0.1639 |0.2500 |0.2800 |0.3261 |0.2075 |0.1667 | 0.0346 0.0288 0.0403 0.0403 0.0432 0.0317 0.0115 0.2647 0.1029 0.1176 0.1618 0.1324 0.1912 0.0294 0.20 |0.3158 |0.1148 |0.1429 |0.2200 |0.1957 |0.2453 |0.0833 | 0.0519 0.0202 0.0231 0.0317 0.0259 0.0375 0.0058 1AM |11 |24 |15 |12 |12 |4 |5 |83 0.1325 |0.2892 |0.1807 |0.1446 |0.1446 |0.0482 |0.0602 |0.24 |0.1930 |0.3934 |0.2679 |0.2400 |0.2609 |0.0755 |0.2083 | 0.0317 0.0692 0.0432 0.0346 0.0346 0.0115 0.0144 2AM |10 |10 |15 |10 |3 |9 |9 |66 0.1515 |0.1515 |0.2273 |0.1515 |0.0455 |0.1364 |0.1364 |0.19 |0.1754 |0.1639 |0.2679 |0.2000 |0.0652 |0.1698 |0.3750 | 0.0288 0.0288 0.0432 0.0288 0.0086 0.0259 0.0259 3AM | 6 |10 | 4 | 3 | 7 |16 | 4 |50 0.1200 |0.2000 |0.0800 |0.0600 |0.1400 |0.3200 |0.0800 |0.14 0.1053 0.1639 0.0714 0.0600 0.1522 0.3019 0.1667 |0.0173 |0.0288 |0.0115 |0.0086 |0.0202 |0.0461 |0.0115 | ColTot1|57 |61 |56 |50 |46 |53 |24 |347 1 |0.164 |0.176 |0.161 |0.144 |0.133 |0.153 |0.069 | __+____+ Test for independence of all factors Chi^2 = 51.37488 d.f.= 24 (p=0.0009432535) Yates' correction not used Some expected values are less than 5, don't trust stated p-value > get.crosstabs.percenterror.ont(~hour+year) [,1] [,2] [,3] [,4] [,5] [,6] [,7] $\begin{bmatrix} 1, 1 \end{bmatrix} -9 -29 & 8 & 21 & 41 & -10 & -28 \\ \begin{bmatrix} 2, 1 \end{bmatrix} 61 & -41 & -27 & 12 & 0 & 25 & -57 \\ \begin{bmatrix} 3, 1 \end{bmatrix} -19 & 64 & 12 & 0 & 9 & -68 & -13 \\ \begin{bmatrix} 4, 1 \end{bmatrix} -8 & -14 & 41 & 5 & -66 & -11 & 97 \\ \end{bmatrix}$ [5,] -27 14 -50 -58 6 110 16

Manitoba

The crosstabs are included for Manitoba but since the numbers are so small we can't make any reliable conclusion.

> crosstabs(~hour+year, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude) Call: crosstabs(~ hour + year, data = tirf.df, subset = province == "Manitoba", na.action = na.exclude) 53 cases in table

+ N N/RowT N/ColT N/Tota +	otal									
hour	year 1992	1993	1994	1995	1996	1997	1998	RowTotl		
 11PM		0.118 0.222		1 0.059 0.200 0.019	2 0.118 0.250 0.038		0 0.000 0.000 0.000	17 0.32 		
12AM		0.000 0.000		2 0.222 0.400 0.038	1 0.111 0.125 0.019		2 0.222 0.500 0.038	9 0.17 		
 1AM		0.455 0.556		1 0.091 0.200 0.019	2 0.182 0.250 0.038		0 0.000 0.000 0.000	11 0.21 		
2AM		0.143 0.111		0 0.000 0.000 0.000	1 0.143 0.125 0.019		1 0.143 0.250 0.019	7 0.13 		
ЗАМ		0.111 0.111		1 0.111 0.200 0.019	2 0.222 0.250 0.038		1 0.111 0.250 0.019	9 0.17 		
ColTotl			10 0.189	5 0.094	8 0.151	12 0.226	+ 4 0.075	53 		
<pre>Test for independence of all factors Chi^2 = 35.74037 d.f.= 24 (p=0.05815202) Yates' correction not used Some expected values are less than 5, don't trust stated p-value > get.crosstabs.percenterror.man(~hour+year) [,1] [,2] [,3] [,4] [,5] [,6] [,7] [1,] -100 -31 87 -38 -22 56 -100 [2,] 18 -100 77 136 -26 -100 194 [3,] -4 168 -100 -4 20 -20 -100 [4,] -100 -16 -24 -100 -5 89 89 [5,] 253 -35 -100 18 47 -51 47 ></pre>										

Wkgrp and month are associated in Ontario (p-value $\approx 1\%$)

Ontario

There are increased fatalities in ThuSat-Feb (65%) and ThuSat-Dec (58%) and decreased fatalities in ThuSat-May (42%) and ThuSat-Nov (56%).

<pre>> crosstabs(~wkgrp+month, data=tirf.df, subset=province=="Ontario", na.action=na.exclude) Call: crosstabs(~ wkgrp + month, data = tirf.df, subset = province == "Ontario", na.action = na.exclude) 375 cases in table ++ N N/RowTotal N/RowTotal N/ColTotal N/Total ++ </pre>													
wkgrp		Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	RowTotl
SunWed	9 0.039 0.562 0.024	9 0.039 0.375 0.024	13 0.056 0.542 0.035	16 0.069 0.667 0.043	25 0.107 0.781 0.067	27 0.116 0.692 0.072	24 0.103 0.632 0.064	29 0.124 0.659 0.077	19 0.082 0.514 0.051	27 0.116 0.643 0.072	25 0.107 0.833 0.067	10 0.043 0.400 0.027	233 0.62
ThuSat	7 0.049 0.438 0.019	15 0.106 0.625 0.040	11 0.077 0.458 0.029	8 0.056 0.333 0.021	7 0.049 0.219 0.019	12 0.085 0.308 0.032	14 0.099 0.368 0.037	15 0.106 0.341 0.040	18 0.127 0.486 0.048	15 0.106 0.357 0.040	5 0.035 0.167 0.013	15 0.106 0.600 0.040	142 0.38
ColTotl	L 16 0.043	24 0.064	24 0.064	24 0.064	32 0.085	39 0.104	38 0.101	44 0.117	37 0.099	42 0.112	30 0.080	25 0.067	375
Chi Yat > get.c [, [1,]	<pre> 0.043 0.064 0.064 0.085 0.104 0.101 0.117 0.099 0.112 0.080 0.067 Test for independence of all factors Chi^2 = 24.726 d.f.= 11 (p=0.00999656) Yates' correction not used > get.crosstabs.percenterror.ont(~wkgrp + month) [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11] [,12] [1,] -9 -40 -13 7 26 11 2 6 -17 3 34 -36</pre>												

Manitoba

The numbers are too small to draw any conclusion.

<pre>> crosstabs(~wkgrp+month, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude) Call: crosstabs(~ wkgrp + month, data = tirf.df, subset = province == "Manitoba", na.action = na.exclude) 58 cases in table ++ N N/RowTotal N/ColTotal N/Total ++</pre>													
wkgrp	month Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	RowTotl
	4 0.098 1.000 0.069	2 0.049 1.000 0.034	1 0.024 1.000 0.017	1 0.024 0.500 0.017	5 0.122 0.833 0.086	2 0.049 1.000 0.034	6 0.146 0.857 0.103	6 0.146 0.667 0.103	4 0.098 0.800 0.069	4 0.098 0.500 0.069	5 0.122 0.500 0.086	1 0.024 0.500 0.017	41 0.71
	0 0.000 0.000 0.000	1	0 0.000 0.000 0.000	1 0.059 0.500 0.017	1 0.059 0.167 0.017	0 0.000 0.000 0.000	1 0.059 0.143 0.017	3 0.176 0.333 0.052	1 0.059 0.200 0.017	4 0.235 0.500 0.069	5 0.294 0.500 0.086	1 0.059 0.500 0.017	17 0.29
ColTotl	4 0.069	2 0.034	1 0.017			2 0.034	7 0.121	9 0.155	5 0.086	8 0.138	10 0.172	2 0.034	58
Chi Yat Som > get.c [, [1,]	<pre>Test for independence of all factors Chi^2 = 9.781977 d.f.= 11 (p=0.5500987) Yates' correction not used Some expected values are less than 5, don't trust stated p-value > get.crosstabs.percenterror.man(~wkgrp + month) [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11] [,12]</pre>												

Wkgrp and year are not associated

Ontario

```
> crosstabs(~wkgrp+year, data=tirf.df, subset=province=="Ontario", na.action=na.exclude)
Call:
crosstabs ( ~ wkgrp + year, data = tirf.df, subset = province == "Ontario", na.action = na.exclude)
375 cases in table
+----+
IN
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
wkgrp |year
      |1992 |1993 |1994 |1995 |1996 |1997 |1998 |RowTotl|
SunWed |37 |41 |42 |30 |30 |34 |19 |233
|0.159 |0.176 |0.180 |0.129 |0.129 |0.146 |0.082 |0.62
                                                                        |0.607 |0.631 |0.677 |0.556 |0.625 |0.586 |0.704 |
|0.099 |0.109 |0.112 |0.080 |0.080 |0.091 |0.051 |
ThuSat |24 |24 |20 |24 |18 |24 | 8 |142

      |0.169
      |0.169
      |0.141
      |0.169
      |0.127
      |0.169
      |0.056
      |0.38

      |0.393
      |0.369
      |0.323
      |0.444
      |0.375
      |0.414
      |0.296
      |

      |0.064
      |0.064
      |0.053
      |0.064
      |0.064
      |0.064
      |0.021
      |

ColTotl|61 |65 |62 |54 |48 |58 |27 |375
|0.163 |0.173 |0.165 |0.144 |0.128 |0.155 |0.072 |
                                                                        1
Test for independence of all factors
   Chi^2 = 2.988713 d.f.= 6 (p=0.8102621)
   Yates' correction not used
> get.crosstabs.percenterror.ont(~wkgrp+year)
    [,1] [,2] [,3] [,4] [,5] [,6] [,7]
[1,] -2 2 9 -11 1 -6 13
[2,] 4 -2 -15 17 -1 9 -22
```

Manitoba

> crosstabs(~wkgrp+year, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude) Call: crosstabs (~ wkgrp + year, data = tirf.df, subset = province == "Manitoba", na.action = na.exclude) 58 cases in table +----+ | N 1 |N/RowTotal| |N/ColTotal| |N/Total | +----+ wkgrp |year |1992 |1993 |1994 |1995 |1996 |1997 |1998 |RowTotl| SunWed | 5 | 5 | 7 | 4 | 7 | 9 | 4 | 41 1 0.122 0.122 0.171 0.098 0.171 0.220 0.098 0.71 |0.833 |0.556 |0.636 |0.667 |0.875 |0.692 |0.800 | |0.086 |0.086 |0.121 |0.069 |0.121 |0.155 |0.069 | _____+ ThuSat |1 |4 |4 |2 |1 |4 |1 |17

 |0.059
 |0.235
 |0.235
 |0.118
 |0.059
 |0.235
 |0.059
 |0.29

 |0.167
 |0.444
 |0.364
 |0.333
 |0.125
 |0.308
 |0.200
 |

 |0.017
 |0.069
 |0.034
 |0.017
 |0.069
 |0.017
 |

 ColTotl|6 |9 |11 |6 |8 |13 |5 |58 | |0.103 |0.155 |0.190 |0.103 |0.138 |0.224 |0.086 | Test for independence of all factors $Chi^2 = 3.082463 d.f. = 6 (p=0.798428)$ Yates' correction not used Some expected values are less than 5, don't trust stated p-value > get.crosstabs.percenterror.man(~wkgrp+year) [,1] [,2] [,3] [,4] [,5] [,6] [,7] [1,] 18 -21 -10 -6 24 -2 13 [2,] -43 52 24 14 -57 5 -32 >

C. TIRF Dataset: Late Night Subset TIME SERIES

Summary

TIRF data set contains 400 observations of car-driver fatality accidents from January 1992 to December 1998. There were actually 434 cases in the original SPSS file but after deleting all cases where the hour variable was missing we were left with 400 cases. Of these 400 cases, 347 were in Ontario and 53 were in Manitoba.

Figure 1 and Table 1 show time series of total car fatality accidents per month for January 1992 to December 1998. In Manitoba there are fewer fatalities and there no noticeable change over time. In Ontario there is an increasing trend for about the first two years and this is followed by a decreasing trend starting around 1994. The table below indicates the death rate has been halved.

Determined in the cost structureBeginning of 1992End of 1998Percentage ChangeOntario4.52.349Manitoba0.320.37-13

Estimated mean death rate/month from loess smooth.

Figure 2 shows that in Ontario there was a marked downward trend starting around 1995 in fatalities with drink=yes. There also appears a slight decline in Manitoba for fatalities with drink=yes but the numbers are smaller and there is a lot of variability. The number of fatalities in Ontario with drink=no has also started to decline since about 1995 or 1996.

Figure 3a&b. The comparisons between driving with drink and without drink within drink factor, and between ThuSat and SunWed within wkgrp factor are shown in Fig 3 (a) and (b). There more fatalities with drink=yes and there are are more accidents in SunWed. In Ontario, both SunWed and ThuSat have a downward trend.

<u>Figure 6a,b-I,b-ii</u>, fatalities by hour, drink and province. In each hour slot the accidents are larger in drinking alcohol group than no drinking alcohol group. In <u>Figure 4b-ii</u>, there is an increasing trend in Ontario in the 2AM slot and decreasing trends at 11PM, 12AM and 1AM.

Figure 4a and 4b are the fatalities by wkgrp and hour. In each hour slot the accidents are larger in SunWed group than in ThuSat group. There are declining trends at 1AM in both weekgroups. At 2AM-ThuSat there is a downward trend. The other panels do not exhibit noticeable trends.

<u>Tables 5a-5d.</u> Monthly time series, deaths by province, hour and wkgrp. Mann-Kendall tests. Ontario 1AMSunWed and 2AMThuSat downward trend (<5%) and Manitoba 2AMSunWed upward trend (<5%).

<u>Tables 7a-d</u>. Annual total fatalities are decomposed by hour, wkgrp and drink for Ontario and Manitoba. Table 6b shows that there has been a shift in fatalities from early evening to late evening starting around 1996. The Mann-Kendall trend test is statistically significant on a two-sided test for Ontario fatalities with drink=yes for Total, 11PM, 1AM, SunWed and ThuSat and in all cases the sign of tau indicates a downward trend. The trend test is not significant for drink=no in Ontario. There are no trends in Manitoba for either drink=yes or drink=no.

No Figure 7 in this report.

Figures 8a,b,c,d. STL analysis for monthly time series with drink=yes and drink=no in Ontario. For drink=yes, a change occurred in 1996 and this is reflected in the trend and seasonal component. The peaks are higher in the seasonal after 1996 and the troughs are lower pre-1996. The shape of the seasonal component has changed. The seasonal component shows fatalities with drink peak in Aug and reach a minimum in Jan. There is a secondary peak in Oct. The trend is upward for Sep, Oct and Jan but is elsewhere downward and or level.

Figures 9a,b,c,d. STL analysis, Ontario, drink=no, downward trend since 1996. Although Mann-Kendall test is not significant the loess trend indicates a downward trend in recent years (Figure 9d) Seasonal component is very irregular and changing. Peak in July and trough in Apr. There is a lot of change over time.

Figures <u>10a,b</u> <u>11a,b</u> STL analysis for monthly time series with drink=yes and drink=no in Manitoba.

Data Visualization 1. All fatalities, Ontario and Manitoba

In the data visualization plots we have used a 60% robust linear smoother.

Table 1a. Monthly fatalities, Ontario

				5									
> tirf	> tirf.ont.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec												
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	1	4	4	4	б	10	3	3	5	б	5	6	
1993:	4	4	4	4	б	9	8	9	2	5	3	3	
1994:	2	5	5	5	4	4	7	5	4	7	5	3	
1995:	1	2	1	3	5	б	5	б	7	7	2	5	
1996:	1	2	5	2	7	3	б	8	3	7	2	0	
1997:	2	4	3	3	1	4	7	8	б	б	б	3	
1998:	2	0	0	2	2	3	1	2	4	3	2	3	
> sum	tirf	E.ont	t.ts) = 34	47								
> Sea	asor	nalM	lann	Ken	dal:	l(ti	irf.	ont	.ts)			
tau :	= -0	. 28	7.	g	1 =0	0.24	197%						
- uu					-			•					

Table 1b. Monthly fatalities, Manitoba

> tirf.man.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec

 1992:
 0
 0
 1
 0
 1
 1
 0
 1
 1
 0

 1993:
 0
 0
 0
 1
 0
 2
 1
 1
 2
 1
 1

 1994:
 0
 0
 0
 0
 1
 4
 1
 3
 1
 0

 1995:
 0
 1
 1
 0
 1
 0
 1
 0
 0
 0
 1

 1996:
 1
 1
 0
 0
 1
 1
 0
 2
 0
 0
 2
 0

 1997:
 3
 0
 0
 1
 0
 2
 1
 1
 2
 0

 1997:
 3
 0
 0
 1
 0
 2
 1
 1
 2
 2
 0

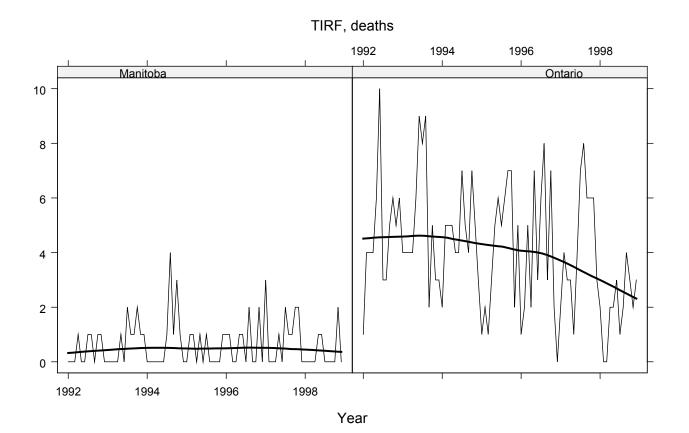
 1998:
 0
 0
 0
 1
 1
 0
 0
 0
 2
 0

 > sum(tirf.man.ts) [1] 53 > SeasonalMannKendall(tirf.man.ts) tau = 0.0324, sl =74.96% > pc.change(tirf.ont.ts) [1] 4.512848 2.311812 48.772666 [1] 3.783476 1.935217 48.850810 > pc.change(tirf.man.ts) [1] 0.3246352 0.3678582 -13.3143284 > aggregate(tirf.ont.ts,1,mean) > aggregate(tirf.ont.ts,1,mean) 1992: 4.750000 5.083333 4.666667 4.166667 3.833333 4.416667 2.000000 > aggregate(tirf.man.ts,1,mean) 1992: 0.4166667 0.7500000 0.8333333 0.4166667 0.6666667 1.0000000 0.3333333 > MannKendall(tirf.ont.ts) [1] -0.20029531 0.01088074 > MannKendall(tirf.man.ts) [1] 0.01823654 0.83558142

Table 1c-i. Estimated mean death rate/month from loess smooth.

	Beginning of 1992	End of 1998	Percentage Change
Ontario	4.5	2.3	49
Manitoba	0.32	0.37	-13

Figure 1. Time series trellis plot of total car fatality accidents per month for January 1992 to December 1998. In Manitoba there are fewer fatalities and there no noticeable change over time. In Ontario there is an increasing trend for about the first two years and this is followed by a decreasing trend starting around 1994. Because of the extremely low death rates for car fatalities in Manitoba, it does not seem to provide any useful information. It just gets in the way of understanding the Ontario data so it will be omitted from further graphical displays.

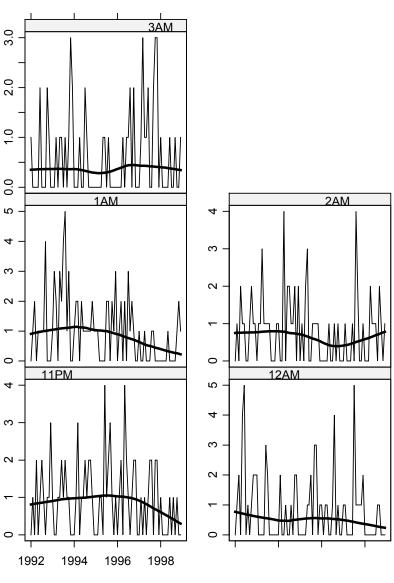


> ont.11PM.ts

2. Fatalities by hour Table 2. Time series tabulation by province and hour

man 11 DM	+ -										
man.11PM.	Feb	Mar	∆ nr	Max	Tun		λυα	Con	Oat	Nov	Dec
1992: 1	0	0	0	0	2	0 41	Aug 0	0	2	1	0
1993: 0	0	1	0	1	1	0	1	0	1	3	2
1994: 0		0	1	0	0	2	1	0	0	0	0
1995: 0		0	0	1	0	0	1	0	0	0	0
1996: 0	0	0	1	0	1	1	2	0	2	0	0
1997: 0	-	3	0	1	2	0	0	2	2	3	Ő
1998: 1		0	0	0	1	0	0	1	0	0	1
> MannKen		(man				-		_	-	-	_
tau = 0.0			=62		/						
> man.12A	,										
Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992: 0	0	0	0	0	0	1	Ō	0	0	0	0
1993: 0	0	0	0	0	0	0	0	0	0	0	0
1994: 0	0	0	0	0	0	1	1	0	1	0	0
1995: 0	0	0	0	0	0	1	0	0	0	0	1
1996: 0	1	0	0	0	0	0	0	0	0	0	0
1997: 0	0	0	0	0	0	0	0	0	0	0	0
1998: 0	0	0	0	0	1	0	0	0	0	1	0
> MannKen	dall	(man	.12AI	M.ts)						
tau = 0.0	306,	s	1 =73	3.949	20						
> man.1AM	.ts										
Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992: 0	0	0	0	0	0	0	0	0	0	1	0
1993: 0	0	0	0	0	0	2	1	0	1	0	1
1994: 0	0	0	0	0	0	0	0	0	0	0	0
1995: 0	0	1	0	0	0	0	0	0	0	0	0
1996: 0	0	0	0	0	0	0	2	0	0	0	0
1997: 0	0	0	0	1	0	0	0	0	0	1	0
1998: 0	0	0	0	0	0	0	0	0	0	0	0
> MannKen	dall										
tau = -0.		, :	sl =4	41.80	58						
> man.2AM											
	Feb										
1992: 0	0	0	0	0	0	0	0	0	0	0	0
1993: 0	0	0	0	0	0	0	0	0	0	1	0
1994: 0		0	0	0	0	0	0	0	0	1	0
1995: 0		0	0	0	0	0	0	0	0	0	0
1996: 0		0	0	0	0	0	0	0	0	1	0
1997: 2		0	0	0	0	0	0	0	1	0	0
1998: 0	0	0	0	0	0	0	0	0	0	1	0
> MannKen											
tau = 0.1 > man.3AM		SI	=18	./6							
	. ts Feb	Mar	∆ nr	Max	Tun		λυα	Con	Oat	Nov	Dec
1992: 0	0	Mai 0	дрі 1	May 0	0	0 0 0	Aug 1	998 0	1	0	Dec 0
1992: 0 1993: 0		0	0	0	0	0	0	0	1	0	0
1994: 0		0	0	0	0	0	0	0	0	0	0
1994: 0 1995: 0		0	0	0	0	0	0	0	0	0	0
1995: 0 1996: 0		0	0	1	0	0	0	0	0	1	0
1997: 0		0	0	0	0	0	0	1	0	0	0
1998: 0		0	0	1	0	0	0	0	0	0	0
> MannKen					0	0	0	0	0	0	0
tau = -0.			sl =!		58						
544 0.		, .	'	• 5							

Figure 2a. Time series line plot by hour

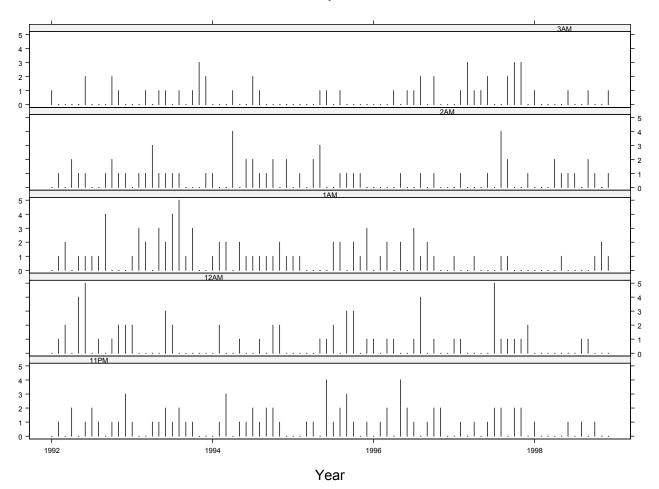


TIRF, deaths by hour, Ontario.

Year

Figure 2b. Time series loess analysis by hour

TIRF, deaths by hour, Ontario.

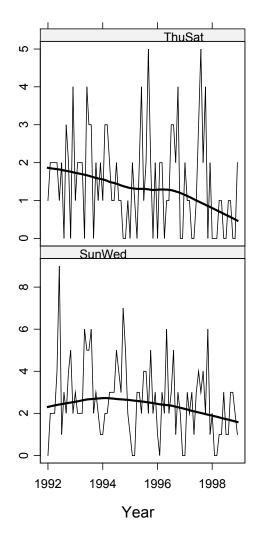


3. Fatalities by wkgrp

> ont	> ont.ThuSat.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	1	2	2	2	2	1	2	0	3	2	0	4
1993:	1	2	2	2	0	4	3	3	0	2	1	2
1994:	1	3	3	2	1	1	2	1	1	0	0	1
1995:	0	2	1	0	2	4	1	2	5	2	0	2
1996:	0	2	2	0	1	1	3	3	2	4	0	0
1997:	2	1	1	0	0	1	3	5	2	4	0	2
1998:	0	0	0	1	1	0	0	1	1	0	0	2
> Sea	sonal	LManı	Kend	lall	(ont	. Thu:	Sat.	ts)				
tau =	-0.2	243,	s	L =1.	.29%							
> ont	.Sun	Ved.t	s									
> ont				Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
> ont 1992:				Apr 2	May 4	Jun 9	Jul 1	Aug 3	Sep 2	Oct 4	Nov 5	Dec 2
	Jan	Feb	Mar									
1992:	Jan 0	Feb 2	Mar 2	2	4	9	1	3	2	4	5	2
1992: 1993:	Jan 0 3	Feb 2 2	Mar 2 2	2 2	4 6	9 5	1 5	3 6	2 2	4 3	5 2	2 1
1992: 1993: 1994:	Jan 0 3 1	Feb 2 2 2	Mar 2 2 2	2 2 3	4 6 3	9 5 3	1 5 5	3 6 4	2 2 3	4 3 7	5 2 5	2 1 2
1992: 1993: 1994: 1995:	Jan 0 3 1 1	Feb 2 2 2 0	Mar 2 2 2 0	2 2 3 3	4 6 3 3	9 5 3 2	1 5 5 4	3 6 4 4	2 2 3 2	4 3 7 5	5 2 5 2	2 1 2 3
1992: 1993: 1994: 1995: 1996:	Jan 0 3 1 1	Feb 2 2 2 0 0	Mar 2 2 2 0 3	2 2 3 3 2	4 6 3 3 6	9 5 3 2 2	1 5 4 3	3 6 4 4 5	2 2 3 2 1	4 3 7 5 3	5 2 5 2 2	2 1 2 3 0
1992: 1993: 1994: 1995: 1996: 1997:	Jan 0 3 1 1 1 0 2	Feb 2 2 0 0 3 0	Mar 2 2 0 3 2 0	2 2 3 2 3 2 3 1	4 6 3 6 1	9 5 2 2 3 3	1 5 4 3 4 1	3 6 4 5 3	2 2 3 2 1 4	4 3 7 5 3 2	5 2 5 2 6	2 1 2 3 0 1
1992: 1993: 1994: 1995: 1996: 1997: 1998: > Seas	Jan 0 3 1 1 1 0 2	Feb 2 2 0 0 3 0	Mar 2 2 0 3 2 0 Xenc	2 2 3 2 3 1 1 all	4 6 3 6 1	9 5 2 2 3 3 3	1 5 4 3 4 1	3 6 4 5 3	2 2 3 2 1 4	4 3 7 5 3 2	5 2 5 2 6	2 1 2 3 0 1

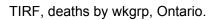
> sum(ont.SunWed.ts+ont.ThuSat.ts)
[1] 347

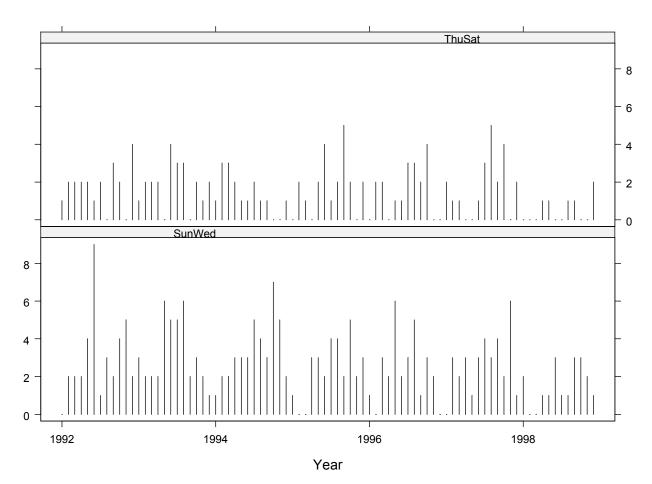
Figure 3a. Time series line plot by wkgrp



TIRF, deaths by wkgrp, Ontario.

Figure 3b. Time series line plot.

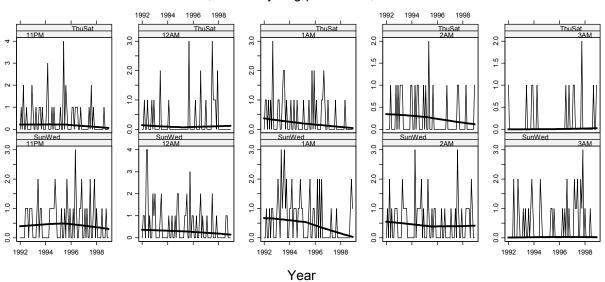




4. Fatalities by hour and wkgrp

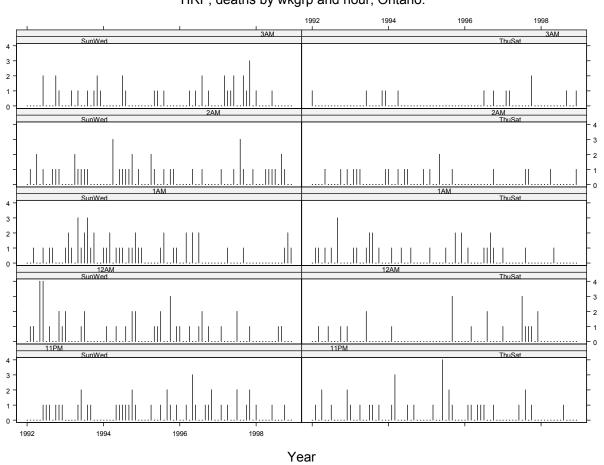
Figure 4a and 4b are the fatalities by wkgrp and hour. There are declining trends at 1AM in both weekgroups. At 2AM-ThuSat there is a downward trend. The other panels do not exhibit noticeable trends.

Figure 4a. Loess analysis



TIRF, deaths by wkgrp and hour, Ontario.

Figure 4b. Line plot of time series



TIRF, deaths by wkgrp and hour, Ontario.

Table 5a to 5d

Monthly time series, deaths by province, hour and wkgrp. Mann-Kendall tests. Ontario 12AMSunWed, 1AMSunWed and 1AMThuSat downward trend (<5%) and Manitoba 2AMSunWed upward trend (<5%).

Table 4a. Ontario-SunWed

> ont	11 DN	ເຊັນກາ	t hav									
- 0110		Feb			Max	Jun	.Tul	Aug	Sep	Oat	Nov	Dec
1992:	0 211	0	0	0	0	1	1	Lug 1	99C 0	1	1	1 1
1993:	0	0	0	0	1	2	0	1	1	0	0	0
1993:	0	0	0	0	1	1	1	1	1	2	1	0
1994:	0	0	0	1	1 0	1 0	1	0	2	2	0	1
1995:	0	0	1	0	3	1	1 0	0	1	1	2	0
	0	1	0	1	0	1 0	2	0	0	1	2	0
1997:												0
1998:	1	0	0	0	0	1	0	0	0	1	0	0
> Seas							MSum	wea.	_S)			
tau =					55.2	28						
> ont						-	- 1		~	<u> </u>		-
1000.			Mar	_	_				_			
1992:	0	1	1	0	4	4	0	1	0	0	2	1
1993:	2	0	0	0	0	1	2	0	0	0	0	0
1994:	0	1	0	0	1	0	0	1	0	2	2	0
1995:	0	0	0	0	1	1	2	0	0	3	0	1
1996:	1	0	0	1	0	0	1	2	0	1	0	0
1997:	0	1	0	0	0	0	2	0	0	0	1	0
1998:	0	0	0	0	0	0	0	1	1	0	0	0
> Sea							MSun	Ned.	ts)			
tau =					4.16	5%						
> ont						_		_	_	. .		_
1000.			Mar						_			Dec
1992:	0	0	1	0	0	1	0	1	1	0	0	0
1993:	1	2	1	0	3	1	2	3	1	2	0	0
1994:	1	1	2	0	1	1	1	0	1	1	2	1
1995:	1	0	0	0	0	0	1	2	0	0	1	1
1996:	0	0	2	0	2	0	2	0	0	0	0	0
1997:	0	0	0	1	0	0	0	0	1	0	0	0
1998:	0	0	0	0	0	0	0	0	0	1	2	1
> Sea					-		Sunwe	ed.ts	5)			
tau =					1.72	7%						
> ont						T	T 1		0	0		D
1000.			Mar	-	-			-	-		Nov	
1992:	0	1	0	2	0	1	0	0	1	1	1	0
1993:	0	0	0	2	1	1	1	1	0	0	0	0
1994:	0	0	0	3	0	1	1	1	1	2	0	1
1995:	0	0	0	2	1	0	0	1	0	1	1	0
1996:	0	0	0	0	1	0	0	1	0	0	0	0
1997:	0	1	0	0	0	1	0	3	1	0	0	1
1998:	0	0	0	1	1	1	1	0	2	1	0	0
> Sea							Sunwe	ed.ts	5)			
tau =					59.5	L≷						
> ont					Marr	T	T 1	7	0	0	Marr	Der
1002.			Mar	_	_				_			
1992:	0	0	0	0	0	2	0	0	0	2	1	0
1993:	0	0	1	0	1	0	0	1	0	1	2	1
1994:	0	0	0	0	0	0	2	1	0	0	0	0
1995:	0	0	0	0	1	1	0	1	0	0	0	0
1996:	0	0	0	1	0	1	0	2	0	1	0	0
1997:	0	0	2	1	1	2	0	0	2	1	3	0
1998:	1	0	0	0	0	1	0	0	0	0	0	0
> Seas						. 3AM	sunWe	ea.ts	5)			
tau =	0.02	Ω4/,	sl	- =81	⊥./∛							

Table 5b. Ontario-ThuSat

ont.1	1PMTł	nuSat	.ts									
				Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	1	0	2	0	0	1	Ō	0	0	0	2
1993:	1	0	0	1	0	0	1	1	0	1	0	0
1994:	0	1	3	0	0	0	1	0	1	0	0	0
1995:	0	0	1	0	0	4	0	2	1	0	0	0
1996:	0	1	1	0	1	1	1	0	0	1	0	0
1997:	0	0	0	0	0	1	0	2	0	1	0	0
1998:	0	0	0	0	0	0	0	1	0	0	0	0
> Mani								-		Ũ		0
tau =				l =1!			,					
> ont												
					Mav	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	0	1	0	0	1	0	0	0	1	0	1
1993:	0	0	0	0	0	2	0	0	0	0	0	0
1994:	0	1	0	0	0	0	0	0	0	0	0	0 0
1995:	0	0	0	0	0	0	0	0	3	0	0	0
1996:	0	0	1	0	0	0	0	2	0	0	0	0
1997:	1	0	0	0	0	0	3	1	1	1	0	2
1998:	0	0	0	0	0	0	0	0	0	0	0	0
> Man								0	0	0	0	0
tau =					97.22		10/					
> ont					91.24	<u> </u>						
> 0110		Feb			Maw	Tun		Aua	Sen	Oct	Nov	Dec
1992:	0	1	1	0	1	0	1	0	3	0	0	0
1993:	0	1	1	0	0	1	2	2	0	1	0	0
1994:	0	1	0	0	1	0	0	1	0	0	0	0
1995:	0	1	0	0	0	0	1	0	0	2	0	2
1995:	0	1	0	0	0	0	1	1	2	1	0	0
1990:	1	0	0	0	0	0	0	1	0	0	0	0
1998:	0	0	0	0	1	0	0	0	0	0	0	0
> Man								0	0	0	0	0
tau =					.867 ⁹		• /					
> ont					••••	•						
> 0110					Mav	สมท	.T11]	Aug	Sep	Oct	Nov	Dec
1992:	0	0	0	0	1	0	0	0	0 0	1	0	1
1993:	0	1	1	1	0	0	0	0	0	0	0	1
1994:	1	0	0	1	0	1	1	0	0	0	0	1
1995:	0	1	0	0	2	0	0	0	1	0	0	0
1996:	0	0	0	0	0	0	0	0	0	1	0	0
1990:	0	0	0	0	0	0	0	1	1	0	0	0
1998:	0	0	0	1	0	0	0	0	0	0	0	1
> Man								0	0	0	0	Ŧ
tau =							5)					
> ont		,			J. 10'	0						
> 0110					Maw	Tun		Aug	Sen	Oct	Nov	Dec
1992:	1	0	Mai 0	0	May 0	0	0 0 0	Aug 0	998 0	0000	0	0
1992:	0	0	0	0	0	1	0	0	0	0	1	1
1993:	0	0	0	1	0	0	0	0	0	0	0	0
1994:	0	0	0	0		0	0	0	0		0	
	0				0		1			0 1		0
1996:		0	0	0	0	0		0	0	1 2	0	0
1997:	0	1 0	1 0	0	0	0	0	0	0		0	0
1998:	0			0 27 M	0 Thu C	0	0	0	1	0	0	1
> Mani							5)					
tau =	0.00	, 280	s.	L =44	4.35	õ						

Table 5c. Manitoba-SunWed

man.1	1PMS:	unWeo	d.ts									
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	0	0	0	0	0	0	0	0	0	0	0
1993:	0	0	0	0	1	0	0	0	1	0	0	0
1994:	0	0	0	0	0	0	0	3	1	1	0	0
1995:	0	0	0	0	1	0	0	0	0	0	0	0
1996:	1	0	0	0	0	1	0	0	0	0	0	0
1997:	1	0	0	0	0	0	2	0	0	0	0	0
1998:	0	0	0	0	0	0	0	0	0	0	0	0
> Manı	nKend	dall	(man	.11PI	ISun	Ved.t	s)					
tau =	-0.0	0179	, :	sl =8	34.69)						
> man	.12AI	MSun	Wed.	ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	0	0	0	0	0	1	0	0	0	0	0
1993:	0	0	0	0	0	0	0	0	0	0	0	0
1994:	0	0	0	0	0	0	1	0	0	1	0	0
1995:	0	0	0	0	0	0	0	0	0	0	0	1
1996:	0	1	0	0	0	0	0	0	0	0	0	0
1997:	0	0	0	0	0	0	0	0	0	0	0	0
1998:	0	0	0	0	0	1	0	0	0	0	1	0
> Manı	nKend	dall	(man	.12AI	ISun	ved.t	s)					
tau =				1 =59								
> man	.1AM	SunWe	ed.t:	5								
					May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	0	0	0	0	0	0	Ō	0	0	0	0
1993:	0	0	0	0	0	0	2	1	0	0	0	0
1994:	0	0	0	0	0	0	0	0	0	0	0	0
1995:	0	0	1	0	0	0	0	0	0	0	0	0
1996:	0	0	0	0	0	0	0	1	0	0	0	0
1997:	0	0	0	0	1	0	0	0	0	0	1	0
1998:	0	0	0	0	0	0	0	0	0	0	0	0
> Manı												
tau =				1 =83			- ,					
> man												
					Mav	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	0	0	0	0	0	0	0	0	0	0	0
1993:	0	0	0	0	0	0	0	0	0	0	0	0
1994:	0	0	0	0	0	0	0	0	0	0	0	0
1995:	0	0	0	0	0	0	0	0	0	0	0	0
1996:	0	0	0	0	0	0	0	0	0	0	1	0
1997:	2	0	0	0	0	0	0	0	0	1	0	0
1998:	0	0	0	0	0	0	0	0	0	0	1	0
> Manı	nKend					ed.ts	3)					
tau =												
> man												
					Mav	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	0	0	1	0	0	0	1	۹۵ <i>۵</i>	1	0	0
1993:	0	0	0	0	0	0	0	0	0	0	0	0
1994:	0	0	0	0	0	0	0	0	0	0	0	0
1995:	0	1	0	0	0	0	0	0	0	0	0	0
1996:	0	0	0	0	1	0	0	0	0	0	1	0 0
1997:	0	0	0	0	0	0	0	0	1	0	0	0
1998:	0	0	0	0	1	0	0	0	0	0	0	0
> Manı								0	5	0	5	Ŭ
tau =				sl ='			- /					
	5.0		, ,			- •						

Table 5d. Manitoba-ThuSat

man.11	man.11PMThuSat.ts												
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	0	0	0	0	0	0	0	0	0	
1993:	0	0	0	0	0	0	0	0	0	0	0	0	
1994:	0	0	0	0	0	0	0	0	0	1	0	0	
1995:	0	0	0	0	0	0	0	0	0	0	0	0	
1996:	0	0	0	0	0	0	0	0	0	0	0	0	
1997:	0	0	0	0	0	0	0	1	0	1	1	0	
1998:	0	0	0	0	0	0	0	0	0	0	0	0	
> Manı	nKend	dall	(man	.11PI	MThu	Sat.t	s)						
tau =	0.13	38,	sl	=12	.78%								
> man	.12AI	MThu	Sat.	ts									
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	0	0	0	0	0	0	0	0	0	
1993:	0	0	0	0	0	0	0	0	0	0	0	0	
1994:	0	0	0	0	0	0	0	1	0	0	0	0	
1995:	0	0	0	0	0	0	1	0	0	0	0	0	
1996:	0	0	0	0	0	0	0	0	0	0	0	0	
1997:	0	0	0	0	0	0	0	0	0	0	0	0	
1998:	0	0	0	0	0	0	0	0	0	0	0	0	
> Manı	nKend	dall	(man	.12AI	MThu:	Sat.t	s)						
tau =	-0.0	0265	, :	sl ='	78.0	58							
> man													
	Jan		Mar	Apr	May	Jun	Jul	Aug	Sep			Dec	
1992:	0	0	0	0	0	0	0	0	0	0	1	0	
1993:	0	0	0	0	0	0	0	0	0	1	0	1	
1994:	0	0	0	0	0	0	0	0	0	0	0	0	
1995:	0	0	0	0	0	0	0	0	0	0	0	0	
1996:	0	0	0	0	0	0	0	1	0	0	0	0	
1997:	0	0	0	0	0	0	0	0	0	0	0	0	
1998:	0	0	0	0	0	0	0	0	0	0	0	0	
> Manı							3)						
tau =					3.53	6							
> man						_	_ 1	_	~			_	
1000				-	-			Aug	-				
1992:	0	0	0	0	0	0	0	0	0	0	0	0	
1993:	0	0	0	0	0	0	0	0	0	0	1	0	
1994:	0	0	0	0	0	0	0	0	0	0	1	0	
1995:	0	0	0	0	0	0	0	0	0	0	0	0	
1996:	0	0	0	0	0	0	0	0	0	0	0	0	
1997:	0	0	0	0	0	0	0	0	0	0	0	0	
1998:	0	0	0	0 2 7 M	0 Thu C	0	0	0	0	0	0	0	
> Manı)714					5)						
					±3.02	26							
> man					Most	Tum	T11]	Aug	Con	Oat	Nou	Dog	
1992:	0	сер 0	Mar 0				0			0000	0		
1992:	0	0	0	0 0	0 0	0 0	0	0 0	0 0	1	0	0 0	
1993:	0	0	0	0	0	0	0	0	0	0	0	0	
1994:	0	0	0	0	0	0	0	0	0	0	0	0	
1995:	0	0	0	0	0	0	0	0	0	0	0	0	
1990: 1997:	0	0	0	0	0	0	0	0	0	0	0	0	
1998:	0	0	0	0	0	0	0	0	0	0	0	0	
> Manı								0	0	0	0	U	
tau =					40.9		- /						
cuu -	0.0	5,02	, ,	JT	- · · · · .								

6. Fatalities by drink

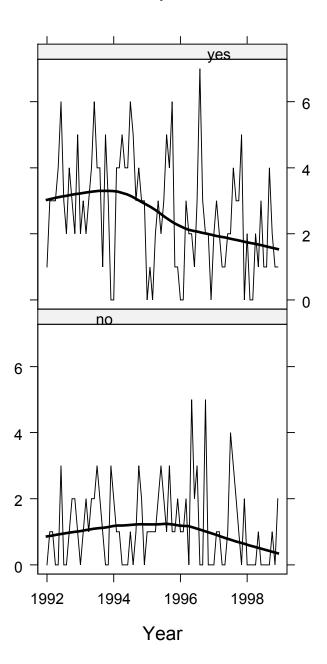
Table 6. Time series tabulation and Mann-Kendall trend test for fatalities by province (ont or man) and drinking class (yes or no) > ont.ves.ts

> ont.yes.ts												
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	1	3	3	3	4	6	3	2	4	3	2	5
1993:	2	3	2	3	4	6	4	4	1	5	3	0
1994:	0	4	4	5	4	4	6	5	3	4	3	3
1995:	0	1	0	2	3	2	3	5	4	6	1	1
1996:	0	0	3	2	2	1	3	7	3	2	2	0
1997:	2	3	2	1	1	2	2	4	3	3	5	0
1998:	2	0	0	2	1	3	1	1	4	2	1	1
> SeasonalMannKendall(ont.yes.ts)												
tau = -0.336, $sl = 0.0459%$												
> ont.no.ts												
	Jan	Feb	Mar	Apr	May	Jun	Jul	Auq	Sep	Oct	Nov	Dec
1992:	0	1	1	0	0	3	0	Ő	1	2	2	1
1993:	0	1	2	1	2	2	3	2	1	0	0	3
1994:	2	1	1	0	0	0	1	0	1	3	2	0
1995:	1	1	1	1	2	3	2	1	3	1	1	2
1996:	1	1	2	0	5	2	3	0	0	5	0	0
1997:	0	1	1	0	0	1	4	3	2	1	0	2
1998:	0	0	0	0	1	0	0	0	0	1	0	2
> Sea	sona	l Manı	- Kend	all	(ont	.no.t	s)	-	-	_	•	_
tau =				l =1:			,					
> man		-										
, man	-		Mar	Apr	Mav	สมท	.Tu l	Αιια	Sep	Oct	Nov	Dec
1992:	0	0	0	1	0	0	1	1	٩ <u></u> ٥٥	1	0	0
1993:	0	0	0	0	1	0	2	1	1	2	0	1
1994:	0	0	0	0	0	0	1	3	0	3	1	0
1995:	0	1	1	0	1	0	0	0	0	0	0	1
1996:	0	0	0	0	1	1	0	1	0	0	2	0
1997:	2	0	0	0	0	0	1	1	0	2	2	0
1998:	0	0	0	0	1	1	0	0	0	0	2	0
> Sea	-				_	_	-			Ũ	-	Ū
tau =					7.77		,					
> man				- ,		0						
			Mar	Apr	Mav	สมท	Jul	Αιια	Sep	Oct	Nov	Dec
1992:	0	0	0	0	0	0	0	0	qي 0	0	1	0
1993:	0	0	0	0	0	0	0	0	0	0	1	0
1994:	0	0	0	0	0	0	0	1	1	0	0	0
1995:	0	0	0	0	0	0	1	0	0	0	0	0
1996:	1	1	0	0	0	0	0	1	0	0	0	0
1997:	1	0	0	0	1	0	1	0	1	0	0	0
1998:	0	0	0	0	0	0	0	0	0	0	0	0
> Sea:	-	-	-	-	-	-	-	0	0	0	0	U
							101					
tau = 0.0843, sl =53.38%												

Figure 5.

Figure 5 shows that in Ontario there was a marked downward trend starting around 1994 in fatalities with drink=yes. The number of fatalities in Ontario with drink=no has also started to decline since about 1995 or 1996.

TIRF deaths by drink, Ontario



6. Fatalities by drink and hour

Figure 6, fatalities by hour, drink and province. In each hour slot the accidents are larger in drinking alcohol group than no drinking alcohol group. There is an increasing trend in Ontario in the 2AM slot and decreasing trends at 11PM, 12AM and 1AM.

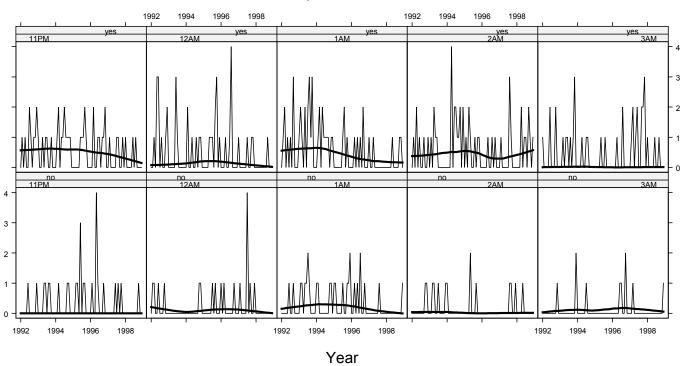
	the 2			anu	ucci	Cash	ig in	ciius	at I	11 101	, 127	1111 6	un
> ont	.11PM	/no.t	s										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	0	0	1	0	0	0	0	0	1	
1993:	0	0	0	0	1	1	0	1	1	0	0	0	
1994:	0	0	1	0	0	0	0	0	1	1	0	0	
1995:	0	0	1	1	0	3	0	1	1	0	0	0	
1996:	0	1	0	0	4	1	0	0	0	1	0	0	
1997:	0	0	0	0	0	1	0	1	0	1	0	0	
1998:	0	0	0	0	0	0	0	0	0	1	0	0	
						-	Ū	0	0	-	Ŭ	Ŭ	
<pre>> MannKendall(ont.11PMno.ts) tau = -0.0276, sl =76.11%</pre>													
> ont.12AMno.ts													
/ 0110				7.000	Morr	Tum	т., 1	7.1.0	Com	Oat	Nor	Dog	
1002.		Feb											
1992:	0	1	1	0	0	1	0	0	0	1	0	0	
1993:	0	0	0	0	0	0	0	0	0	0	0	0	
1994:	0	0	0	0	0	0	0	0	0	1	1	0	
1995:	0	0	0	0	0	0	1	0	1	0	0	0	
1996:	1	0	1	0	0	0	0	0	0	1	0	0	
1997:	0	1	0	0	0	0	4	0	1	0	0	1	
1998:	0	0	0	0	0	0	0	0	0	0	0	0	
> Manı	nKend	dall(ont	.12AI	Mno.1	ts)							
tau =	-0.0)277,	5	sl ='	76.1	78							
> ont	.1AMr	no.ts	5										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	0	0	1	0	0	1	0	0	0	
1993:	0	1	1	0	1	1	2	1	0	0	0	0	
1994:	1	1	0	0	0	0	0	0	0	1	1	0	
1995:	1	1	0	0	0	0	1	0	0	1	1	2	
1996:	0	0	1	0	1	0	2	0	0	1	0	0	
1997:	0	0	0	0	0	0	0	1	0	0	0	0	
1998:	0	0	0	0	0	0	0	0	0	0	0	1	
> Mani							0	0	0	0	0	-	
tau =				1 =1:	5.379	6							
> ont				7		T	T 1	7	a	0 +		D	
1000.		Feb		-	-			-	-				
1992:	0	0	0	0	0	0	0	0	0	1	1	0	
1993:	0	0	1	1	0	0	1	0	0	0	0	1	
1994:	1	0	0	0	0	0	0	0	0	0	0	0	
1995:	0	0	0	0	2	0	0	0	1	0	0	0	
1996:	0	0	0	0	0	0	0	0	0	0	0	0	
1997:	0	0	0	0	0	0	0	1	1	0	0	1	
1998:	0	0	0	0	1	0	0	0	0	0	0	0	
> Manı	nKend	dall(ont	. 2AMı	no.ts	з)							
tau =	-0.0)659,	5	sl =4	46.62	18							
> ont													
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	0	Ō	0	0	0	Ō	0	1	0	
1993:	0	0	0	0	0	0	0	0	0	0	0	2	
1994:	0	0	0	0	0	0	1	0	0	0	0	0	
1995:	0	0	0	0	0	0	0	0	0	0	0	0	
1996:	0	0	0	0	0	1	1	0	0	2	0	0	
1990:	0	0	1	0	0	0	0	0	0	0	0	0	
1997:	0	0	0	0	0	0	0	0	0	0	0	1	
							U	U	U	U	0	Ŧ	
> Mani													
tau =	0.05	οσ ι ,	s.	L = 5!	אי . כ	б							

tau = 0.0531, sl =55.78%

- 11 5 0													
ont.1	-			7.000	More	Tum	T.,]	7.110	Com	Oat	Nor	Dog	
1992:	0 an	гер 1	Mar 0	Apr 1	May 0	0 UII	3 U I	Aug 1	sep 0	1	NOV 1	Dec 2	
1992:	1	0	0	1	0	1	1	0	0	1	0	0	
1993:	0	1	2	0	1	1	2	1	1	1	1	0	
1994:	0	0	2	0	0	1	1	1	2	1	0	0	
1995:	0	0	2	0	0	1	1	1 0	2	1	2	0	
1998:	0	1	2	0	0	1 0	1	1	1 0	1 0	2	0	
1997:	1	1 0	0	0	0	1	1 0	1	0	0	1 0	0	
> Man						_	0	T	0	0	0	0	
tau =					.147 ⁹								
> ont					• /	•							
, 0110		-		Apr	Mav	สมท	[נוד	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	1	0	3	3	0	1	٩ <u></u> ٥٥	0	1	2	
1993:	0	0	0	0	0	3	1	0	0	0	0	0	
1994:	0	2	0	0	1	0	0	1	0	1	1	0	
1995:	0	0	0	0	1	1	1	0	2	3	0	1	
1996:	0	0	0	1	0	0	1	4	0	0	0	0	
1997:	1	0	0	0	0	0	1	0	0	1	1	0	
1998:	0	0	0	0	0	0	0	0	1	0	0	0	
> Mani			-		-	-	Ũ	Ũ	-	Ũ	Ũ	0	
tau =					3.79								
> ont						0							
0110				Apr	Mav	สมท	[נוד	Aug	Sep	Oct	Nov	Dec	
1992:	0	1	2	0	1	0	1	0	3	0	0	0	
1993:	1	2	1	0	2	0	2	3	1	3	0	0	
1994:	0	1	2	0	2	1	1	1	1	0	1	1	
1995:	0	0	0	0	0	0	1	2	0	1	0	0	
1996:	0	0	1	0	1	0	1	0	2	0	0	0	
1997:	1	0	0	1	0	0	0	0	0	0	0	0	
1998:	0	0	0	0	1	0	0	0	0	1	1	0	
> Man	nKend	lall	(ont	.1AM	ves.	ts)							
tau =					. 2899								
> ont	.2AM	/es.t	s										
				Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	1	0	2	0	1	0	Ō	1	0	0	1	
1993:	0	1	0	2	1	1	0	0	0	0	0	0	
1994:	0	0	0	4	0	2	2	1	1	2	0	2	
1995:	0	1	0	2	1	0	0	1	0	1	1	0	
1996:	0	0	0	0	1	0	0	1	0	1	0	0	
1997:	0	1	0	0	0	0	0	3	1	0	0	0	
1998:	0	0	0	2	0	1	1	0	2	1	0	1	
> Mani	nKend	dall	(ont	.2AM	yes.t	ts)							
tau =	-0.0	0267	, :	sl ='	.04 76.04	48							
> ont	. 3AMy	/es.t	s										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	1	0	0	0	0	2	0	0	0	2	0	0	
1993:	0	0	1	0	1	1	0	1	0	1	3	0	
1994:	0	0	0	1	0	0	1	1	0	0	0	0	
1995:	0	0	0	0	1	0	0	1	0	0	0	0	
1996:	0	0	0	1	0	0	0	2	0	0	0	0	
1997:	0	1	2	0	1	2	0	0	2	2	3	0	
1998:	1	0	0	0	0	1	0	0	1	0	0	0	
> Man	nKend	lall	(ont	. 3AM	yes.	ts)							
		L95.		1 =83									

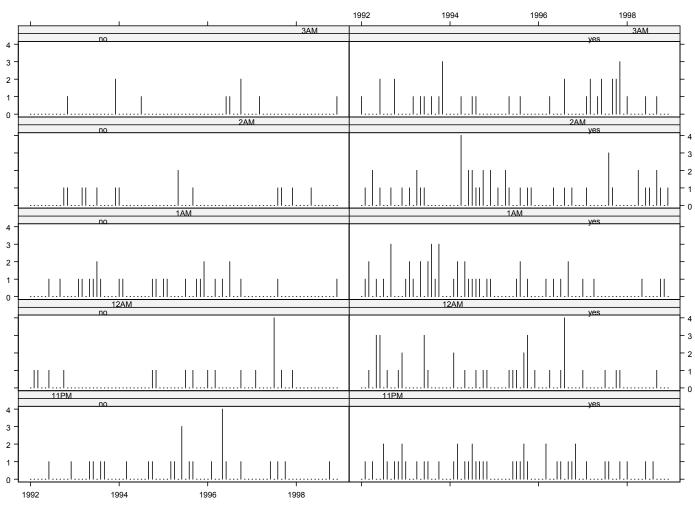
tau = 0.0195, sl =82.67%

Figure 6a. Loess Analysis



TIRF, deaths by hour and drink, Ontario.

Figure 6b. Line plot of time series



TIRF, deaths by hour and drink, Ontario.

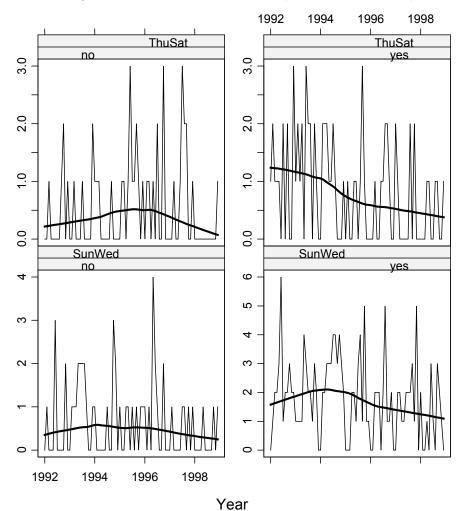
Year

7. Fatalities by drink and wkgrp

> ont				7	Mes	T	T 1	7	0.0	Oct	New	Do -
1000.		Feb		-	May			Aug	Sep	Oct		Dec
1992:	0	1	0	0	0	3	0	0	0	0	2	0
1993:	0	1	1	1	2	2	2	2	1	0	0	1
1994:	1	0	0	0	0	0	1	0	0	3	2	0
1995:	1	0	0	1	1	0	1	0	1	0	1	1
1996:	1	0	1	0	4	2	1	0	0	2	0	0
1997:	0	1	0	0	0	0	1	1	0	1	0	1
1998:	0	0	0	0	1	0	0	0	0	1	0	1
> Manı)					
tau =				sl =:	30.53	38						
> ont												
	Jan	Feb		_	_	Jun		Aug	Sep	Oct		Dec
1992:	0	0	1	0	0	0	0	0	1	2	0	1
1993:	0	0	1	0	0	0	1	0	0	0	0	2
1994:	1	1	1	0	0	0	0	0	1	0	0	0
1995:	0	1	1	0	1	3	1	1	2	1	0	1
1996:	0	1	1	0	1	0	2	0	0	3	0	0
1997:	0	0	1	0	0	1	3	2	2	0	0	1
1998:	0	0	0	0	0	0	0	0	0	0	0	1
> Manı	nKend	dall	(ont	. Thu	Satno	o.ts)					
tau =	0.01	129,	s	L =88	8.56 ⁹	20						
> ont	.1AMı	no.ts	3									
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	0	0	0	0	1	0	0	1	0	0	0
1993:	0	1	1	0	1	1	2	1	0	0	0	0
1994:	1	1	0	0	0	0	0	0	0	1	1	0
1995:	1	1	0	0	0	0	1	0	0	1	1	2
1996:	0	0	1	0	1	0	2	0	0	1	0	0
1997:	0	0	0	0	0	0	0	1	0	0	0	0
1998:	0	0	0	0	0	0	0	0	0	0	0	1
ont.Su	unWed	lyes	.ts									
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	1	2	2	3	6	1	2	2	3	2	2
1993:	1	1	1	1	4	3	2	2	1	3	2	0
1994:	0	2	2	3	3	3	4	4	3	4	3	2
1995:	0	0	0	2	2	1	3	4	1	5	1	1
1996:	0	0	2	2	2	0	2	5	1	1	2	0
1997:	0	2	2	1	1	2	2	2	3	1	5	0
1998:	2	0	0	1	0	3	1	0	3	2	1	0
> Manı	nKend	lall	(ont	. Sunī	Vedy	es.ta	3)					
tau =	-0.1	L47,	s	L =7	.105	8						
> ont			es.ts	5								
		Feb			May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	1	2	1	1	1	0	2	0	2	0	0	3
1993:	1	2	1	2	0	3	2	2	0	2	1	0
1994:	0	2	2	2	1	1	2	1	0	0	0	1
1995:	0	1	0	0	1	1	0	1	3	1	0	0
1996:	0	0	1	0	0	1	1	2	2	1	0	0
1997:	2	1	0	0	0	0	0	2	0	2	0	0
1998:	0	0	0	1	1	0	0	1	1	0	0	1
> Manı								-	-	5	5	-
tau =							- /					

tau = -0.239, sl =0.4885%

Figure 7. The comparisons between driving with drink and without drink within drink factor, and between ThuSat and SunWed within wkgrp factor are shown in Fig 3. There are more fatalities with drink=yes and there are are more accidents in SunWed. In Ontario, both SunWed and ThuSat have a downward trend.



TIRF deaths by WKGRP and Drink, Ontario, panels scaled independently.

8. Annual Time Series

Table 8a. Annual fatalities, Ontario and Manitoba. There are 347 deaths in Ontario and 53 in Manitoba. The death rate for Total in Ontario has declined from 57 to 24 which represents an average annual rate of decrease of about 13%.

> tii	rf.annı	ual.or	nt						
	Total	PM11	AM12	AM1	AM2	AM3	Sun₩ed	ThuSat	
1992	57	12	18	11	10	6	36	21	
1993	61	10	7	24	10	10	39	22	
1994	56	14	8	15	15	4	40	16	
1995	50	14	11	12	10	3	29	21	
1996	46	15	9	12	3	7	28	18	
1997	53	11	13	4	9	16	32	21	
1998	24	4	2	5	9	4	18	б	
> Mar	nnKenda	allMat	rix(t	irf.	.annı	ual.	ont)		
-	Fotal	PM11	AN	112	AN	41	AM2 A	M3 SunWec	l ThuSat
tau -	-0.71	-0.10) -0.	.14	-0.4	49 ·	-0.48	0 -0.52	-0.41
sl%	3.55	87.93	876.	.39	17.1	16 2	20.40 1	00 13.31	27.23

tirf	.annua	l.man							
	Total	PM11	AM12	AM1	AM2	AM3	Sun₩ed	ThuSat	
1992	5	0	1	1	0	3	4	1	
1993	9	2	0	5	1	1	5	4	
1994	10	б	3	0	1	0	7	3	
1995	5	1	2	1	0	1	4	1	
1996	8	2	1	2	1	2	7	1	
1997	12	б	0	2	3	1	9	3	
1998	4	0	2	0	1	1	4	0	
> Mai	nnKend	allMat	crix(t	cirf.	.ann	ual.r	man)		
г	Fotal	PM11	AM12	I	AM1	AM2	2 AM3	3 SunWed	ThuSat
tau	0	0.10	0	-0.	.10	0.4	7 -0.17	0.26	-0.37
sl%	100	87.64	100	87.	.64	23.49	5 73.77	52.54	34.08

Annual rate of decline in Total fatalities in Ontario > $1-(24/57)^{(1/6)}$ [1] 0.1342562 **Tables 8b(i)-(iv).** Annual total fatalities are decomposed by hour, wkgrp and drink for Ontario and Manitoba. Table 8b might suggest that there has been a shift in fatalities from early evening to late evening starting around 1996. The Mann-Kendall trend test is statistically significant on a two-sided test for Ontario fatalities with drink=yes for Total, 11PM, 1AM, SunWed and ThuSat and in all cases the sign of tau indicates a downward trend. The trend test is not significant for drink=no in Ontario. There are no trends in Manitoba for either drink=yes or drink=no.

Table 8b(i). Annual Fatalities, Ontario, drink=no

			-	- 1		11	-)						
> ti	rf.drin	1k.no	.annua	a1.01	ntl,-	- T]							
	Total	PM11	AM12	AM1	AM2	AM3	SunWed	ThuSa	at				
1992	11	2	4	2	2	1	6		5				
1993	17	4	0	7	4	2	13		4				
1994	11	3	2	4	1	1	7		4				
1995	19	7	2	7	3	0	7		L2				
1996	19	7	3	5	0	4	11		8				
1997	15	3	7	1	3	1	5	-	LO				
1998	4	1	0	1	1	1	3		1				
> Mar	nKendal	lAnnu	al(tir	f.dr:	ink.n	o.anr	ual.ont	[,-1])					
	Tc	tal		PM11		AM1	.2	AM1		AM2	AM3	SunWed	ThuSat
tau -	-0.05006	262 -	0.0500	6262	0.05	00626	2 -0.350)4383	-0.25	503131	-0.05634362	-0.3903600	0
sl	1.00000	000	1.0000	0000	1.00	00000	0 0.356	54410	0.53	387009	1.00000000	0.2876112	1

Table 8b(ii). Annual Fatalities, Ontario, drink=yes

> ti:	rf.drii	nk.yes	s.annu	.lau	ont[,-1]		-				
	Total	PM11	AM12	AM1	AM2	AM3	SunWed	ThuSat	2			
1992	39	9	11	8	6	5	26	13	3			
1993	37	5	4	15	5	8	21	10	5			
1994	45	11	б	11	14	3	33	12	2			
1995	28	6	9	4	7	2	20	8	3			
1996	25	8	б	5	3	3	17	8	3			
1997	28	4	4	2	5	13	21		7			
1998	18	3	1	3	8	3	13	ĩ	5			
> Man	nKendal	lAnnua	al(tir	f.dr:	ink.y	es.ar	nual.ont	:[,-1])				
	Tot		-	M11		AM12	-	AM1 A		AM3	SunWed	ThuSat
	0.68313						-0.6190		0			-0.878310025
sl	0.04828	61 0	.06904	757	0.12	43064	0.0301	15876	1	0.8753926	0.09471773	0.009809151

Table 8b(iii). Annual Fatalities, Manitoba, drink=no

	()					,		,		
> ti	rf.drinl	c.no	.annua	ıl.ma	an[,-	1]				
	Total 1	PM11	AM12	AM1	AM2	AM3	SunWe	d ThuSa	t	
1992	1	0	0	1	0	0		0	1	
1993	1	0	0	0	1	0		0	1	
1994	2	2	0	0	0	0		2	0	
1995	1	0	1	0	0	0		0	1	
1996	3	1	1	1	0	0		3	0	
1997	4	1	0	1	1	1		4	0	
1998	0	0	0	0	0	0		0	0	
> Mar	nKendall	Annua	al(tir	f.dri	lnk.no	.ann	ual.ma	n[,-1])		
	Total		PM11		AM12	2 AM1	AM2	AM3	SunWed	ThuSat
tau (.2057378	0.11	166424	0.13	880131	L 0	0 0	.3563483	0.3944053	-0.6299408
sl (0.6380424	0.80	551347	0.84	164506	51	1 0	.4532547	0.3150597	0.1116118

Table 8b(iv). Annual Fatalities, Manitoba, drink=yes

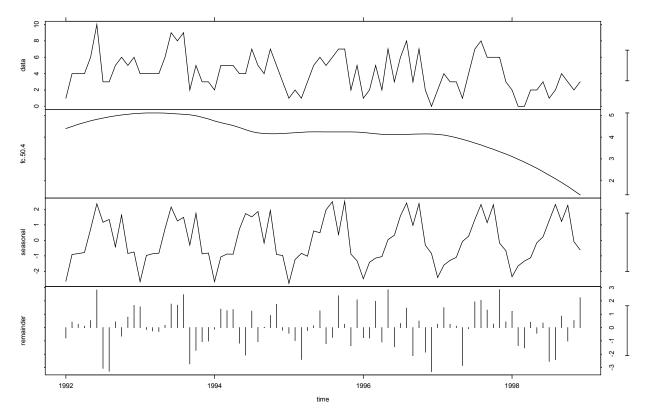
> ti	rf.dri	nk.yes	s.annu	.al.r	man[,-1]						
	Total	PM11	AM12	AM1	AM2	AM3	SunWed	ThuSat	Ē			
1992	4	0	1	0	0	3	4	(C			
1993	8	2	0	5	0	1	5	:	3			
1994	8	4	3	0	1	0	5		3			
1995	4	1	1	1	0	1	4	(C			
1996	5	1	0	1	1	2	4		1			
1997	8	5	0	1	2	0	5	:	3			
1998	4	0	2	0	1	1	4	(C			
> Man	nKendal	lAnnua	al(tir	f.dr	ink.y	es.an	nual.mai	n[,-1])				
	To	otal	P	M11		AM1	2	AM1	AM2	AM3	SunWed	ThuSat
										-0.2646281		-0.05634362
sl	1.00000	0000 1	.00000	000	1.00	00000	0 1.00	000000	0.1001784	0.5253584	0.8596838	1.00000000

9. STL Analysis

R-sq = 53.9 %

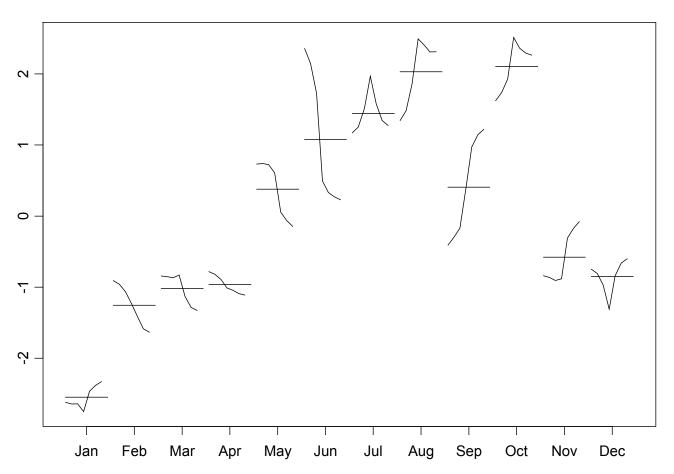
Figure 9a.





ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2

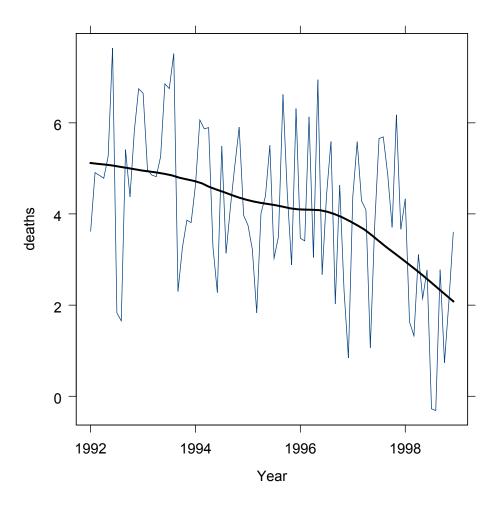
Figure 9b. Monthplot of Seasonal Component. May through October are high and November through April are relatively lower. Upward trends exist in August and September and to a lesser extent in January and November.



Monthplot, seasonal, TIRF deaths Ontario

Figure 9c. Loess trend analysis of deseasonalized TIRF.

tau = -0.28, sl =0.01655% The loess trend line has decreased from a monthly death-rate of 4.32 in January 1992 to 1.78 in December 1998. This corresponds to an annual rate of decrease of 12%. > pc.change(tirf.deseasonalized.ts) [1] 5.117620 2.086254 59.233909 > 1-(2.086254/5.117620)^(1/7) [1] 0.1203124



TIRF, deseasonalilzed deaths

C. TIRF All Hours - Crosstabs

Frequency Analysis of Drink Variable

In Ontario out of a total of 375 cases, there are 44 cases where drink=unknown and only 2 cases in Manitoba. Thus in Ontario there remain 331 cases.

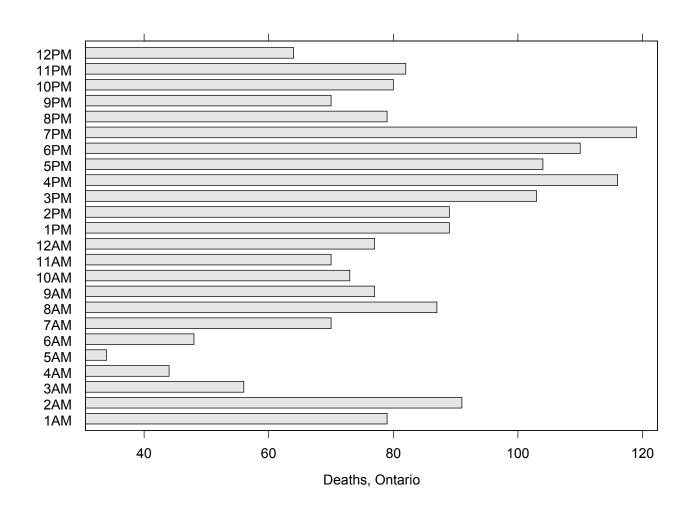
```
>crosstabs(~drink, data=tirf.df, subset=province=="Ontario", na.action=na.exclude)
Call:
crosstabs(formula = ~ drink, data = tirf.df, subset = province == "Ontario",
    na.action = na.exclude)
375 cases in table
+---+
| N
      |N/Total|
+----+
drink |
----+
unknown| 44
           |0.12 |
____+
no |107 |
    0.29
           - 1
----+
           yes |224
    10.6
            ----+
> crosstabs(~drink, data=tirf.df, subset=province=="Manitoba", na.action=na.exclude)
Call:
crosstabs(formula = ~ drink, data = tirf.df, subset = province == "Manitoba", na.action
= na.exclude)
58 cases in table
+----+
| N
    _____
|N/Total|
+----+
drink |
----+
unknown| 2
            0.034
----+
no |12 |
|0.21 |
----+
yes |44 |
    |0.76 |
----+
```

Frequency Analysis of Hour Variable Ontario

> barchart.hour()

 1AM
 2AM
 3AM
 4AM
 5AM
 6AM
 9AM
 10AM
 11AM
 12AM
 1PM
 2PM
 3PM
 4PM
 5PM
 6PM
 9PM
 10PM
 11PM
 12PM

 79
 91
 56
 44
 34
 48
 70
 87
 73
 70
 77
 89
 89
 103
 116
 104
 110
 119
 79
 70
 80
 82
 64



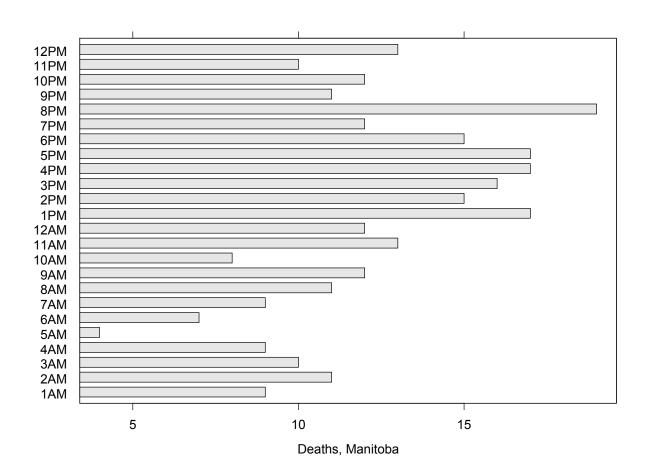


Deaths, Ontario

TIRF All Hours Crosstabs

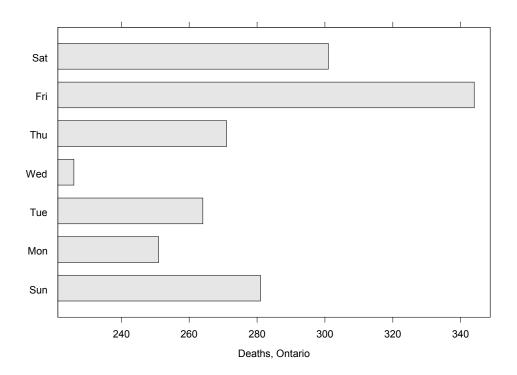
Frequency Analysis, hour, Manitoba

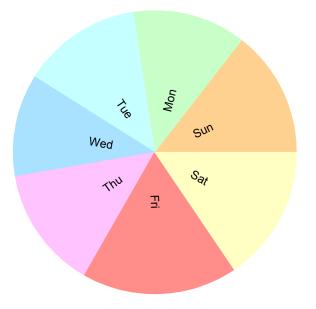
> ba	rchai	rt.ho	our.r	nan())																		
1AM	2AM	ЗАМ	4 AM	5AM	6AM	7AM	8AM	9AM	10AM	11AM	12AM	1PM	2PM	ЗPМ	4PM	5PM	6PM	7PM	8PM	9PM	10PM	11PM	12PM
9	11	10	9	4	7	9	11	12	8	13	12	17	15	16	17	17	15	12	19	11	12	10	13



Frequency Analysis of Day Variable, Ontario

> barchart.day()
Sun Mon Tue Wed Thu Fri Sat
281 251 264 226 271 344 301

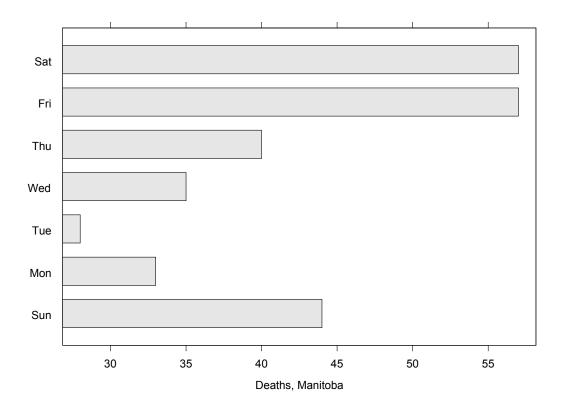




Deaths, Ontario

Frequency Analysis of Day Variable, Manitoba

```
> barchart.day.man()
Sun Mon Tue Wed Thu Fri Sat
44 33 28 35 40 57 57
```





Deaths, Manitoba

Frequency Analysis of Drink Variable, Ontario

In Ontario 53.2% of drivers who are killed had been drinking while in Manitoba, 83.5% of drivers who are killed had been drinking.

Drink and hour are not independent

Ontario

It is clear that drinking is more involved with late-night accidents.

```
> crosstabs (~hour+drink, data=tirfall.df, subset=prov=="Ont", na.action=na.exclude)
Call:
crosstabs(formula = ~ hour + drink, data = tirfall.df, subset = prov == "Ont", na.action =
  na.exclude)
1911 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----
        -+
hour |drink
    |unknown|no |yes |RowTotl|
1AM | 9 |21 |49 |79
    0.11 0.27 0.62 0.041
0.031 0.02 0.087
     |0.031 |0.02 |0.087 |
|0.0047 |0.011 |0.026 |
                       1
2AM | 7 | 28 | 56 | 91
|0.077 | 0.31 | 0.62 | 0.04
|0.024 | 0.026 | 0.099 |
                       |0.048
     |0.0037 |0.015 |0.029 |
3AM | 5 |12 |39 |56
|0.089 |0.21 |0.7 |0.029
    |0.017 |0.011 |0.069 |
     |0.0026 |0.0063 |0.02 |
4AM | 2 |10 |32 |44
    0.045 |0.23 |0.73 |0.023
|0.0069 |0.0095 |0.057 |
     |0.001 |0.0052 |0.017 |
| 2 |13 |19 |34
|0.059 |0.38 |0.56 |0.018
5AM
     |0.0069 |0.012 |0.034 |
     |0.001 |0.0068 |0.0099 |
| 3 | 28 | 17 | 48
6AM
     |0.062 |0.58 |0.35 |0.025
     |0.01 |0.026 |0.03 |
|0.0016 |0.015 |0.0089 |
                       1
|13 |44 |13 |70
|0.19 |0.63 |0.19 |0.037
7AM
     0.045 0.042 0.023
     |0.0068 |0.023 |0.0068 |
_____+
   |14 |62 |11 |87
8AM
     |0.16 |0.71 |0.13
                       0.046
     |0.048 |0.059 |0.02 |
|0.0073 |0.032 |0.0058 |
                        9AM |16 |58 | 3 |77
|0.21 |0.75 |0.039 |0.04
|0.055 |0.055 |0.0053 |
     |0.0084 |0.03 |0.0016 |
_____+
10AM |16 |49 | 8
                       |73
     0.22 0.67 0.11 0.038
     0.055 |0.046 |0.014 |
     |0.0084 |0.026 |0.0042 |
```

	0.26 0.062	0.64 0.043	7 0.1 0.012 0.0037	70 0.037
	0.1 0.028		6 0.078 0.011 0.0031	77 0.04
	0.28 0.086		5 0.056 0.0089 0.0026	89 0.047
	0.25 0.076	0.055	9 0.1 0.016 0.0047	89 0.047
	0.16	0.7 0.068	+ 15 0.15 0.027 0.0078	++ 103 0.054
	0.19 0.076	0.64 0.07	20 0.17 0.035 0.01	116 0.061
	0.15 0.055	0.69	+ 16 0.15 0.028 0.0084	++ 104 0.054
	0.2 0.076	0.63	+ 19 0.17 0.034 0.0099	++ 110 0.058
	0.1 0.041	0.6	+ 36 0.3 0.064 0.019	++ 119 0.062
	0.2 0.055	0.35	+ 35 0.44 0.062 0.018	++ 79 0.041
	0.043	0.54 0.036	29 0.41 0.051 0.015	++ 70 0.037
	0.15 0.041	0.35 0.026		80 0.042
	0.085 0.024	0.44 0.034	+ 39 0.48 0.069 0.02	++ 82 0.043
	0.062 0.014 0.0021	0.018 0.0099	+ 41 0.64 0.073 0.021	++ 64 0.033
ColTotl			+	++ 1911

0.15 0.55 0.3
Test for independence of all factors Chi^2 = 453.8353 d.f.= 46 (p=0) Yates' correction not used
<pre>> x<-get.crosstabs.contrib.ont(~hour+drink)</pre>
Contribution to chi-sq:
<pre>> dimnames(x) list(levels(tirfall.df\$hour), levels(tirfall.df\$drink))</pre>
× X
unknown no yes
1AM -0.86 -3.43 5.32
2AM -1.83 -3.15 5.62
3AM -1.20 -3.41 5.53
4AM -1.81 -2.91 5.28
5AM -1.39 -1.34 2.83
6AM -1.59 0.28 0.75
7AM 0.73 0.85 -1.69
8AM 0.22 2.00 -2.90
9AM 1.26 2.36 -4.14
LOAM 1.48 1.36 -2.92
LIAM 2.26 1.01 -3.01
L2AM -1.08 3.13 -3.51
1PM 3.13 1.39 -4.15
2PM 2.31 1.25 -3.37
3PM 0.09 1.99 -2.79
4PM 1.05 1.23 -2.43
5PM 0.05 1.91 -2.65
6PM 1.30 1.05 -2.36
7PM -1.43 0.64 0.15
8PM 1.16 -2.37 2.42
9PM -2.34 -0.12 1.84
LOPM -0.04 -2.44 3.37
11PM -1.54 -1.39 3.01
L2PM -1.83 -2.76 5.09

>

|0.029 |0.027 |

- 1

Manitoba

Accidents involving alcohol tend to happen more in late night hours in Manitoba too.

We exclude drink="unknown" since there are so few cases.

```
> crosstabs(~hour+drink, data=tirfall.df, subset=(prov=="Man")&drink!="unknown",
  na.action=na.exclude)
Call:
crosstabs(formula = ~ hour + drink, data = tirfall.df, subset = (prov == "Man") & drink !=
  "unknown", na.action = na.exclude)
252 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
hour |drink
    |no |yes
               |RowTot]|
----+
1AM
   | 2 | 7 | 9
                      1
    |0.22 |0.78
                |0.036 |
     |0.014 |0.062
                 |0.0079 |0.028
                | 3
          | 8
2 A M
                111
          0.73
     0.27
                 |0.044
     |0.022 |0.071
                |0.012 |0.032 |
_____+
   | 2
         | 8
ЗАМ
                |10
     |0.2 |0.8
                |0.04
     |0.014 |0.071
                |0.0079 |0.032
                 +----
           +----
   | 2 | 7
                |9
4AM
     0.22 |0.78
0.014 |0.062
                 0.036
                |0.0079 |0.028
                ----+
5AM
    0 4 4
          1
     0
                0.016
     0
         |0.035 |
     0
          |0.016 |
____
    _+___+
     | 2 | 5
                |7
6AM
          |0.71
     0.29
                0.028
     |0.014 |0.044
                 |0.0079 |0.02
                7AM
     | 6
          | 3
                |9
          |0.33
     0.67
                 |0.036
     |0.043 |0.027
                |0.024 |0.012 |
-+-
   | 6
          | 2
                18
8AM
     0.75
          0.25
                 |0.032
     |0.043 |0.018
                |0.024 |0.0079 |
----+----+-----+-----+-----
                |10
9AM
   | 8
          | 2
     0.8
          0.2
                0.04
     |0.058 |0.018 |
     |0.032 |0.0079 |
----+---+----+----+---
10AM | 4 | 3 |7
                      1
     0.57 0.43 0.028
```

	0.016	0.012	
	0.67 0.058	0.33 0.035	12 0.048
	0.8 0.058	0.2 0.018	10 0.04
	0.8 0.086	0.2	15 0.06
	0.8 0.058		10 0.04
3PM	11 0.85 0.079	0.15	13 0.052
	0.82 0.065	0.18 0.018	11 0.044
	0.77 0.072	0.23 0.027	13 0.052
	0.73 0.058	0.27	11 0.044
	0.73 0.058	0.27	11 0.044
	7 0.44 0.05 0.028	0.56 0.08	16 0.063
	5 0.45 0.036	6 0.55	11 0.044
	5 0.42 0.036	0.58 0.062	12 0.048
	0.3 0.022	0.7 0.062 0.028	10 0.04
	2 0.15	11 0.85 0.097	13 0.052

November 11, 2001, 2001 TIRF All Hours Crosstabs

----+ ColTotl|139 |113 |252 | |0.55 |0.45 | | -----+ Test for independence of all factors Chi^2 = 62.88204 d.f.= 23 (p=0.00001448521) Yates' correction not used Some expected values are less than 5, don't trust stated p-value > x<-get.crosstabs.contrib.man(~hour+drink) Contribution to chi-sq: > dimnames(x)_list(levels(tirfall.df\$hour), levels(tirfall.df\$drink)[2:3]) > x yes no 1AM -1.33 1.48 2AM -1.25 1.38 3AM -1.50 1.66 4AM -1.33 1.48 5AM -1.49 1.65 6AM -0.95 1.05 7AM 0.46 -0.52 8AM 0.76 -0.84 9AM 1.06 -1.17 10AM 0.07 -0.08 11AM 0.54 -0.60 12AM 1.06 -1.17 1PM 1.30 -1.44 2PM 1.06 -1.17 3PM 1.43 -1.59 4PM 1.19 -1.32 5PM 1.06 -1.17 6PM 0.78 -0.87 7PM 0.78 -0.87 8PM -0.61 0.68 9PM -0.43 0.48 10PM -0.63 0.70 11PM -1.07 1.19 12PM -1.93 2.14>

Drink and day are associated

Ontario

There are more fatalities on Saturday and Sunday and fewer on Wednesday.

```
> crosstabs(~drink+day, data=tirfall.df, subset=(prov=="Ont")&drink!="unknown",
na.action=na.exclude)
Call:
crosstabs(formula = ~ drink + day, data = tirfall.df, subset = (prov == "Ont") & drink !=
"unknown", na.action = na.exclude)
1635 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
drink |day
     |Sun |Mon |Tue |Wed |Thu |Fri |Sat |RowTotl|
no |128 |157 |178 |141 |160 |171 |132 |1067 |
|0.12 |0.15 |0.17 |0.13 |0.15 |0.16 |0.12 |0.65 |
|0.53 |0.74 |0.78 |0.77 |0.69 |0.62 |0.51 | |
|0.078 |0.096 |0.11 |0.086 |0.098 |0.1 |0.081 | |
yes |114 | 55 | 50 | 42 | 71 |107 |129 |568
|0.2 |0.097 |0.088 |0.074 |0.12 |0.19 |0.23 |0.35
|0.47 |0.26 |0.22 |0.23 |0.31 |0.38 |0.49 |
                                                              1
                                                              |0.07 |0.034 |0.031 |0.026 |0.043 |0.065 |0.079 |
           ____+
ColTotl|242 |212 |228 |183 |231 |278 |261 |1635 |
    |0.15 |0.13 |0.14 |0.11 |0.14 |0.17 |0.16 |
Test for independence of all factors
     Chi^2 = 79.46879 d.f.= 6 (p=4.551914e-015)
      Yates' correction not used
> fix(get.crosstabs.contrib.ont)
> x<-get.crosstabs.contrib.ont(~drink+day)</pre>
Contribution to chi-sq:
> dimnames(x)_list(levels(tirfall.df$drink)[2:3],levels(tirfall.df$day))
> x
     Sun Mon Tue Wed Thu Fri Sat
no -2.38 1.59 2.39 1.97 0.75 -0.77 -2.94
yes 3.26 -2.17 -3.28 -2.71 -1.03 1.06 4.03
```

Manitoba

There are more fatalities on Saturday and fewer on Wednesday.

```
> crosstabs(~drink+day, data=tirfall.df, subset=(prov=="Man")&drink!="unknown",
   na.action=na.exclude)
Call:
crosstabs(formula = ~ drink + day, data = tirfall.df, subset = (prov == "Man") & drink !=
    "unknown", na.action = na.exclude)
255 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
drink |day
             |Mon |Tue |Wed |Thu |Fri |Sat |RowTotl|
      lSun
no |19 |15 |12 |25 |22 |27 |19 |139
|0.14 |0.11 |0.086 |0.18 |0.16 |0.19 |0.14 |0.55
|0.49 |0.56 |0.6 |0.81 |0.61 |0.52 |0.38 |
                                                               |139
                                                                        1
                                                               10.55
                                                                        0.075 0.059 0.047 0.098 0.086 0.11 0.075
yes |20 |12 | 8 | 6 |14 |25 |31 |116

      |0.17
      |0.1
      |0.069
      |0.052
      |0.12
      |0.22
      |0.27
      |0.45

      |0.51
      |0.44
      |0.4
      |0.19
      |0.39
      |0.48
      |0.62
      |

      |0.078
      |0.047
      |0.031
      |0.024
      |0.055
      |0.098
      |0.12
      |

                                                                        ColTotl|39 |27 |20 |31 |36 |52 |50 |255 |
|0.15 |0.11 |0.078 |0.12 |0.14 |0.2 |0.2 |
Test for independence of all factors
Chi^2 = 15.59117 d.f.= 6 (p=0.01612495)
Yates' correction not used
> x<-get.crosstabs.contrib.man(~drink+day)</pre>
Contribution to chi-sq:
> dimnames(x) list(levels(tirfall.df$drink)[2:3],levels(tirfall.df$day))
> x
      Sun Mon Tue Wed Thu Fri Sat
no -0.49 0.07 0.33 1.97 0.54 -0.25 -1.58
yes 0.54 -0.08 -0.36 -2.16 -0.59 0.28 1.73
```

C. TIRF Dataset: All Hours Dataset **TIME SERIES**

Monthly Time Series

There is an estimated 44% decline in death-rate per month in Ontario but no decline in Manitoba.

Ontario

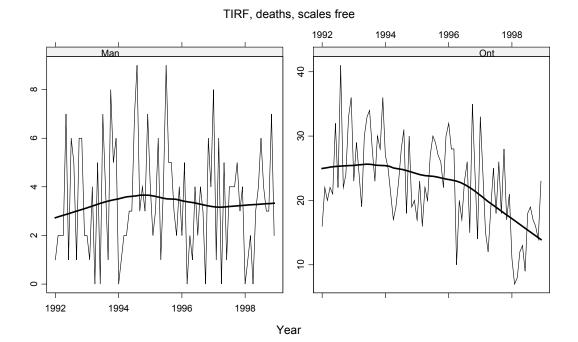
> tirf	.ont	t.ts										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	16	22	20	22	21	32	22	41	22	24	33	36
1993:	23	29	24	19	30	33	34	28	23	30	28	36
1994:	27	25	21	17	19	23	28	31	18	30	19	20
1995:	17	23	16	22	20	27	30	29	27	26	22	30
1996:	32	28	28	10	20	17	23	26	15	35	24	14
1997:	33	24	15	12	19	25	18	26	18	28	17	21
1998:	11	7	8	12	13	9	18	19	17	16	14	23
> sum(tirf	E.ont	t.ts)									
[1] 19	10											
> Seas	onal	LManı	nKend	dall	(tirf	E.ont	t.ts))				
tau =	-0.4	124,	s	L =4	.4046	≥-4%						
> pc.c	hang	ge(t:	irf.d	ont.t	cs)							
[1] 25	.034	110 1	13.91	L354	44.4	12162	2					
Manito	ba											

Manitoba

tirf.r	nan.t	.s										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	1	2	2	2	7	1	6	5	1	6	6	2
1993:	2	1	4	0	5	0	7	4	1	8	5	6
1994:	-		2		-	3	7	9			3	
1995:				6	1	4	9	5		-	2	4
1996:	2	5	0	2	1	4	2	4	3	0	6	4
1997:	8	1	6	0	5	1	4	4	4	5	3	4
1998:	0	1	2	0	3	4	6	4	3	3	7	2
> sum	(tir	E.mai	n.ts))								
[1] 28	39											
> Seas	sonal	lManı	nKend	dall	(tir:	E.Mar	n.ts))				
Proble	em: (Obje	ct "t	tirf	.Man	.ts"	not	four	nd			
> Seas	sonal	lManı	nKend	dall	(tir:	E.mar	n.ts))				
tau =	-0.0	0833	, 3	sl =3	38.4	78						
> pc.0	chang	ge(t:	irf.r	nan.	ts)							
[1]	2.73	33143	1 3	3.32	6042	-21	. 6930	07				
>												

Estimated mean death rate/month from loess smooth.

	Beginning of 1992	End of 1998	Percentage Change
Ontario	25.0	13.9	44
Manitoba	0.32	0.37	-13



Drink Time Series, Ontario

All series show declines.

```
drink = no
 > ont.no.ts
          Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec
 1992: 11 10 7 10 4 18 10 19 14 13 20 18
 1993: 12 13 12 10 18 16 15
                                                                7 8 12 14 27
1994: 19 13 7 8 8 13 16 17 13 16 11 10
1995: 14 14 13 14 10 16 16 15 15 14 18 18
1996: 27 20 14 6 11 12 12 9 7 24 15 10
1997:201798712131212157121998:62467411121111916
> SeasonalMannKendall(ont.no.ts)
tau = -0.187, sl = 4.367%
 > sum(ont.no.ts)
[1] 1056
 drink = yes
 > ont.yes.ts
         Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec

      Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec

      1992:
      3
      7
      8
      10
      13
      11
      9
      13
      7
      6
      9
      10

      1993:
      6
      9
      7
      7
      7
      8
      11
      12
      10
      15
      7
      2

      1994:
      4
      10
      7
      8
      6
      8
      12
      11
      5
      10
      8
      6

      1995:
      3
      3
      7
      9
      6
      7
      9
      10
      10
      3
      4

      1996:
      1
      5
      10
      3
      7
      5
      8
      11
      7
      9
      5
      2

      1997:
      7
      5
      4
      1
      5
      5
      3
      10
      5
      9
      8
      6

      1998:
      4
      1
      4
      5
      4
      4
      6
      4
      6
      3
      2
      4

                                                                                             2
6
> SeasonalMannKendall(ont.yes.ts)
tau = -0.491, sl =1.488e-5%
 > sum(ont.yes.ts)
 [1] 564
 drink = unknown
 > ont.unknown.ts
          Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec

      1992:
      2
      5
      5
      2
      4
      3
      3
      9
      1
      5
      4
      8

      1993:
      5
      7
      5
      2
      5
      9
      8
      9
      5
      3
      7
      7

      1994:
      4
      2
      7
      1
      5
      2
      0
      3
      0
      4
      0
      4

1995: 0 6 0 1 1 5 7 5 2 2 1 8
1996: 4 3 4 1 2 0 3 6 1 2 4 2
1997: 6 2 2 3 7 8 2 4 1 4 2 3
1998: 1 4 0 1 2 1 1 3 0 2 3 3
> SeasonalMannKendall(ont.unknown.ts)
tau = -0.315, sl = 0.08059\%
 > sum(ont.unknown.ts)
[1] 290
```

Drink Time Series, Manitoba

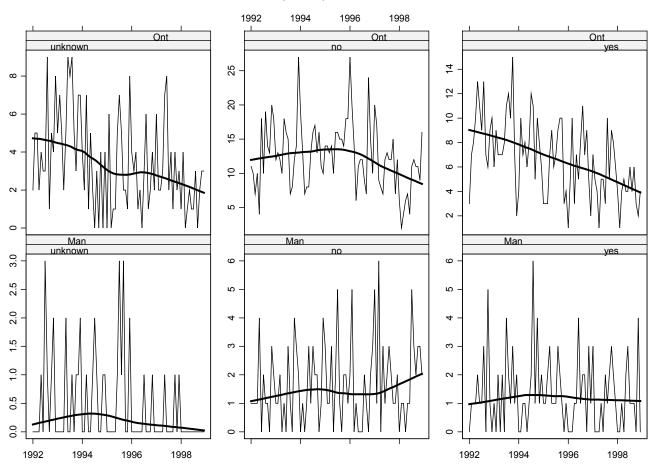
drink = noman.no.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1 1 1 1 4 0 2 1 1 0 3 2 1993: 1 1 2 0 1 0 3 1 0 4 3 2 1994: 0 1 0 1 3 1 3 2 2 0 1 4 1995: 3 1 1 3 0 2 5 1 0 2 2 1 1996: 2 5 0 1 0 0 0 2 1 0 2 3

 1997:
 5
 1
 6
 0
 3
 1
 2
 3
 2
 1
 1
 2

 1998:
 0
 1
 1
 0
 1
 1
 5
 3
 2
 3
 3
 2

 > SeasonalMannKendall(man.no.ts) tau = 0.125, sl =20.16% > sum(man.no.ts) [1] 139 drink = yes> man.yes.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec San Feb Mar Apr May Sun Sur Aug Sep Oct NOV De1992:011121130511993:102020421311994:001101261411995:112311132101996:000114122031997:3001021232 0 2 1998: 0 0 1 0 2 3 1 1 1 0 4 0 > SeasonalMannKendall(man.yes.ts) tau = -0.125, sl = 20.23%> sum(man.yes.ts) [1] 113 drink = unknown > man.unknown.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 0 0 0 0 1 0 3 1 0 1 2 0 1993: 0 0 0 0 2 0 0 1 0 1 1 2 1994: 0 0 1 0 0 1 2 1 0 0 1 1 1995: 0 0 0 0 0 1 3 1 3 0 0 2 1996: 0 0 0 0 0 0 1 0 0 1 0 1997: 0 0 0 0 1 0 0 0 1 0 1 0 0 0 0 0 0 0 0 0 0 1998: > SeasonalMannKendall(man.unknown.ts) WARNING: Error exit, tauk2. IFAULT = 12 WARNING: Error exit, tauk2. IFAULT = 12 WARNING: Error exit, tauk2. IFAULT = 12 tau = -0.383, sl =0.1075% > sum(man.unknown.ts) [1] 37

Clear downward trends in Ontario. The downward trend starts around 1996 for drink=no in Ontario. There may be a small downward trend for drink=unknown in Manitoba and a possible recent upward swing in drink=no in Manitoba.

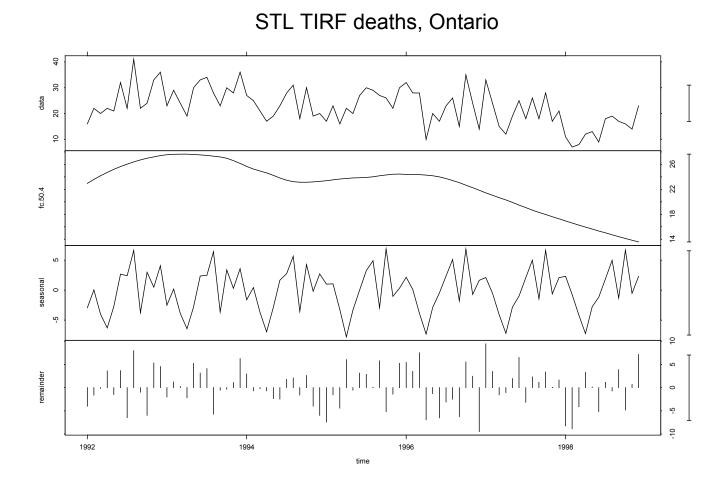


TIRF, drink, scales free

Year

STL Analysis

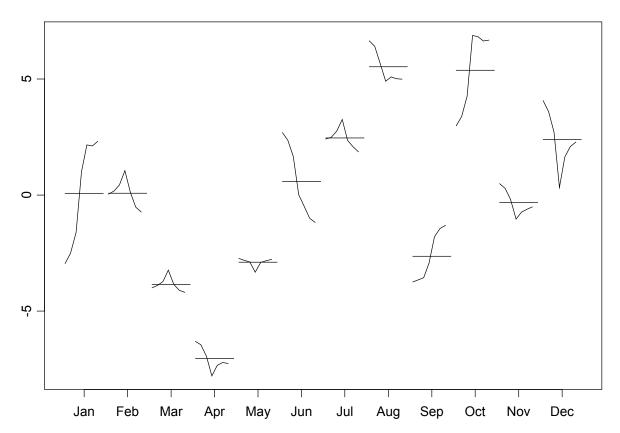
R-sq = 63.4%



ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2

Monthplot of Seasonal Component

Decreasing trend in June and increasing trends in January, September and October. Seasonal minimum in April and local minimum in September. Seasonal maximum in August and in October.

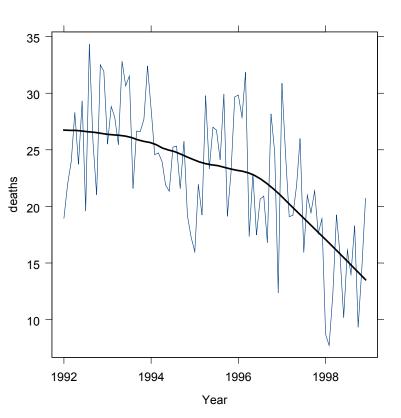


Monthplot, seasonal, TIRF deaths Ontario

Loess Trend Analysis of Deasonalized Series

tau = -0.412, sl = 2.964e - 6%

In summary there appears to be a gradual decline starting in 1992 which turns into a steeper drop starting in the first quarter of 1996.



TIRF, deseasonalilzed deaths

D. MTO Dataset

Summary

The data set of Ministry of Transportation Collision of Ontario, from January 1992 to December 1998, was analyzed by using crosstabs and data visualization.

There are 26,026 records. Each record corresponds to an injured driver in a traffic accident on Ontario's highways (only non-emergency, motorized, highway vehicles). There are four injury classes: mimimal (48.3%), minor (39.8%), major (9.7%) and fatal (2.3%). The four injury classes should be analyzed separately.

All injury classes show increases from 1992 up to 1995 and then a decline. In the fatal accidents the decline post 1995 is much less steep than for the other three injury classes. See <u>fatal</u>, <u>major</u>, <u>minor</u> and <u>minimal</u> time series plots.

There is an increase in deaths in the 2AM-ThuSat window and corresponding decreases in 1AM-ThuSat and 12AM-ThuSat. Also there is a decrease in deaths in 1AM-SunWed window. <u>See trellis plot</u> and <u>Mann-Kendall tests</u>.

Major Injuries decreased in the 12AM and 1AM for both weekgroups but the decrease was stronger in the ThuSat window. <u>See trellis plot</u> and <u>Mann-Kendall tests</u>.

Minor injuries showed decreased slightly in the 12AM-SunWed window. There were strong decreases also in the 1AM SunWed and ThuSat windows. There was a strong increase in the 2AM SunWed window. <u>See trellis plot</u> and <u>Mann-Kendall tests</u>.

Minimal injuries decreased strongly in the 1AM window for both SunWed and ThuSat. There was a strong increase in the 2AM SunWed window and a slight increase in the 2AM ThuSat window. <u>See trellis plot</u> and <u>Mann-Kendall tests</u>.

The mixture of increasing and decreasing trends in the Minor and Minimial injuries could possibly be caused by a change in drinking habits.

STL Analysis of Fatilities suggests a sharply increasing trend to 1995 followed by a decline. The seasonal component quite variables and changing. July and August have the highest deaths. The trend is mildly down for July and sharply increasing for August. See <u>STL</u> and <u>Monthplot</u>.

STL Analysis of Major Injuries shows an increasing trend to about 1995 followed a shart drop to a new low. The seasonal pattern is fairly stable. Highest number of injuries are in July and August. There is an increasing trend in April and decreasing trend in March. See <u>STL</u> and <u>Monthplot</u>.

Frequencies

			Injury		
		Frequency	Percent	Valid Percent	Cumulative Percent
Valid	Minimal	12576	48.3	48.3	48.3
	Minor	10348	39.8	39.8	88.1
	Major	2512	9.7	9.7	97.7
	Fatal	590	2.3	2.3	100.0
	Total	26026	100.0	100.0	

Vehicle Province or State

		Frequency	Percent	Valid Percent	Cumulative Percent
Valid	Alberta	38	.1	.1	.1
	British Columbia	33	.1	.1	.3
	Manitoba	38	.1	.1	.4
	New Brunswick	18	.1	.1	.5
	Newfoundland	5	.0	.0	.5
	Nova Scotia	19	.1	.1	.6
	Ontario	25072	96.3	96.3	96.9
	Prince Edward Island	3	.0	.0	96.9
	Quebec	270	1.0	1.0	98.0
	Saskatchewan	13	.0	.0	98.0
	Yukon and North West Territories	2	.0	.0	98.0
	U.S.A.	284	1.1	1.1	99.1
	Other Foreign	1	.0	.0	99.1
	Unknown	230	.9	.9	100.0
	Total	26026	100.0	100.0	

Crosstabs

Count

		Minimal	Minor	Major	Fatal	Total
Vehicle	Alberta	16	16	5	1	38
Province	British Columbia	10	17	6		33
or State	Manitoba	22	14	1	1	38
	New Brunswick	7	7	4		18
	Newfoundland	3	1	1		5
	Nova Scotia	9	7	2	1	19
	Ontario	12146	9969	2392	565	25072
	Prince Edward Island		2	1		3
	Quebec	129	109	27	5	270
	Saskatchewan	7	4	2		13
	Yukon and North West Territories	1		1		2
	U.S.A.	138	111	27	8	284
	Other Foreign			1		1
	Unknown	88	91	42	9	230
Total		12576	10348	2512	590	26026

Vehicle Province or State * Injury Crosstabulation

Injury * Late Night Time Period Crosstabulation

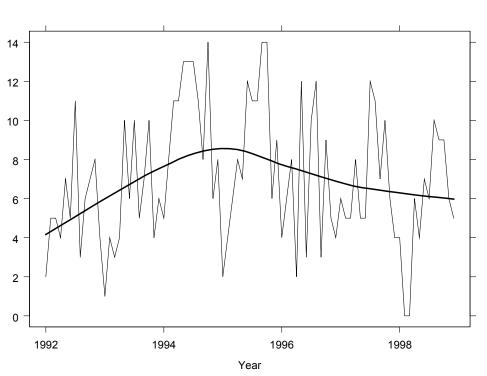
Count								
		Midnight One am Two am Eleven to						
	to 1 am to 2 am to 3 am Midnight							
Injury	Minimal	2836	3052	2451	4237	12576		
	Minor	2308	2684	2165	3191	10348		
	Major	580	708	591	633	2512		
	Fatal	130	162	144	154	590		
Total		5854	6606	5351	8215	26026		

Count

Fatalities

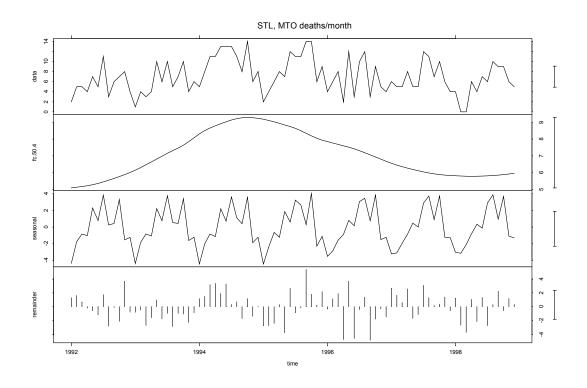
Monthly Overall Deaths

> mto	4.ts											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	2	5	5	4	7	5	11	3	6	7	8	4
1993:	1	4	3	4	10	6	10	5	7	10	4	6
1994:	5	8	11	11	13	13	13	11	8	14	б	8
1995:	2	4	б	8	7	12	11	11	14	14	б	9
1996:	4	б	8	2	12	3	10	12	3	9	5	4
1997:	б	5	5	8	5	5	12	11	7	10	б	4
1998:	4	0	0	б	4	7	б	10	9	9	б	5



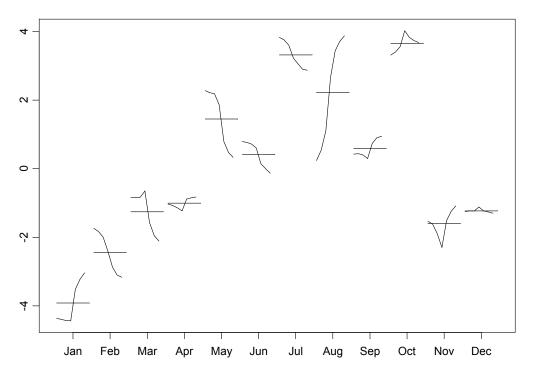
MTO deaths/month

STL Analysis: Fatalities



ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2 , lambda = 1

Seasonal Component Plot



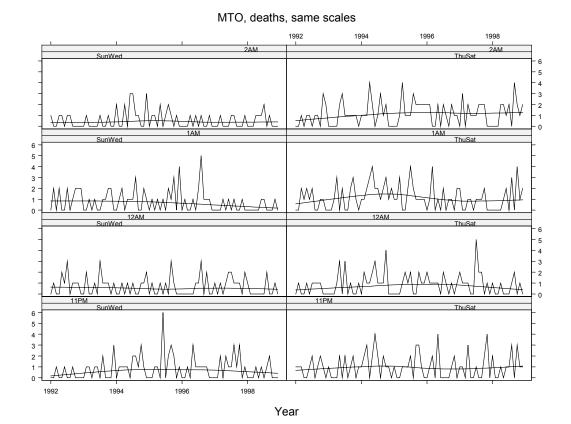
Seasonal Component, MTO, deaths/month

MTO

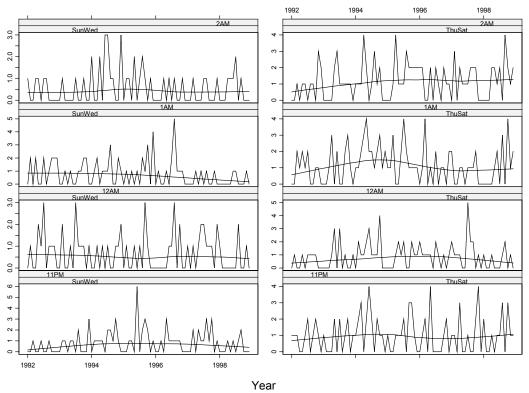
Weekgroup & Hour Time Series: Fatalities

> 11104.	.111	MSur	nWed	.ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	0	1	0	0	1	0	0	1	0	0	0
1993:	0	1	1	0	1	1	0	2	0	0	0	3
1994:	0	1	1	1	1	0	2	2	1	3	1	0
1995:	0	0	1	1	0	б	0	2	3	2	0	1
1996:	0	0	1	0	3	1	1	1	1	1	0	0
1997:	0	0	2	0	2	1	1	3	1	3	0	1
1998:	0	0	0	1	0	1	0	1	2	0	0	0
> Mannł	Kend	dall	(mto	4.111	PMSui	nWed	.ts)					
tau = (0.07	745,	s	l =38	8.249	20						
> mto4.	.127	MSur	nWed	.ts								
· ·	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	1	0	0	2	1	3	0	1	1	1	0
1993:	0	2	0	0	1	0	3	1	1	1	0	1
1994:	0	0	1	0	1	0	1	0	0	1	1	2
1995:	0	1	0	0	1	0	1	0	3	1	0	0
1996:	0	0	0	0	0	1	1	3	0	2	0	0
1997:	1	0	1	0	1	2	2	1	1	1	0	2
1998:	1	0	0	0	0	0	0	2	0	0	1	0
> Mannł	Kend	dall	mto	4.122	AMSui	nWed	.ts)					
tau = -	-0.0)582,	, :	sl =!	50.03	38						
> mto4.	.1AM	ISunV	Ved.t	S								
	-											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	Jan O	Feb 2	Mar 0	Apr 2	May 0	Jun 0	Jul 2	Aug 0	Sep 1	Oct 2	Nov 2	Dec 2
				-	_			-	_			
1992:	0	2	0	2	0	0	2	0	1	2	2	2
1992: 1993:	0 0	2 0	0 1	2 0	0 1	0 0	2 0	0 1	1 1	2 2	2 2	2 0
1992: 1993: 1994:	0 0 0	2 0 1	0 1 2	2 0 0	0 1 1	0 0 1	2 0 1	0 1 3	1 1 0	2 2 0	2 2 2	2 0 1
1992: 1993: 1994: 1995:	0 0 0 0	2 0 1 1	0 1 2 0	2 0 0 1	0 1 1 0	0 0 1 1	2 0 1 0	0 1 3 2	1 1 0 1	2 2 0 3	2 2 2 0	2 0 1 4
1992: 1993: 1994: 1995: 1996:	0 0 0 0	2 0 1 1 1	0 1 2 0 0	2 0 0 1 0	0 1 1 0 1	0 0 1 1 0	2 0 1 0 2	0 1 3 2 5	1 1 0 1 1	2 2 0 3 1	2 2 2 0 1	2 0 1 4 0
1992: 1993: 1994: 1995: 1996: 1997:	0 0 0 0 0 0	2 0 1 1 0 0	0 1 2 0 0 0 0	2 0 1 0 1 0	0 1 1 0 1 0	0 0 1 1 0 1	2 0 1 0 2 0 1	0 1 3 2 5 1	1 1 0 1 1 0	2 2 0 3 1 0	2 2 0 1 1	2 0 1 4 0 0
1992: 1993: 1994: 1995: 1996: 1997: 1998:	0 0 0 0 0 0 0	2 0 1 1 0 0 0 all(0 1 2 0 0 0 0 (mto	2 0 1 0 1 0 4.1AI	0 1 1 0 1 0	0 1 1 0 1 1 Wed.1	2 0 1 0 2 0 1	0 1 3 2 5 1	1 1 0 1 1 0	2 2 0 3 1 0	2 2 0 1 1	2 0 1 4 0 0
1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannH	0 0 0 0 0 0 0 0 0 0	2 0 1 1 0 0 dall(0 1 2 0 0 0 0 (mto s	2 0 1 0 1 0 4.1AM L =4	0 1 0 1 0 0 0 MSunT	0 1 1 0 1 1 Wed.1	2 0 1 0 2 0 1	0 1 3 2 5 1	1 1 0 1 1 0	2 2 0 3 1 0	2 2 0 1 1	2 0 1 4 0 0
1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannH tau = - > mto4	0 0 0 0 0 0 0 0 0 0 0 2 2 0	2 0 1 1 0 0 dall(.74 , (Sun)	0 1 2 0 0 0 (mto s	2 0 1 0 1 0 4.1Ar L =4 5	0 1 1 0 1 0 MSunI •123	0 0 1 0 1 1 8 Ved.1	2 0 1 0 2 0 1 1 5)	0 1 3 2 5 1 0	1 1 0 1 1 0 0	2 2 0 3 1 0 0	2 2 0 1 1	2 0 1 4 0 0 0
1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannH tau = - > mto4	0 0 0 0 0 0 0 0 0 0 0 2 2 0	2 0 1 1 0 0 dall(.74 , (Sun)	0 1 2 0 0 0 (mto s	2 0 1 0 1 0 4.1Ar L =4 5	0 1 1 0 1 0 MSunI •123	0 0 1 0 1 1 8 Ved.1	2 0 1 0 2 0 1 1 5)	0 1 3 2 5 1 0	1 1 0 1 1 0 0	2 2 0 3 1 0 0	2 2 0 1 1	2 0 1 4 0 0 0
1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manne tau = - > mto4	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	2 0 1 1 0 0 dall(74 , KsunW Feb	0 1 2 0 0 0 (mto4 5 Ved.1 Mar	2 0 1 0 4.1AN L =4 CS Apr	0 1 0 0 MSun .123 May	0 1 1 0 1 1 Wed.1	2 0 1 0 2 0 1 1 5) Jul	0 1 3 2 5 1 0 8	1 1 0 1 1 0 0 Sep	2 2 0 3 1 0 0	2 2 0 1 1 1 Nov	2 0 1 4 0 0 0 0
1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannH tau = - > mto4	0 0 0 0 0 (end -0.1 .2AM Jan 1	2 0 1 1 0 0 1 1 0 0 1 1 0 1 5 1 0 1 5 1 0	0 1 2 0 0 0 (mto s Ved.t Mar 0	2 0 1 0 4.1AN 1 =4 cs Apr 1	0 1 0 0 MSun .123 May 1	0 0 1 1 0 1 1 8 Ved.1	2 0 1 0 2 0 1 5 5) Jul 1	0 1 3 2 5 1 0 8 4ug 1	1 1 0 1 1 0 0 0 Sep 0	2 2 0 3 1 0 0 0	2 2 0 1 1 1 Nov	2 0 1 4 0 0 0 0 0 0 0 0 0
1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannH tau = - > mto4 1992: 1993:	0 0 0 0 0 0 0 0 0 0 2 A M Jan 1 0	2 0 1 1 0 0 4all(74 , (SunV Feb 0 1	0 1 2 0 0 0 0 (mto s Ned.t Mar 0 0	2 0 1 0 1 0 4.1AM L =4 CS Apr 1 0	0 1 1 0 1 0 0 MSun 1 .123 2 May 1 0	0 1 1 1 Wed.1 8 Jun 0 0	2 0 1 0 2 0 1	0 1 3 2 5 1 0 8 4 0 8 4 0 9	1 0 1 1 0 0 0 Sep 0 0	2 0 3 1 0 0 0 0 0 0 0 0	2 2 0 1 1 1 1 Nov 0 0	2 0 1 4 0 0 0 0 0 0 0 0
1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannH tau = - > mto4 1992: 1993: 1994:	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	2 0 1 1 0 0 dall(74 , ISunW Feb 0 1 0	0 1 2 0 0 0 0 0 (mto s Ved.t Mar 0 0 0	2 0 0 1 0 4.1AN L =4 5 Apr 1 0 2	0 1 1 0 1 0 0 MSun 1 .123 5 May 1 0 0	0 0 1 1 1 Wed.1 5 Jun 0 3	2 0 1 0 2 0 1 1 5 5) Jul 1 1 3	0 1 3 2 5 1 0 4 ug 1 0 1	1 0 1 1 0 0 0 Sep 0 0 1	2 2 0 3 1 0 0 0 0 0 0 0 1 0	2 2 0 1 1 1 1 Nov 0 0 0	2 0 1 4 0 0 0 0 2 0 3
1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannH tau = - > mto4 1992: 1993: 1994: 1995:	0 0 0 0 2 AM Jan 1 0 2 0	2 0 1 1 0 0 dall(74 , ISunW Feb 0 1 0 1	0 1 2 0 0 0 0 (mto s Ned.t Mar 0 0 1	2 0 0 1 0 4.1AN L =4 5 Apr 1 0 2 0	0 1 1 0 1 0 0 MSunM .123 ⁵ May 1 0 0 2	0 0 1 1 1 Wed.1 5 Jun 0 3 0	2 0 1 2 0 1 :s) Jul 1 3 1	0 1 3 2 5 1 0 4 4 0 1 2	1 0 1 1 0 0 0 Sep 0 0 1 1	2 2 0 3 1 0 0 0 0 0 0 1 0 0	2 2 0 1 1 1 1 Nov 0 0 0 1	2 0 1 4 0 0 0 0 0 3 0
<pre>1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannH tau = - > mto4 1992: 1993: 1994: 1995: 1996:</pre>	0 0 0 0 0 0 0 0 0 0 0 0 0 1 2 0 0 0	2 0 1 1 0 0 4 all(74, 1 SunV Feb 0 1 0 1 0	0 1 2 0 0 0 (mto s Ved.t Mar 0 0 0 1 0	2 0 0 1 0 1 0 4.1AM 1 =4 5 Apr 1 0 2 0 1	0 1 0 0 MSun 123 May 1 0 0 2 0	0 0 1 0 1 Ved.1 5 Jun 0 3 0 1	2 0 1 2 0 1 5 5) Jul 1 3 1 0	0 1 3 2 5 1 0 Aug 1 0 1 2 1	1 1 0 1 1 0 0 0 5 ep 0 0 1 1 0	2 2 0 3 1 0 0 0 0 1 0 0 0 0 0	2 2 0 1 1 1 1 Nov 0 0 0 1 1	2 0 1 4 0 0 0 0 3 0 0 0
<pre>1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannH tau = - > mto4 1992: 1993: 1994: 1995: 1996: 1997:</pre>	0 0 0 0 0 0 0 0 0 0 0 0 1 2 2 0 0 0 0 0	2 0 1 1 0 0 1 4 1 0 1 0 1 0 0 0 0 0 0 0	0 1 2 0 0 0 (mto s Ved.1 Mar 0 0 0 1 0 1 0	2 0 0 1 0 4.1AN 1 =4 5 S Apr 1 0 2 0 1 0 1	0 1 1 0 0 MSunW .123 12 12	0 0 1 1 Ned.1 5 Jun 0 0 3 0 1 0 1	2 0 1 2 0 1 5 5) Jul 1 3 1 0 1 2	0 1 3 2 5 1 0 4 ug 1 2 1 2 1 1	1 1 0 1 1 0 0 0 5 ep 0 0 1 1 0 0	2 2 0 3 1 0 0 0 0 1 0 0 0 0 0 0 0	2 2 0 1 1 1 1 0 0 0 0 1 1 0	2 0 1 4 0 0 0 0 3 0 0 1

mto4.11PM	ThuSa	at.ts	5									
Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992: 1	1	1	0	0	1	2	0	1	2	1	0	
1993: 1	0	0	0	2	0	2	0	1	2	0	1	
1994: 1	2	3	0	2	4	2	0	2	1	1	0	
1995: 1	0	0	1	1	1	2	0	3	3	1	0	
1996: 0	1	2	0	4	0	0	0	0	1	2	0	
1997: 1	1	0	3	0	0	1	0	0	2	4	0	
1998: 2	0	0	1	0	1	1	3	0	3	1	1	
> MannKen	dall	(mto4	4.111	PMThu	JSat	.ts)						
tau = 0.0	0375	, :	sl =9	96.74	48							
> mto4.12	AMTh	uSat	.ts									
Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992: 0	1	0	0	1	0	1	1	1	1	0	0	
1993: 0	0	0	1	3	0	3	0	1	0	0	1	
1994: 0	2	1	1	2	3	1	1	1	4	0	0	
1995: 0	0	0	1	2	1	2	0	0	2	1	1	
1996: 2	1	1	1	1	0	2	1	0	1	0	1	
1997: 2	1	1	1	0	0	5	2	2	0	1	0	
1998: 1	0	0	1	0	0	0	1	2	0	1	0	
> MannKen	dall	(mto	4.122	AMThu	JSat	.ts)						
tau = 0.0	321,	s	L =70).79 ⁹	20							
> mto4.1A	MThu	Sat.	S									
Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992: 0	0	2	1	2	1	2	0	0	1	1	0	
1993: 0	0	1	3	0	2	0	0	2	3	1	0	
1994: 1	1	2	3	4	2	2	1	2	3	1	2	
1995: 1	1	3	0	0	2	4	2	1	1	1	1	
1996: 0	1	4	0	1	0	2	0	1	1	0	2	
1997: 2	0	0	1	1	0	1	1	1	2	0	0	
1998: 0	0	0	0	1	2	0	3	0	4	1	2	
> MannKen	dall	(mto4	4.1AM	MThu	Sat.	ts)						
tau = -0.	0472	, :	sl =!	57.12	28							
> mto4.2A	MThu	Sat.1	S									
Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992: 0	0	1	0	1	1	0	1	1	0	3	2	
1993: 0	0	0	0	2	3	1	1	1	1	1	0	
1994: 1	1	1	4	2	0	1	3	1	2	0	0	
1995: 0	0	1	4	1	1	1	3	2	2	2	2	
1996: 2	2	0	0	2	0	2	1	0	2	1	1	
1997: 0	3	0	2	1	1	1	2	2	2	0	0	
1998: 0		0	2	2	1	2	0	4	2	1	2	
> MannKen	dall	(mtoʻ	4.2AI	MThu	Sat.	ts)						
tau = 0.1	32,	sl	=11	.38%								



Time series trellis plot. Fatalities. Same scales.

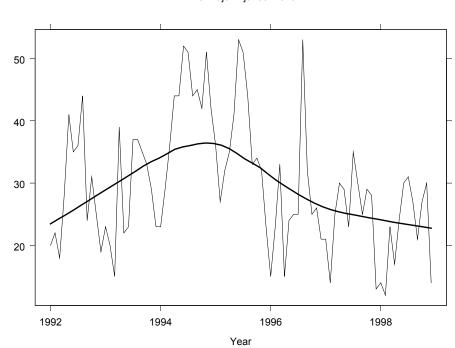


Time series trellis plot. Fatalities. Scales free.

MTO, deaths, scales free

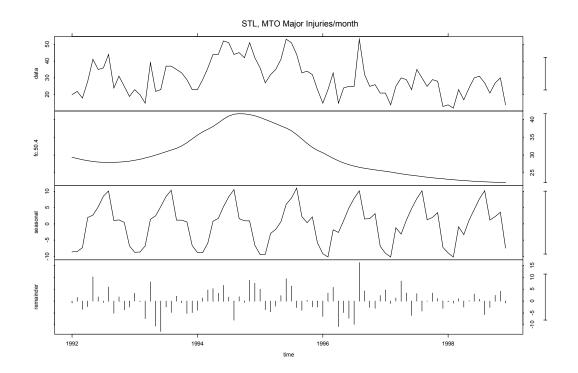
Major Injuries Monthly Overall Major Injuries

> mto	3.ts											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug			Nov	Dec
1992:	20	22	18	28	41	35	36	44	24	31	25	19
1993:	23	20	15	39	22	23	37	37	35	33	29	23
1994:	23	29	36	44	44	52	51	44	45	42	51	42
1995:	36	27	32	35	41	53	51	44	33	34	32	23
1996:	15	23	33	15	24	25	25	53	32	25	26	21
1997:	21	14	25	30	29	23	35	30	25	29	28	13
1998:	14	12	23	17	24	30	31	27	21	27	30	14



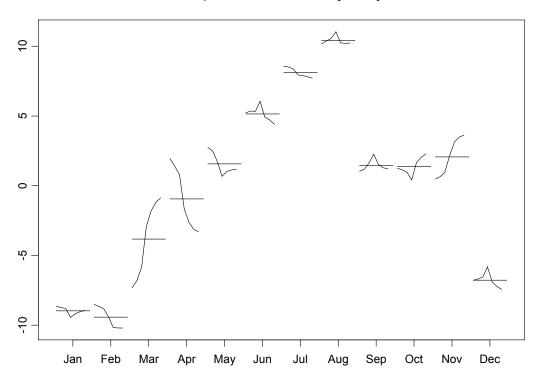
MTO major injuries/month

STL: Major Injuries



ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2 , lambda = 1

Seasonal Component: Major Injuries

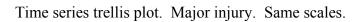


Seasonal Component, MTO, Major Injuries/month

Weekgroup & Hour Time Series: Major Injuries

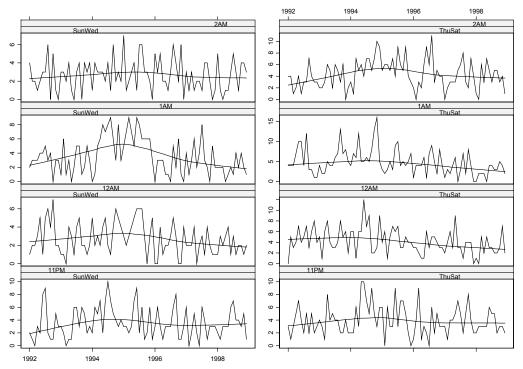
> mto3.	11PM	ISun	wed.	ts								
J	an I	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	2	1	0	3	2	8	9	2	1	1	5	3
1993:	3	2	0	1	1	6	6	3	6	5	2	3
1994:	2	6	5	7	2	7	10	7	5	4	3	4
1995:	3	3	2	3	7	9	2	6	1	2	6	1
1996:	3	6	3	1	3	3	3	6	8	1	1	3
1997:	6	0	2	5	2	0	6	4	1	3	3	3
1998:	2	1	3	3	3	6	7	4	4	3	5	1
> MannK	enda	all(mto3	3.111	PMSui	nWed	.ts)					
tau = 0	.022	26,	s	L =7'	7.76	20						
> mto3.	12AN	MSun	wed.	ts								
J	an I	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	1	2	2	3	5	1	5	6	4	7	2	2
1993:	1	1	0	4	3	1	4	5	2	2	1	2
1994:	5	2	3	2	4	5	2	1	4	б	5	4
1995:	3	2	3	4	5	6	6	6	3	0	3	5
1996:	0	0	4	1	2	2	3	5	3	0	0	4
1997:	2	2	2	4	0	2	3	1	4	2	1	1
1998:	1	3	2	3	4	1	3	1	2	2	1	2
> MannK	enda	all(mto3	3.124	AMSui	nWed	.ts)					
tau = -	0.13	33,	s	L =9.	. 555 ⁹	8						
> mto3.	1AMS	SunW	led.t	s								
J	an I	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	2	3	3	3	4	4	5	3	4	0	3	3
1993:	1	6	1	3	0	2	5	5	1	2	5	3
1994:	0	1	5	6	8	7	8	9	6	3	8	3
1995:	5	7	9	7	5	9	8	6	6	6	6	3
1996:	0	3	3	3	1	1	0	5	2	6	1	0
1997:	4	1	3	6	2	4	8	3	2	0	5	2
1998:	2	2	2	0	1	2	1	4	2	4	2	1
> MannK	enda	all(mto3	3.1AM	ISun	Wed.t	s)					
tau = -					L.979							
> mto3.	2AMS	SunW	led.t	s								
J	an I	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	4	2	2	1	2	3	3	б	0	5	1	0
1993:	3	3	2	4	1	0	3	4	0	4	3	4
1994:	1	4	3	3	3	2	4	1	6	2	3	2
1995:	7	1	2	3	4	1	6	6	3	3	2	0
1996:	5	3	5	2	2	1	3	6	4	2	6	0
1997:	3	1	3	1	2	2	1	4	4	2	0	1
1998:	5	1	0	1	1	3	5	3	1	4	4	3
> MannK	enda	all(mto3	3.2AM	ISun	Wed.t	s)					
tau = -					79.84							

< #												
> # mto3.111	יייאר	huga	+ + a									
			Mar .	۸nr	Max	Tun		λυα	Con	Oat	Nov	Dec
1992:	3	1	3	прі 5	may 7	4	2	5	3ep 2	5	2	3
1993:	4	3	1	8	4	5	4	4	2	2	4	2
1994:	2	2	6	3	10	10	7	5	9	5	3	6
1995:	6	0	6	3	3	9	3	7	7	5	2	0
1996:	1	4	9	1	3	2	0	, 6	2	5	3	3
1997:	3	1	4	4	5	7	5	2	5	8	4	2
1998:	2	3	3	3	3	6	5	5	2	3	3	2
> MannKe	_	-	-	-	-	-	-	5	-	5	5	2
tau = -(76.01		,					
> mto3.1						2.0						
			Mar .		Mav	Jun	Jul	Αιια	Sep	Oct	Nov	Dec
1992:	0	5	3	4	7	4	5	7	3	6	8	4
1993:	5	1	6	8	4	3	3	4	8	5	6	2
1994:	6	6	1	6	6	12	7	9	2	2	3	9
1995:	4	6	1	5	7	6	7	3	3	5	4	4
1996:	3	3	2	1	1	6	3	5	5	4	3	3
1997:	2	4	6	3	9	3	3	1	4	4	4	0
1998:	1	0	5	2	5	3	3	2	2	3	7	2
> MannKe	enda											
tau = -0					.536		,					
> mto3.1												
	L'AM.	ThuS	at.t	S								
					May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
			at.t Mar 4		May 10	Jun 10	Jul 4	Aug 12	Sep 3	Oct 3	Nov 1	Dec 1
Ja	an 1	Feb	Mar .	Apr	_			_	_			
Ja 1992:	an 1 4	Feb 4	Mar . 4	Apr 7	10	10	4	12	3	3	1	1
Ja 1992: 1993:	an 1 4 4	Feb 4 2	Mar . 4 2	Apr 7 5	10 4	10 4	4 6	12 7	3 13	3 7	1 8	1 5
Ja 1992: 1993: 1994:	an 1 4 4 4	Feb 4 2 7	Mar . 4 2 6	Apr 7 5 12	10 4 5	10 4 5	4 6 6	12 7 5	3 13 8	3 7 13	1 8 16	1 5 5
Ja 1992: 1993: 1994: 1995:	an 1 4 4 4 3	Feb 4 2 7 2	Mar . 4 2 6 3	Apr 7 5 12 5	10 4 5 3	10 4 5 9	4 6 6 10	12 7 5 4	3 13 8 5	3 7 13 4	1 8 16 5	1 5 5 7
Ja 1992: 1993: 1994: 1995: 1996:	an 1 4 4 4 3 1	Feb 4 2 7 2 4	Mar . 4 2 6 3 4	Apr 7 5 12 5 4	10 4 5 3 6	10 4 5 9 1	4 6 10 7	12 7 5 4 9	3 13 8 5 5	3 7 13 4 2	1 8 16 5 8	1 5 7 4
Ja 1992: 1993: 1994: 1995: 1996: 1997:	an 1 4 4 3 1 0	Feb 4 2 7 2 4 3 2	Mar . 2 6 3 4 2 2	Apr 7 5 12 5 4 4 2	10 4 5 3 6 6 0	10 4 5 9 1 0 4	4 6 10 7 3 4	12 7 5 4 9 7	3 13 8 5 5 2	3 7 13 4 2 8	1 8 16 5 8 4	1 5 7 4 0
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998:	an 1 4 4 3 1 0 end	Feb 4 2 7 2 4 3 2 all(Mar . 4 2 6 3 4 2 2 mto3	Apr 7 5 12 5 4 4 2 .1AM	10 4 5 3 6 6 0	10 4 5 9 1 0 4 Sat.1	4 6 10 7 3 4	12 7 5 4 9 7	3 13 8 5 5 2	3 7 13 4 2 8	1 8 16 5 8 4	1 5 7 4 0
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe	an 1 4 4 3 1 0 enda	Feb 4 2 7 2 4 3 2 all(14 ,	Mar . 4 2 6 3 4 2 2 mto3 sl	Apr 7 5 12 5 4 2 .1AN =0	10 4 5 3 6 6 0 4Thus	10 4 5 9 1 0 4 Sat.1	4 6 10 7 3 4	12 7 5 4 9 7	3 13 8 5 5 2	3 7 13 4 2 8	1 8 16 5 8 4	1 5 7 4 0
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe tau = -(> mto3.2	an 1 4 4 3 1 1 0 enda 2 2AM	Feb 4 2 7 2 4 3 2 all(14 , ThuS	Mar . 4 2 6 3 4 2 2 mto3 sl	Apr 7 5 12 5 4 4 2 .1AN =0	10 4 5 3 6 6 0 4 Thus 595 4	10 4 5 9 1 0 4 Sat.1	4 6 10 7 3 4 55)	12 7 5 4 9 7 3	3 13 8 5 5 2 3	3 7 13 4 2 8 5	1 8 16 5 8 4 4	1 5 7 4 0 2
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe tau = -(> mto3.2 Ja 1992:	an : 4 4 4 3 1 1 0 end: 2 2 AM ² 4	Feb 4 2 7 2 4 3 2 all(14 , ThuS Feb 4	Mar . 4 2 6 3 4 2 2 mto3 sl at.t	Apr 7 5 12 5 4 4 2 .1AN =0	10 4 5 3 6 6 0 4 Thus 595 4	10 4 5 9 1 0 4 5at.1 1 % Jun 1	4 6 10 7 3 4 55)	12 7 5 4 9 7 3 8 Aug 3	3 13 8 5 2 3 Sep 7	3 7 13 4 2 8 5	1 8 16 5 8 4 4	1 5 7 4 0 2 Dec 3
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe tau = -(> mto3.2 Ja	an 1 4 4 3 1 1 0 enda 2 2 AM	Feb 4 2 7 2 4 3 2 all(14 , ThuS Feb	Mar . 4 2 6 3 4 2 2 mto3 sl at.t Mar .	Apr 7 12 4 4 2 .1AN =0 S	10 4 5 3 6 0 4 Thus 5954 May	10 4 5 9 1 0 4 Sat.1 4 %	4 6 10 7 3 4 5s)	12 7 5 4 9 7 3 8	3 13 8 5 2 3 Sep	3 7 13 4 2 8 5 0ct	1 8 16 5 8 4 4	1 5 7 4 0 2 Dec
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe tau = -(> mto3.2 Ja 1992:	an : 4 4 4 3 1 1 0 end: 2 2 AM ² 4	Feb 4 2 7 2 4 3 2 all(14 , ThuS Feb 4	Mar . 4 2 6 3 4 2 2 mto3 sl at.t Mar . 1	Apr 7 5 12 5 4 4 2 .1AN =0 5 Apr 2	10 4 5 3 6 0 4 Thus 5954 May 4	10 4 5 9 1 0 4 5at.1 1 % Jun 1	4 6 10 7 3 4 55) Jul 3	12 7 5 4 9 7 3 8 Aug 3	3 13 8 5 2 3 Sep 7	3 7 13 4 2 8 5 0ct 4	1 8 16 5 8 4 4 4 8 8 8 4 3	1 5 7 4 0 2 Dec 3
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe tau = -(> mto3.2 Ja 1992: 1993: 1994: 1995:	an : 4 4 3 1 1 0 2 2 AM ² 2 2 AM ² 3 5	Feb 4 2 7 2 4 3 2 all(14 , ThuS Feb 4 2	Mar . 4 2 6 3 4 2 2 mto3 sl 3 at.t Mar . 1 3 7 6	Apr 7 5 12 5 4 4 2 .1AN =0. 5 5 5 5	10 4 5 3 6 0 4 Thus 5 954 May 4 5	10 4 5 9 1 0 4 Sat.1 1 2	4 6 10 7 3 4 55) Jul 3 6	12 7 5 4 9 7 3 8 Aug 3 5	3 13 8 5 2 3 3 Sep 7 3 5 5	3 7 13 4 2 8 5 0 ct 4 6	1 8 16 5 8 4 4 4 8 9 8 9 8 9 9 0	1 5 7 4 0 2 Dec 3 2
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe tau = -(> mto3.2 Ja 1992: 1993: 1994:	an : 4 4 4 3 1 1 0 2 2 AM ⁴ 2 3	Feb 4 2 7 4 3 2 all(14 , ThuS Feb 4 2 1	Mar . 4 2 6 3 4 2 2 mto3 sl 3 at.t Mar . 1 3 7	Apr 7 5 12 5 4 4 2 .1AN =0 5 5 5 2	10 4 5 3 6 0 4 Thus 5 5 5 5 4 5 6 7 6	10 4 5 9 1 0 4 Sat.1 1 2 4	4 6 10 7 3 4 55) Jul 3 6 7	12 7 5 4 9 7 3 8 4 9 7 3 8 5 7	3 13 8 5 2 3 Sep 7 3 5 5 3	3 7 13 4 2 8 5 0ct 4 6 7 9 5	1 8 16 5 8 4 4 4 8 Nov 3 0 10 4 4	1 5 7 4 0 2 Dec 3 2 9
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe tau = -(> mto3.2 Ja 1992: 1993: 1994: 1995:	an : 4 4 1 1 0 end: 2 2 AM 2 3 5 2 0	Feb 4 2 7 4 3 2 all(14 , ThuS Feb 4 2 1 6	Mar . 4 2 6 3 4 2 2 mto3 sl 3 7 6 3 3	Apr 7 5 12 5 4 4 2 .12 5 4 4 2 .12 5 5 4 5 5 5 5 5 2 3	10 4 5 6 0 4 Thus 5 5 5 5 4 5 6 7 6 3	10 4 5 9 1 0 4 Sat.1 1 2 4 4 9 5	4 6 10 7 3 4 5 5) Jul 3 6 7 9 6 6	12 7 5 4 9 7 3 8 Aug 3 5 7 6 11 8	3 13 8 5 2 3 Sep 7 3 5 5 3 3 3	3 7 13 4 2 8 5 0ct 4 6 7 9 5 2	1 8 16 5 8 4 4 4 8 9 10 10 4	1 5 7 4 0 2 Dec 3 2 9 3 4 4
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe tau = -(> mto3.2 Ja 1992: 1993: 1994: 1995: 1996:	an : 4 4 3 1 1 0 end 3 0.2 2 AM 4 2 3 5 2	Feb 4 2 7 2 4 3 2 all(14 , ThuS Feb 4 2 1 6 0	Mar . 4 2 6 3 4 2 2 mto3 sl 3 4 1 3 7 6 3	Apr 7 5 12 5 4 4 2 .1AN =0 5 5 5 2	10 4 5 3 6 0 4 Thus 5 5 5 5 4 5 6 7 6	10 4 5 9 1 0 4 Sat.1 1 2 4 4 9	4 6 10 7 3 4 5 5) Jul 3 6 7 9 6	12 7 5 4 9 7 3 8 Aug 3 5 7 6 11	3 13 8 5 2 3 Sep 7 3 5 5 3	3 7 13 4 2 8 5 0ct 4 6 7 9 5	1 8 16 5 8 4 4 4 8 Nov 3 0 10 4 4	1 5 7 4 0 2 Dec 3 2 9 3 4
Ja 1992: 1993: 1994: 1995: 1996: 1997: 1998: > MannKe tau = -(> mto3.2 Ja 1992: 1993: 1994: 1995: 1996: 1997:	an : 4 4 3 1 1 0 22AM 4 2 2AM 4 2 3 5 2 0 1 2 0	Feb 4 2 7 2 4 3 2 all(14 , ThuS Feb 4 2 1 6 0 2 0 all(Mar . 4 2 6 3 4 2 2 mto3 sl 3 4 1 3 7 6 3 6 3 6 mto3	Apr 7 5 12 5 4 4 2 .1AN 2 6 5 5 2 3 3 .2AN	10 4 5 6 0 4 Thus 5 954 May 4 5 6 7 6 3 7	10 4 5 9 1 0 4 5 5 5 5 5 5 5 5 5 5 5 5 5 1.1	4 6 10 7 3 4 5 5) Jul 3 6 7 9 6 6 3	12 7 5 4 9 7 3 8 Aug 3 5 7 6 11 8	3 13 8 5 2 3 Sep 7 3 5 5 3 3 3	3 7 13 4 2 8 5 0ct 4 6 7 9 5 2	1 8 16 5 8 4 4 4 8 8 9 10 10 4 7	1 5 7 4 0 2 Dec 3 2 9 3 4 4



2AM 2AM SunWed ThuSa 11PM ThuS unW/ Γ. Year

MTO, Major Injury, Same Scales



Time series trellis plot. Major injury. Scales free.

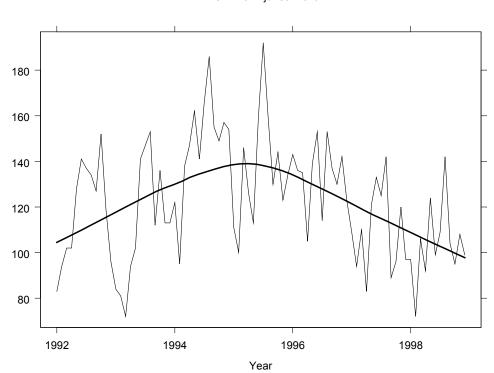
MTO, Major Injury, Scales Free

Year

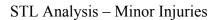
Minor Injuries

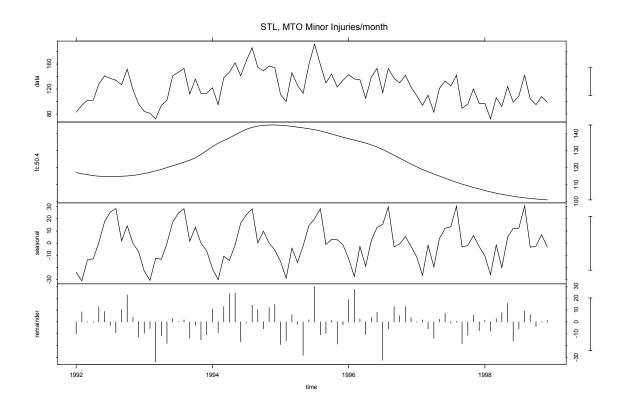
Monthly Overall Minor Injuries

> mto:	2.ts											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	83	94	102	102	128	141	137	134	127	152	119	96
1993:	84	81	72	94	102	141	147	153	112	136	113	113
1994:	122	95	138	147	162	141	166	186	155	149	157	154
1995:	111	100	146	126	113	159	192	159	130	144	123	134
1996:	143	136	135	105	139	153	114	153	137	130	142	123
1997:	109	94	110	83	121	133	125	142	89	96	120	97
1998:	97	72	106	92	124	99	109	142	104	95	108	99



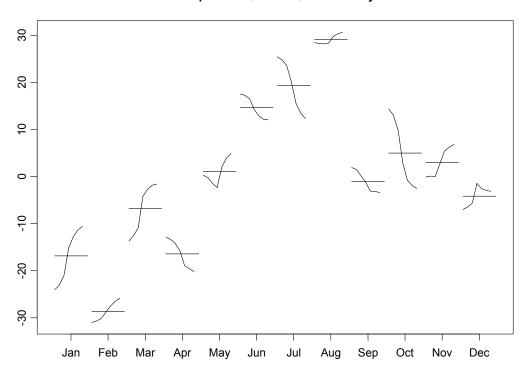
MTO minor injuries/month





ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2 , lambda = 1

Seasonal Component - Minor Injuries



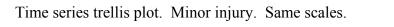
Seasonal Component, MTO, Minor Injuries/month

MTO

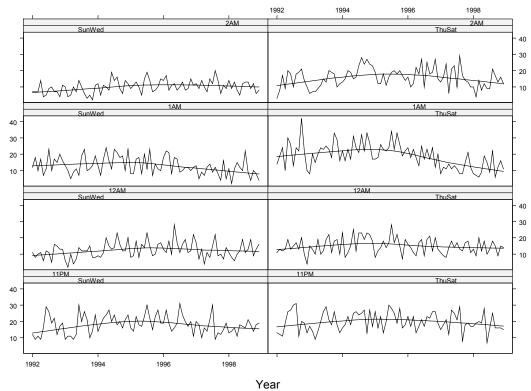
Weekgroup & Hour Time Series: Minor Injuries

> mto2.1	1PMSu	nWed	.ts								
Ja	in Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992: 1	.2 9	11	9	16	29	26	19	22	12	17	19
1993:	9 11	11	9	12	30	20	26	21	10	14	24
1994: 1	4 18	22	24	27	20	24	18	20	16	21	24
1995: 1	6 14	23	18	17	23	30	22	15	23	27	19
1996: 1	9 27	19	14	16	19	31	24	20	17	20	11
1997: 2	0 18	10	16	14	21	30	9	13	12	15	19
1998: 1	.3 14	16	11	18	17	16	21	18	14	18	19
> MannKe	endall	(mto2	2.11	PMSui	nWed	.ts)					
tau = -0	.0112	, 8	sl =	88.63	1%						
> mto2.1	2AMSu	nWed	.ts								
Ja	in Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992: 1	.1 8	10	11	б	12	11	5	16	15	13	13
1993:	6 2	10	4	7	14	12	12	12	15	8	8
1994:	9 8	11	20	15	13	14	23	16	12	12	20
1995:	8 9	18	8	16	14	23	15	9	10	16	17
1996: 1	.3 11	18	10	28	17	11	16	19	11	12	22
1997: 1	.7 7	13	4	11	8	15	22	9	10	7	14
1998: 1	.1 8	6	10	12	19	11	13	19	9	13	16
> MannKe	ndall	(mto2	2.12	AMSui	nWed	.ts)					
tau = 0	124,	sl	=10	.53%							
> mto2.1	AMSun	Wed.t	ts								
Ja	in Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
	.2 18		17	7	11	23	10	17	15	20	16
1993: 1	.3 10	5	9	11	7	19	23	10	11	13	19
1994: 1	.3 7	17	24	16	11	23	21	18	19	10	23
1995:	8 8	17	18	8	20	11	23	7	14	11	10
1996: 1	8 22	21	8	18	17	9	10	12	12	10	12
1997: 1	.3 5	9	7	12	14	9	12	8	4	16	5
1998: 1	.1 2	11	15	12	10	22	9	4	10	8	4
> MannKe	ndall	(mto2	2.1A	MSuni	Wed.	ts)					
tau = -0	.203,	s	1 =0	.778	2 %						
> mto2.2			ts								
Ja	in Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	7 7	7	14	4	5	9	10	7	7	4	11
1993: 1	.0 4		10	7	14	10	6	3	5	2	11
	.2 5	11	10	8	19	14	16	9	6	14	12
	9 11		10	5	15	19	14	7	8	10	15
	.4 17		11	12	8	10	13	8	9	15	11
	.2 9		10	7	15	12	20	14	6	11	9
	.4 9		8	5	12	13	13	9	12	6	8
> MannKe											
tau = 0		sl =(

	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	D
1992:	13	12	11	19	26	27	30	31	11	20	19	
1993:	17	14	9	14	20	26	29	21	18	26	18	
1994:	16	18	23	18	25	26	24	25	27	13	27	
1995:	26	16	23	13	20	30	28	25	19	28	16	
1996:	23	17	23	13	17	21	17	23	22	26	26	
1997:	15	18	21	12	27	25	21	29	8	18	19	
1998:	17	19	20	19	28	7	16	25	16	16	16	
> Mann	Kend	dall	(mto)	2.11		uSat	.ts)					
tau =				sl =								
> mto2				.ts								
					Mav	Jun	Jul	Auq	Sep	Oct	Nov	D
1992:	11	13	12	13	19	13	15	17	13	20	10	
1993:	15	13	10	19	10	12	19	22	11	18	19	
1994:	23	8	11	15	25	12	23	23	19	23	22	
1995:	8	12	14	18	14	17	28	17	22	15	7	
1996:	16	14	11	9	17	19	9	19	21	12	20	
1997:	11	9	8	14	18	14	17	13	12	17	18	
1998:	13	9	20	12	18	10	13	18	15	9	15	
> Mann		dall	(mto	2.12	AMTh	uSat	.ts)					
tau =	-0.0	0346	, :	sl =	65.3	1%						
> mto2	2.1A	MThus	Sat.	ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	D
1992:	14	19	24	10	30	26	13	24	22	42	22	
1993:	8	20	15	20	24	23	25	23	18	33	29	
1994:	22	16	23	17	31	24	21	32	22	33	27	
1995:	17	18	26	23	22	24	34	23	33	26	19	
1996:	24	18	23	18	13	25	17	23	17	24	12	
1997:	8	12	11	14	11	13	11	8	8	15	21	
1998:	8	7	6	9	19	15	9	22	6	12	16	
> Mann	Kend	dall	(mto	2.1A	MThu	Sat.	ts)					
tau =	-0.3	327,	s	1 =0	.001	491 %						
> mto2	2.2AI	MThus	Sat.	ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	D
1992:	3	8	17	9	20	18	10	18	19	21	14	
1993:	6	7	7	9	11	15	13	20	19	18	10	
1994:	13	15	20	19	15	16	23	28	24	27	24	
1995:	19	12	12	18	11	16	19	20	18	20	17	
1996:	16	10	12	22	18	27	10	25	18	19	27	
1997:	13	16	24	б	21	23	10	29	17	14	13	
1998:	10	4	14	8	12	9	9	21	17	13	16	
					MThu							



MTO, Minor Injury, Same Scales



Time series trellis plot. Minor injury. Scales free.

1992 1994 1996 1998 2AM 2AM 8 20 10 15 20 25 15 9 ß 2 40 5 10 15 20 30 20 9 SunWe 5 10 15 20 25 10 15 20 25 ァ 2 11PM 11PM 15 20 25 30 30 15 20 25 9 10 1994 , 1996 1992 1998

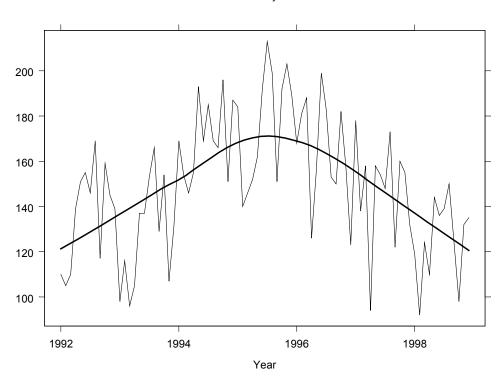
Year

MTO, Minor Injury, Scales Free

Minimal Injuries

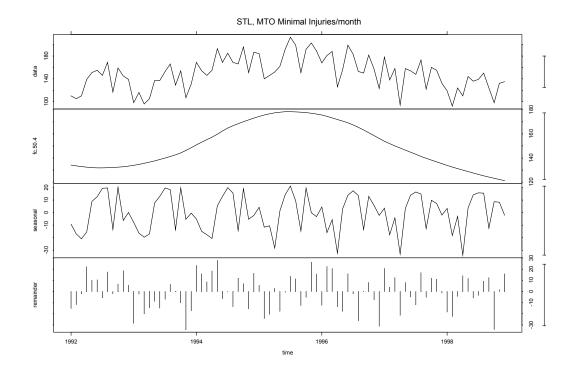
Monthly Overall Minimal Injuries

> mtol	l.ts											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	110	105	110	139	151	155	146	169	117	159	145	139
1993:	98	116	96	105	137	137	153	166	129	154	107	131
1994:	169	154	146	155	193	169	185	169	166	196	151	187
1995:	184	140	146	152	162	192	213	199	151	192	203	189
1996:	168	181	188	126	156	199	183	153	150	182	157	123
1997:	178	138	158	94	158	154	148	173	122	160	155	132
1998:	119	92	124	110	144	136	139	150	124	98	132	135



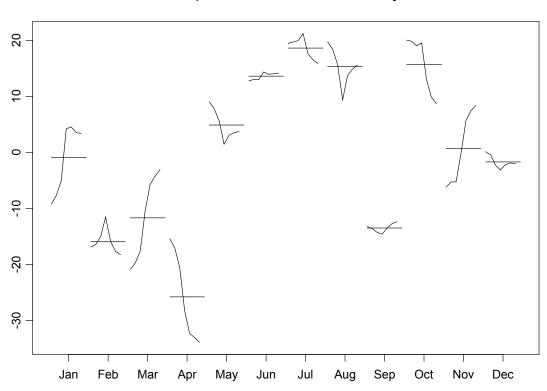
MTO minimal injuries/month

STL: Minimal Injuries



ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2 , lambda = 1

Seasonal Component: Minimal Injuries



Seasonal Component, MTO, Minimal Injuries/month

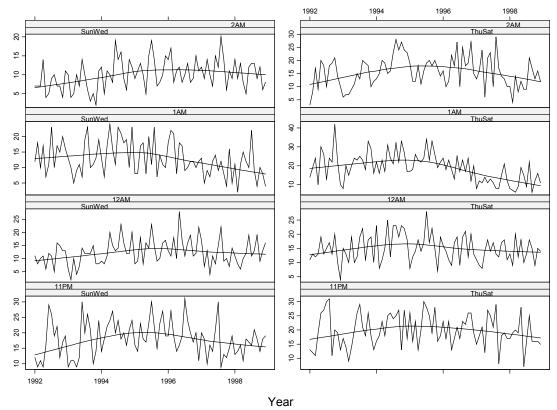
Weekgroup & Hour Time Series: Minimal Injuries

> mtol	.111	MSur	nWed.	.ts									
					May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	11	16	15	28	13	26	25	25	10	18	24	32	
1993:	13	12	12	18	26	16	18	27	18	16	13	27	
1994:	29	33	27	21	28	27	24	25	20	30	24	27	
1995:	19	33	30	23	31	25	41	47	23	26	46	25	
1996:	25	26	40	18	29	32	44	23	16	26	30	14	
1997:	50	16	24	15	18	26	31	30	19	17	17	28	
1998:	18	19	21	16	20	36	28	22	17	14	18	26	
> Mannl	Kend	lall	(mtol	1.111	MSu	nWed	.ts)						
tau =	0.08	394,	s	L = 23	3.79	010							
> mtol	.127	MSur	wed.	.ts									
ı	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	10	8	7	14	10	12	11	18	7	17	25	22	
1993:	14	11	24	15	9	20	17	18	14	12	8	21	
1994:	15	10	17	12	18	18	18	22	18	10	18	24	
1995:	13	19	15	21	20	19	23	23	20	14	29	21	
1996:	11	17	27	15	17	18	23	18	7	19	12	13	
1997:	23	11	15	11	8	11	23	17	12	34	21	11	
1998:	16	8	12	14	9	12	19	10	9	11	12	17	
> Mannl	Kend	lall	(mtol	L.122	AMSui	nWed	.ts)						
tau =	0.00)619	, s	sl =9	93.82	28							
> mtol	.1AM	ISun	ved.t	s									
ı	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	17	13	9	15	11	22	15	17	13	14	13	12	
1993:	9	8	9	12	8	15	20	12	16	17	4	14	
1994:	19	21	17	18	20	23	24	22	17	17	15	18	
1995:	15	14	21	14	20	25	17	22	14	6	17	20	
1996:	21	25	17	15	8	13	12	18	10	23	13	8	
1997:	19	6	20	11	8	15	11	13	10	10	11	13	
1998:	12	3	13	13	6	11	13	4	4	12	8	13	
> Mannl	Kend	lall	(mto]	L.1AM	ISun	Wed.t	s)						
tau =	-0.1	18,	sl	=1.8	326%								
> mtol	.2AN	ISunV	ved.t	s									
,	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	8	3	9	- 4	6	6	7	7	8	10	5	13	
1993:	7	6	5	10	12	11	11	9	13	5	11	9	
1994:	11	14	14	12	6	10	17	11	12	15	8	9	
1995:	11	7	10	12	11	15	13	15	6	11	15	12	
1996:	12	11	13	10	16	11	18	24	11	17	14	17	
1997:	9	14	9	11	14	21	16	18		15	11	12	
1998:	13	9	8	11	11	6	11	13	7	6	11	9	
> Manni													
tau =					L0329		,						
		- 1											

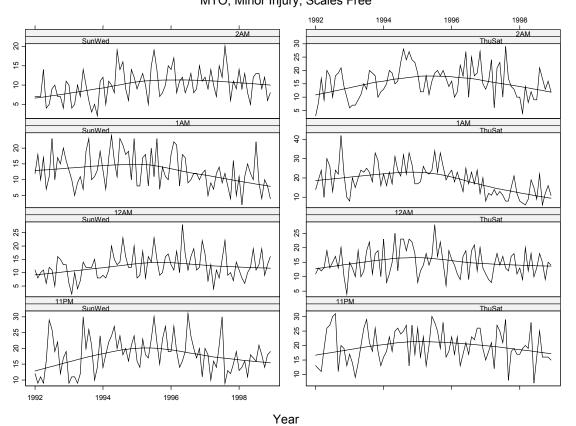
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	D
1992:	23	15	25	30	25	33	26	32	19	26	24	
1993:	17	23	13	15	25	18	33	26	17	32	20	
1994:	32	25	17	30	38	28	31	27	34	34	27	
1995:	37	26	21	24	25	34	40	21	26	25	32	
1996:	26	26	30	22	32	49	26	19	33	31	23	
1997:	36	31	27	13	34	30	29	32	18	25	23	
1998:	25	25	21	20	43	27	23	25	36	13	28	
> Mann	Kend	dall	(mto:	1.11	PMTh	uSat	.ts)					
tau =	0.0	64,	sl	=40	.07%							
> mtol	.122	AMThu	ıSat	.ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	D
1992:	7	13	10	12	26	22	17	24	7	21	16	
1993:	21	16	9	7	18	15	21	25	19	25	19	
1994:	23	14	13	20	30	18	20	22	21	28	14	
1995:	15	14	12	19	27	26	26	27	19	32	16	
1996:	25	15	16	9	21	22	15	14	18	23	23	
1997:	13	18	19	7	27	12	13	18	14	12	16	
1998:	11	8	15	13	20	23	16	25	22	16	23	
> Mann	Kend	dall	(mto:	1.12	AMTh	uSat	.ts)					
tau =	0.03	111,	s	1 =88	8.61	00						
> mtol	.1AI	MThu	Sat.	ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	D
1992:	17	27	22	22	35	15	24	32	32	34	20	
1993:	11	26	14	16	24	25	20	26	20	29	20	
1994:	21	23	25	23	33	27	26	24	25	41	25	
1995:	27	12	23	30	14	25	34	26	21	45	28	
1996:	33	37	21	24	16	26	28	20	30	25	19	
1997:	11	18	23	13	20	25	10	21	19	21	27	
1998:	12	8	24	13	14	12	16	23	10	15	10	
> Mann	Kend	dall	(mto:	1.1AM	MThu	Sat.	ts)					
tau =				1 =0	.281	8%						
> mtol												
				-	-			_	_		Nov	D
1992:	17	10	13	14		19	21	14	21	19		
1993:	6	14	10	12	15	17	13	23	12	18	12	
1994:	19	14	16	19	20	18	25	16	19	21	20	
1995:	47	15	14	9	14	23	19	18	22	33	20	
1996:	15	24	24	13	17	28	17	17	25	18	23	
1997:	17	24	21	13	29	14	15	24	22	26	29	
1998:	12	12	10	10	21	9	13	28	19	11	22	

Time series trellis plot. Minimal injury. Same scales.

MTO, Minor Injury, Scales Free



Time series trellis plot. Minimal injury. Scales free.



MTO, Minor Injury, Scales Free

D. MTO DRC Dataset: Crosstabs

Summary

MTODRC dataset is the MTO dataset with the added variable, driver, which has two values, sober or drunk. The database given has 26,026 records. Each record contains the variable injury which is the driver injury in one of four categories: minimal, minor, major, killed. Another variable, prov, indicates the license plate on the vehicle. The <u>frequency tabulation</u> indicates the number of accidents involving each value of the prov variable. After Ontario the next two are USA and "unknown" code = -1.

Driver+Injury: Drunk drivers tend not to be involved in minimal injury accidents and are more likely to be involved in major injury or fatal accidents. Overall, for about 30% of all accidents, the driver is drunk. Overall about 48%, 40%, 10% and 2% of accidents are respectively minimal injury, minor injury, major injury and fatal.

Driver+Wkgrp: As expected drunk drivers are relatively more common in the ThuSat weekgroup. Overall, about 43% and 57% of all accidents occur respectively in the SunWed and ThuSat weekgroups.

Driver + Hour: Drunk drivers are more prevalent at 1AM and 2AM. Overall about 32%, 22%, 25% and 21% of accidents in the respective hour windows beginning at 11PM, 12AM, 1AM and 2AM.

<u>Driver + Province</u>: USA and Quebec drivers involved in accidents tend to be relatively more sober. Drivers in the "unknown" vehicle province category tend to be relatively more drunk.

Frequencies: Province

After Ontario the next two are USA and "unknown" code = -1.

Vehicle Province or State

		Frequency	Percent
Valid	-1	230	.9
	Alberta	38	.1
	British	33	.1
	Columbia		
	Manitoba	38	.1
	New	18	.1
	Brunswick		
	Newfoundl	5	.0
	and		
	Nova	19	.1
	Scotia		
	Ontario	25072	96.3
	Prince	3	.0
	Edward		
	Island		
	Quebec	270	1.0
	Saskatche	13	.0
	wan		
	Yukon and	2	.0
	North		
	West		
	Territories		
	U.S.A.	284	1.1
	Other	1	.0
	Foreign		
	Total	26026	100.0

Driver + Injury

Drunk drivers tend not to be involved in minimal injury accidents and are more likely to be involved in major injury or fatal accidents.

```
Call:
crosstabs(formula = ~ driver + injury, data = mto.df, na.action =
na.exclude)
26026 cases in table
+----+
N
N/RowTotal
N/ColTotal
N/Total
+---+
driver |injury
 |minimal|minor |major |killed |RowTotl|
----+
sober |9628 |6912 |1284 | 197 |18021 |

      0.53
      0.38
      0.071
      0.011
      0.69

      0.77
      0.67
      0.51
      0.33
      |

      0.37
      0.27
      0.049
      0.0076
      |

----+
drunk 2948 3436 1228 393 8005
0.37 0.43 0.15 0.049 0.31
0.23 0.33 0.49 0.67
0.11 0.13 0.047 0.015
----+
ColTotl|12576 |10348 |2512 |590 |26026 |
 0.48 0.4 0.097 0.023
-----+
Test for independence of all factors
     Chi<sup>2</sup> = 1088.853 d.f.= 3 (p=0)
     Yates' correction not used
 Chi-sq decomposition: (obs-exp)/sqrt(exp)
    minimal minor major killed
sober 9.86 -2.99 -10.92 -10.47
drunk -14.79 4.49 16.38 15.70
```

Driver + Wkgrp

As expected drunk drivers are relatively more common in the ThuSat weekgroup.

```
Call:
crosstabs(formula = ~ driver + wkgrp, data = mto.df, na.action =
na.exclude)
26026 cases in table
+----+
N
N/RowTotal
N/ColTotal
N/Total
+----+
driver |wkgrp
|SunWed |ThuSat |RowTotl|
----+
sober80519970180210.450.550.69
     0.72 |0.67
     0.31 0.38
----+
drunk |3191 |4814 |8005
|0.4 |0.6 |0.31
|0.28 |0.33 |
|0.12 |0.18 |
----+
ColTotl | 11242 | 14784 | 26026 |
 0.43 0.57 |
----+
Test for independence of all factors
     Chi<sup>2</sup> = 52.3307 d.f.= 1 (p=4.689582e-013)
     Yates' correction not used
Chi-sq decomposition: (obs-exp)/sqrt(exp)
    SunWed ThuSat
sober 3.02 -2.64
drunk -4.54 3.96
```

Driver + Hour

Drunk drivers are more prevalent at 1AM and 2AM.

```
Call:
crosstabs(formula = ~ driver + hour, data = mto.df, na.action =
na.exclude)
26026 cases in table
+----+
N
N/RowTotal
N/ColTotal
N/Total
+----+
driver |hour
 |11PM |12AM |1AM |2AM |RowTotl|

      sober
      6581
      4192
      4069
      3179
      18021
      |

      0.37
      0.23
      0.23
      0.18
      0.69
      |

      0.8
      0.72
      0.62
      0.59
      |
      |

      0.25
      0.16
      0.16
      0.12
      |

drunk |1634 |1662 |2537 |2172 |8005 |
|0.2 |0.21 |0.32 |0.27 |0.31 |
|0.2 |0.28 |0.38 |0.41 |
|0.063 |0.064 |0.097 |0.083 |
----+
ColTot1 8215 | 5854 | 6606 | 5351 | 26026 |
 0.32 0.22 0.25 0.21
----+
Test for independence of all factors
      Chi<sup>2</sup> = 895.2281 d.f.= 3 (p=0)
      Yates' correction not used
Chi-sq decomposition: (obs-exp)/sqrt(exp)
       11PM 12AM 1AM 2AM
sober 11.84 2.18 -7.47 -8.64
drunk -17.76 -3.27 11.21 12.97
```

Driver + Province

USA and Quebec drivers involved in accidents tend to be relatively more sober. Drivers in the "unknown" vehicle province category tend to be relatively more drunk.

```
Call:
crosstabs(formula = ~ prov + driver, data = mto.df, na.action = na.exclude)
26026 cases in table
+---+
N
N/RowTotal
N/ColTotal
N/Total
+----+
prov |driver
|sober
          drunk |RowTotl|
_____
            +----
                   _+___+
Al | 24 | 14 | 38 |
|0.63 | 0.37 | 0.0015 |
----+
      ____
                  --+----
          | 9
  24
BC
                 33
         0.27
    0.73
                 0.0013
_____+
Man 31
            7
                  38
                  0.0015
    0.82
           0.18
8
  | 10
                   18
NB
           0.44
    0.56
                  6.9e-4
          ----+------
-----+-----
                  NWFLD | 4 | 1
|0.8 | 0.2
                   5
                   1.9e-4
NS
     | 10
           | 9 |19
    0.53
           0.47
                   7.3e-4
    +-----
            +----
____
Ont |173397733 |25072 |
    0.69 0.31 0.96
PEI | 1 | 2 | 3
|0.33 | 0.67 | 1.2e-4
    -+----
            -+-
Que | 212 | 58
|0.79 | 0.21
                 270
0.01
                   0.01
      _ _ _ _
            +----
                   _ + _ _ _ _
Sask | 11 | 2 | 13
|0.85 | 0.15 | 5.e-4
_____
             ____
                    +---
YNWT | 1 | 1
                  2
    0.5 0.5
                  7.7e-5
----+-----
            +----
                   -+----
          | 59
USA 225
                  284
    0.79 0.21
                  0.011
Other 0
         | 1
                  1
    0
                   3.8e-5
            1
  ----+-----
            -+----
Unknown | 129 | 101
|0.56 | 0.44
                   230
                   0.0088
ColTotal|18021 |8005
                  26026
  0.69
           0.31
----+---+----+----+----+----+---
```

```
Test for independence of all factors
      Chi<sup>2</sup> = 56.71319 d.f. = 13 (p=2.017336e-007)
       Yates' correction not used
      Some expected values are less than 5, don't trust stated p-value
 Chi-sq decomposition: (obs-exp)/sqrt(exp)
    sober drunk
Al -0.45 0.68
    BC 0.24 -0.36
    Man 0.91 -1.37
    NB -0.70 1.05
  NWFLD 0.29 -0.43
    NS -0.87 1.31
    Ont -0.16 0.24
   PEI -0.75 1.12
   Que 1.83 -2.75
   Sask 0.67 -1.00
   YNWT -0.33 0.49
  USA 2.02 -3.03
Other -0.83 1.25
Unknown -2.40 3.60
```

D. MTO DRC Dataset: Time Series

Summary

Fatalities: **Data** and **Trellis**

Upward trend in 2AMSunWedSober and **downward** in 1AMSatThuDrunk. Data are small counts, mostly 0 some 1's 2's 3's and 4's.

Major Injury: **Data** and **Trellis**

Downward trends: 12AMSunWed.drunk.ts, 1AMSunWed.drunk.ts, 12AMThuSat.drunk.ts, 1AMThuSat.drunk.ts.

Minor Injury: **Data** and **Trellis**

Upward trend: 12AMSunWed.sober.ts, 2AMSunWed.sober.ts Downward trend: 12AMThuSat.sober.ts, 1AMThuSat.sober.ts, 1AMSunWed.drunk.ts, 11PMThuSat.drunk.ts, 12AMThuSat.drunk.ts, 1AMThuSat.drunk.ts

Minimal Injury **<u>Data</u>** and <u>Trellis</u>

upward trend: 11PMSunWed.sober.ts, 2AMSunWed.sober.ts, 11PMThuSat.sober.ts, 2AMThuSat.sober.ts, 2AMSunWed.drunk.ts downward trend: 1AMThuSat.sober.ts, 11PMSunWed.drunk.ts, 12AMSunWed.drunk.ts, 1AMSunWed.drunk.ts, 11PMThuSat.drunk.ts, 12AMThuSat.drunk.ts, 1AMThuSat.drunk.ts

Fatalities

Upward trend in 2AMSunWedSober and downward in 1AMSatThuDrunk. Data are small counts, mostly 0 some 1's 2's 3's and 4's.

Data Listing

> mto	4.111	PMSur	nWed	.sobe	er.t	5							
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	0	0	1	0	0	0	0	0	0	
1993:	0	0	1	0	1	0	0	1	0	0	0	1	
1994:	0	0	0	1	0	0	0	1	1	1	0	0	
1995:	0	0	1	0	0	3	0	1	0	0	0	0	
1996:	0	0	0	0	2	1	0	1	0	0	0	0	
1997:	0	0	0	0	2	1	0	1	1	3	0	1	
1998:	0	0	0	1	0	0	0	1	2	0	0	0	
> Manı	nKend	dall	(mto4	1.111	PMSui	nWed	.sobe	er.ts	3)				
tau =	0.12	25,	sl	=15	.8%								
> mto4.12AMSunWed.sober.ts													
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	0	0	0	0	0	1	1	0	0	
1993:	0	0	0	0	1	0	2	0	0	0	0	0	
1994:	0	0	0	0	0	0	1	0	0	0	1	1	
1995:	0	0	0	0	0	0	0	0	1	0	0	0	
1996:	0	0	0	0	0	1	0	1	0	1	0	0	
1997:	0	0	1	0	0	2	2	1	1	0	0	1	
1998:	1	0	0	0	0	0	0	0	0	0	0	0	
> Manı	nKend	dall	(mto4	4.122	AMSui	nWed	.sobe	er.ts	5)				
tau =	0.08	825,	s	L = 35	5.629	00							
> mto	4.1AM	¶Sun≬	Ved.s	sobei	r.ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	0	0	0	1	0	0	0	0	0	
1993:	0	0	0	0	0	0	0	0	0	0	1	0	
1994:	0	1	0	0	1	1	0	0	0	0	0	0	
1995:	0	1	0	0	0	0	0	0	0	1	0	3	
1996:	0	1	0	0	0	0	1	1	0	1	0	0	
1997:	0	0	0	0	0	1	0	1	0	0	1	0	
1998:	0	0	0	0	0	0	0	0	0	0	0	0	
<pre>> MannKendall(mto4.1AMSunWed.sober.ts)</pre>													
tau =	0.03	35,	sl	=70	.03%								
> mto	4.2AM	¶Sun	Ved.s	sobei	r.ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	0	0	0	0	1	0	0	0	0	
1993:	0	0	0	0	0	0	0	0	0	0	0	0	
1994:	1	0	0	0	0	0	1	0	0	0	0	0	
1995:	0	1	0	0	2	0	0	0	1	0	1	0	
1996:	0	0	0	1	0	0	0	0	0	0	0	0	
1997:	0	0	1	0	0	0	0	0	0	0	0	1	
1998:	0	0	0	0	1	1	1	0	0	0	0	0	
> Manı	nKend	dall	(mto4	1.2AM	MSun	Wed.s	sobei	r.ts)				
tau =				=21									

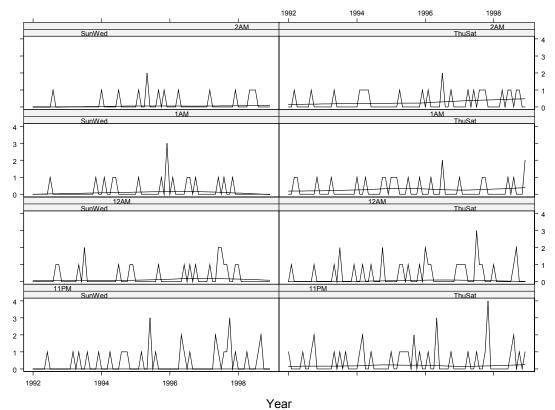
> #												
mto4.1												
	Jan											
1992:	1	0	0	0	0	1	0	0	1	2	0	0
1993:	0	0	0	0	1	0	1	0	1	0	0	0
1994:	0	1	2	0	0	1	0	0	0	0	0	0
1995: 1996:	1 0	0 1	0 0	1 0	1 3	1 0	1 0	0 0	2 0	0 1	1 0	0 0
1996: 1997:	0	1 0	0	1	3 0	0	1	0	0	1 1	4	0
1997:	0	0	0	1	0	0	1	2	0	1 1	4	1
	-	-	-		-	-	_		-	T	0	Т
<pre>> MannKendall(mto4.11PMThuSat.sober.ts) tau = 0.0348, sl =69.44%</pre>												
> mto4		,										
	Jan							Aug	Sen	Oct	Nov	Dec
1992:	0 211 .	1	0	0	0	0 0 0	0 0 1	Aug 0	0 0	1	0	0
1993:	0	0	0	0	1	0	2	0	0	0	0	1
1994:	0	0	1	0	0	1	0	0	0	2	0	0
1995:	0	0	0	1	1	0	1	0	0	0	1	0
1996:	2	1	1	0	0	0	0	0	0	0	0	1
1997:	1	1	1	0	0	0	3	1	1	0	0	0
1998:	1	0	0	0	0	0	0	1	2	0	0	0
> Mann	Kend	all(mto4	4.12/	AMThu	ıSat	.sobe	er.ts	3)			
tau =	0.09	5,	sl	=28	.25%							
> mto4	.1AM	ThuS	Sat.s	sobei	r.ts							
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	0	0	1	1	0	0	0	0	0	0	1	0
1993:	0	0	0	1	0	0	0	0	0	0	0	0
1994:	1	0	0	1	0	0	0	0	0	1	1	0
1995:	1	1	1	0	0	0	1	0	0	0	1	0
1996:	0	0	1	0	0	0	2	0	0	0	0	0
1997:	1	0	0	0	0	0	0	0	1	1	0	0
1998:	0	0	0	0	0	1	0	1	0	0	0	2
	> MannKendall(mto4.1AMThuSat.sober.ts)											
tau =					5.53	0						
> mto4						_		_	~	. ·		_
	Jan			-	-			-	-			
1992:	0	0	1	0	0	0	0	0	1	0	0	0
1993:	0	0	0	0	1	0	0	0	0	0	0	0
1994:	0	1 0	1	1 1	1 0	0	0	0	0 0	0	0	0
1995: 1996:	0	0 1	0	1	-	0	0 2	0	-	0	0	1 0
1996: 1997:	0 0	1 0	0 0	0	0 0	0 1	∠ 0	0 1	0 1	1 1	0	0
1997: 1998:	0	0	0	1	0	1 1	1	1	1	1 1	0	0
> Mann										Т	U	U
tau =				=5.		ac.	-onei		,			
cau -	···/	÷,	91		0.00.0							

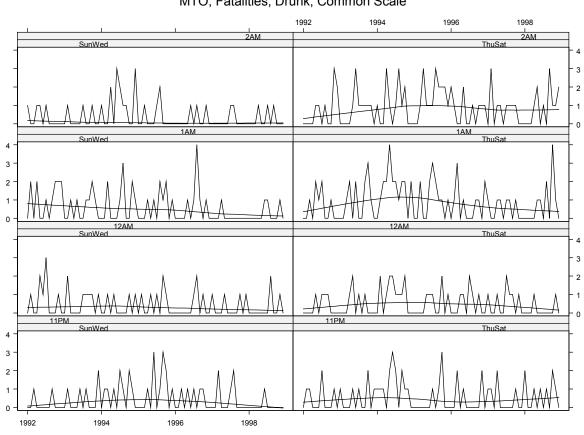
> #													
mto4.1	1PMS	SunWe	ed.dr	runk	.ts								
ı	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	1	0	0	0	0	0	1	0	0	0	
1993:	0	1	0	0	0	1	0	1	0	0	0	2	
1994:	0	1	1	0	1	0	2	1	0	2	1	0	
1995:	0	0	0	1	0	3	0	1	3	2	0	1	
1996:	0	0	1	0	1	0	1	0	1	1	0	0	
1997:	0	0	2	0	0	0	1	2	0	0	0	0	
1998:	0	0	0	0	0	1	0	0	0	0	0	0	
> Mannl	Kend	dall((mto4	1.111	MSur	nWed	.dru	nk.ts	5)				
tau = ·	-0.0)479,	, :	sl =5	58.54	18							
> mto4	.127	AMSur	nWed.	dru	nk.ts	5							
,	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	1	0	0	2	1	3	0	0	0	1	0	
1993:	0	2	0	0	0	0	1	1	1	1	0	1	
1994:	0	0	1	0	1	0	0	0	0	1	0	1	
1995:	0	1	0	0	1	0	1	0	2	1	0	0	
1996:	0	0	0	0	0	0	1	2	0	1	0	0	
1997:	1	0	0	0	1	0	0	0	0	1	0	1	
1998:	0	0	0	0	0	0	0	2	0	0	1	0	
> Mannl	Kend	dall((mto4	1.12/	AMSur	nWed	.dru	nk.ts	5)				
tau =	-0.1	24,	s	=15	5.76	0							
> mto4	.1AM	ISun∛	ved.c	lrunł	.ts								
ı	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	2	0	2	0	0	1	0	1	2	2	2	
1993:	0	0	1	0	1	0	0	1	1	2	1	0	
1994:	0	0	2	0	0	0	1	3	0	0	2	1	
1995:	0	0	0	1	0	1	0	2	1	2	0	1	
1996:	0	0	0	0	1	0	1	4	1	0	1	0	
1997:	0	0	0	1	0	0	0	0	0	0	0	0	
1998:	0	0	0	0	0	1	1	0	0	0	1	0	
> Manni	Kend	lall((mto4	1.1A	ISun	ved.	irunl	t.ts)				
tau =	-0.2	203,	s	L =1.	.8729	6							
> mto4	.2AN	ISun≀	ved.d	lrunł	.ts								
ı	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	1	0	0	1	1	0	1	0	0	0	0	0	
1993:	0	1	0	0	0	0	1	0	0	1	0	0	
1994:	1	0	0	2	0	3	2	1	1	0	0	3	
1995:	0	0	1	0	0	0	1	2	0	0	0	0	
1996:	0	0	0	0	0	1	0	1	0	0	1	0	
1997:	0	0	0	0	0	0	1	1	0	0	0	0	
1998:	0	0	0	1	0	0	1	0	1	0	0	0	
> Manni	Kend					ved.o	drun	.ts)				
tau =									,				

tau = -0.0967, sl =27.39%

Trellis Time Series Plots. Same Scale on Each Page.

MTO, Fatalities, Sober, Common Scale





MTO, Fatalities, Drunk, Common Scale

Year

Major Injury

Downward trends: 12AMSunWed.drunk.ts, 1AMSunWed.drunk.ts, 12AMThuSat.drunk.ts, 1AMThuSat.drunk.ts

Data Listing

> mto3	3.111	PMSui	nWed	.sob	er.t:	5						
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	2	1	0	1	0	4	9	1	1	1	4	2
1993:	3	1	0	0	1	2	3	1	3	2	2	3
1994:	0	2	4	5	2	5	б	5	3	1	1	2
1995:	2	1	1	2	4	9	2	3	1	1	4	0
1996:	2	5	3	0	3	1	2	3	7	1	1	2
1997:	5	0	2	4	2	0	5	3	0	2	2	3
1998:	1	0	1	2	3	6	5	3	1	2	4	0
> Manr	nKend	dall	(mto:	3.11	PMSui	nWed	.sobe	er.ts	3)			
tau =	0.00	59,	sl	=38	.87%							
> mto3			nWed	.sob	er.t	5						
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	1	2	0	2	3	0	3	2	2	4	1	1
1993:	0	0	0	3	3	1	2	3	1	1	0	1
1994:	4	1	3	0	2	3	2	0	2	3	3	3
1995:	3	1	1	1	3	1	4	5	2	0	2	5
1996:	0	0	3	0	1	1	3	1	2	0	0	3
1997:	1	1	2	2	0	1	3	1	3	2	0	1
1998:	0	1	1	1	2	0	2	1	2	1	0	1
> Manr	nKend	dall	(mt.o.	3.12	AMSui	nWed	.sobe	er.ts	3)			
tau =				=22.					- /			
> mto3		,										
						Jun	Jul	Auq	Sep	Oct	Nov	Dec
1992:	2	0	1	1	1	2	2	3	3	0	1	2
1993:	1	1	0	1	0	0	3	1	1	0	3	0
1994:	0	1	1	2	6	2	2	6	2	1	2	1
1995:	2	4	4	4	2	4	3	2	1	1	1	2
1996:	0	1	0	0	0	0	0	3	2	5	1	0
1997:	2	0	0	0	1	1	7	0	1	0	4	2
1998:	1	0	0	0	0	1	1	3	1	2	2	1
> Manr	_	-	-		-	_	_	-				-
tau =					39.4		00000		/			
> mto3			,									
						מווה	Jul	Αμσ	Sen		Nov	Dec
1992:	2	2	0	дрі 1	0	1	1	Aug 3	0	1	0	0
1993:	2	0	0	2	0	0	0	2	0	3	1	0
1994:	0	2	1	1	0	1	3	1	2	0	2	2
1995:	5	1	1	0	1	0	3	5	2	0	1	2 0
1995:	3	2	1	2	0	0	2	3	1	1	1	0
1996:	5 1	2 1	1	2 0	0	2	2 0	2	1	1	1 0	0
	1	1	1	0	1	2 1	4	2 1	1	⊥ 2	1	1
1998:	_	-	-		_	_	-	-	-	2	T	Т
> Manr			•		97.4		sopei	L.LS	/			
tau =	-0.0	10204	±,	ST	-91.1	τõ						

> #													
mto3.2	11PMS	SunWe	ed.dr	runk	.ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	0	2	2	4	0	1	0	0	1	1	
1993:	0	1	0	1	0	4	3	2	3	3	0	0	
1994:	2	4	1	2	0	2	4	2	2	3	2	2	
1995:	1	2	1	1	3	0	0	3	0	1	2	1	
1996:	1	1	0	1	0	2	1	3	1	0	0	1	
1997:	1	0	0	1	0	0	1	1	1	1	1	0	
1998:	1	1	2	1	0	0	2	1	3	1	1	1	
> Manı	nKend	dall	(mto3	3.111	MSur	nWed	.drur	nk.ts	3)				
tau =	-0.0)62,	s	=45	5.63	0							
> mto:	3.124	AMSur	wed.	dru	nk.ts	3							
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:	0	0	2	1	2	1	2	4	2	3	1	1	
1993:	1	1	0	1	0	0	2	2	1	1	1	1	
1994:	1	1	0	2	2	2	0	1	2	3	2	1	
1995:	0	1	2	3	2	5	2	1	1	0	1	0	
1996:	0	0	1	1	1	1	0	4	1	0	0	1	
1997:	1	1	0	2	0	1	0	0	1	0	1	0	
1998:	1	2	1	2	2	1	1	0	0	1	1	1	
> Manı	nKend	lall	(mto3	3.12	AMSui	wed	.dru	nk.ts	3)				
tau =					.0249								
Lau -	-0.1	L H / /	51	L = O .	. 0241	б							
						б							
> mto:	3.1AN	/SunV	Ved.c	lrunł	.ts		Jul	Auq	Sep	Oct	Nov	Dec	
	3.1AN		Ved.c	lrunł	.ts		Jul 3	Aug 0	Sep 1	Oct 0	Nov 2	Dec 1	
> mto:	3.1AN Jan	ISun∛ Feb	Ved.d Mar	lrun) Apr	k.ts May	Jun							
> mto: 1992:	3.1AN Jan 0	ISunV Feb 3	Ved.c Mar 2	lrun) Apr 2	k.ts May 3	Jun 2	3	0	1	0	2	1	
> mto: 1992: 1993:	3.1AN Jan 0 0	ISunV Feb 3 5	Ved.d Mar 2 1	lrun Apr 2 2	c.ts May 3 0	Jun 2 2	3 2	0 4	1 0	0 2	2 2	1 3	
> mto: 1992: 1993: 1994:	3.1AN Jan 0 0 0	ISuni Feb 3 5 0	Ved.d Mar 2 1 4	lrun Apr 2 2 4 3	K.ts May 3 0 2	Jun 2 2 5	3 2 6	0 4 3	1 0 4	0 2 2	2 2 6	1 3 2	
> mto: 1992: 1993: 1994: 1995:	3.1AN Jan 0 0 0 3	1Suni Feb 3 5 0 3	Ved.c Mar 2 1 4 5	lrun Apr 2 2 4	K.ts May 3 0 2 3	Jun 2 2 5 5	3 2 6 5	0 4 3 4	1 0 4 5	0 2 2 5	2 2 6 5	1 3 2 1	
> mto: 1992: 1993: 1994: 1995: 1996:	3.1AN Jan 0 0 0 3 0	ISuni Feb 3 5 0 3 2	Ved.d Mar 2 1 4 5 3	lrun Apr 2 2 4 3 3	x.ts May 3 0 2 3 1	Jun 2 5 5 1	3 2 6 5 0	0 4 3 4 2	1 0 4 5 0	0 2 2 5 1	2 2 6 5 0	1 3 2 1 0	
<pre>> mto3 1992: 1993: 1994: 1995: 1996: 1997: 1998:</pre>	3.1AN Jan 0 0 3 0 2 1	4Sun Feb 3 5 0 3 2 1 2	Ved.c Mar 2 1 4 5 3 3 2	lrun) Apr 2 4 3 3 6 0	C.ts May 3 0 2 3 1 1 1	Jun 2 5 5 1 3 1	3 2 6 5 0 1 0	0 4 3 4 2 3 1	1 0 4 5 0 1	0 2 5 1 0	2 2 5 0 1	1 3 2 1 0 0	
<pre>> mto3 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manual</pre>	3.1AN Jan 0 0 3 0 2 1 Kenc	4Sun Feb 3 5 0 3 2 1 2 1 2	Ved.c Mar 2 1 4 5 3 3 2 (mto 3	lrun) Apr 2 4 3 3 6 0 8.1A	(.ts May 3 0 2 3 1 1 1 1	Jun 2 5 5 1 3 1 Ved.	3 2 6 5 0 1 0	0 4 3 4 2 3 1	1 0 4 5 0 1	0 2 5 1 0	2 2 5 0 1	1 3 2 1 0 0	
<pre>> mto3 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau =</pre>	3.1AN Jan 0 0 0 3 0 2 1 Kenc -0.1	4Sun Feb 3 5 0 3 2 1 2 1 2 4a11 (Ved.c Mar 2 1 4 5 3 3 2 (mto 3 2 (mto 3	lrun) Apr 2 4 3 6 0 8.1A1 L =7	(.ts May 0 2 3 1 1 1 (Sun)	Jun 2 5 5 1 3 1 Ved.	3 2 6 5 0 1 0	0 4 3 4 2 3 1	1 0 4 5 0 1	0 2 5 1 0	2 2 5 0 1	1 3 2 1 0 0	
<pre>> mto3 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manual</pre>	3.1AN Jan 0 0 3 0 2 1 NKenc 3.2AN	4Sun Feb 3 5 0 3 2 1 2 1 2 1 2 1 2 1 2 4 1 1 2 4 1	Ved.c Mar 2 1 4 5 3 2 (mto3 s] Ved.c	lrun) Apr 2 4 3 3 6 0 3.1AI 1 = 7 1run)	<pre>x.ts May 3 0 2 3 1 1 1 </pre> <pre>Solution </pre>	Jun 2 5 5 1 3 1 Ved.	3 2 5 0 1 0 1 0	0 4 3 4 2 3 1 c.ts	1 0 4 5 0 1 1	0 2 5 1 0 2	2 6 5 0 1 0	1 3 2 1 0 0 0	
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau = > mto:</pre>	3.1AN Jan 0 0 3 0 2 1 Kenc 3.2AN Jan	4SunW Feb 3 5 0 3 2 1 2 1 2 1 2 1 2 4SunW Feb	Ved.c Mar 2 1 4 5 3 2 (mto3 s] Ved.c Mar	lrun) Apr 2 4 3 6 0 3.1A 5 1 1 1 1 4 7	<pre>c.ts May 3 0 2 3 1 1 1 573 c.ts May</pre>	Jun 2 5 5 1 3 1 Ved. 8	3 2 6 5 0 1 0 irun Jul	0 4 3 4 2 3 1 c.ts Aug	1 0 4 5 0 1 1	0 2 5 1 0 2 Oct	2 6 5 0 1 0 Nov	1 3 2 1 0 0 0	
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau = > mto: 1992:</pre>	3.1AN Jan 0 0 3 0 2 1 NKenc -0.1 3.2AN Jan 2	4SunW Feb 3 5 0 3 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 5 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 5 0 7 7 5 0 7 7 1 7 5 0 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7	Ved.c Mar 2 1 4 5 3 2 (mto 3 2 (mto 3 8] Ved.c Mar 2	lrun Apr 2 4 3 6 0 3.1 1 1 1 1 4 7 1 1 1 1 1 1 0	May May 0 2 3 1 1 1 573 5 c.ts May 2	Jun 2 5 5 1 3 1 Ved. 6 Jun 2	3 2 6 5 0 1 0 1 0 1 0 1 1 2	0 4 3 4 2 3 1 c.ts Aug 3	1 0 4 5 0 1 1 Sep 0	0 2 5 1 0 2 Oct 4	2 6 5 0 1 0 Nov	1 3 2 1 0 0 0 0 0 Dec 0	
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau = > mto: 1992: 1993:</pre>	3.1AN Jan 0 0 3 0 2 1 Kenc 3.2AN Jan 2 1	4SunW Feb 3 5 0 3 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2	Ved.c Mar 2 1 4 5 3 2 (mto3 8 (mto3 8 Ved.c Mar 2 2	<pre>lrunl Apr 2 2 4 3 3 6 0 3.1A</pre>	<pre>x.ts May 3 0 2 3 1 1 1 xsunt 573 x.ts May 2 1</pre>	Jun 2 5 5 1 3 1 Ved. 8 Jun 2 0	3 2 6 5 0 1 0 1 1 0 1 1 0 1 1 2 3	0 4 3 4 2 3 1 c.ts 3 2	1 0 4 5 0 1 1 Sep 0 0	0 2 5 1 0 2 Oct 4 1	2 6 5 0 1 0 8 Nov 1 2	1 3 2 1 0 0 0 0 0 Dec 0 4	
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau = > mto: 1992: 1992: 1993: 1994:</pre>	3.1AN Jan 0 0 3 2 1 Kenc 3.2AN Jan 2 1 1	4SunW Feb 3 5 0 3 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 1 2 5 5 0 7 5 0 7 7 7 7 7 7 9 7 9 7 9 7 9 7 9 7 9 7 9	Ved. c Mar 2 1 4 5 3 2 (mto3 8] Ved. c Mar 2 2 2 2	<pre>Irun! Apr 2 2 4 3 3 6 0 3.1A I = 7 Irun! Apr 0 2 2 2</pre>	c.ts May 0 2 3 1 1 1 573 c.ts May 2 1 3	Jun 2 5 5 1 3 1 Ved. 6 Jun 2	3 2 6 5 0 1 0 drun 3 Jul 2 3 1	0 4 3 4 2 3 1 c.ts 3 2 0	1 0 4 5 0 1 1 1 Sep 0 0 4	0 2 5 1 0 2 0 ct 4 1 2	2 6 5 0 1 0 Nov	1 3 2 1 0 0 0 0 0 2 0 4 0	
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Mann tau = > mto: 1992: 1992: 1993: 1994: 1995:</pre>	3.1AN Jan 0 0 3 0 2 1 Kenc -0.1 3.2AN Jan 2 1 1 2	4SunV Feb 3 5 0 3 2 1 2 1 2 1 2 4 2 1 2 4 2 1 5 0 3 2 0 0	Ved.c Mar 1 4 5 3 2 (mto3 8] Ved.c Mar 2 2 1	<pre>Irunl Apr 2 2 4 3 3 6 0 3.1A I =7 Irunl Apr 0 2 3 3</pre>	c.ts May 0 2 3 1 1 1 573 c.ts May 2 1 3 3	Jun 2 5 5 1 3 1 Ved. 8 Jun 2 0 1	3 2 6 5 0 1 0 drun Jul 2 3 1 3	0 4 3 4 2 3 1 c.ts 3 2 0 1	1 0 4 5 0 1 1 1 Sep 0 0 4 1	0 2 5 1 0 2 0 ct 4 1 2 3	2 6 5 0 1 0 Nov 1 2 1 1	1 3 2 1 0 0 0 0 0 2 0 4 0 0	
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau = > mto: 1992: 1992: 1993: 1994: 1995: 1996:</pre>	3.1AN Jan 0 0 3 0 2 1 Kenc -0.1 3.2AN Jan 2 1 1 2 2	ISUNV Feb 3 5 0 3 2 1 2 1 2 1 2 1 2 1 2 1 1 5 5 0 3 2 0 1	Ved.o Mar 2 1 4 5 3 2 (mto3 (mto3 8) Ved.o Mar 2 2 1 4	<pre>Irunl Apr 2 2 4 3 6 0 3.1An 1 =7 Irunl Apr 0 2 3 0 3 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0</pre>	c.ts May 0 2 3 1 1 573 c.ts May 2 1 3 3 2	Jun 2 5 5 1 3 1 Ved. 6 Jun 2 0 1 1 1	3 2 6 5 0 1 0 1 1 2 3 1 3 1 3	0 4 3 4 2 3 1 c.ts 2 0 1 3	1 0 4 5 0 1 1 1 Sep 0 0 4 1 3	0 2 5 1 0 2 0 ct 4 1 2 3 1	2 6 5 0 1 0 Nov 1 2 1 1 5	1 3 2 1 0 0 0 0 2 2 0 4 0 0 0 0	
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Mann tau = > mto: 1992: 1992: 1993: 1994: 1995: 1996: 1997:</pre>	3.1AN Jan 0 0 3 0 2 1 Kenc -0.1 3.2AN Jan 2 1 1 2 2 2	4SunV Feb 3 5 0 3 2 1 2 4 1 2 4 2 4 2 4 5 1 0 3 2 0 1 0	Ved.o Mar 2 1 4 5 3 2 (mto3 (mto3 8) Ved.o Mar 2 2 1 4 2	<pre>drund Apr 2 2 4 3 3 6 0 0 3.1ar L =7 drund Apr 0 2 3 0 1</pre>	c.ts May 3 0 2 3 1 1 573 c.ts May 2 1 3 3 2 2	Jun 2 2 5 5 1 3 1 3 1 Ved. 6 8 7 0 1 1 1 0	3 2 6 5 0 1 0 drun 2 3 1 3 1 1	0 4 3 4 2 3 1 c.ts 3 2 0 1 3 2	1 0 4 5 0 1 1 1 Sep 0 0 4 1 3 3	0 2 5 1 0 2 0 ct 4 1 2 3 1 1	2 6 5 0 1 0 Nov 1 2 1 1 5 0	1 3 2 1 0 0 0 0 4 0 0 0 1	
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau = > mto: 1992: 1992: 1993: 1994: 1995: 1996:</pre>	3.1AN Jan 0 0 3 0 2 1 Kenc -0.1 3.2AN Jan 2 1 1 2 2 2 4	ISUNV Feb 3 5 0 3 2 1 2 1 2 1 2 1 2 1 4 2 1 5 0 3 2 0 1 0 1	Ved.o Mar 2 1 4 5 3 2 (mto3 (mto3 8) Ved.o Mar 2 2 1 4 2 0	<pre>drund Apr 2 2 4 3 3 6 0 0 3.1ar L =7 drund Apr 0 2 3 0 1 1</pre>	c.ts May 0 2 3 1 1 573 c.ts May 2 1 3 3 2 2 0	Jun 2 2 5 5 1 3 1 3 1 Ved. 6 Ved. 7 0 1 1 1 0 2	3 2 6 5 0 1 0 1 1 2 3 1 3 1 3 1 1	0 4 3 4 2 3 1 C.ts 3 2 0 1 3 2 2 2	1 0 4 5 0 1 1 1 8 9 0 4 1 3 3 1	0 2 5 1 0 2 0 ct 4 1 2 3 1	2 6 5 0 1 0 Nov 1 2 1 1 5	1 3 2 1 0 0 0 0 2 2 0 4 0 0 0 0	

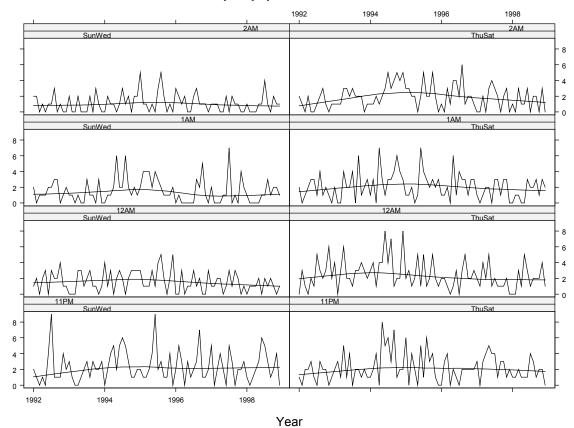
tau = -0.0241, sl = 77.07%

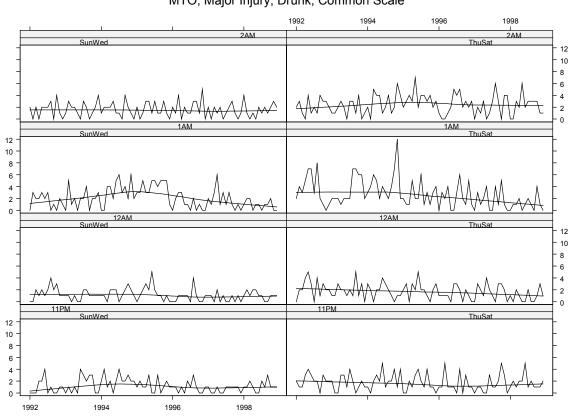
> #												
mto3.1	11 DM1	Thuse	at dr	nink	tq							
		Feb				Jun	.T11]	Διια	Sep	Oct	Nov	Dec
1992:	2	1	1	3	4	3	2	2	٩ <u></u> ٥٥	3	2	2
1993:	2	0	0	3	3	1	4	2	0	1	2	1
1994:	1	0	2	3	2	5	1	2	2	4	1	4
1995:	0	0	2	1	3	4	2	1	4	1	1	0
1996:	1	1	5	1	1	1	0	4	0	3	1	1
1997:	0	1	1	0	0	3	1	1	2	5	3	0
1998:	1	2	1	2	2	5	1	2	1	1	1	2
> Manı	nKend	dall((mto3	3.111	PMThu	ıSat	.dru	nk.ts	3)			
tau =					28.45				- /			
> mto:												
		Feb					Jul	Auq	Sep	Oct	Nov	Dec
1992:	0	2	2	4	5	3	0	4	1	3	2	2
1993:	1	1	3	2	2	1	2	1	5	1	3	0
1994:	3	2	0	2	2	4	3	2	1	0	1	1
1995:	2	3	0	3	2	5	2	2	1	0	2	3
1996:	1	1	1	1	0	3	3	2	0	2	1	0
1997:	0	3	2	1	4	2	1	0	3	3	2	0
1998:	1	0	2	1	0	0	2	0	0	1	3	1
> Manı	nKend	dall(mto3	.12	AMThu	Sat.	dru	nk.ts	3)			
tau =	-0.1	L97,			. 5559							
tau = > mto:			sl	=1	.555%							
	3.1AN		sl Sat.c	=1	. <mark>555</mark> % k.ts	6				Oct	Nov	Dec
	3.1AN	4ThuS	sl Sat.c	=1	. <mark>555</mark> % k.ts	6				Oct 1	Nov 0	Dec 1
> mto	3.1AN Jan	MThus Feb	sl Sat.c Mar	=1 lrunł Apr	.555 9 K.ts May	Jun	Jul	Aug	Sep			
> mto: 1992:	3.1AN Jan 2	AThus Feb 4	sat.c Mar 3	=1 lrun Apr 5	.555 c.ts May 7	Jun 7	Jul 3	Aug 8	Sep 2	1	0	1
> mto: 1992: 1993:	3.1AN Jan 2 2	IThus Feb 4 2	sat.c Mar 3 2	lrun Apr 5 1	5559 c.ts May 7 2	Jun 7 2	Jul 3 2	Aug 8 7	Sep 2 7	1 6	0 6	1 2
> mto: 1992: 1993: 1994:	3.1AM Jan 2 2 3	AThus Feb 4 2 4	Sat.c Mar 3 2 6	= 1 lrun Apr 5 1 5	5555 c.ts May 7 2 2	Jun 7 2 4	Jul 3 2 3	Aug 8 7 2	Sep 2 7 4	1 6 7	0 6 12	1 2 2
> mto: 1992: 1993: 1994: 1995:	3.1AN Jan 2 2 3 2	AThuS Feb 4 2 4 1	Sat.c Mar 3 2 6 1	= 1 lrun Apr 5 1 5 5	5559 x.ts May 7 2 2 2	Jun 7 2 4 2	Jul 3 2 3 6	Aug 8 7 2 1	Sep 2 7 4 3	1 6 7 1	0 6 12 3	1 2 2 4
> mto: 1992: 1993: 1994: 1995: 1996:	3.1AN Jan 2 2 3 2 0	4ThuS Feb 4 2 4 1 3	sat .c Mar 3 2 6 1 2	=1 Apr 5 1 5 5 4	5559 May 7 2 2 2 0	Jun 7 2 4 2 0	Jul 3 2 3 6 3	Aug 8 7 2 1 6	Sep 2 7 4 3 2	1 6 7 1	0 6 12 3 5	1 2 2 4 1
> mto: 1992: 1993: 1994: 1995: 1996: 1997:	3.1AN Jan 2 2 3 2 0 0 0 0	4Thus Feb 4 2 4 1 3 3 1	s at.c Mar 3 2 6 1 2 1 1	=1 Apr 5 1 5 5 4 2 2	5555 May 7 2 2 2 2 0 4 0	Jun 7 2 4 2 0 0	Jul 3 2 3 6 3 0 2	Aug 8 7 2 1 6 4 1	Sep 2 7 4 3 2 1 0	1 6 7 1 5	0 6 12 3 5 1	1 2 4 1 0
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998:</pre>	3.1AN Jan 2 2 3 2 0 0 0 0	4 Feb 4 2 4 1 3 3 1 1 1	s at.c Mar 3 2 6 1 2 1 1 (mto 3	=1 Apr 5 1 5 4 2 2 3.1A	5555 May 7 2 2 2 2 0 4 0	Jun 7 2 4 2 0 0 1 Sat.	Jul 3 2 3 6 3 0 2	Aug 8 7 2 1 6 4 1	Sep 2 7 4 3 2 1 0	1 6 7 1 5	0 6 12 3 5 1	1 2 4 1 0
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu</pre>	3.1AN Jan 2 2 3 2 0 0 0 0 0 0 0 0 0	4Thus Feb 4 2 4 1 3 3 1 1 1 1 1 1 1 1 1 1 	sl Sat.c Mar 3 2 6 1 2 1 1 (mto3 sl	=1 Apr 5 1 5 4 2 2 3.1A1	5555 May 7 2 2 2 0 4 0 4 0 140	Jun 7 2 4 2 0 0 1 Sat.	Jul 3 2 3 6 3 0 2	Aug 8 7 2 1 6 4 1	Sep 2 7 4 3 2 1 0	1 6 7 1 5	0 6 12 3 5 1	1 2 4 1 0
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau =</pre>	3.1AN Jan 2 2 3 2 0 0 0 0 0 0 0 0 0 3.2AN	4Thus Feb 4 2 4 1 3 3 1 1 1 1 1 1 1 1 1 1 	s Sat.c Mar 3 2 6 1 2 1 1 (mto3 s Sat.c	=1. Apr 5 1 5 4 2 2 3.1A1 5 4 2 2 1 4 2 2 8.1A1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	555 May 7 2 2 2 0 4 0 4 0 4 0 5.ts	Jun 7 2 4 2 0 1 5at.0	Jul 3 3 6 3 0 2 1run	Aug 8 7 2 1 6 4 1 c.ts	Sep 2 7 4 3 2 1 0	1 6 7 1 5 4	0 6 12 3 5 1 1	1 2 4 1 0 0
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau =</pre>	3.1AN Jan 2 2 3 2 0 0 0 0 0 0 0 8 .2AN	4Thus Feb 4 2 4 1 3 1 1 1 1 1 1 1 1 1 1 	s Sat.c Mar 3 2 6 1 2 1 1 (mto3 s Sat.c	=1. Apr 5 1 5 4 2 2 3.1A1 5 4 2 2 8.1A1 5 4 2 2 8.1A1 5 5 4 2 2 8.1A1	555 May 7 2 2 2 0 4 0 4 0 4 0 5.ts	Jun 7 2 4 2 0 1 5at.0	Jul 3 3 6 3 0 2 1run	Aug 8 7 2 1 6 4 1 c.ts	Sep 2 7 4 3 2 1 0	1 6 7 1 5 4	0 6 12 3 5 1 1	1 2 4 1 0 0
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Manu tau = > mto:</pre>	3.1AN Jan 2 2 3 2 0 0 0 0 0 NKenc 3.2AN Jan	4Thus Feb 4 2 4 1 3 1 1 1 1 1 1 1 1 1 1 	s at.c Mar 3 2 6 1 2 1 1 (mto3 s at.c Mar	=1 Apr 5 1 5 4 2 2 3.1AI 6 -=0 Apr	5555 May 7 2 2 2 2 0 4 0 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5	Jun 7 2 4 2 0 1 5at.0)1% Jun	Jul 3 2 3 6 3 0 2 irun Jul	Aug 8 7 2 1 6 4 1 c.ts Aug	Sep 2 7 4 3 2 1 0 Sep	1 6 7 1 1 5 4 Oct	0 6 12 3 5 1 1 1 Nov	1 2 4 1 0 0
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Mann tau = > mto: 1992:</pre>	3.1AN Jan 2 2 3 2 0 0 0 0 nKenc 3.2AN Jan 2	4Thus Feb 4 2 4 1 3 1 1 1 1 1 1 1 1 1 1 	s Mar 3 2 6 1 2 1 1 (mto3 Sat.c Mar 1	=1. Apr 5 1 5 4 2 2 3.1A1 Apr 0	5555 May 7 2 2 2 0 4 0 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5	<pre>Jun 7 2 4 2 0 0 1 5at.0)1% Jun 1</pre>	Jul 3 2 3 6 3 0 2 irun 3 1 2 1 2 1 2	Aug 8 7 2 1 6 4 1 c.ts Aug 1	Sep 2 7 4 3 2 1 0 Sep 4	1 6 7 1 1 5 4 0ct 3	0 6 12 3 5 1 1 8 Nov 3	1 2 4 1 0 0 0 Dec 2
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Mann tau = > mto: 1992: 1993:</pre>	3.1AN Jan 2 2 3 2 0 0 0 NKenc 3.2AN Jan 2 1	4Thus Feb 4 2 4 1 3 1 1 1 1 1 1 1 1 1 1 	sat.c Mar 3 2 6 1 2 1 1 5 at.c Mar 1 2	=1. Apr 5 1 5 4 2 2 3.1AI Apr 0 3	5555 May 7 2 2 0 4 0 1 2 2 2 0 4 0 4 0 4 0 4 0 1 4 0 1 4 0 1 4 0 1 4 0 1 4 0 1 4 0 1 1 1 1 1 1 1 1 1 1 1 1 1	Jun 7 2 4 2 0 1 Sat.)1% Jun 1 0	Jul 3 2 3 6 3 0 2 1 run Jul 2 3	Aug 8 7 2 1 6 4 1 c.ts Aug 1 3	Sep 2 7 4 3 2 1 0 Sep 4 1	1 6 7 1 1 5 4 Oct 3 4	0 6 12 3 5 1 1 8 8 8 9 8 0	1 2 4 1 0 0 0 Dec 2 1
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Mann tau = > mto: 1992: 1992: 1993: 1994:</pre>	3.1AN Jan 2 2 3 2 0 0 0 0 NKenc 3.2AN Jan 2 1 2	4 Feb 4 2 4 1 3 1 1 1 1 1 1 1 1 1 1 	sat.c Mar 3 2 6 1 2 1 1 (mto3 Sat.c Mar 1 2 5	=1 Apr 5 5 4 2 2 3.1A 4 2 2 3.1A Apr 0 3 4	5555 May 7 2 2 2 0 4 0 4 0 4 0 4 0 14 (.ts May 4 2 4 0 4 4 0 4 4 2 2 0 4 4 0 4 4 2 2 0 4 4 4 2 2 0 4 4 4 4 4 4 4 4 4 4 4 4 4	Jun 7 2 4 2 0 1 Sat. 0 1 Sat. 0 1 8 1 0 1 1	Jul 3 3 6 3 0 2 1 run Jul 2 3 2	Aug 8 7 2 1 6 4 1 c.ts 3 4	Sep 2 7 4 3 2 1 0 Sep 4 1 1	1 6 7 1 5 4 Oct 3 4 2	0 6 12 3 5 1 1 1 Nov 3 0 6	1 2 4 1 0 0 0 Dec 2 1 4
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Mann tau = > mto: 1992: 1992: 1993: 1994: 1995:</pre>	3.1AN Jan 2 2 3 2 0 0 0 0 NKenc -0.3 3.2AN Jan 2 1 2 2	4Thus Feb 4 2 4 1 3 1 1 1 1 1 1 1 1 1 1 	sat.c Mar 3 2 6 1 2 1 1 (mto3 Sat.c Mar 1 2 5 4	. =1 drund Apr 5 1 5 5 4 2 2 2 3.1An . =0 drund Apr 0 3 4 3	555 May 7 2 2 2 0 4 0 4 0 1114 5 014 5 014 5 0 14 5 0 14 5 0 14 5 0 14 5 0 14 5 0 14 5 0 14 5 0 14 5 14 14 14 14 14 14 14 14 14 14	<pre>Jun 7 2 4 2 0 0 1 5at.0)1% Jun 1 0 1 2</pre>	Jul 3 2 3 6 3 0 2 1 1 1 1 1 2 3 2 4	Aug 8 7 2 1 6 4 1 c.ts 3 4 4	Sep 2 7 4 3 2 1 0 8 5 9 4 1 3	1 6 7 1 5 4 Oct 3 4 2 4	0 6 12 3 5 1 1 1 Nov 3 0 6 2	1 2 4 1 0 0 0 Dec 2 1 4 3
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Mann tau = > mto: 1992: 1993: 1994: 1995: 1996:</pre>	3.1AN Jan 2 2 3 2 0 0 0 0 NKenc -0.3 3.2AN Jan 2 1 2 2 1	4Thus Feb 4 2 4 1 3 1 1 1 1 1 1 1 1 1 1 	sat.co Mar 3 2 6 1 2 1 1 (mto3 Sat.co Mar 1 2 5 4 0	. =1 drund Apr 5 1 5 5 4 2 2 2 3.1Ar - =0 drund Apr 0 3 4 3 1 1 1 2 2 1 1 1 5 5 5 1 1 5 5 5 1 1 5 5 5 5 1 1 5 5 5 5 5 5 5 5 5 5 5 5 5	555 May 7 2 2 2 0 4 0 0 1 1 1 2 2 0 4 0 1 1 2 2 2 0 4 0 1 1 2 2 2 0 4 0 1 2 2 2 0 4 0 1 2 2 2 0 4 0 1 2 2 2 0 4 0 1 1 1 1 1 1 1 1 1 1 1 1 1	Jun 7 2 4 2 0 1 Sat. 0 1 Sat. 0 1 3 1 2 5	Jul 3 2 3 6 3 0 2 2 drunl 3 2 4 4 4	Aug 8 7 2 1 6 4 1 c.ts 3 4 4 5	Sep 2 7 4 3 2 1 0 8 5 9 4 1 3 2	1 6 7 1 5 4 0ct 3 4 2 4 3	0 6 12 3 5 1 1 1 Nov 3 0 6 2 2	1 2 4 1 0 0 0 Dec 2 1 4 3 3
<pre>> mto: 1992: 1993: 1994: 1995: 1996: 1997: 1998: > Mann tau = > mto: 1992: 1993: 1994: 1995: 1996: 1997:</pre>	3.1AN Jan 2 2 3 2 0 0 0 0 Kenc -0.3 3.2AN Jan 2 1 2 2 1 0 0	4Thus Feb 4 2 4 1 3 1 1 1 1 1 1 1 1 1 1 	sat.co Mar 3 2 6 1 2 1 1 (mto3 Sat.co Mar 1 2 5 4 0 1 3	<pre>. =1 Arunl Apr 5 1 5 4 2 2 3.1Ar . =0 irunl Apr 0 3 4 3 1 3 2</pre>	555 May 7 2 2 2 0 4 0 4 0 1110 	Jun 7 2 4 2 0 1 Sat. 0 1 Sat. 1 2 5 1 2	Jul 3 2 3 6 3 0 2 3 0 2 drunl 2 3 2 4 4 3 3 2	Aug 8 7 2 1 6 4 1 c.ts 1 3 4 4 5 6 3	Sep 2 7 4 3 2 1 0 8 5 9 4 1 3 2 3 3	1 6 7 1 5 4 0 ct 3 4 2 4 3 0	0 6 12 3 5 1 1 1 Nov 3 0 6 2 2 4	1 2 4 1 0 0 0 Dec 2 1 4 3 3 4

tau = 0.0399, sl =62%

Trellis Time Series Plots: Common scale on each page.

MTO, Major Injury, Sober, Common Scale





MTO, Major Injury, Drunk, Common Scale

Year

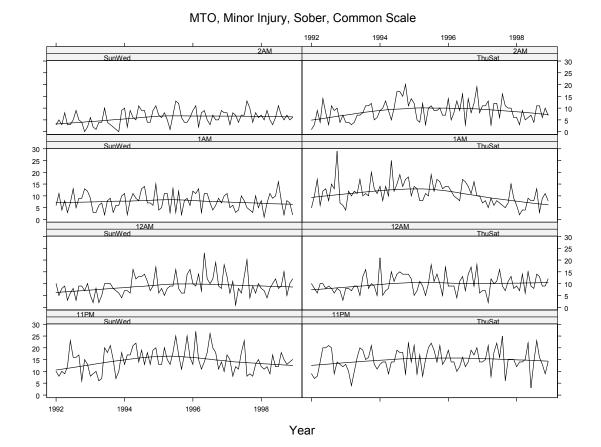
Minor Injury

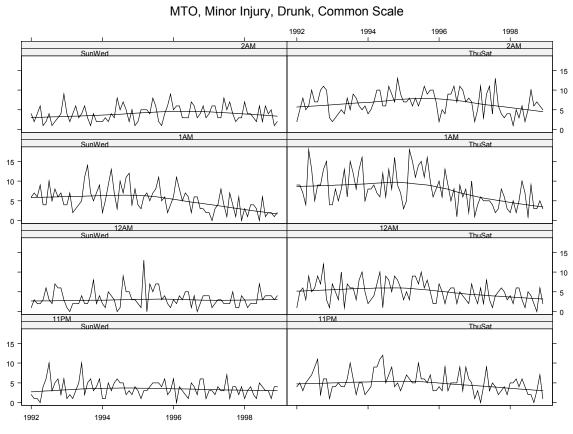
tau = 0.0881, sl =25.11%

```
tau = 0.103, sl =19.35%
```

tau = -0.0346, sl = 65.77%

Trellis Time Series Plots







Minimal Injury

tau = 0.173, sl =2.367%

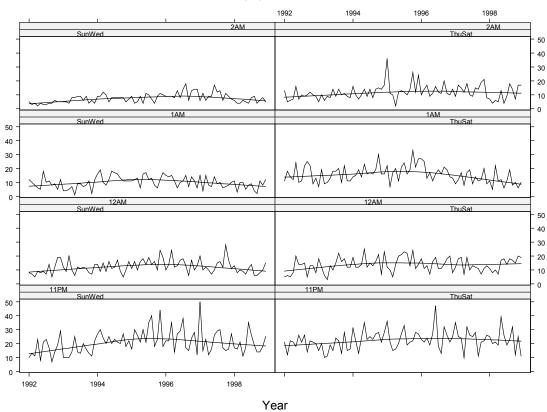
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> #												
mto1.11	PMS	SunWe	ed.dr	runk	.ts							
J	an	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	1	3	4	5	5	5	2	8	3	4	5	3
1993:	3	2	2	4	1	2	5	7	2	3	2	5
1994:	2	3	3	1	3	7	1	4	0	3	6	5
1995:	3	3	4	2	1	4	6	7	5	3	2	7
1996:	3	4	5	4	2	3	б	6	1	2	2	2
1997:	0	1	1	3	2	0	2	0	2	2	1	1
1998:	1	3	0	5	3	1	2	3	3	0	0	1
> MannK	enċ	lall	(mtol	.11	PMSui	wed	.dru	nk.ts	3)			
tau = -	0.2	273,	s]	=0	.056	3 9 %						
> mtol.	12 <i>F</i>	MSur	nWed.	dru	nk.ts	5						
J	an	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	2	1	2	5	2	3	4	4	2	5	6	3
1993:	2	2	4	5	3	8	б	6	3	4	1	7
1994:	1	1	2	3	2	7	б	5	9	5	8	8
1995:	2	3	5	5	2	5	4	9	4	4	5	2
1996:	1	3	3	3	4	1	5	5	1	2	2	0
1997:	3	1	3	1	1	1	8	1	0	б	4	0
1998:	3	0	2	5	2	1	5	4	3	3	2	2
> MannK	enċ	lall	(mtol	.12	AMSui	wed	.dru	nk.ts	5)			
tau = -	0.1	.56,	s]	=4	.74 %							
> mtol.	1AM	ISunV	ved.c	lrunl	k.ts							
J	an	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:	5	3	1	9	6	4	5	6	5	5	8	0
1993:	5	4	4	5	7	6	11	5	8	6	2	2
1994:	3	2	7	10	8	5	7	6	4	6	4	7
1995:	4	2	8	8	4	8	5	9	6	0	5	3
1996:	6	11	2	4	2	3	1	3	3	9	2	2
1997:	4	2	5	3	1	1	2	3	4	3	3	3
1998:	1	0	3	3	1	3	4	0	2	1	0	1
> MannK			•				drun	k.ts)			
tau = -					.086	e-4%						
> mtol.												
			Mar	-	-			_	_			
1992:	3	0	5	2	2	3	4	3	4	4	0	8
1993:	2	1	1	2	4	2	2	3	4	1	5	5
1994:	2	5	2	2	1	2	9	3	4	7	3	2
1995:	2	3	5	3	7	4	3	8	2	2	4	2
1996:	3	2	4	2	3	3	4	6	5	4	0	3
						4	4	-	2	4	2	5
1997:	3	3	3	2	5	4	-	5	2	-	2	
1998:	7	5	4	5	6	2	3	4	3	4	∠ 3	5 4
	7 enc	5 lall	4 (mtol	5	6 MSuni	2	3	4	3	-		

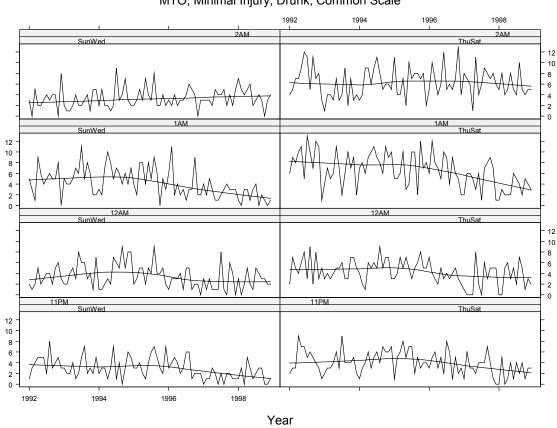
> MannKendall(mto1.2AMThuSat.drunk.ts)

tau = 0.0118, sl =88.24%

Trellis Time Series Plots



MTO, Minimal Injury, Sober, Common Scale



MTO, Minimal Injury, Drunk, Common Scale

E. MVA Dataset

Summary

The MVA dataset is comprised of 84 records over the period May 1992 to April 1999 of serious trauma caused by motor vehicle accidents in Ontario. The data are obtained from the Ontario Trauma Registry. About 45% of OTR cases occur in the 3AMPlus group. Since we are primarily interested in the other groups, the data corresponding to this subgroup was analysed separately.

<u>HOUR</u>: The uniformity hypothesis is rejected by a chi-square test at less than 0.04%. The major departure is due to fewer than expected traumas in the 12AM and 1AM windows.

<u>June-to-Sep</u> June-to-Sep account for 45% so these months have a much higher share. Assuming cases are uniformly distributed over the months, yields an observed chi-squared goodness-of-fit statistic of 73.79 on 11 df which is statistically significant at less than 10^{-10} .

Hour and month are not associated

Hour and year are may be associated The largest departures from expected occur in the hourly windows in 1999.

Figure 1a. MVA counts/month (excluding 3AMPlus) show an initial upward trend followed by a downward trend starting about 1995

Figure 1b. MVA counts/month 3AMPlus show an initial upward trend to about 1995 and a downward trend post-1995.

Figure 2a and Figure 2b show there has been a strong downward trend since about 1995 for each hour 11AM, 12PM, 1AM, 2AM.

Table 2. Counts by hour – monotonic trend not detected at 12AM and 2AM.

Figure 3a STL, MVA. A peak is reached in 1995 is followed by a downward trend.

Figure 3b. MVA Seasonal The seasonal has a well defined peak in August, smaller peak in Nov and minimum in February. There is a downward trend in Apr and an upward trend Jul. and Nov

Figure 4a. STL, MVA, 3AMPlus. Increasing trend to 1995 and decreasing thereafter. The seasonal pattern changes in 1996 with the Jul-Aug peak becoming less prominent and the smaller Oct peak increasing.

Figure 4b. MVA, 3AMPlus Seasonal. Well defined peaks in Mar, Aug and Oct and minimum in Feb, Apr and Sep. The Oct peak has increased since 1996 and the Aug peak has decreased since 1996.

Introduction

The data set of Ontario Trauma Registry (OTR) contains the both motor vehicle accident and non motor vehicle accident cases (ISS Scale > 12, severe injury) from May 1992 to April 1999.

The hour factor variable includes 11:00pm-12:00am, 12:00am-1:00am, 1:00am-2:00am, 2:00am - 3:00am and 3:00am -10:59pm.

Variables of interest:

counts	number of mva cases with ISS Scale > 12 (severe injury)
year	ordered factor, 8 levels: 1992 to 1999
month	ordered factor, 12 levels: January to December
hour	ordered factor with 5 levels: "11PM", "12AM", "1AM", "2AM", "3AMPlus"

About 44% of MVA cases occur in the 3AMPlus group. Since we are primarily interested in the other groups, the data corresponding to this subgroup was analysed separately. Also notice that there is no wkgrp variable for this data.

```
> crosstabs(counts~hour, data=mva.df)
Call:
crosstabs(counts ~ hour, data = mva.df)
1964 cases in table
+----+
| N
     |N/Total|
+----+
hour |
----+
11PM |285
          0.15
          ----+
12AM |220
          0.11
          ____+
1AM
    |318
          0.16
         ----+
2AM
    |271
          |0.14
          ----+
3AMPlus|870
          0.44
----+
```

Crosstabs Analysis

Hour

The uniformity hypothesis is rejected by a chi-square test at less than 0.04%. The major departure is due to fewer than expected traumas in the 12AM and 1AM windows.

```
> crosstabs(counts~hour, data=mval.df)
Call:
crosstabs(counts ~ hour, data = mval.df)
1094 cases in table
+----+
| N
      |N/Total|
+----+
hour
     ----+
11PM |285
             0.26
             ----+
    |220
12AM
             0.20
             ----+
1AM
     |318
             0.29
             ----+
     |271
2AM
             0.25
             ----+
> cs.test(c(285,220,318,271))
$Xsq:
[1] 18.21207
$Pval:
[1] 0.0003976996
$Decomp:
[1] 0.48354662 10.46526508 7.24040219 0.02285192
```

MVA

Hour and month are not associated (p-value=0.8024353)

June-to-Sep account for 45% so these months have a much higher share. Assuming cases are uniformly distributed over the months, yields an observed chi-squared goodness-of-fit statistic of 73.79 on 11 df which is statistically significant at less than 10^{-10} .

There is no dependence between hour and month.

N N/Row N/Col N/Tot	Total Total												
hour	month Jan			· •	· -								RowTotl
11PM	15 0.0526 0.2632 0.0137	19 0.0667 0.3519 0.0174	17 0.0596 0.2429 0.0155	16 0.0561 0.1951 0.0146	22 0.0772 0.2558 0.0201	29 0.1018 0.2636 0.0265	33 0.1158 0.2481 0.0302	30 0.1053 0.2174 0.0274	32 0.1123 0.2857 0.0293	23 0.0807 0.2706 0.0210	28 0.0982 0.3111 0.0256	21 0.0737 0.2727 0.0192	İ
12AM	14 0.0636 0.2456 0.0128	5 0.0227 0.0926 0.0046	14 0.0636 0.2000 0.0128	15 0.0682 0.1829 0.0137	22 0.1000 0.2558 0.0201	20 0.0909 0.1818 0.0183	23 0.1045 0.1729 0.0210	28 0.1273 0.2029 0.0256	20 0.0909 0.1786 0.0183	17 0.0773 0.2000 0.0155	23 0.1045 0.2556 0.0210	19 0.0864 0.2468 0.0174	220 0.20
1AM	15 0.0472 0.2632 0.0137	16 0.0503 0.2963 0.0146	21 0.0660 0.3000 0.0192	36 0.1132 0.4390 0.0329	23 0.0723 0.2674 0.0210	33 0.1038 0.3000 0.0302	44 0.1384 0.3308 0.0402	35 0.1101 0.2536 0.0320	32 0.1006 0.2857 0.0293	24 0.0755 0.2824 0.0219	19 0.0597 0.2111 0.0174	20 0.0629 0.2597 0.0183	318 0.29
2AM	13 0.0480 0.2281 0.0119	14 0.0517 0.2593 0.0128	18 0.0664 0.2571 0.0165	15 0.0554 0.1829 0.0137	19 0.0701 0.2209 0.0174	28 0.1033 0.2545 0.0256	33 0.1218 0.2481 0.0302	45 0.1661 0.3261 0.0411	28 0.1033 0.2500 0.0256	21 0.0775 0.2471 0.0192	20 0.0738 0.2222 0.0183	17 0.0627 0.2208 0.0155	271 0.25
ColTot	1 57 0.052	54 0.049	70 0.064	82 0.075	86 0.079	110 0.101	133 0.122	138 0.126	112 0.102	85 0.078	90 0.082	77 0.070	-+ 1094
<pre>Test f > get. [1,] [2,] [3,] [4,] contri [1,] [2,] [3,] -</pre>	or indepe Chi^2 = 2 Yates' cc crosstabs ,1] [,2] 1 35	endence of 29.64457 prrection 3.percent [,3] [,4 -7 -2 -1 - 3 5 4 -2 ochi-sq: 20 chi-sq: 21 [,3] 32 -0.29 78 -0.02 98 0.14	of all fa d.f.= 33 n not use cerror(cc l] [,5] [25 -2 -9 27 51 -8 26 -11 [,4] [-1.16 -0 -0.37 1 2.49 -0	ctors (p=0.63 d unts~hou ,6] [,7] 1 -5 -10 -14 3 14 3 0 ,5] [,6 .09 0.0 .13 -0.4 .40 0.1	49629) [,8] [, -17 1 - -13 32 [,7] 6 -0.28 5 -0.72 8 0.86	9] [,10] 10 4 11 -1 -2 -3 1 0 [,8] [-0.99 0 0.05 -0 -0.81 -0	[,11] [19 27 -27 -10 ,9] [,10 .52 0.1 .53 -0.0	5,12] 523 -11 -11 0] [,11] 8 0.94 02 1.15 .4 -1.40	[,12] 0.21 0.89 -0.50	+		+	-+

MVA

Hour and year are may be associated (p-value about 10%)

The largest departures from expected occur in the hourly windows in 1999.

```
> crosstabs(counts~hour+year, data=mva.df, subset=!mva.df$hour=="3AMPlus")
Call:
crosstabs(counts ~ hour + year, data = mva.df, subset = !(mva.df$hour == "3AMPlus"))
1094 cases in table
+----+
N I
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
hour |year
 |1992 |1993 |1994 |1995 |1996 |1997 |1998 |1999 |RowTotl|
11PM |30 |50 |47 |44 |44 |35 |27 | 8 |285
    |0.10526|0.17544|0.16491|0.15439|0.15439|0.12281|0.09474|0.02807|0.26
     0.30928|0.31056|0.23383|0.22564|0.27329|0.22436|0.25472|0.47059|
     |0.02742|0.04570|0.04296|0.04022|0.04022|0.03199|0.02468|0.00731|
_____+
12AM |13 |27 |39 |46 |41 |34 |19 | 1 |220
     0.059090.122730.177270.209090.186360.154550.086360.004550.20
     0.13402|0.16770|0.19403|0.23590|0.25466|0.21795|0.17925|0.05882|
     0.011880.024680.035650.042050.037480.031080.017370.00091
1AM |30 |47 |68 |62 |38 |40 |26 |7 |318 |
    0.094340.147800.213840.194970.119500.125790.081760.022010.29
    0.30928 0.29193 0.33831 0.31795 0.23602 0.25641 0.24528 0.41176
    0.027420.042960.062160.056670.034730.036560.023770.00640
2AM |24 |37 |47 |43 |38 |47 |34 | 1 |271 |
     0.08856|0.13653|0.17343|0.15867|0.14022|0.17343|0.12546|0.00369|0.25
     0.24742|0.22981|0.23383|0.22051|0.23602|0.30128|0.32075|0.05882|
     0.02194 0.03382 0.04296 0.03931 0.03473 0.04296 0.03108 0.00091
_____+
                                                        --+---+
ColTotl|97 |161 |201 |195 |161 |156 |106 |17 |1094 |
  |0.089 |0.147 |0.184 |0.178 |0.147 |0.143 |0.097 |0.016 |
_____+
Test for independence of all factors
  Chi^2 = 30.08949 d.f.= 21 (p=0.09020436)
  Yates' correction not used
  Some expected values are less than 5, don't trust stated p-value
> get.crosstabs.percenterror(counts~hour+year)
   [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8]
 \begin{bmatrix} 1 & 1 & 9 & 19 & -10 & -13 & 5 & -14 & -2 & 81 \\ \begin{bmatrix} 2 & 1 & -33 & -17 & -4 & 17 & 27 & 8 & -11 & -71 \\ \begin{bmatrix} 3 & 1 & 6 & 0 & 16 & 9 & -19 & -12 & -16 & 42 \\ \begin{bmatrix} 4 & 1 & 0 & -7 & -6 & -11 & -5 & 22 & 29 & -76 \\ \end{bmatrix} 
contribution to chi-sq:
   [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8]
[1,] 0.94 1.24 -0.74 -0.95 0.32 -0.88 -0.12 1.70
[2,] -1.47 -0.94 -0.22 1.08 1.52 0.47 -0.50 -1.31
[3,] 0.34 0.03 1.25 0.71 -1.29 -0.79 -0.87 0.93
[4,] -0.01 -0.46 -0.40 -0.76 -0.30 1.34 1.51 -1.56
```

Data Visualization

In our time series plots we use a 60% robust locally linear smoother to visualize the trend except in the STL analysis where a 60% robust locally quadratic smoother is used for trend. Monotonic trend test significant in both cases.

Table 1a.

OTR monthly time series, excluding 3AMPlus window.

```
> mva.tot.ts
       Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec
1992:
                           13
                                              12
                                15
                                     13
                                          10
                                                   12
                                                         7
                                                             15
1993:
         5
             13
                           15
                                          21
                                              22
                                                   10
                                                              9
                   4
                      18
                                18
                                     12
                                                        14
         7
1994:
              7
                  14
                      21
                           15
                                12
                                     25
                                          24
                                              19
                                                   20
                                                        17
                                                             20
1995:
        17
                           14
                                25
                                          29
                                                    7
                                                              9
              8
                  17
                      12
                                     28
                                              16
                                                        13
1996:
                                                              7
        7
             10
                  16
                       8
                           13
                                11
                                     20
                                          24
                                              18
                                                   12
                                                        15
1997:
        10
              6
                  14
                      13
                           11
                                18
                                     16
                                          19
                                              13
                                                   11
                                                        15
                                                             10
         5
              5
                        5
                            5
                                                              7
1998:
                   4
                                11
                                          11
                                                   13
                                                         9
                                     19
                                              12
1999:
              5
                        5
         6
                   1
> SeasonalMannKendall(mva.tot.ts)
tau = -0.262,
                   sl =0.4841%
```

Table 1b

MVA monthly time series, 3AMPlus window.

```
> mva0.tot.ts
       Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec
1992:
                           14
                                11
                                    16
                                         17
                                              15
                                                   16
                                                       14
                                                            16
1993:
                 11
                      15
                           15
                                16
                                    23
                                         22
                                                   17
                                                       20
                                                            16
        13
             12
                                              16
1994:
        15
                 14
                           19
                                22
                                    20
                                         25
                                              21
                                                   23
                                                            16
             13
                      15
                                                       19
        18
1995:
             17
                 20
                      18
                           19
                                23
                                    25
                                         25
                                              21
                                                   20
                                                       19
                                                            20
1996:
        16
            15
                 21
                      15
                           21
                                23
                                    22
                                         23
                                              20
                                                   20
                                                       22
                                                            16
1997:
            17
                                20
                                                   22
                                                            17
        15
                 17
                      12
                           19
                                    21
                                         21
                                              18
                                                       14
1998:
        15
            14
                 14
                      18
                           19
                               20
                                    23
                                         22
                                              21
                                                   23
                                                       17
                                                            20
1999:
        16
            12
                 12
                      14
> SeasonalMannKendall(mva0.tot.ts)
tau = 0.249,
                 sl =0.8697%
```

Figure 1a. MVA counts/month (excluding 3AMPlus) show an initial upward trend followed by a downward trend starting about 1995.

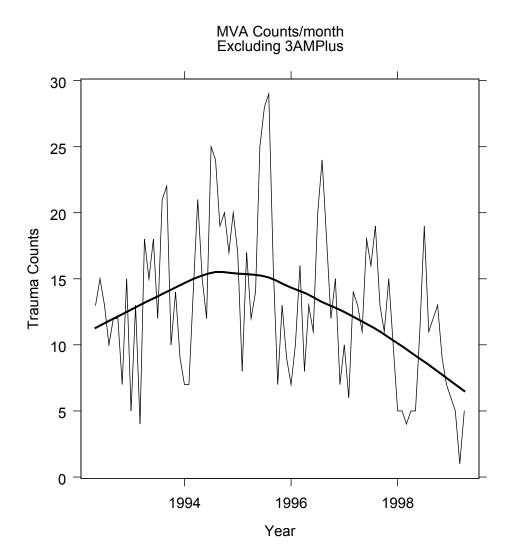
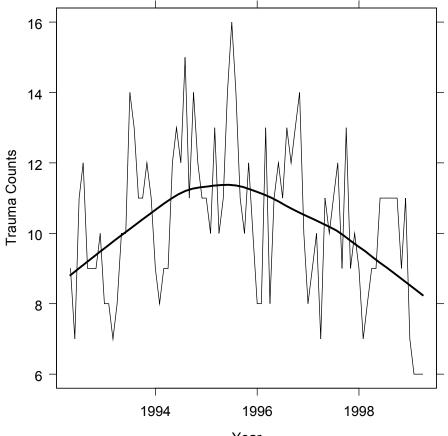


Figure 1a.

Figure 1b. MVA counts/month 3AMPlus show an initial upward trend to about 1995 and a downward trend post-1995.

Figure 1b.

MVA Counts/month for 3AMPlus





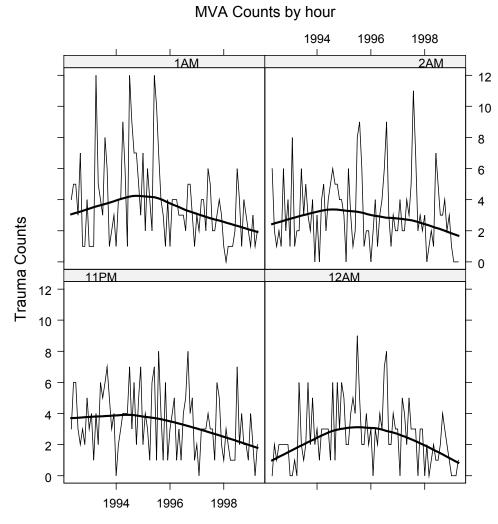
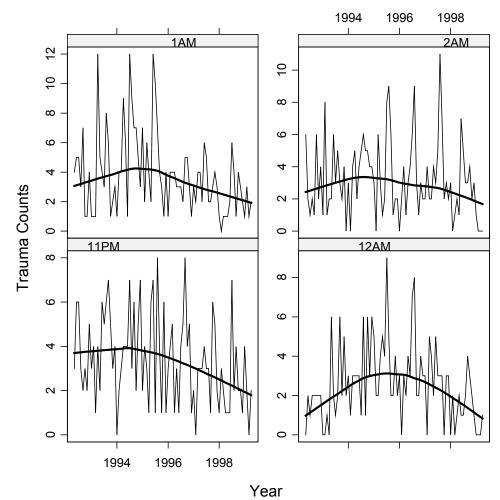


Figure 2a. MVA counts by hour. Common scale. The downward trend does not appear as step in the 2AM window.

Year

Figure 2b.



MVA Counts by hour

Independent vertical scale.

Table 2. Tabulations and Seasonal Mann-Kendall Tests

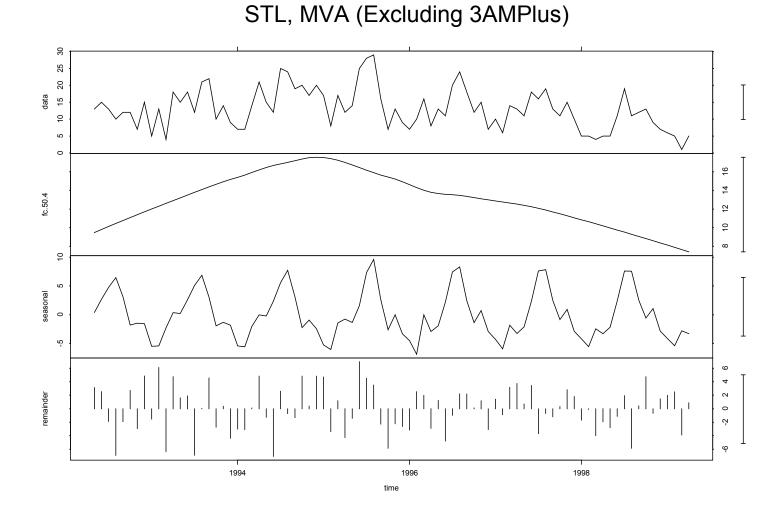
Monotonic trend test significant for 11PM and 1AM at <5% but not for 12AM and 2AM.

	nt.11	LPM.t	ts									
				Apr	May	Jun	Jul	Auq	Sep	Oct	Nov	Dec
1992:				1	3	6	6	3	2	3	2	5
1993:	3	4	1	4	2	6	5	6	7	5	3	4
1994:	0	2	3	4	4	4	7	3	6	2	5	7
1995:	2	4	3	1	6	7	1	8	4	1	6	1
1996:	3	4	5	1	3	1	4	5	8	4	5	1
1997:	2	0	3	3	3	4	3	3	1	6	5	2
1998:	1	3	2	1	1	1	7	2	4	2	2	1
1999:	4	2	0	2								
> Seas	sonal	LManr					LPM.	ts)				
tau =	-0.1	L95,	s	1 =3	.875	5						
> cour												
	Jan	Feb	Mar	Apr	May		Jul		Sep			Dec
1992:					0	2	1	2	2	2	2	2
1993:	0	0	1	0	6	2	1	2	6	2	5	2
1994:	3	1	3	3	3	3	1		1	6	3	6
1995:	5	2	2	4	-	4	9	-	2	2	4	
1996:	3	0	3	2		3	7	-	2	2	4	-
1997:	3	0	5	4		5	3		3	0	3	
1998:	0	2	0	1	2	1	1	2	4	3	2	1
1999:	0	0	0	1								
> Seas							2AM.1	ts)				
tau =	-0.0)132,	, :	s1 =8	39.00	28						
\ ~~···	а <u>н</u> 17		~									
> cour												
				∆ nr	Most	Turn	T., 1	7110	Son	Oat	Nou	Dog
1002.				Apr	May							
1992:	Jan	Feb	Mar		4	5	5	3	7	1	1	4
1993:	Jan 1	Feb 1	Mar 1	12	4 5	5 4	5 3	3 8	7 6	1 1	1 2	4 3
1993: 1994:	Jan 1 1	Feb 1 4	Mar 1 4	12 9	4 5 6	5 4 1	5 3 12	3 8 9	7 6 7	1 1 7	1 2 5	4 3 3
1993: 1994: 1995:	Jan 1 1 7	Feb 1 4 2	Mar 1 4 6	12 9 4	4 5 6 2	5 4 1 12	5 3 12 10	3 8 9 7	7 6 7 4	1 1 7 3	1 2 5 1	4 3 3 4
1993: 1994: 1995: 1996:	Jan 1 1 7 1	Feb 1 4 2 4	Mar 1 4 6 4	12 9 4 4	4 5 6 2 3	5 4 1 12 3	5 3 12 10 3	3 8 9 7 2	7 6 7 4 5	1 1 7 3 5	1 2 5 1 3	4 3 3 4 1
1993: 1994: 1995: 1996: 1997:	Jan 1 7 1 3	Feb 1 4 2 4 2	Mar 1 4 6 4 4	12 9 4 4	4 5 2 3 2	5 4 12 3 6	5 3 12 10 3 5	3 9 7 2 2	7 6 7 4 5 2	1 7 3 5 3	1 2 5 1 3 4	4 3 4 1 3
1993: 1994: 1995: 1996: 1997: 1998:	Jan 1 7 1 3 1	Feb 1 4 2 4 2 0	Mar 1 4 6 4 4 1	12 9 4 4 4 1	4 5 6 2 3	5 4 1 12 3	5 3 12 10 3	3 9 7 2 2	7 6 7 4 5	1 1 7 3 5	1 2 5 1 3	4 3 3 4 1
1993: 1994: 1995: 1996: 1997: 1998: 1999:	Jan 1 1 7 1 3 1 1	Feb 1 4 2 4 2 0 3	Mar 1 4 6 4 4 1 1	12 9 4 4 1 2	4 5 2 3 2 1	5 4 12 3 6 2	5 3 12 10 3 5 6	3 9 7 2 4	7 6 7 4 5 2	1 7 3 5 3	1 2 5 1 3 4	4 3 4 1 3
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas	Jan 1 7 1 3 1 1 50000	Feb 1 4 2 4 2 0 3 1Man	Mar 1 4 6 4 1 1 0.Keno	12 9 4 4 1 2 dall	4 5 2 3 2 1	5 4 12 3 6 2	5 3 12 10 3 5 6	3 9 7 2 4	7 6 7 4 5 2	1 7 3 5 3	1 2 5 1 3 4	4 3 4 1 3
1993: 1994: 1995: 1996: 1997: 1998: 1999:	Jan 1 7 1 3 1 1 50000	Feb 1 4 2 4 2 0 3 1Man	Mar 1 4 6 4 1 1 0.Keno	12 9 4 4 1 2 dall	4 5 2 3 2 1	5 4 12 3 6 2	5 3 12 10 3 5 6	3 9 7 2 4	7 6 7 4 5 2	1 7 3 5 3	1 2 5 1 3 4	4 3 4 1 3
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau =	Jan 1 7 1 3 1 1 sonal -0.2	Feb 1 4 2 0 3 IMann 231 ,	Mar 1 4 6 4 1 1 5	12 9 4 4 1 2 dall	4 5 2 3 2 1	5 4 12 3 6 2	5 3 12 10 3 5 6	3 9 7 2 4	7 6 7 4 5 2	1 7 3 5 3	1 2 5 1 3 4	4 3 4 1 3
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas	Jan 1 7 1 3 1 5 0.2 nt.22	Feb 1 4 2 0 3 IMann 231, AM.ts	Mar 1 4 6 4 1 1 nKeno s	12 9 4 4 1 2 dall 1 =1	4 5 2 3 2 1 (cour . 457 ?	5 4 12 3 6 2 nt.12	5 3 12 10 3 5 6 AM.ts	3 9 7 2 2 4 s)	7 6 7 4 5 2 1	1 7 3 5 3 4	1 2 5 1 3 4 3	4 3 4 1 3 2
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > cour	Jan 1 7 1 3 1 5 0.2 nt.22	Feb 1 4 2 0 3 IMann 231, AM.ts	Mar 1 4 6 4 1 1 nKeno s	12 9 4 4 1 2 dall 1 =1	4 5 2 3 2 1 (cour . 457 ?	5 4 12 3 6 2 nt.12 8	5 3 12 10 3 5 6 AM.t:	3 8 9 7 2 2 4 s)	7 6 7 4 5 2 1 Sep	1 7 3 5 3 4 Oct	1 2 5 1 3 4 3 Nov	4 3 4 1 3 2 Dec
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > cour 1992:	Jan 1 7 1 3 1 5 5 0.2 1 1 5 0.2 Jan	Feb 1 4 2 4 2 0 3 IMann 231 , Feb	Mar 1 4 6 4 1 1 nKeno s Mar	12 9 4 4 1 2 dall 1 =1 Apr	4 5 2 3 2 1 (cour . 457 9 May 6	5 4 12 3 6 2 nt.12 5	5 3 12 10 3 5 6 AM.ts Jul 1	3 8 9 7 2 2 4 s) Aug 2	7 6 7 4 5 2 1 Sep 1	1 1 7 3 5 3 4 0ct	1 2 5 1 3 4 3 8 Nov 2	4 3 4 1 3 2 Dec 4
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > cour 1992: 1993:	Jan 1 7 1 3 1 sonal -0.2 Jan 1	Feb 1 4 2 4 2 0 3 IMann 231, AM.ts Feb 8	Mar 1 4 6 4 1 1 nKeno s Mar 1	12 9 4 4 1 2 dall 1 =1 Apr 2	4 5 2 3 2 1 (cour .457 5 May 6 2	5 4 12 3 6 2 nt.12 5	5 3 12 10 3 5 6 AM.ts Jul 1 3	3 8 9 7 2 2 4 s) Aug 2 5	7 6 7 4 5 2 1 \$ \$ep 1 3	1 1 7 3 5 3 4 0ct 6 2	1 2 5 1 3 4 3 Nov 2 4	4 3 4 1 3 2 Dec 4 0
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > cour 1992:	Jan 1 7 1 3 1 5 5 0.2 1 1 5 0.2 Jan	Feb 1 4 2 4 2 0 3 IMann 231 , Feb	Mar 1 4 6 4 1 1 nKeno s Mar	12 9 4 4 1 2 dall 1 =1 Apr	4 5 2 3 2 1 (cour . 457 9 May 6	5 4 12 3 6 2 nt.12 5	5 3 12 10 3 5 6 AM.ts Jul 1	3 8 9 7 2 2 4 s) Aug 2	7 6 7 4 5 2 1 Sep 1	1 1 7 3 5 3 4 0ct	1 2 5 1 3 4 3 8 Nov 2 4 4	4 3 4 1 3 2 Dec 4 0 4
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > coun 1992: 1993: 1994:	Jan 1 7 1 3 1 5 5 0 2 Jan 1 3	Feb 1 4 2 0 3 IMann 231, AM.ts Feb 8 0	Mar 1 4 6 4 1 1 mKeno s Mar 1 4	12 9 4 4 1 2 dall 1 =1 Apr 2 5	4 5 2 3 2 1 (cour .4579 May 6 2 2	5 4 12 3 6 2 nt.12 8 Jun 2 6 4	5 3 12 10 3 5 6 AM.t: Jul 1 3 5	3 8 9 7 2 2 4 5 8 9 7 2 4 5 6	7 6 7 2 1 Sep 1 3 5	1 1 7 3 5 3 4 0ct 6 2 5	1 2 5 1 3 4 3 Nov 2 4	4 3 4 1 3 2 Dec 4 0 4 2
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > coun 1992: 1993: 1994: 1995:	Jan 1 7 1 3 1 5 5 0.2 Jan 1 3 3 3	Feb 1 4 2 0 3 LMann 231, AM.ts Feb 8 0 0	Mar 1 4 6 4 1 1 Mar 5 Mar 1 4 6	12 9 4 4 1 2 dall 1 =1 Apr 2 5 3	4 5 2 3 2 1 (cour .4579 May 6 2 2 1	5 4 12 3 6 2 nt.12 5 Jun 2 6 4 2	5 3 12 10 3 5 6 AM.t: Jul 1 3 5 8	3 8 9 7 2 2 4 5 8 9	7 6 7 4 5 2 1 Sep 1 3 5 6	1 1 7 3 5 3 4 Oct 6 2 5 1	1 2 5 1 3 4 3 8 Nov 2 4 4 2	4 3 4 1 3 2 Dec 4 0 4
1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > coun 1992: 1993: 1994: 1995: 1996:	Jan 1 7 1 3 1 5 5 0 2 Jan 1 3 3 0	Feb 1 4 2 0 3 LMann 231, Feb 8 0 0 2	Mar 1 4 6 4 1 1 mKeno s Mar 1 4 6 4	12 9 4 1 2 dall 1 =1 Apr 2 5 3 1	4 5 2 3 2 1 (cour .457? May 6 2 2 1 3	5 4 12 3 6 2 nt.12 5 Jun 2 6 4 2 4	5 3 12 10 3 5 6 AM.t: Jul 1 3 5 8 6	3 8 9 7 2 4 4 5 5 6 9 9	7 6 7 4 5 2 1 Sep 1 3 5 6 3	1 1 7 3 5 3 4 Oct 6 2 5 1 1	1 2 5 1 3 4 3 Nov 2 4 4 2 3	4 3 4 1 3 2 Dec 4 0 4 2 2
<pre>1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > cour 1992: 1993: 1994: 1995: 1996: 1997:</pre>	Jan 1 7 1 3 1 5 5 0 2 1 3 0 2	Feb 1 4 2 4 2 0 3 LMann 231, Feb 8 0 0 2 4	Mar 1 4 6 4 1 1 mKeno s Mar 1 4 6 4 2	12 9 4 4 1 2 dall 1 =1 Apr 2 5 3 1 2	4 5 6 2 3 1 (cour .4579 May 6 2 2 1 3 4	5 4 12 3 6 2 11.12 5 Jun 2 6 4 2 4 3	5 3 12 10 3 5 6 AM.t: Jul 1 3 5 8 6 5	3 8 9 7 2 4 3 5 6 9 9 11	7 6 7 4 5 2 1 Sep 1 3 5 6 3 7	1 1 7 3 5 3 4 Oct 6 2 5 1 1 2	1 2 5 1 3 4 3 8 Nov 2 4 4 2 3 3	4 3 4 1 3 2 Dec 4 0 4 2 2 2
<pre>1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > cour 1992: 1993: 1994: 1995: 1996: 1997: 1998:</pre>	Jan 1 7 1 3 1 1 5 5 0 2 3 1 3 0 2 3 1	Feb 1 4 2 4 2 0 3 LMann 231, Feb 8 0 0 2 4 0 0 2 4 0 0 0 2 4 0 0 3 LMann 2 4 2 0 3 LMann 2 4 2 0 3 LMann 2 4 2 0 3 LMann 2 4 2 0 3 LMann 2 4 2 0 3 LMann 2 4 2 0 3 LMann 2 4 2 0 3 LMann 2 4 0 0 3 LMann 2 4 0 0 0 0 0 0 0 0 0 0 0 0 0	Mar 1 4 6 4 1 1 mKeno s Mar 1 4 6 4 2 1 0	12 9 4 4 1 2 dall 1 =1 Apr 2 5 3 1 2 2 0	4 5 6 2 3 2 1 (cour .4579 May 6 2 2 1 3 4 1	5 4 12 3 6 2 nt.12 5 Jun 2 6 4 2 4 3 7	5 3 12 10 3 5 6 AM.t: Jul 1 3 5 8 6 5 5 5	3 8 9 7 2 2 4 5 5 6 9 9 11 3	7 6 7 4 5 2 1 Sep 1 3 5 6 3 7	1 1 7 3 5 3 4 Oct 6 2 5 1 1 2	1 2 5 1 3 4 3 8 Nov 2 4 4 2 3 3	4 3 4 1 3 2 Dec 4 0 4 2 2 2
<pre>1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > cour 1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999:</pre>	Jan 1 1 7 1 3 1 50001 -0.2 Jan 1 3 0 2 3 1 50001 3 3 0 2 3 1 50001 3 3 0 2 3 1 50001 3 1 50001 5000 5000 1 500 1 5000 1 5000 1 1 5000 1 1 1 1 1 1 1 1 1 1 1 1 1	Feb 1 4 2 4 2 0 3 IMann 231, Feb 8 0 0 2 4 0 0 LMann 20 1 1 4 2 1 4 2 0 3 1 4 2 1 4 2 0 3 1 1 4 2 1 4 2 0 3 1 1 1 1 1 1 1 1 1 1 1 1 1	Mar 1 4 6 4 1 1 1 Mar 5 Mar 1 4 6 4 2 1 0 0 MKeno	12 9 4 4 1 2 dall 1 =1 Apr 2 5 3 1 2 2 0 dall	4 5 6 2 3 2 1 (cour .4579 6 2 2 1 3 4 1 (cour .4579	5 4 12 3 6 2 nt.12 5 Jun 2 6 4 2 4 3 7 nt.22	5 3 12 10 3 5 6 AM.t: Jul 1 3 5 8 6 5 5 5	3 8 9 7 2 2 4 5 5 6 9 9 11 3	7 6 7 4 5 2 1 Sep 1 3 5 6 3 7	1 1 7 3 5 3 4 Oct 6 2 5 1 1 2	1 2 5 1 3 4 3 8 Nov 2 4 4 2 3 3	4 3 4 1 3 2 Dec 4 0 4 2 2 2

A peak is reached in 1995 is followed by a downward trend. There is a change in the shape of the seasonal component occuring in 1996.

Figure 3a

R-sq = 69.2 %



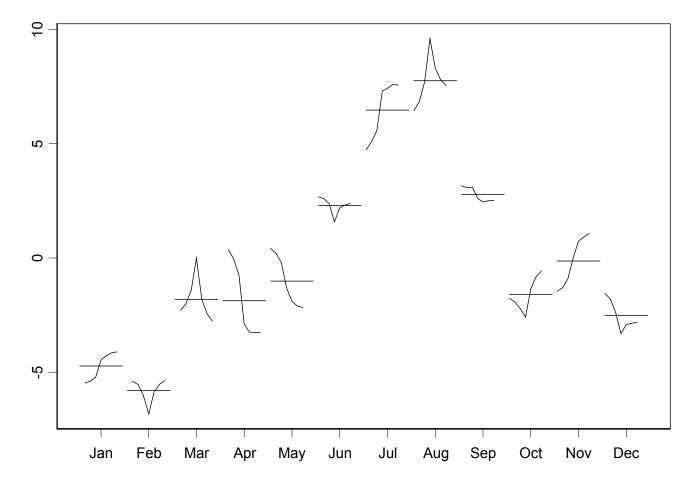
ss.window = 7 ,ss.robust = TRUE , fc window = 50.4 fc degree = 2

12

Figure 3b. The seasonal has a well defined peak in August, smaller peak in Nov and minimum in February. There is a downward trend in Apr and an upward trend Jul. and Nov



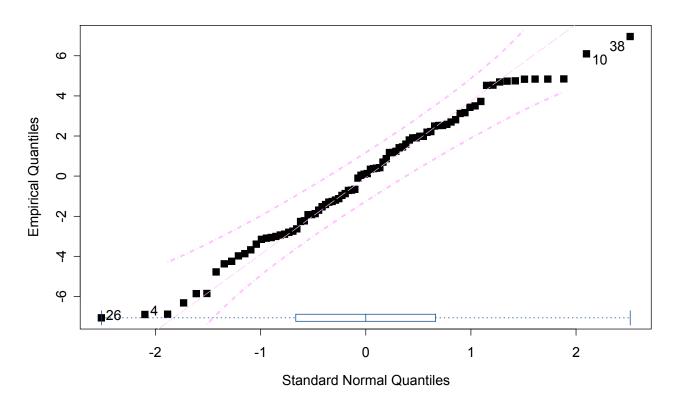
Seasonal Component, MVA Excluding 3AMPlus



MVA

Figure 3c. Normal probability plot.

Reasonably normal. No strong indication of outliers.

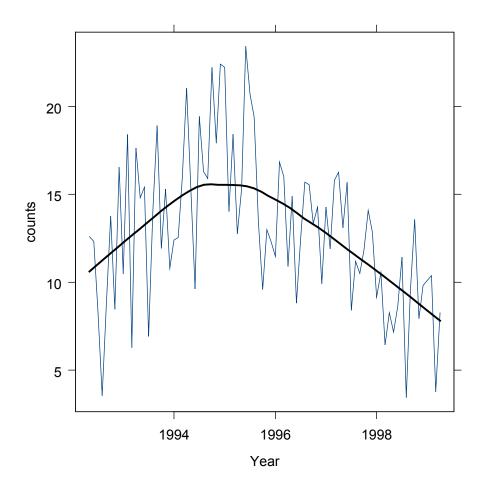


Skewness Coefficient: $g_1 = -0.1394417$, s.l. = 0.288189 Michael's Statistic: $D_sp = 0.05841193$, s.l. = 0.4598983 Wilk-Shapiro Statistic: W = 0.9680623, s.l. = 0.1442419

Figure 3d.

Deseasonalized series. Strong downward trend since 1995. Mann-Kendall test highly significant.

tau = -0.226, sl = 0.2146%

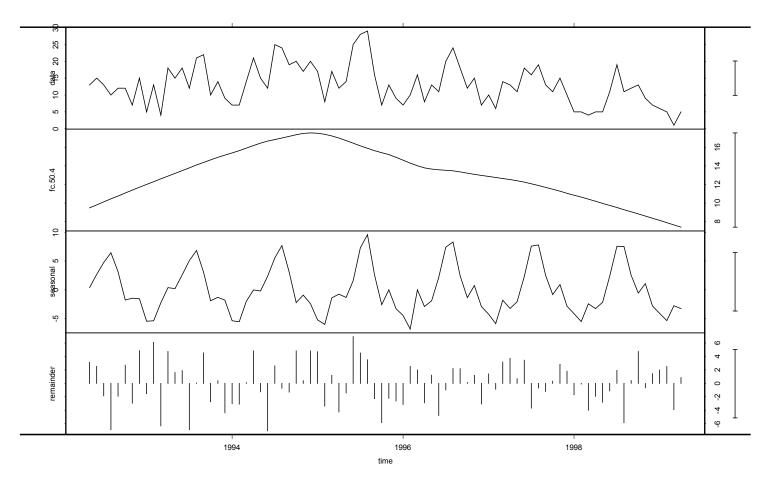


MVA counts (excluding 3AM), deseasonalilzed

STL Decomposition, OTR Excluding 3AMPlus

Figure 4a. A peak is reached in 1995 is followed by a downward trend. There is a change in the shape of the seasonal component occuring in 1996.

Figure 4a.



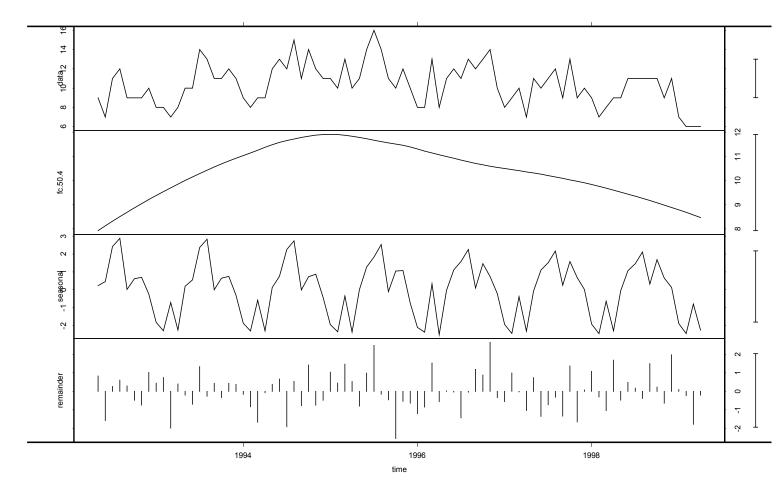
STL, MVA (Excluding 3AMPlus)

ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2

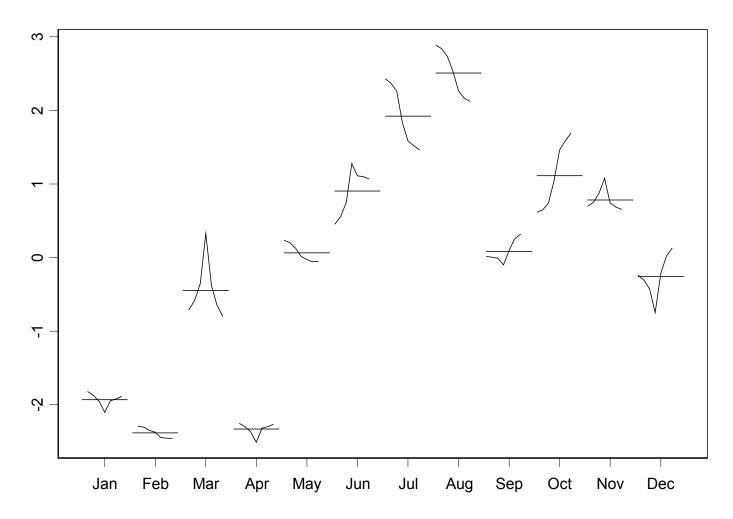
STL Decomposition, OTR 3AMPlus

Figure 4a. Increasing trend to 1995 and decreasing thereafter. The seasonal pattern changes in 1996 with the Jul-Aug peak becoming less prominent and the smaller Oct peak increasing.

STL, MVA 3AMPlus



ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2 Figure 4b. Well defined peaks in Mar, Aug and Oct and minimum in Feb, Apr and Sep. The Oct peak has increased since 1996 and the Aug peak has decreased since 1996.



Seasonal Component, MVA, 3AMPlus

F. OTR Dataset

Summary

The data of both motor vehicle accident and non-motor vehicle trauma cases in Ontario from May 1992 to April 1999. About 45% of OTR cases occur in the 3AMPlus group. Since we are primarily interested in the other groups, the data corresponding to this subgroup was analysed separately.

<u>HOUR</u>: The uniformity hypothesis is rejected by a chi-square test at less than 0.1%. The major departure is due to fewer than expected trauma's in the 12AM window.

<u>June-to-Sep</u> account for 42% so these months have a much higher share. Assuming cases are uniformly distributed over the months, yields an observed chi-squared goodness-of-fit statistic of 73.79 on 11 df which is statistically significant at less than 10^{-10} .

Hour and month are not associated

Hour and year are not associated

<u>Figure 1a</u>. OTR counts/month (excluding 3AMPlus) show a convex pattern of an initial upward trend followed by a downward trend starting about 1995. No monotonic trend so Mann-Kendall not significant as expected - <u>Table 1a</u>.

<u>Figure 1b</u>. OTR counts/month 3AMPlus show an initial upward trend to 1996 and a flattening out. Seasonal Mann-Kendall test is significant at less than 1% indicating an upward trend – <u>Table 1b</u>.

<u>Table 2</u>. Time series disaggregated into hourly windows. No monotonic trends indicated by Seasonal Mann-Kendall tests.

Figure 2a and Figure 2b: nearly flat for 2AM but others show an upward movement followed by downward.

Figure 3a. STL, OTR. A peak is reached in 1995 is followed by a downward trend. There is a change in the shape of the seasonal component occuring in 1996.

Figure 3b. OTR. The seasonal has a well defined peak in August and minimum in February. There is a downward trend in Apr and an upward trend Jul. and Nov This results in two peaks after about 1996 – the second peak being in Nov.

Figure 4a. OTR, 3AMPlus. A gradual increase until it levels off in 1996. Slight change in seasonal shape after 1996

Figure 4b. OTR, 3AMPlus. Well defined peaks in Aug and Oct and minimum in Feb. The Oct peak has increased since 1996 and the Aug peak has decreased since 1996.

Figure 4d – loess analysis of deseasonalized OTR, 3AMPlus is similar to Figure 1b as expected.

Introduction

The data set of Ontario Trauma Registry (OTR) contains the both motor vehicle accident and non motor vehicle accident cases (ISS Scale > 12, severe injury) from May 1992 to April 1999.

The hour factor variable includes 11:00pm-12:00am, 12:00am-1:00am, 1:00am-2:00am, 2:00am - 3:00am and 3:00am -10:00am

Variables of interest:

counts	number of mva cases with ISS Scale > 12 (severe injury)
year	ordered factor, 8 levels: 1992 to 1999
month	ordered factor, 12 levels: January to December
hour	ordered factor with 5 levels: "11PM", "12AM", "1AM", "2AM", "3AMPlus"

About 45% of OTR cases occur in the 3AMPlus group. Since we are primarily interested in the other groups, the data corresponding to this subgroup was analysed separately. Also notice that there is no wkgrp variable for this data.

```
> crosstabs(counts~hour, data=otr.df)
Call:
crosstabs(counts ~ hour, data = otr.df)
3368 cases in table
+---+
| N
     |N/Total|
+----+
hour
    ----+
11PM | 492
           0.15
           ----+
12AM
    | 394
           0.12
           ----+
    | 500
1AM
           |0.15
           ----+
    | 473
           2AM
     10.14
           ----+
3AMPlus|1509
           0.45
           ____+
```

Crosstabs Analysis

Hour

The uniformity hypothesis is rejected by a chi-square test at less than 0.1%. The major departure is due to fewer than expected traumas in the 12AM window.

```
> crosstabs(counts~hour, data=otr1.df)
Call:
crosstabs(counts ~ hour, data = otrl.df)
1859 cases in table
+----+
| N
   |N/Total|
+---+
hour |
----+
   |492
11PM
           0.26
           ____+
12AM
    |394
           0.21
           ----+
    |500
1AM
           0.27
           ----+
2AM
    |473
          0.25
           _____
```

OTR

Hour and month are not associated (p-value=0.8024353)

June-to-Sep account for 42% so these months have a much higher share. Assuming cases are uniformly distributed over the months, yields an observed chi-squared goodness-of-fit statistic of 73.79 on 11 df which is statistically significant at less than 10^{-10} .

There is no dependence between hour and month.

<pre>> crosstabs(counts~hour+month, data=otr.df, subset=!mva.df\$hour=="3AMPlus") Call: crosstabs(counts ~ hour + month, data = otr.df, subset = !(mva.df\$hour == "3AMPlus")) 1859 cases in table ++ N N/RowTotal N/RowTotal N/ColTotal N/Total ++</pre>													
hour	month Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	RowTotl
 11PM	39 0.079 0.300 0.021	34 0.069 0.318 0.018	28 0.057 0.207 0.015	37 0.075 0.278 0.020	39 0.079 0.247 0.021	48 0.098 0.286 0.026	49 0.100 0.241 0.026	49 0.100 0.222 0.026	54 0.110 0.310 0.029	32 0.065 0.225 0.017	46 0.093 0.324 0.025	37 0.075 0.253 0.020	492 0.26
12AM	32 0.081 0.246 0.017	19 0.048 0.178 0.010	30 0.076 0.222 0.016	24 0.061 0.180 0.013	35 0.089 0.222 0.019	34 0.086 0.202 0.018	45 0.114 0.222 0.024	50 0.127 0.226 0.027	33 0.084 0.190 0.018	32 0.081 0.225 0.017	32 0.081 0.225 0.017	28 0.071 0.192 0.015	394 0.21
1AM	31 0.062 0.238 0.017	30 0.060 0.280 0.016	39 0.078 0.289 0.021	46 0.092 0.346 0.025	46 0.092 0.291 0.025	43 0.086 0.256 0.023	60 0.120 0.296 0.032	53 0.106 0.240 0.029	42 0.084 0.241 0.023	38 0.076 0.268 0.020	30 0.060 0.211 0.016	42 0.084 0.288 0.023	500 0.27
2AM	28 0.059 0.215 0.015	24 0.051 0.224 0.013	38 0.080 0.281 0.020	26 0.055 0.195 0.014	38 0.080 0.241 0.020	43 0.091 0.256 0.023	49 0.104 0.241 0.026	69 0.146 0.312 0.037	45 0.095 0.259 0.024	40 0.085 0.282 0.022	34 0.072 0.239 0.018	39 0.082 0.267 0.021	473 0.25
ColTotl	130 0.070	-+ 107 0.058 -+	135 0.073	133 0.072	158 0.085	168 0.090	203 0.109	221 0.119	174 0.094 +	142 0.076	142 0.076	146 0.079	1859 ++

Test for independence of all factors

Chi^2 = 28.33842 d.f.= 33 (p=0.6985065)

Yates' correction not used

Hour and year are independent

There are 62% more in the 11PM-1999 window than might be expected under independence but the result is only significant at about 12.5%.

```
> crosstabs(counts~hour+year, data=otr.df, subset=!otr.df$hour=="3AMPlus")
Call:
crosstabs(counts ~ hour + year, data = otr.df, subset = !(otr.df$hour ==
  "3AMPlus"))
1859 cases in table
+----+
| N |
|N/RowTotal|
[N/ColTotal]
|N/Total |
+----+
hour |year
    _____+
11PM |37 |75 |79 |88 |73 |66 |53 |21 |492 |
     0.0752 0.1524 0.1606 0.1789 0.1484 0.1341 0.1077 0.0427 0.26
     0.2450 |0.3036 |0.2516 |0.2716 |0.2786 |0.2324 |0.2325 |0.4286 |
     |0.0199 |0.0403 |0.0425 |0.0473 |0.0393 |0.0355 |0.0285 |0.0113 |
_____+
12AM |31 |45 |70 |73 |63 |56 |47 | 9 |394 |
     0.0787 |0.1142 |0.1777 |0.1853 |0.1599 |0.1421 |0.1193 |0.0228 |0.21
     0.2053 |0.1822 |0.2229 |0.2253 |0.2405 |0.1972 |0.2061 |0.1837 |
                                                               0.0167 0.0242 0.0377 0.0393 0.0339 0.0301 0.0253 0.0048
                                                               _____+
1AM |42 |70 |90 |95 |60 |77 |54 |12 |500 |
     0.0840 0.1400 0.1800 0.1900 0.1200 0.1540 0.1080 0.0240 0.27
                                                             0.2781 0.2834 0.2866 0.2932 0.2290 0.2711 0.2368 0.2449
     |0.0226 |0.0377 |0.0484 |0.0511 |0.0323 |0.0414 |0.0290 |0.0065 |
_____+
     |41 |57 |75 |68 |66 |85 |74 | 7 |473
2AM
                                                              0.0867 0.1205 0.1586 0.1438 0.1395 0.1797 0.1564 0.0148 0.25
     0.2715 |0.2308 |0.2389 |0.2099 |0.2519 |0.2993 |0.3246 |0.1429 |
     0.0221 0.0307 0.0403 0.0366 0.0355 0.0457 0.0398 0.0038
_____+
ColTot1|151 |247 |314 |324 |262 |284 |228 |49 |1859 |
  |0.081 |0.133 |0.169 |0.174 |0.141 |0.153 |0.123 |0.026 |
                                                              _____+
Test for independence of all factors
  Chi^2 = 28.56446 d.f.= 21 (p=0.1248566)
  Yates' correction not used
> get.crosstabs.percenterror(counts~hour+year)
  [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8]
 \begin{bmatrix} 1 & -7 & 15 & -5 & 3 & 5 & -12 & -12 & 62 \\ [2,] & -3 & -14 & 5 & 6 & 13 & -7 & -3 & -13 \\ [3,] & 3 & 5 & 7 & 9 & -15 & 1 & -12 & -9 \\ [4,] & 7 & -9 & -6 & -18 & -1 & 18 & 28 & -44 \\ \end{bmatrix} 
contribution to chi-sq:
[,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8]
[1,] -0.47 1.19 -0.45 0.24 0.44 -1.06 -0.95 2.23
[2,] -0.18 -1.02 0.42 0.52 1.00 -0.54 -0.19 -0.43
[3,] 0.22 0.44 0.60 0.84 -1.25 0.07 -0.94 -0.32
[4,] 0.42 -0.74 -0.55 -1.59 -0.08 1.50 2.10 -1.55
```

Data Visualization

In our time series plots we use a 60% robust locally linear smoother to visualize the trend except in the STL analysis where a 60% robust locally quadratic smoother is used for trend. Monotonic trend test is significant at <1% for 3AMPlus for not for aggregated 11PM-to-2AM windows. Upward trend for 3AMPlus.

Table 1a.

OTR monthly time series, excluding 3AMPlus window.

> otr	> otr.tot.ts											
	Jan	Feb	Mar	Apr	Мау	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					21	23	18	19	20	19	11	20
1993:	12	20	6	27	22	24	23	26	31	18	20	18
1994:	20	13	26	28	25	15	35	40	29	32	20	31
1995:	25	19	30	23	27	40	37	43	21	17	23	19
1996:	15	21	24	11	22	22	27	38	28	17	21	16
1997:	22	11	27	22	23	29	23	31	24	21	28	23
1998:	20	11	14	9	18	15	40	24	21	18	19	19
1999:	16	12	8	13								
> Seas	> SeasonalMannKendall(otr.tot.ts)											
tau =	-0.()567	, 5	sl =5	548							

Table 1b.

OTR monthly time series, 3AMPlus window.

```
> otr0.tot.ts
       Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov
                                                         Dec
1992:
                          14
                               11
                                    16
                                        17
                                             15
                                                  16
                                                      14
                                                           16
1993:
        13
            12
                 11
                      15
                          15
                               16
                                    23
                                        22
                                             16
                                                  17
                                                      20
                                                           16
1994:
        15
            13
                 14
                      15
                          19
                               22
                                    20
                                        25
                                             21
                                                  23
                                                      19
                                                           16
1995:
        18
            17
                 20
                      18
                          19
                               23
                                    25
                                        25
                                             21
                                                  20
                                                      19
                                                           20
1996:
        16
            15
                 21
                      15
                          21
                               23
                                    22
                                        23
                                             20
                                                  20
                                                      22
                                                           16
1997:
        15
            17
                 17
                      12
                          19
                               20
                                    21
                                        21
                                             18
                                                  22
                                                      14
                                                           17
        15
            14
                      18
                          19
                               20
                                        22
                                             21
                                                  23
                                                           20
1998:
                 14
                                    23
                                                      17
1999:
        16
            12
                 12
                      14
> SeasonalMannKendall(otr0.tot.ts)
tau = 0.249,
                 sl =0.8697%
```

Figure 1a. OTR counts/month (excluding 3AMPlus) show a convex pattern of an initial upward trend followed by a downward trend starting about 1995.

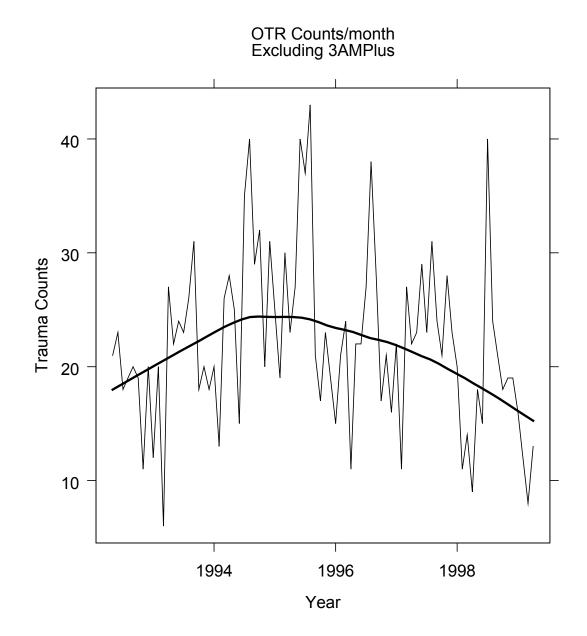
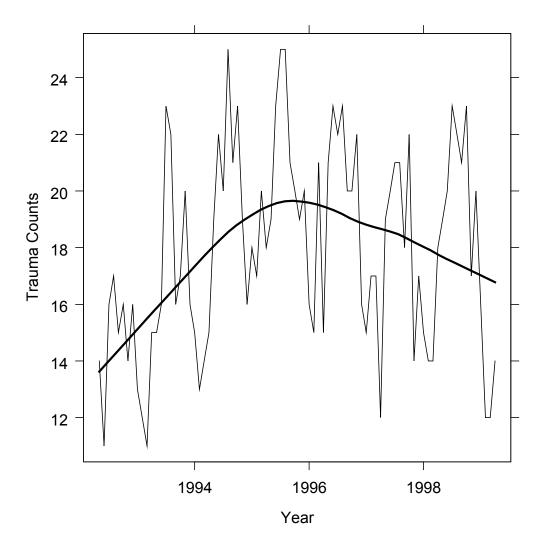
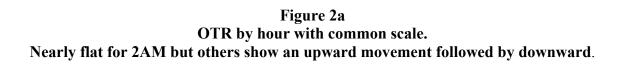
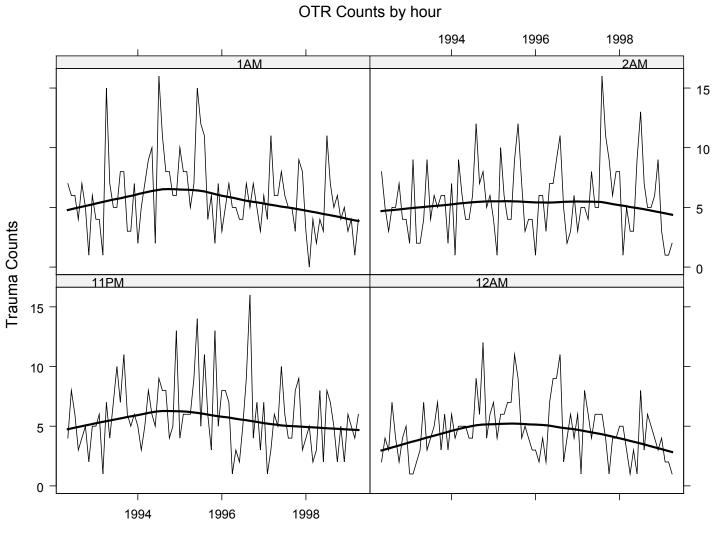


Figure 1b. OTR counts/month 3AMPlus show an initial upward trend to 1996 and a less step downward trend post-1996.



OTR Counts/month for 3AMPlus

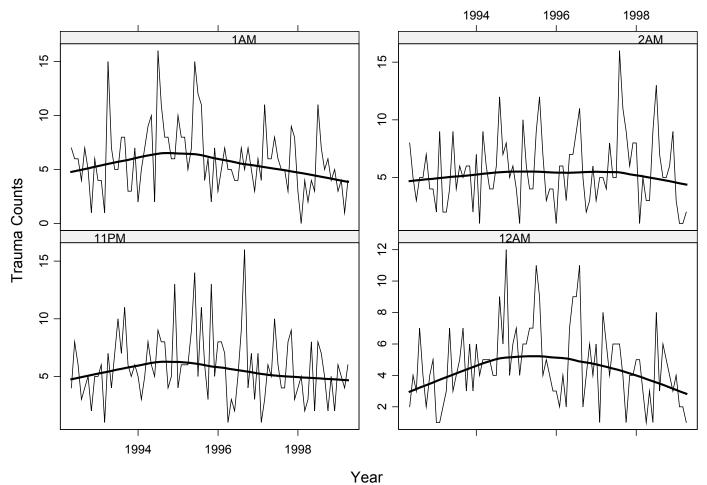




Year

9

Figure 2b. Independent scaling.



OTR Counts by hour

Independent scales

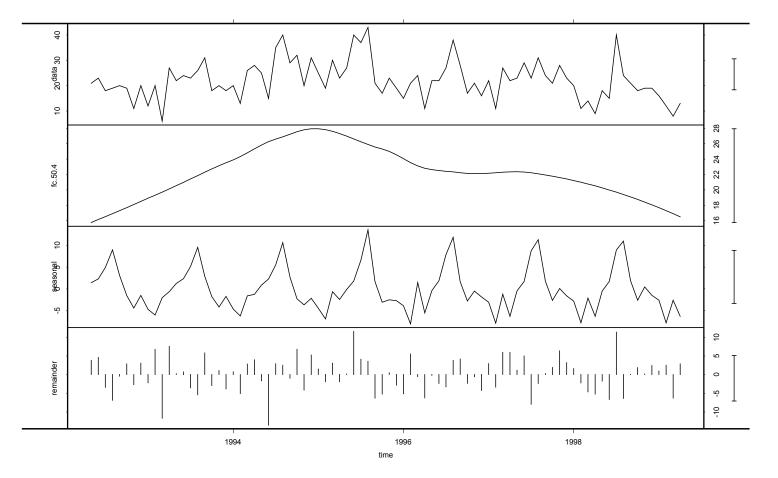
OTR

> count.11PM.ts												
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					4	8	6	3	4	5	2	5
1993:	5	6	1	7	4	7	10	7	11	6	5	6
1994:	5	3	5	8	6	5	9	8	8	4	5	13
1995:	4	6	6	6	9	14	5	11	6	3	13	5
1996:	8	8	7	1	3	2	5	9	16	4	7	3
1997:	7	1	3	6	5	10	6	4	4	8	9	3
1998:	4	5	2	3	8	2	8	7	5	2	5	2
1999:	6	5	4	6								
> Seas					(cou	nt.11	1PM.	ts)				
tau =								,				
> cour												
0001				Apr	Mav	מווד	ر 11 ¹	Aug	Sen		Nov	Dec
1992:	Juii	1.00	11011	L	2	4	3	7 Tug	зср 4	2	4	5
1993:	1	1	2	3	7	3	4	, 5	- 7	3	г 6	3
1993.	1 6	4	5	5	5		4		6	12	4	6
1995:	7	4		6	7		11	9	4	5	4	3
1996:	3	2	4	2	7	9	9		2	4		4
1990.	6	1	8	6	4		6		4	1	4	4
	5	5	° 3	0 1	43	0 1	8		4	1 5	4	43
1998:	-				2	T	0	2	0	5	4	2
1999: 4 2 2 1												
<pre>> SeasonalMannKendall(count.12AM.ts) tau = -0.0299, sl =75.19%</pre>												
				51 —	/3.1	96						
> cour				7	Moss	T	T 1	7	0.0.00	0~+	Nore	Dee
1000.	Jan	гер	Mar	Apr				Aug				
1992:	л	1	1	1 5	7	6	6	4	7	5	1	6
1993:	4	4	1	15	7	5	5	8	8	3	3	7
1994:	2	5	7	9	10	2	16		8	8	6	6
1995:	10	8		5	7	-	12		4	6		7
1996:	3	5	7	5	5	4	4		5	7	-	3
1997:	6	4	11	6	6	8	6		5	3	9	8
1998:	3	0	4	2	4	3	11	7	5	6	4	5
1999:	3	4	1	4								
> Seas							AM.t:	s)				
tau =				L =9	.095	01						
> cour							_					
	Jan	Feb	Mar	Apr				Aug				
1992:					8	5	3	5	5	7	4	4
1993:	2	9	2	2	4	9	4		5	6	6	2
1994:	7	1	9	6	4	4	6		7	8	5	6
1995:	4	1	10	6	4	4	9		7	3	4	4
1996:	1	6	6	3	7	7	9	11	5	2	3	6
1997:	3	5	5	4	8	5	5	16	11	9	6	8
1998:	8	1	5	3	3	9	13	7	5	5	6	9
1999:	3	1	1	2								
> Seas				dall	(coui	nt.27	AM.t:	s)				
tau =					9.99 ⁹							
		•										

Table 2. Data tabulations and Seasonal Mann-Kendall testsNo monotonic trend significant at 5%

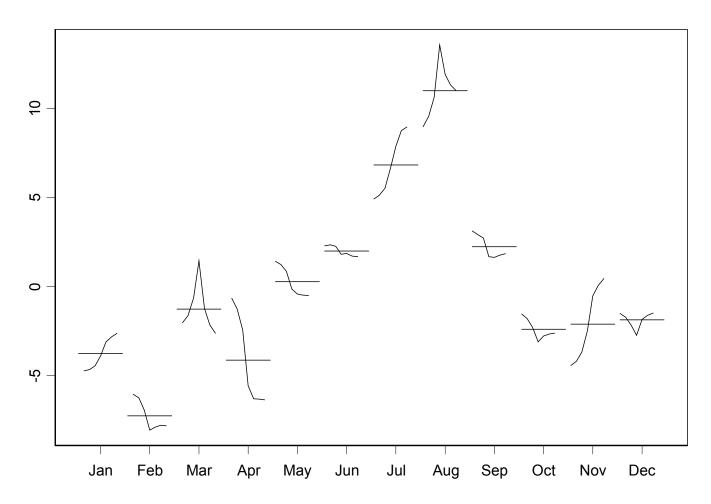
STL Decomposition, OTR Excluding 3AMPlus

Figure 3a. A peak is reached in 1995 is followed by a downward trend. There is a change in the shape of the seasonal component occuring in 1996. R-sq = 62.5 %

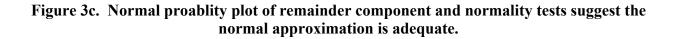


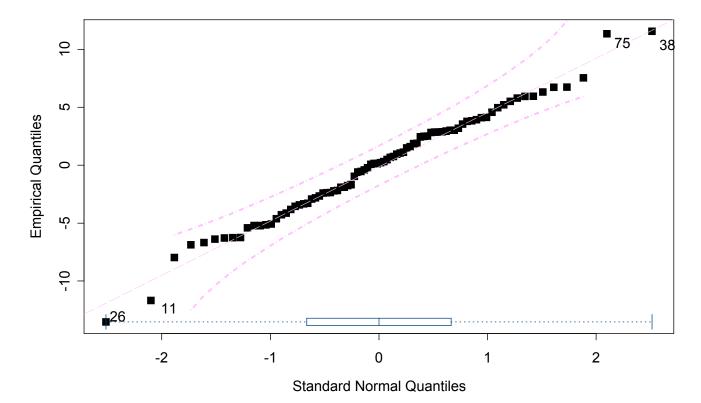
STL, OTR (Excluding 3AMPlus)

ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2 **Figure 3b**. The seasonal has a well defined peak in August and minimum in February. There is a downward trend in Apr and an upward trend Jul. and Nov This results in two peaks after about 1996 – the second peak being in Nov.



Seasonal Component, Excluding 3AMPlus

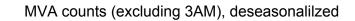


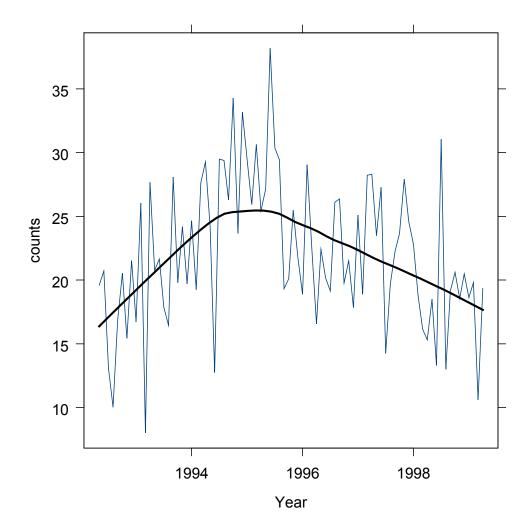


Skewness Coefficient: $g_1 = -0.1499071$, s.l. = 0.2741649 Michael's Statistic: D_sp = 0.04209311, s.l. = 0.5012842 Wilk-Shapiro Statistic: W = 0.9856928, s.l. = 0.8398335

Figure 3d. Loess trend analysis of deseasonalized data and Mann-Kendall test.

tau = -0.0666, sl = 36.92%



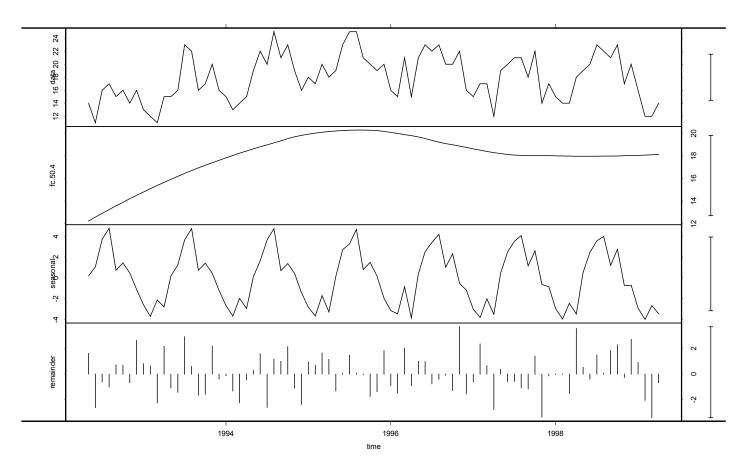


OTR

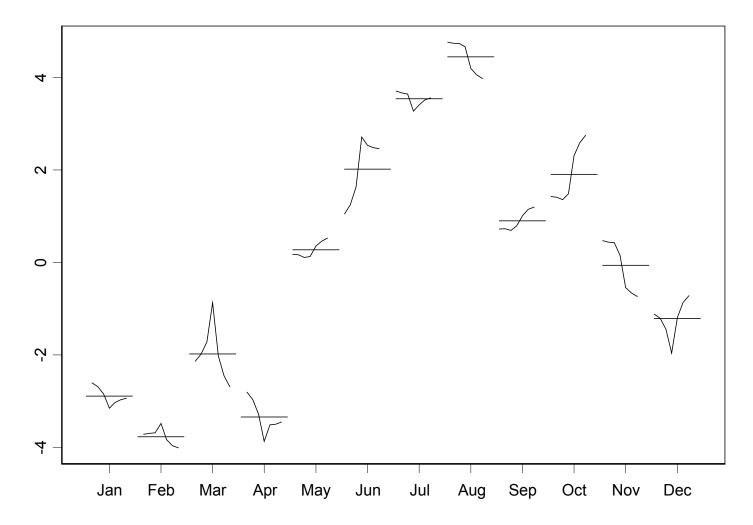
STL Decomposition, OTR 3AMPlus

Figure 4a. A gradual increase until it levels off in 1996. Slight change in seasonal shape after 1996. Seasonal peak has decreased. R-sq = 80.3 %. There is a significant upward trend.

STL, OTR 3AMPlus

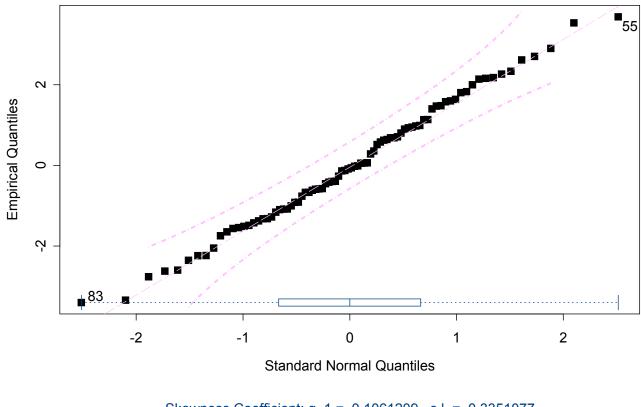


ss.window = 7 ,ss.robust = TRUE , fc.window = 50.4 , fc.degree = 2 **Figure 4b.** Well defined peaks in Aug and Oct and minimum in Feb. The Oct peak has increased since 1996 and the Aug peak has decreased since 1996.



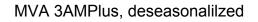
Seasonal Component, 3AMPlus

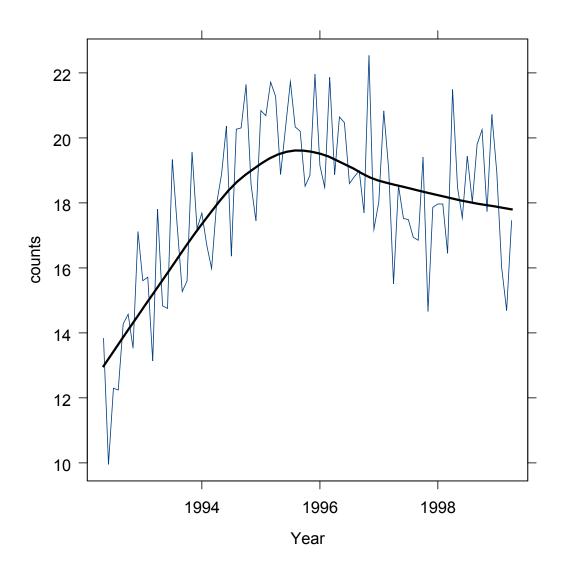
Figure 4c. Normal probability plot – no outiers or non-normality evident.



Skewness Coefficient: $g_1 = 0.1061209$, s.l. = 0.3351077 Michael's Statistic: $D_sp = 0.03207909$, s.l. = 0.06057737 Wilk-Shapiro Statistic: W = 0.9789253, s.l. = 0.5259117

Figure 4d.Trend analysis of deaseasonalized series. Upward trend detected but it levels off.tau = 0.212, sl = 0.4091%





G. LPS Dataset

Introduction

London Police Data set contains 8920 observations regarding with person being charged due to assault or impaired driving from May 1992 to May 1999. There are 3257 assault charges and 5663 impaired driving charges.

And only cases from 11:00pm to 4:00am are analyzed. The time of charge happened, such as hour, week, month and year, was also recorded.

Variables of interests in the data set

persons	person number (1 for first person, 2 for second, etc)
counts	number of charges
type	factor with 2 levels: "assault", "impaired"
wkgrp	ordered factor with 2 levels: "SunWed", "ThuSat"
hour	ordered factor with 5 levels: "11PM", "12AM", "1AM", "2AM", "3AM"
year	ordered factor with 8 levels: 1992 to 1999
month	ordered factor with 12 levels: January to December

The distribution of the number of person charged is shown in the table below.

> table(lps.df\$persons,lps.df\$type)

persons	assault	impaired
1	3079	5654
2	138	9
3	35	0
4	4	0
5	1	0

The following table shows the 4 incidents where 4 persons were charged with assault. > lps.df[indl&ind2,]

	persons	counts	type	concord	wkgrp	hour	month	day	year	charges
7517	4	1	assault	1420	SunWed	12AM	Feb	Thu	1996	4
8093	4	2	assault	1430	ThuSat	12AM	Sep	Sat	1998	8
8094	4	2	assault	1430	ThuSat	12AM	Sep	Sat	1998	8
8102	4	1	assault	1420	SunWed	ЗАМ	Sep	Mon	1998	4

The response variable chosen to analyze was counts.

G. LPS Dataset: Assaults

Summary

The LPS assault data include sexual assault. There were 3257 assault cases in London from May of 1992 to May of 1999. <u>Table 1</u>. There was an extreme outlier in <u>March 1999</u>. To improve data visualization a log transformation has been used.

<u>Crosstabs</u>. Assaults are more frequent ThuSat (62%) than in SunWed (38%). Assaults from 1:00am to 2:00am count for most cases (24%) while 3:00am to 4:00am only 13%. Assaults are more frequent ThuSat (63%) than in SunWed (37%). In SunWed slot, there are 17% more assaults than expected in 11PM and 11% and 17% fewer than expected in the 2AM and 3AM slots respectively. As might be expected from the loess analysis, hour and year are related. See Figure 3b.

Figure 1a. and Figure 1b. Assaults decreased slighly to about mid-1994 and have remained flat since.

Figure 2. Since 1996 there has been a slight decline in SunWed and slight increase in ThuSat.

<u>Figure 3b</u>. Loess analysis by hour show there is an upward trend in 2AM and 3AM and a downward trend in 1AM and 12AM. Monotonic trend tests are significant for 2AM and 1AM. The trend at 1AM accelerates after 1996.

Figure 4b From Figure 4b we see that since 1996 there has been an upward shift at 2AM and 3AM and a downward trend at 1AM for both SunWed and ThuSat. For 11PM-SunWed the downward trend in progress before 1996 flattened out after 1996 while for 11PM-ThuSat this downward trend switched to an increase after 1996. For 12AM-SunWed there is a slight downward trend for post-1996 and an apparent shift downward for 12AM-ThuSat.

Figure 5a. STL: The logged data is used. The seasonal component is very unstable and increases in amplitude after 1996. R-sq = 52.5%

<u>Figure 5b</u>. The overall seasonal pattern is with a March maximum and with November minimum. Large trend upward in March and large downward in May. Smaller downward trends in October and November. July and August are relatively high and don't vary much.

Figure 5c. Tests for normality are fine. The two largest and smallest values in the normal probability plot are indicated -- see also <u>Table 1</u>.

Crosstabs Analysis

Hour and wkgrp are associated

Assaults are more frequent ThuSat (62%) than in SunWed (38%). Assaults from 1:00am to 2:00am count for most cases (24%) while 3:00am to 4:00am only 13%. Assaults are more frequent ThuSat (63%) than in SunWed (37%). In SunWed slot, there are 17% more assaults than expected in 11PM and 11% and 17% fewer than expected in the 2AM and 3AM slots respectively.

```
Call:
crosstabs(formula = counts ~ wkgrp + hour, data = lps.df, subset = lps.df$type ==
"assault",
    na.action = na.exclude)
3301 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
wkgrp |hour
    |11PM
          |12AM |1AM |2AM |3AM
                                 |RowTotl|
SunWed |265 |321 |287 |231 |136 |1240 |
    |0.21 |0.26 |0.23 |0.19 |0.11
                                 10.38
                                       0.44 0.4 0.37 0.34 0.31
                                  0.08 0.097 0.087 0.07 0.041
_____+
ThuSat |340 |477 |486 |458 |300 |2061
    0.16 0.23 0.24 0.22 0.15
                                0.62
    |0.56 |0.6 |0.63 |0.66 |0.69 |
    0.1
          |0.14 |0.15 |0.14 |0.091 |
ColTotl|605 |798 |773 |689 |436 |3301
|0.18 |0.24 |0.23 |0.21 |0.13 |
                                       1
Test for independence of all factors
    Chi^2 = 24.84403 d.f.= 4 (p=0.00005407599)
    Yates' correction not used
> get.crosstabs.percenterror(counts~wkgrp+hour)
Crosstabs - percentage error: 100*(Obs-Exp)/Exp,
    [,1] [,2] [,3] [,4] [,5]
    17 7 -1 -11 -17
[1,]
            1 6 10
[2,] -10 -4
> contrib(counts~wkgrp+hour)
contribution to chi-sq:
    [,1] [,2] [,3] [,4] [,5]
[1,] 2.50 1.23 -0.20 -1.73 -2.17
[2,] -1.94 -0.95 0.15 1.34 1.68
```

Hour and Year are associated

As might be expected from the loess analysis, hour and year are related. See Figure 3b.

> crosstabs(counts~hour+year, data=lps.df, na.action=na.exclude, subset=lps.df\$type=="assault") Call: crosstabs(formula = counts ~ hour + year, data = lps.df, subset = lps.df\$type == "assault", na.action = na.exclude) 3301 cases in table +----+ | N |N/RowTotal| |N/ColTotal| IN/Total | +----+ hour |year |1993 |1994 |1995 |1996 |1997 |1998 |1999 |RowTotl| 11992
 11PM
 I
 66
 83
 80
 70
 1106
 64
 1102
 34
 605

 I0.11
 I0.14
 I0.13
 I0.12
 I0.18
 I0.11
 I0.17
 I0.056
 I0.18

 I0.2
 I0.17
 I0.19
 I0.16
 I0.22
 I0.15
 I0.22
 I0.15
 I
 0.02 0.025 0.024 0.021 0.032 0.019 0.031 0.01 _____+ 12AM | 62 | 112 | 109 | 119 | 131 | 90 | 98 | 77 | 798 0.078 0.14 0.14 0.15 0.16 0.11 0.12 0.096 0.24 |0.19|0.23|0.26|0.27|0.27|0.21|0.21|0.35||0.019|0.034|0.033|0.036|0.04|0.027|0.03|0.023| _____+
 1AM
 |103
 |138
 | 95
 |114
 |109
 | 95
 | 88
 | 31
 | 773

 |0.13
 |0.18
 |0.12
 |0.15
 |0.14
 |0.12
 |0.11
 |0.04
 |0.23

 |0.31
 |0.28
 |0.23
 |0.26
 |0.22
 |0.19
 |0.14
 |
 0.031 0.042 0.029 0.035 0.033 0.029 0.027 0.0094
 2AM
 56
 91
 79
 97
 95
 116
 112
 43
 689

 |0.081
 |0.13
 |0.11
 |0.14
 |0.14
 |0.17
 |0.16
 |0.062
 |0.21

 |0.17
 |0.18
 |0.19
 |0.22
 |0.19
 |0.27
 |0.24
 |0.19
 |

 |0.017
 |0.028
 |0.024
 |0.029
 |0.029
 |0.035
 |0.034
 |0.013
 |

 _____+

 | 48
 | 69
 | 57
 | 44
 | 47
 | 61
 | 73
 | 37
 | 436

 | 0.11
 | 0.16
 | 0.13
 | 0.11
 | 0.14
 | 0.17
 | 0.085
 | 0.13

 | 0.14
 | 0.14
 | 0.14
 | 0.14
 | 0.15
 | 0.17
 |

 | 0.015
 | 0.021
 | 0.017
 | 0.013
 | 0.014
 | 0.018
 | 0.022
 | 0.011

 3AM |0.085 |0.13 _____+
 ColTotl|335
 |493
 |420
 |444
 |488
 |426
 |473
 |222
 |3301

 |0.1
 |0.15
 |0.13
 |0.15
 |0.13
 |0.15
 |0.13
 |0.15
 |0.13
 |0.15
 |0.13
 |0.15
 |0.13
 |0.15
 |0.13
 |0.14
 |0.067
 |
 Test for independence of all factors Chi^2 = 90.20904 d.f.= 28 (p=1.838906e-008) Yates' correction not used > get.crosstabs.percenterror(counts~hour+year) Crosstabs - percentage error: 100*(Obs-Exp)/Exp, [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] $\begin{bmatrix} 1, 1 \end{bmatrix} & 7 & -8 & 4 & -14 & 19 & -18 & 18 & -16 \\ \begin{bmatrix} 2, 1 \end{bmatrix} & -23 & -6 & 7 & 11 & 11 & -13 & -14 & 43 \\ \begin{bmatrix} 3, 1 \end{bmatrix} & 31 & 20 & -3 & 10 & -5 & -5 & -21 & -40 \\ \begin{bmatrix} 4, 1 \end{bmatrix} & -20 & -12 & -10 & 5 & -7 & 30 & 13 & -7 \\ \end{bmatrix}$ 8 6 3 -25 -27 8 17 26 [5,] > contrib(counts~hour+year) contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [1,] 0.59 -0.77 0.34 -1.26 1.75 -1.59 1.64 -1.05 [2,] -2.11 -0.66 0.74 1.13 1.20 -1.28 -1.53 **3.18** [3,] **2.77** 2.10 -0.34 0.98 -0.49 -0.48 -2.16 -2.91 [4,] -1.67 -1.17 -0.93 0.45 -0.68 **2.87** 1.34 -0.49 [5,] 0.56 0.48 0.20 -1.91 -2.17 0.63 1.33 1.42

4

LPS Assault Data

There is a mild decrease from 1992 to mid-1994 and then leveling off. There is an exception outlier occurring in March of 1999. The two largest and smallest values in the normal probability plot of the remainder component in the STL analysis of the logged series are shown in bold -- see <u>Figure 5c</u>.

Table 1.

> assa	> assault.ts											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					57	38	44	45	38	42	39	32
1993:	49	32	32	34	34	55	42	49	49	44	30	43
1994:	32	29	35	32	40	33	51	31	36	35	34	32
1995:	28	31	37	35	56	33	45	52	28	29	28	42
1996:	37	51	63	30	28	44	34	40	43	43	31	44
1997:	31	22	48	40	27	44	50	47	21	30	30	36
1998:	32	35	45	46	36	47	43	46	39	33	32	39
1999:	45	35	76	33	33							
> Seas	> SeasonalMannKendall(assault.ts)											
tau =	0.01	195,	s	L =83	3.07%	5						

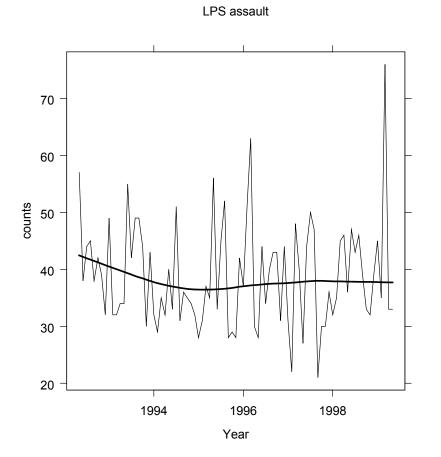
March 1999

Investigating why March 1999 had 76 assault charges, we found that there were numerous cases of large numbers of assaults occuring on ThuSat-12AM slot. There were a total of 74 records corresponding to March 1999 in the original database and 50 of these records correspond to ThuSat. The distribution is summarized below:

counts	1	2	6	9
# records	44	2	2	2

Because ThuSat-12AM-March 1999 is such a large outlier, all crosstabs involving these categories are highly significant at much less than 1%.

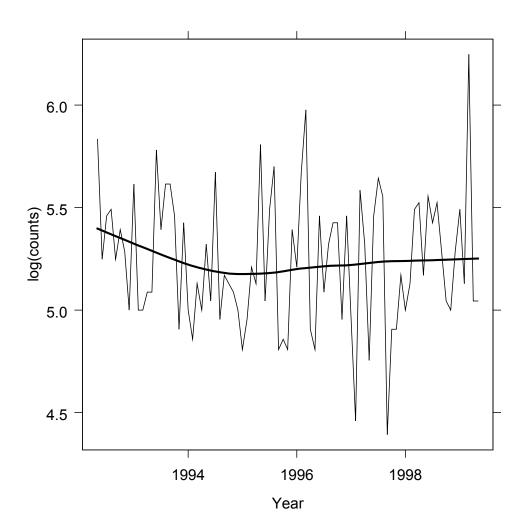
Figure 1.



6

Figure 1b. Logged series

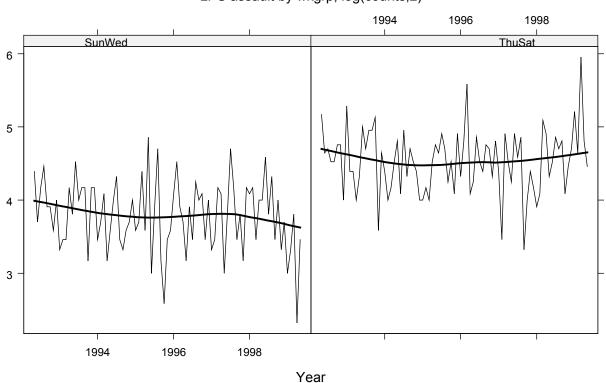
The outlier in March 1999 has the effect of squishing the data into a smaller area on the graph. Graphical resolution is improved by using a log transformation. Specifically a log to the base 2 transformation. This particular version of the log transformation means that we can interpret a one unit change as a doubling.



LPS assault, log(counts,2)

Figure 2. By wkgrp.

There is a very small downward trend in SunWed group.



LPS assault by wkgrp, log(counts,2)

Table 2.

> assault.SunWed.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1993: 10 11 11 1994: 11 13 17 1995: 12 29 1996: 17 23 15 13 9 15 11 19 16 17 1997: 10 11 18 17 8 14 26 18 11 14 18 11 16 16 24 14 20 11 16 10 1998: 17 1999: 8 10 14 5 11 > SeasonalMannKendall(assault.SunWed.ts) tau = -0.131, sl = 15.57% > assault.ThuSat.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 23 23 1993: 39 21 16 20 1994: 21 18 23 20 26 1995: 16 16 23 27 26 19 1996: 20 28 48 17 19 21 27 26 30 23 19 30 24 29 10 16 1997: 21 11 1998: 15 17 34 30 20 23 29 26 28 17 22 1999: 37 25 62 28 22 > SeasonalMannKendall(assault.ThuSat.ts) tau = 0.0277, sl = 76.33%

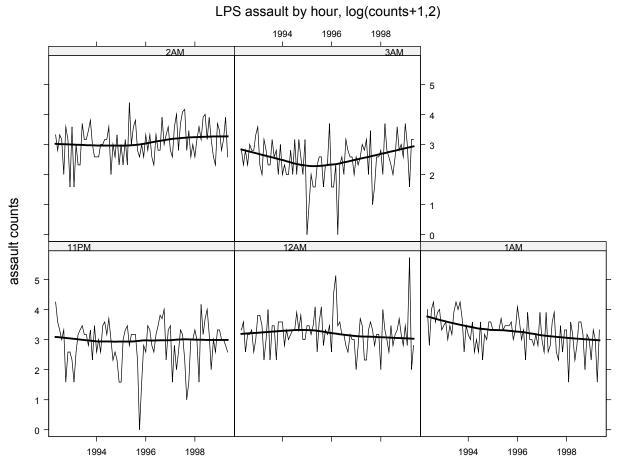


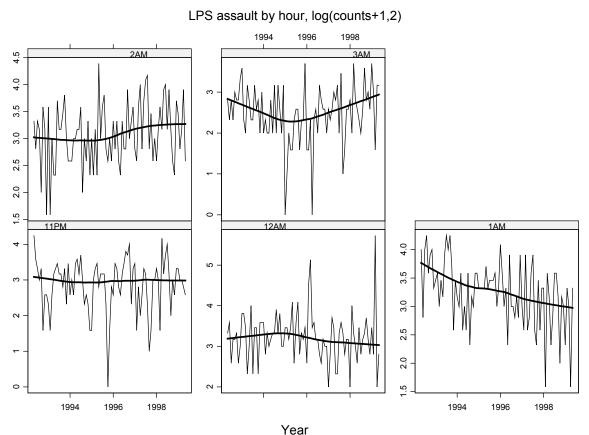
Figure 3a. By hour. Common Scaling

Year

LPS ASSAULTS

Figure 3b. By hour. Independent vertical scaling.

Loess analysis by hour show there is an upward trend in 2AM and 3AM and a downward trend in 1AM and 12AM. Monotonic trend tests are significant for 2AM and 1AM. The trend at 1AM accelerates after 1996.

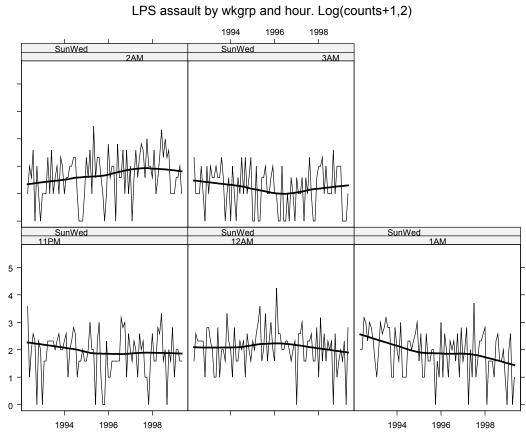


free vertical scaling

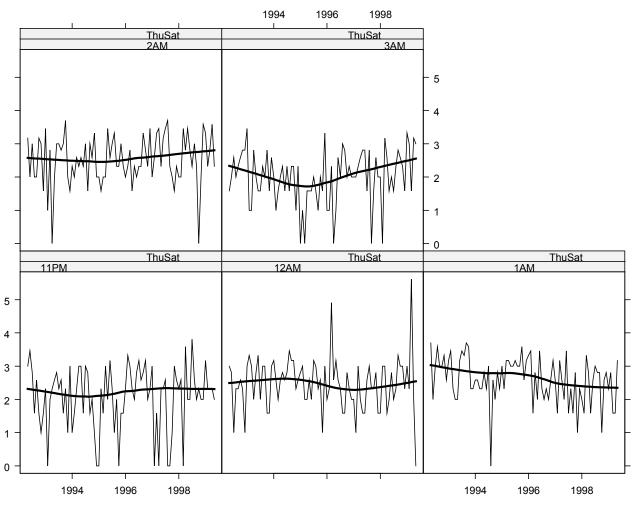
Tabulation and Mann-Kendall Tests

						18	abula	ation	i and	I IVI8	inn-I	Ken	18
> ass													
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct		Dec	
1992:					18	11	9	7	9	2	5	5	
1993:	4	2	5	8	9	10	8	8	6	9	4	10	
1994:	5	7	5	10	11	8	12	7	4	5	4	2	
1995:	2	7	9	10	6	8	8	8	4	0	2	6	
1996:	5	10	9	6	5	8	10	13	12	15	4	9	
1997:	10	2	6	3	5	9	8	4	1	2	7	7	
1998:	9	7	2	17	8	12	15	8	3	7	5	9	
1999:	9	7	7	6	5								
> Sea	sonal	lManı	nKend	dall	(assa	ault	.11 PI	M.ts))				
tau =	0.01	166,	s	1 =8	68								
> ass													
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:				-	9	11	5	8	8	9	5	7	
1993:	13	13	10	4	7	15	4	10	10	4	11	11	
1994:	11	6	9	7	8	9	9	14	9	13	7	7	
1995:	10	10	8	10	16	5	11	16	6	9	8	10	
1996:	- 5	22	34	10	11	8	8	- 0	5	8	7	- 0	
1997:	3	6	12	10	4	4	9	11	9	6	8	8	
1998:	3	15	9	7	5	10	6	8	9	12	8	6	
1999:	10	6	52	3	6	τu	0	0	2	12	0	0	
> Sea:							12 מ	vī + e'	\ \				
tau =					(assa 5.72%		• 12AI	M. CO,	,				
> ass				L -Z.	J. / Z1	δ							
/ d55				7	More	Turn	т., 1	7	Com	Oat	Nor	Dog	
1002.	Jan	гер	Mar	Apr	May 15			-	-				
1992:	1.0	1 1	7	1.0		6	15	18	11	14	15	9	
1993:	10	11	7	10	8	14	18	15	18	12	5	10	
1994:	8	7	11	5	7	5	11	4	8	7	11	11	
1995:	9	9	9	9	12	9	10	10	10	11	7	9	
1996:	16	11	7	9	4	14	7	7	6	8	6	14	
1997:	7	5	14	5	6	11	14	5	4	10	5	9	
1998:	9	2	11	8	4	6	11	11	8	3	8	7	
1999:	4	9	7	2	9								
> Sea							.1AM	.ts)					
tau =				=0.0	0233	78							
> ass													
	Jan	Feb	Mar	Apr	May			-	-				
1992:					9	6	9	8	3	11	8	2	
1993:	11	2	7	4	4	12	8	8	10	13	7	5	
1994:	5	5	7	7	8	8	11	3	7	5	9	4	
1995:	7	4	8	4	20	7	11	13	6	5	7	5	
1996:	9	6	9	5	4	9	6	6	14	7	9	11	
1997:	6	5	11	15	6	12	16	17	6	10	5	7	
1998:	5	8	11	8	14	15	8	14	8	5	4	12	
1999:	10	6	8	14	5								
> Sea	sonal	lManr	nKend	dall	(assa	ault	.2AM	.ts)					
tau =	0.24	41,	sl	=0.	93759	6							
> ass	ault	. 3AM	.ts										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:				-	6	4	6	4	-7	6	6	9	
1993:	11	4	3	8	6	4	4	8	5	6	3	7	
1994:	3	4	3	3	6	3	8	3	8	5	3	8	
1995:	0	1	3	2	2	4	5		2	4	4	12	
1996:	2	2	4	0	4	5	3	8	6	5	5	3	
1997:	5	4	5	7	6	8	3	10	1	2		5	
1998:	6	3	12	6	5	4	3	5	11	6	7	5	
1999:	12	7	2	8	8	-	5	0		0	,	0	
> Sea:							ЗДМ	t s)					
tau =					8.04		• • • • • • •	,					
		,		1		-							

Figure 4a. By hour and wkgrp. Common Scaling.



Year



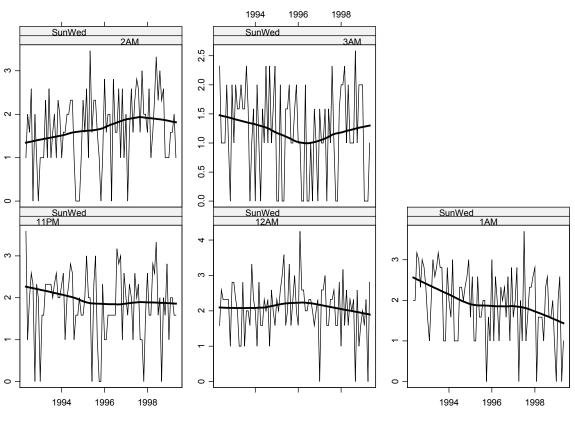
LPS assault by wkgrp and hour. Log(counts+1,2)

Year

LPS ASSAULTS

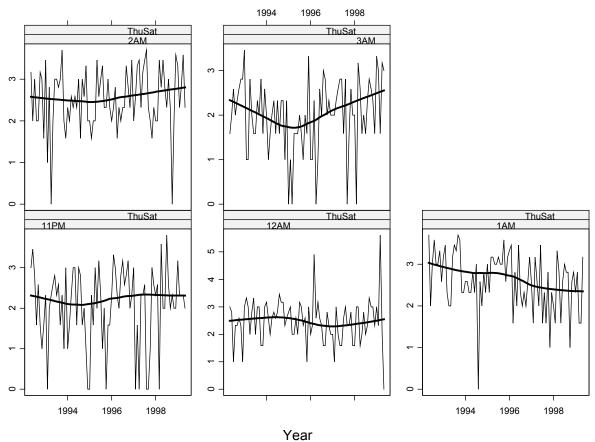
Figure 4b. By hour and wkgrp. Independent Scaling.

From Figure 4b we see that since 1996 there has been an upward shift at 2AM and 3AM and a downward trend at 1AM for both SunWed and ThuSat. For 11PM-SunWed the downward trend in progress before 1996 flattened out after 1996 while for 11PM-ThuSat this downward trend switched to an increase after 1996. For 12AM-SunWed there is a slight downward trend for post-1996 and an apparent shift downward for 12AM-ThuSat.



LPS assault by wkgrp and hour, log(counts+1,2)

Year free vertical scale



LPS assault by wkgrp and hour, log(counts+1,2)

free vertical scale

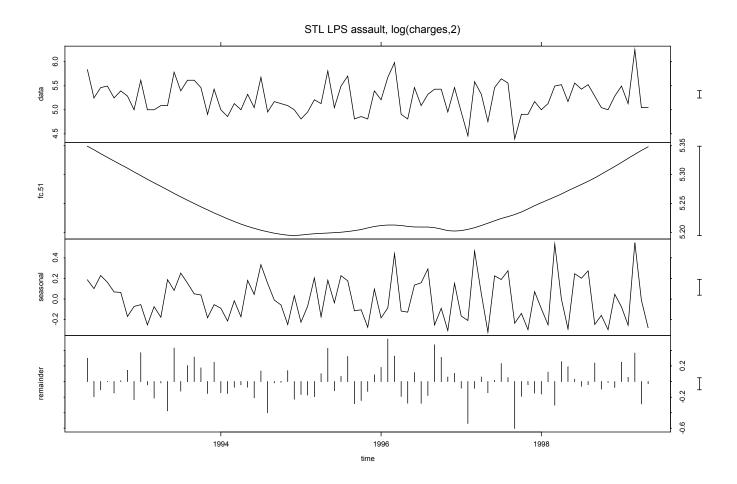
Tabulations and Seasonal Mann-Kendall Tests

	-1-	11	1 C T	ت م م								
> assaı						T	T 1	7	0.000	0	Marr	Dee
	Jan	гер	Mar	Apr	-			Aug	-	Oct		Dec
1992:	0	~	~	4	11	1	3	5	4	0	4	3
1993:	0	2	2	4	4	4	4	3	4	5	3	3
1994:	4	5 3	1	3 3	4	6	5	1	2	2	3	2
1995:	2		7		3	0	4	7	1	0	0	4
1996:	1	1	2	2	2	2	2	8	6	7	1	5
1997:	3	2	4	3	1	5	3		1	1	0	2
1998:	5	2	2	6	5	9	2	3	0	3	2	6
1999:	1	3	3	2	2		11.0			-)		
> Seaso					(assa 33.74		• I I PI	MSum	wea.	LS)		
tau = ·						10						
> assaı						T	т., 1	7	0.0.00	0+	Marr	Dee
	Jan	гер	Mar	Apr	-			Aug	-			Dec
1992:	C	1	2	1	2	5	4	4	4	4 2	1	6
1993:	6	4	3	1	1	6	1	3	3		9	4
1994:	3	1	6	2	2	4	3	4	1	5	3	2
1995:	4	3	5	7	11	2	3	9	2	4	7	3
1996:	2	18	5	5	3	3	4			2	3	4
1997:	0	5	5	7	2	2	4	4	5	2	2	6
1998:	1	8	2	5	2	4	3	4	0	5	1	2
1999:	3	2	4	0	6	_						
> Seaso							.12AI	MSunI	Wed.	ts)		
tau = ·					32.52	28						
> assaı												
	Jan	Feb	Mar	Apr	-			Aug	-			Dec
1992:					3	3	8	7	4	7	6	4
1993:	2	1	3	7	5	6	8	6	6	1	1	6
1994:	3	2	7	1	1	1	4	4	3	4	5	7
1995:	2	5	1	1	5	2	2	3	3	0	2	1
1996:	7	1	5	3	1	4	3	4	2	5	1	6
1997:	2	3	6	0	3	1	12	1	2	4	4	5
1998:	6	0	2	2	2	1	4	5	2	2	3	1
1999:	0	3	5	0	1							
> Seaso							.1AM	Sun₩	ed.t	5)		
tau = ·					. 2954	1 8						
> assaı												
	Jan	Feb	Mar	Apr	May			Aug	Sep	Oct		Dec
1992:					1	3	2	5	0	3	1	0
1993:	1	1	1	4	1	5	1	2	3	1	4	3
1994:	1	2	2	3	3	4	4	1	0	0	0	1
1995:	4	2	5	1	10	2	4		2	1	0	1
1996:	6	2	3	3	0	6	2	2	5	1	5	1
1997:	3	0	2	5	2	4	6	5	2	7	3	3
1998:	2	5	1	2	4	9	4	7	4	5	1	1
1999:	1	2	2	3	1							
> Seaso	onal	Manı	nKend	dall	(assa	ault	.2AM	Sun₩	ed.t	з)		
tau = 0				=2.								
> assat	ult.	. 3AMS	SunWe	ed.t	5							
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					4	1	1	1	3	1	0	3
1993:	1	3	2	2	3	2	2	4	2	0	1	2
1994:	0	3	1	0	2	1	4	1	4	1	2	4
1995:	0	0	3	0	0	2	2	3	1	1	2	3
1996:	1	1	0	0	3	0	0	1	0	2	1	0
1997:	2	1	1	2	0	2	1	4	1	0	0	2
1998:	3	3	4	1	3	1	1	1	5	1	3	3
1999:	3	0	0	0	1							
> Seaso							.3AM	SunWe	ed.t	3)		
tau = ·	-0.0)434,	, 5	sl =	65.08	38						

assaul	1+ 11			- + 0								
assau.			Mar		Matz	Tun	.Tu 1	Aur	Son	Oct	Nov	Dec
1992:	oun	100	mar	прт	7	10	6	2	5 S	2	1	2
1993:	4	0	3	4	5	6	4	5	2	4	1	7
1994:	1	2	4	7	7	2	7	6	2	3	1	0
1995:	0	4	2	7	3	8	4	1	3	0	2	2
1996:	4	9	7	4	3	6	8	5	6	8	3	4
1997:	7	0	2	0	4	4	5	0	0	1	7	5
1998:	4	5	0	11	3	3	13	5	3	4	3	3
1999:	8	4	4	4	3							
> Seas	sonal	LManr	nKend	dall	(assa	ault	.11PI	MThu	Sat.	ts)		
tau =	0.05	577 ,	sl	L =53	3.799	20						
> assa												
	Jan	Feb	Mar	Apr	-	Jun		-	-		Nov	Dec
1992:					7	6	1	4	4	5	4	1
1993:	7	9	7	3	6	9	3	7	7	2	2	7
1994:	8	5	3	5	6	5	6	10	8	8	4	5
1995:	6	7	3	3	5	3	8	7	4	5	1	7
1996:	3	4	29	5	8	5	4	2	2	6	4	3
1997:	3	1	7	3	2	2	5	7	4	4	6	2
1998:	2	7 4	7	2 3	3	6	3	4	9	7	7	4
1999:	7	-	48		0	1 -	1077	um la sa d	7			
> Seas tau =					(assa 29.1 ⁹		• 1 Z AI	MITTIU:	sal.	LS)		
> assa						õ						
/ asso			Mar			Tun	.T11]	Aur	Son	Oct	Nov	Dec
1992:	oun	100	mar	прт	12	3	7	11	50p 7	7	9	5
1993:	8	10	4	3	3	8	10	9	12	11	4	4
1994:	5	5	4	4	6	4	7	0	5	3	6	4
1995:	7	4	8	8	7	7	8	7	7	11	5	8
1996:	9	10	2	6	3	10	4	3	4		5	8
1997:	5	2	8	5	3	10	2	4	2	6	1	4
1998:	3	2	9	6	2	5	7	6	6	1	5	6
1999:	4	6	2	2	8							
> Seas	sonal	LManr	nKend	dall	(assa	ault	.1AM	ThuSa	at.ts	5)		
tau =	-0.2	215,	sl	L =2	.088 ⁹	કે						
> assa	ault.	2AM	FhuSa	at.t:	5							
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					8	3	7	3	3	8	7	2
1993:	10	1	6	0	3	7	7	6	7	12	3	2
1994:	4	3	5	4	5	4	7	2	7	5	9	3
1995:	3	2	3	3	10	5	7	9	4	4	7	4
1996:	3	4	6	2	4	3	4	4	9	6	4	10
1997:	3	5	9	10	4	8	10	12	4	3	2	4
1998:	3	3	10	6	10	6	4	7	4	0	3	11
1999:	9	4	6	11	4		0 7 1 6	T 1 O		- \		
> Seas						auit	• ZAM	rnusa	at.ts	5)		
tau = > assa				=15								
/ d550						Turn	T., 1	7	Com	Oat	Nor	Dee
1992:	udii	гер	Mar	Abt.	May 2	Jun 3	5 JUL	Aug 3	sep 4	5	NOV 6	Dec 6
1992:	10	1	1	6	2	2	2	4	4 3	6	2	5
1993.	3	1	2	3	4	2	4	4	4	4	2	4
1995:	0	1	0	2	2	2	3	2	1	3	2	9
1996:	1	1	4	0	1	5	3	7	6	3	4	3
1997:	3	3	4	5	6	6	2	6	0	2	5	3
1998:	3	0	8	5	2	3	2	4	6	5	4	2
1999:	9	7	2	8	7	0	_	1	0	0	-	-
> Seas						ault	. 3AM'	ThuSa	at.t:	5)		
tau =					9.19							

Figure 5a. STL Analysis

The seasonal component is very unstable and increases in amplitude after 1996. R-sq = 52.5%

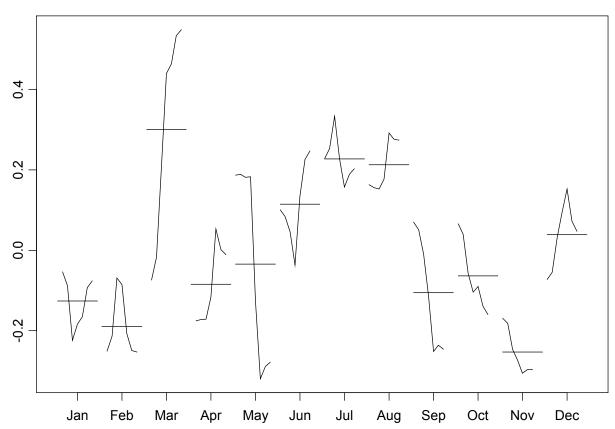


ss.window = 5 ,ss.robust = TRUE , fc.window = 51 , fc.degree = 1

LPS ASSAULTS

Figure 5b. Monthplot of Seasonal Component

The overall seasonal pattern is with a March maximum and with November minimum. Large trend upward in March and large downward in May. Smaller downward trends in October and November. July and August are relatively high and don't vary as much.



LPS assault, seasonal, logged

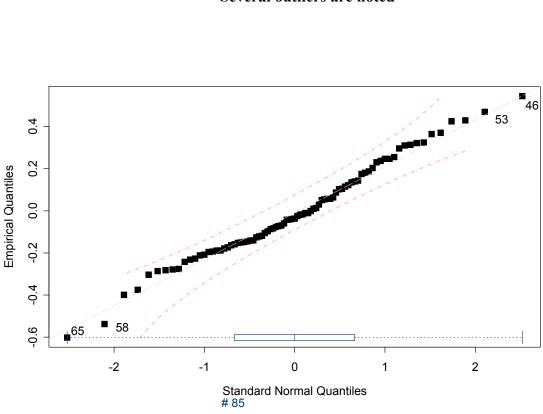
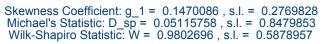
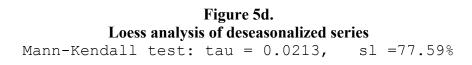


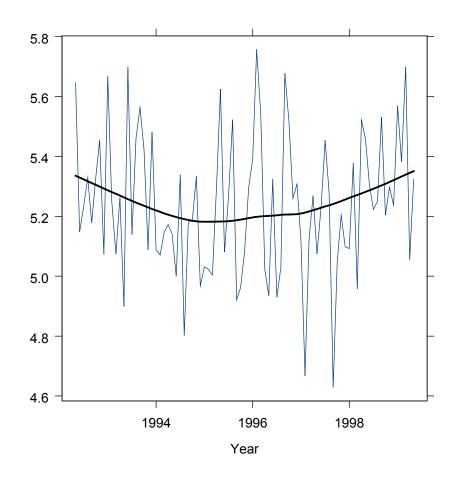
Figure 5c. Normal Probability Plot Several outliers are noted



> cour	nter	.ts										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					1	2	3	4	5	6	7	8
1993:	9	10	11	12	13	14	15	16	17	18	19	20
1994:	21	22	23	24	25	26	27	28	29	30	31	32
1995:	33	34	35	36	37	38	39	40	41	42	43	44
1996:	45	46	47	48	49	50	51	52	53	54	55	56
1997:	57	58	59	60	61	62	63	64	65	66	67	68
1998:	69	70	71	72	73	74	75	76	77	78	79	80
1999:	81	82	83	84	85							
See Ta	able	1.										



LPS assaults, deseasonalilzed, logged



G. LPS Dataset: Impaired

Summary

There were 5663 impaired driving charges in London from May of 1992 to May of 1999. Table 1.

<u>Crosstabs</u>. The ThuSat slot accounts for 65% of the impaired charges. About 33% of charges were in the 1AM slot and only 9.6% in the 3AM

<u>HOUR-WKGP</u>: There is a 9% excess of charges in the 3AM-SunWed group but overall hour-wkgrp are not associated (not statistically significant at 10%).

<u>HOUR-YEAR</u>: As would be expected from our loess analysis in <u>Figure 3b</u> and <u>Figure 3c</u> there has been a swing to later hours post 1996.

Figure 1a. and Figure 1b. Impaired driving charges are declining linearly on a log-scale. The rate of decline is about 13% per year. Because the log transformation simplifies the change to an almost linear decrease, all the data were logged for the remainder of the analysis.

<u>Figure 2</u>. Both SunWed and ThuSat exhibit a nearly linear downward slope on log scale. The downward slope for ThuSat is slightly greater. The table below shows the slope of the linear least squares fit to each group.

Wkgrp	Slope of fit to log(data,2)	Annual rate of decrease %
SunWed	-0.173	11%
ThuSat	-0.229	15%

Figure 3a. For the hour window only 12AM looks linear on a log scale. The curves are all different. It is interesting that 3AM actually increases after 1996 and the rate of decline for 2AM is much slower than at other hours.

<u>Figure 3b</u>. When the scale is free, we see that the 1AM has had a much faster rate of decline since 1996 than that of the data as whole (<u>Figure 1b</u>).

Figure 4b From Figure 4b we see that the SunWed-3AM has had the biggest increase since 1996 For both SunWed and ThuSat there is a shift from early evening to late evening post 1996. This is seen not only in the case of SunWed-3AM where an actual increase has occurred but also in the much smaller slope or rate of decline in the 2AM-SunWed, 2AM-ThuSat and 3AM-ThuSat than indicated for the overall series in Figure 1b. It is intended to model this further using intervention analysis.

Figure 5a. There is a linear downward trend on the log scale. The seasonal component has lots of local maxima and minima.

R-sq = 81.6%

<u>Figure 5b</u>. Local maximum in February, May and October and local minimum in December, April and July. But the seasonal component is variable and shows a lot of changes over time. There is an increasing trend for December, January, May and August.

Figure 5c. The normality tests indicate no significant departure from normality in the remainder component.

Figure 5d. Loess analysis of deseasonalized series. The thicker line shows the linear least-squares fit and the thinner line shows the loess curve. The least-squares line has a slope of 0.200 almost the same as in Figure 1b. A slight lack of fit of the linear least-squares line is evident at the beginning and end of the series. The residuals from the least-squares line are approximately normally distributed with no evident outliers. However further diagnostic checks indicate that the variablility about the line increases with Year.

Crosstabs Analysis

Hour and wkgrp are not associated

The ThuSat slot accounts for 65% of the impaired charges. About 33% of charges were in the 1AM slot and only 9.6% in the 3AM.

There is a 9% excess of charges in the 3AM-SunWed group but overall hour-wkgrp are not associated (not statistically significant at 10%).

```
> crosstabs(charges~wkgrp+hour, data=lps.df, subset=lps.df$type=="impaired",
na.action=na.omit)
Call:
crosstabs(formula = charges ~ wkgrp + hour, data = lps.df, subset = lps.df$type ==
    "impaired", na.action = na.omit)
5674 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total
       1
+----+
wkgrp |hour
          |12AM |1AM |2AM |3AM |RowTotl|
   |11PM
SunWed | 271 | 397 | 620 | 478 | 206 |1972 |
    0.14 0.2 0.31 0.24 0.1
                                    |0.35 |
     |0.35 |0.34 |0.33 |0.36 |0.38
                                   |0.048 |0.07 |0.11 |0.084 |0.036 |
ThuSat | 510 | 769 |1251 | 832 | 340 |3702 |
     0.14 0.21 0.34 0.22 0.092 0.65
    |0.65 |0.66 |0.67 |0.64 |0.62 |
|0.09 |0.14 |0.22 |0.15 |0.06 |
_____+
ColTotl|781 |1166 |1871 |1310 |546 |5674 |
|0.14 |0.21 |0.33 |0.23 |0.096 | |
Test for independence of all factors
     Chi^2 = 6.282868 d.f.= 4 (p=0.178996)
     Yates' correction not used
> get.crosstabs.percenterror(charges~wkgrp+hour)
Crosstabs - percentage error: 100*(Obs-Exp)/Exp,
    [,1] [,2] [,3] [,4] [,5]
[1,] 0 -2 -5 5 9
[2,] 0 1 2 -3 -5
> contrib(charges~wkgrp+hour)
contribution to chi-sq:
    [,1] [,2] [,3] [,4] [,5]
[1,] -0.03 -0.41 -1.19 1.06 1.18
[2,] 0.02 0.30 0.87 -0.78 -0.86
```

Hour and Year are associated: As would be expected from our loess analysis in <u>Figure 3b</u> and <u>Figure 3c</u> there has been a swing to later hours post 1996.

> crosstabs(charges~hour+year, data=lps.df, subset=lps.df\$type=="impaired", na.action=na.omit) Call: crosstabs(formula = charges ~ hour + year, data = lps.df, subset = lps.df\$type == "impaired", na.action = na.omit) 5674 cases in table | N |N/RowTotal| IN/ColTotal| |N/Total | +---+ hour |year |1992 |1993 |1994 |1995 |1996 |1997 |1998 |1999 |RowTotl| _____ _____ _+____ _+___ +----11PM |119 |131 |151 86 |103 | 75 | 82 34 1781 |0.17 10.15 |0.19 |0.11 |0.13 |0.096 |0.1 0.044 10.14 10.14 |0.13 |0.16 |0.12 |0.13 0.12 |0.16 |0.15 1 0.021 |0.023 |0.027 |0.015 |0.018 |0.013 |0.014 |0.006 _____+ 12AM 1181 1249 1174 1144 1166 1112 |101 1 39 11166 |0.21 0.16 0.21 |0.15 0.12 0.14 |0.096 |0.087 |0.033 0.21 0.24 |0.19 0.2 0.21 0.18 0.2 0.17 0.032 0.044 0.031 0.025 0.029 0.02 |0.018 |0.0069 | ___+_____ 1127 1328 1366 1346 1282 1230 1154 1 38 11871 1 A M 0.2 |0.15 |0.18 |0.18 |0.12 |0.082 |0.068 |0.02 10.33 |0.38|0.36|0.37|0.38|0.058|0.065|0.061|0.05 0.3 0.25 0.25 0.041 0.027 0.022 10.17|0.0067 | +----| 70 |1310 |200 |174 |166 |194 |187 |128 2AM |191 |0.15 |0.15 |0.13 |0.13 |0.15 0.14 |0.098 |0.053 10.23 0.23 0.25 |0.31 10.19 0.25 0.31 10.22 10.2 0.034 0.035 0.031 0.029 0.034 0.033 0.023 0.012 _____+ | 51 | 77 | 57 | 74 82 80 | 82 | 43 1546 3AM |0.093 |0.14 0.15 0.1 0.15 0.15 0.14 |0.079 |0.096 0.059 0.075 0.088 0.078 0.1 10.13 10.14 |0.19 0.009 0.014 0.014 0.01 0.014 0.014 0.013 0.0076 _+____+ ColTotl|870 |1023 |927 1735 |773 |610 |512 1224 15674 1 0.14 10.15 0.18 10.16 10.13 10.11 10.09 10.039 ____ _ + _ _ _ _ _ _ _ + _ _ _ _ _ _ _ . . Test for independence of all factors Chi^2 = 185.3234 d.f.= 28 (p=0) Yates' correction not used > get.crosstabs.percenterror(charges~hour+year) Crosstabs - percentage error: 100*(Obs-Exp)/Exp, [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] -3 10 [1,] -7 18 -15 -11 16 -1 5 -11 -9 -5 [2,] 1 18 -4 -15 [3,] 14 8 13 16 -10 -23 -25 -49 9 -5 -15 -19 -2 33 8 35 [4,] -8 -19 -39 -22 [5,] 8 40 50 99 > contrib(charges~hour+year) contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [1,] -0.07 -0.83 2.07 -1.51 -0.33 -0.98 1.37 0.57 [2,] 0.17 2.67 -1.20 -0.57 0.57 -1.19 -0.41 -1.04 [3,] 2.43 1.56 2.31 2.55 -1.56 -3.32 -3.22 -4.17 [4,] -0.70 -2.35 -2.74 -0.28 1.16 3.89 0.90 2.54 [5,] -3.58 -2.16 -0.76 -1.63 0.65 3.04 3.52 4.62

•

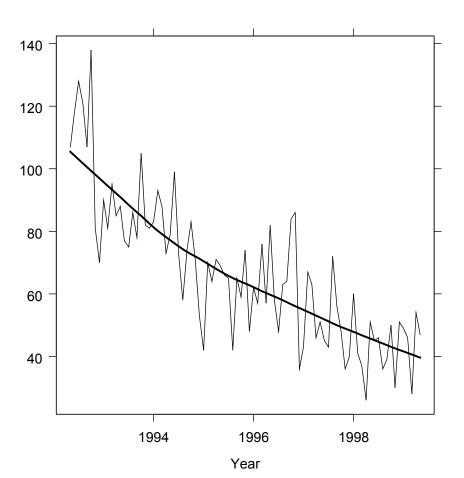
LPS IMPAIRED Data

There is a mild decrease from 1992 to mid-1994 and then leveling off. Possible outliers detected in logged series, <u>Figure 5c</u>, are in bold font below – although they are not significant at 10% on the normality tests for skewness, Michael's test and the Wilk-Shapiro test.

Table 1.

> imp.	ts											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					107	118	128	121	107	138	81	70
1993:	90	81	95	85	88	77	75	86	78	105	82	81
1994:	83	93	88	73	79	99	73	58	74	83	71	53
1995:	42	70	64	71	69	66	65	42	65	59	74	48
1996:	62	57	76	57	82	58	48	63	64	84	86	36
1997:	43	67	63	46	51	45	43	72	56	48	36	40
1998:	60	41	37	26	51	45	46	36	39	50	30	51
1999:	49	46	28	54	47							
> Seas	sona	lManr	nKend	dall	(imp	.ts)						
tau =	-0.7	725,	S	l =1	.7766	e-139	00					

Figure 1.



LPS impaired-driving charges

Figure 1b. Logged series

Figure 1b shows the time series after a log transformation. Specifically a log to the base 2 transformation. This particular version of the log transformation means that we can interpret a one unit change as a doubling and an increase by 0.5 units on the vertical scale corresponds to an increase of about 41% in the original data.

The transformation reveals that the decline in impaired charges is almost linear. Figure 1b shows a thicker line that corresponds to the least squares linear fit. The loess curve is the thinner line which stops at the last observed data value. There is a slight lack of fit of the least-squares line at the very beginning and very end of the series. The slope of the least-squares line is about -0.202 which implies that impaired charges have been declining at the rate of about 13% per year over 1992 to 1999.

Note: $2^{-0.202} \approx 0.869$

LPS impaired, log(charges,2)

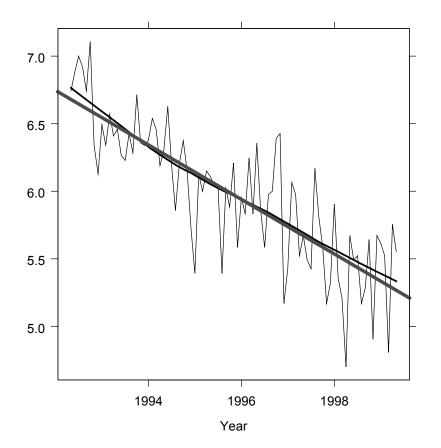
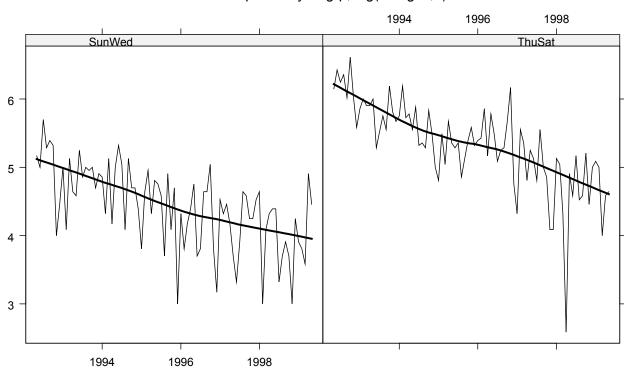


Figure 2. By wkgrp.

Both exhibit a nearly linear downward slope on log scale. The downward slope for ThuSat is slightly greater. The table below shows the slope of the linear least squares fit to each group.

Wkgrp	Slope of fit to log(data,2)	Annual rate of decrease %
SunWed	-0.173	11%
ThuSat	-0.229	15%



LPS impaired by wkgrp, log(charges,2)

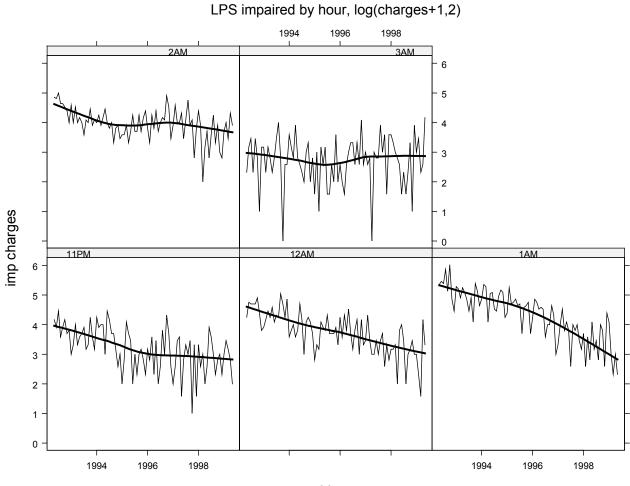
Year

Table 2.

> imp.SunWed.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1993: 1994: 29 1995: 20 28 18 21 27 13 14 25 25 33 1996: 20 14 22 18 13 10 15 25 24 19 1997: 23 1998: 25 8 17 20 21 21 10 13 15 13 8 1999: 15 14 12 30 22 > SeasonalMannKendall(imp.SunWed.ts) tau = -0.508, sl =3.036e-6% > imp.ThuSat.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1993: 58 64 60 1994: 73 53 55 47 40 41 1995: 28 51 41 29 35 42 1996: 42 43 58 36 55 45 38 39 51 1997: 20 47 41 28 38 35 28 47 32 29 1998: 35 33 20 6 30 24 36 23 24 37 32 16 24 25 1999: 34 > SeasonalMannKendall(imp.ThuSat.ts) tau = -0.707, sl = 8.438e - 13%

Figure 3a. By hour. Common Scaling

For the hour window only 12AM looks linear on a log scale. The curves are all different. It is interesting that 3AM actually increases after 1996 and the rate of decline for 2AM is much slower than at other hours.

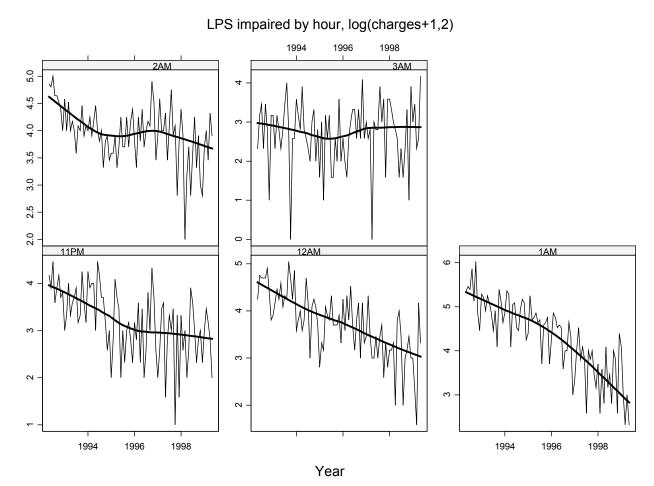


Year

LPS IMPAIRED

Figure 3b. By hour. Independent vertical scaling.

When the scale is free, we see that the 1AM has had a much faster rate of decline since 1996 than that of the data as whole (Figure 1b).

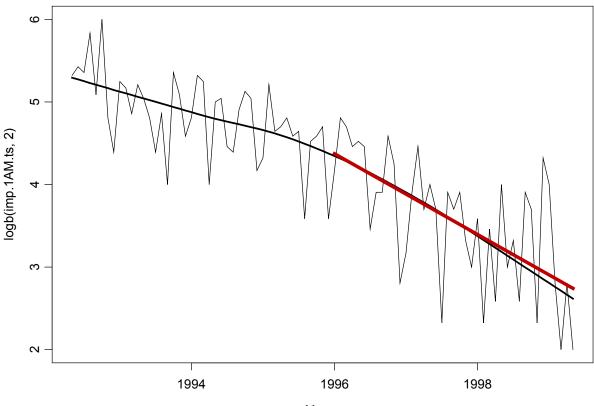


free vertical scaling

Figure 3c.

Least square line fit (thick red line) to data from 1996 to present. The slope of the least squares line to post-1996 is -0.489 which corresponds to a annual rate of decrease of about 28.7%. Since 1998, the actual impaired driving charges are decreasing slightly faster than this. > $1-2^{-0}.489$

[1] 0.2874812

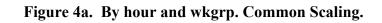


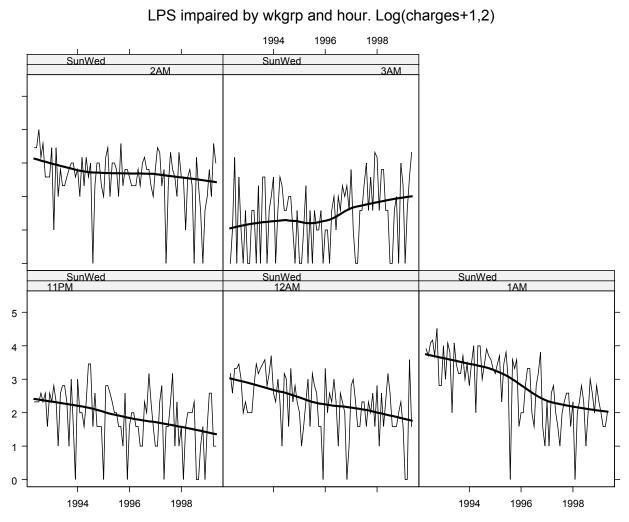
LPS impaired at 1AM, log(charges,2)

Years

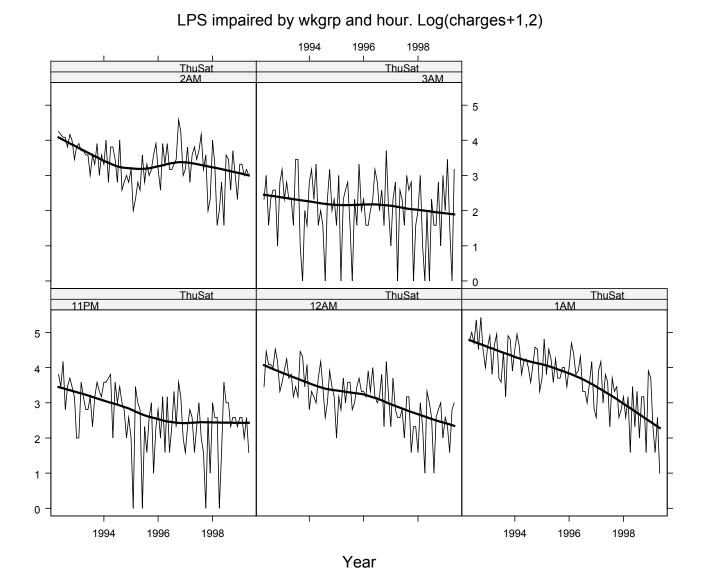
Tabulation and Mann-Kendall Tests

						18	abula	ation	i and	1 IVI8	inn-	Ken	18
> imp													
	Jan	Feb	Mar	Apr	May			Aug	Sep	Oct	Nov	Dec	
1992:					17	14	21	11	14	17	12	13	
1993:	7	9	15	9	11	12	14	8	9	18	11	8	
1994:	18	14	15	15	7	21	17	12	12	8	5	7	
1995:	3	6	16	12	10	3	7	4	7	8	6	4	
1996:	8	6	11	4	10	3	5	13	7	19	12	5	
1997:	3	5	10	11	2	6	9	7	10	1	9	2	
1998:	9	5	7	3	5	14	11	7	4	6	7	4	
1999:	7	10	8	6	3								
> Sea	sona	lManı	nKend	dall	(imp	.11PM	1.ts)					
tau =													
> imp	.12A	M.ts											
_	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:					18	26	25	25	25	29	20	13	
1993:	14	17	21	18	23	16	19	18	32	25	18	28	
1994:	11	13	15	11	13	25	17	7	16	18	16	12	
1995:	6	9	8	16	14	13	19	12	12	12	14	9	
1996:	18	11	19	13	22	12	8	13	17	7	17	9	
1997:	10	19	10	7	7	10	7	10	12	5	9	6	
1998:	8	8	9	3	13	15	10	3	7	8	10	7	
1999:	7	4	2	17	9								
> Sea	sona	lManı		dall	(imp	.12AM	1.ts)					
tau =					.71e								
> imp													
1			Mar	Apr	Mav	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:				-	40	43	41	57	34	64	28	21	
1993:	38	36	29	37	33	28	21	29	16	41	34	24	
1994:	28	40	38	16	32	33	22	21	30	35	33	18	
1995:	20	37	25	26	28	24	25	12	23	24	26	12	
1996:	18	28	26	22	23	22	11	15	15	24	19	7	
1997:	9	15	22	13	16	13	5	15	13	15	10	8	
1998:	12	5	11	6	16	8	10	6	15	13	5	20	
1999:	16	7	4	7	4								
> Sea			nKend		(imp	. 1 A M	ts)						
tau =				1 =0	-		,						
> imp													
T			Mar	Apr	Mav	Jun	Jul	Aug	Sep	Oct.	Nov	Dec	
1992:				1	28	27	31	24	24	22	20	15	
1993:	23	15	22	15	17	15	11	16	15	21	14	16	
1994:	15	18	14	17	21	15	13	15	- 9	13	14	10	
1995:	11	11	14	9	12	18	12	12	18	12	17	20	
1996:	13	9	18	13	20	12	15	17	16	29	22	10	
1997:		23	15					26	14	16	6	13	
1998:	20	14	3		12	6	11	18	9	14	7	6	
1999:	12	15	10	19	14	0	T T	10	2	11	'	0	
> Sea						2ам	ts)						
tau =					.2022		,						
> imp			5.	L 0	• 2 0 2 2	- 0							
> Tub			Mar	Anr	Mav	.Tun	.Tu 1	Aug	Sen	Oct	Nov	Dec	
1992:	Jun	100	TOT	чът	May 4	8	10		10	6	1	Dec 8	
1993:	8	4	8	6	4	6	10	15	6	0	5	5	
1993:	11	8	6	14	4 6	5	4	3	7	9	3	6	
1995:	2	7	1	8	5	8	2	2	5	3	11	3	
1995.	5	3	2	5	7	9	2		9	5	16	5	
1990:	7	5	2 6	0	7	9	9	14	9 7	11	2	11	
1997:	11	5 9	ю 7	6	5	6 2	6 4	14 2	4	9	2	14	
1998:	11	9 10	4	6 5	5 17	2	4	2	4	Э	T	14	
						2714	+ <)						
> Sea tau =				=86		• JAN	. (3)						
cau -	0.0.	- · · /	υ⊥	.00	• 2 0 0								





Year

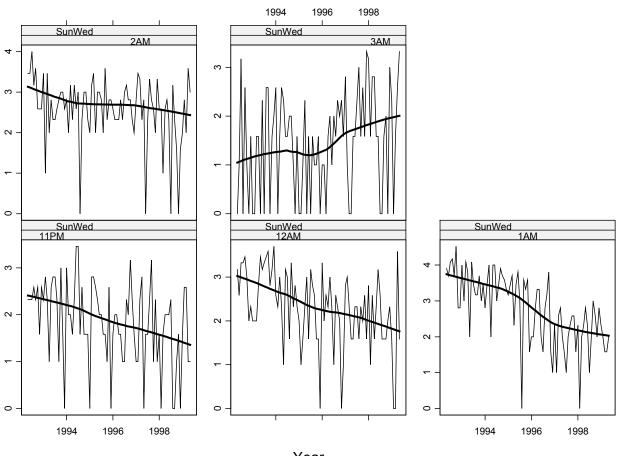


15

LPS IMPAIRED

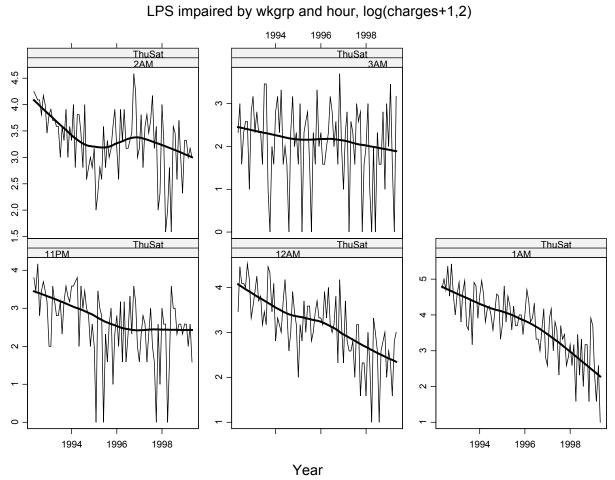
Figure 4b. By hour and wkgrp. Independent Scaling.

From Figure 4b we see that the SunWed-3AM has had the biggest increase since 1996 For both SunWed and ThuSat there is a shift from early evening to late evening post 1996. This is seen not only in the case of SunWed-3AM where an actual increase has occurred but also in the much smaller slope or rate of decline in the 2AM-SunWed, 2AM-ThuSat and 3AM-ThuSat than indicated for the overall series in Figure 1b. It is intended to model this further using intervention analysis.



LPS impaired by wkgrp and hour, log(charges+1,2)

Year free vertical scale



free vertical scale

Tabulations and Seasonal Mann-Kendall Tests

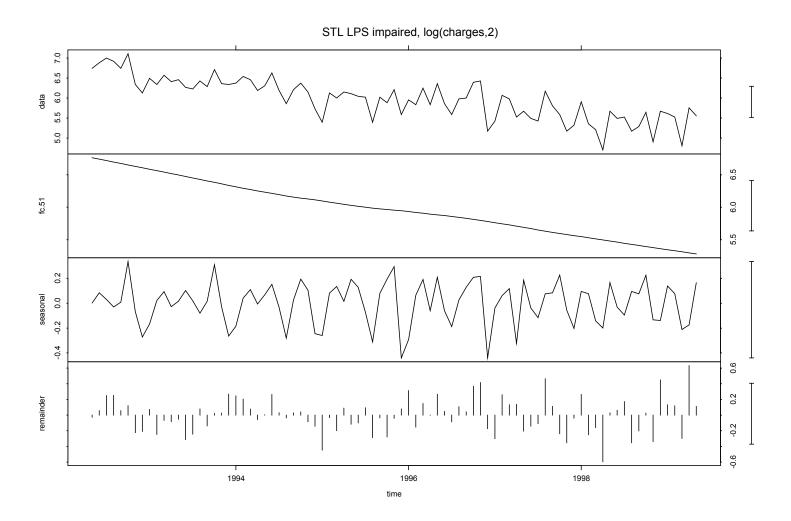
Laba	auro	115 al		caso	11641 1	1411		ciiua				
> imp	.11PM	MSun≬	Wed.t	ts								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					4	4	4	5	4	5	2	5
1993:	4	6	4	1	5	6	6	4	1	7	2	0
1994:	7	3	3	2	4	10	10	2	5	2	2	2
1995:	0	6	6	5	4	3	3	2	2	1	5	0
1996:	2	3	3	2		1	1	4	3	8	4	2
1997:	1	1	4	6	0	2	2		8	1	4	1
1998:	2	0	2	3	3	3	4	0	0	1	2	0
1999:	2	5	5	1	1	110	10					
> Seas							Isun	wea.1	LS)			
tau = > imp		•			5501 ⁹	õ						
> Tilb					Matz	Tun	T11]	Λuα	Son	$\cap c^+$	Nov	Dec
1992:	Jan	гер	Mai	Арт	May 8	5	9 UU	Aug 9	зер 10	7	3	Dec 4
1993:	3	3	3	6	10	8	9	10	11	6	8	12
1994:	5	4	7	4	1	8	7		9	4	6	4
1995:	3	1	2	4	7	2	8		5	2	2	0
1996:	9	3	5	4	7	4	1		4	3	0	1
1997:	6	7	4	2	2	4	4		4	3	5	2
1998:	6	1	5	2	4	8	5		2	2	3	4
1999:	2	0	0	11	2							
> Seas	sonal	lManı	nKend	dall	(imp	.12AM	MSun	Wed.1	cs)			
tau =					.0033				,			
> imp												
_	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					14	12	16	17	12	22	6	6
1993:	15	7	16	13	3	16	10	8	8	12	7	10
1994:	6	10	15	3	15	15	7	10	14	12	11	9
1995:	8	10	12	4	10	13	7	0	11	9	11	2
1996:	3	3	6	9	9	3	2	6	8	13	2	1
1997:	4	1	5	6	3	2	1	3	4		5	2
1998:	4	0	3	4	6	4	1	3	7	5	3	6
1999:	4	3	2	2	3							
> Seas					-		SunWe	ed.ts	5)			
tau =					.0110	∋-8%						
> imp						_			~			_
	Jan	Feb	Mar	Apr				Aug				Dec
1992:	1.0	-	1.0	-	10	10	15	8	11	5	5	5
1993:	10	1	10	3	6	4	4		6	7	7	5
1994:	6	3	8	4	8	5	7		4	7	7	4
1995:	3	8	10	3	7	7	6	3	11	4	6	6
1996:	5	4	4	4	6	4	7	8	6	6	4	3
1997:	6	10	9	4	6	0	4	9	6	5 2	3	9
1998: 1999:	5 3	5 6	1 3	5 11	6 7	4	0	8	4	Z	0	2
> Seas						2 n M a	ราวจพั	od + 1	-)			
tau =					.543 ⁹		Juliw	-u	5)			
> imp					. 545	0						
, Tub					Mav	.Tun	.T11]	Aug	Sen	Oct	Nov	Dec
1992:	oun	100	TIGEL	прт	0	1	8	0	5	1	0	2
1993:	0	0	2	2	0	4	0	5	5	0	2	3
1994:	5	0	2	5	4	2	2	3	3	1	0	2
1995:	0	Ő	1	4		2	0	2	1	1	2	0
1996:	1	1	0	2	3	1	3		4	3	4	2
1997:	6	1	0	0	2	2	4	7	2	5	2	9
1998:	8	2	6	6	2	2	0	0	2	3	0	7
1999:	4	0	2	5	9	_	5	2	_	5	-	
> Seas						. 3AMS	SunWe	ed.ts	5)			
tau =				=7.3					-			

LPS IMPAIRED

Figure 5a. STL Analysis

There is a linear downward trend on the log scale. The seasonal component has lots of local maxima and minima.

R-sq = 81.6%



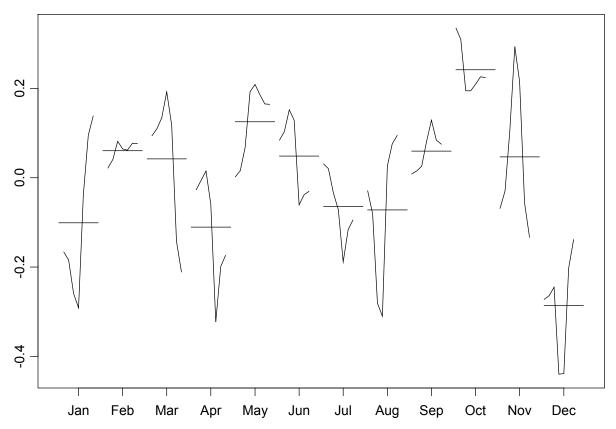
ss.window = 5 ,ss.robust = TRUE , fc.window = 51 , fc.degree = 1

LPS IMPAIRED

Figure 5b. Monthplot of Seasonal Component

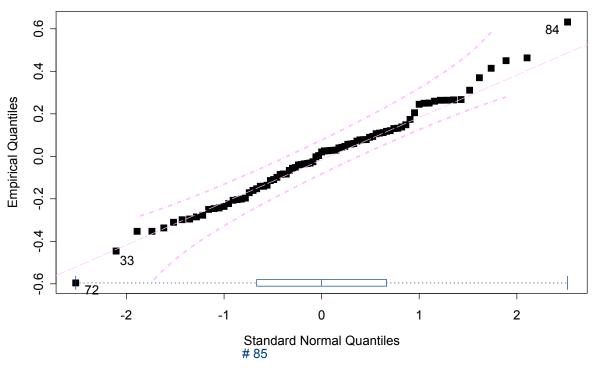
Local maximum in February, May and October and local minimum in December, April and July. But the seasonal component is variable and shows a lot of changes over time. There is an increasing trend for December, January, May and August.

LPS impaired, seasonal, logged



LPS IMPAIRED

Figure 5c. Normal Probability Plot The normality tests indicate no significant departure from normality in the remainder component.



Skewness Coefficient: g 1 = 0.1413422, s.l. = 0.2846118
Michael's Statistic: D sp = 0.03992539, s.l. = 0.3756397
Wilk-Shapiro Statistic: W = 0.9919533 , s.l. = 0.9854363

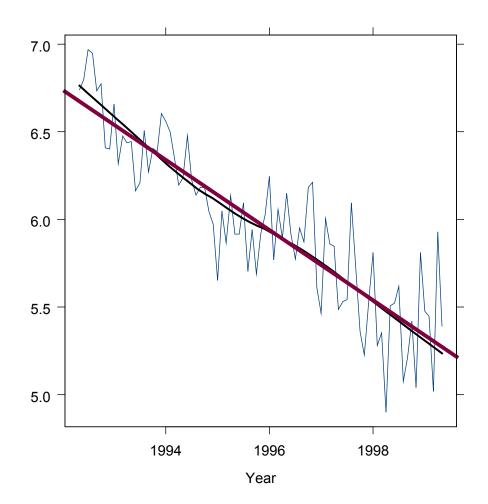
> cour	nter.	.ts										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					1	2	3	4	5	6	7	8
1993:	9	10	11	12	13	14	15	16	17	18	19	20
1994:	21	22	23	24	25	26	27	28	29	30	31	32
1995:	33	34	35	36	37	38	39	40	41	42	43	44
1996:	45	46	47	48	49	50	51	52	53	54	55	56
1997:	57	58	59	60	61	62	63	64	65	66	67	68
1998:	69	70	71	72	73	74	75	76	77	78	79	80
1999:	81	82	83	84	85							
See Ta	able	1.										

Figure 5d. Loess analysis of deseasonalized series

Mann-Kendall test: tau = -0.711, sl =0%

The thicker line shows the linear least-squares fit and the thinner line shows the loess curve. The least-squares line has a slope of -0.200 almost the same as in Figure 1b. A slight lack of fit of the linear least-squares line is evident at the beginning and end of the series. The residuals from the least-squares line are approximately normally distributed with no evident outliers. However further diagnostic checks indicate that the variablility about the line increases with Year.

LPS impaired, deseasonalilzed, logged



H. WPS Dataset

Introduction

WPS (Windsor Police dataset) is comprised of 2943 records covering January 1994 to December 1998. There are 1370 arrests for assault and 1573 arrests for impaired driving. Only data in hourly windows from 11PM to 3:59AM is included.

The WPS dataset differs slightly in structure from the LPS dataset. One difference is that the WPS assaults do not include sexual assault whereas LPS does. Our response variable is simply counts.

Variables of interests in the WPS dataset

counts	number of charges
type	factor with 2 levels: "assault", "impaired"
wkgrp	ordered factor with 2 levels: "SunWed", "ThuSat"
hour	ordered factor with 5 levels: "11PM", "12AM", "1AM", "2AM", "3AM"
year	ordered factor with 8 levels: 1994 to 1999
month	ordered factor with 12 levels: January to December

Note that there are only five complete years of data.

H. WPS Dataset: Assaults

Summary

The WPS assaults do not include sexual assaults. There were 1370 assault cases in the Windsor area from May of 1992 to May of 1999. <u>Table 1</u>.

<u>Crosstabs</u>. In SunWed slot, there are 30% more assaults than expected in 11PM and fewer than expected in the 1AM, 2AM and 3AM slots. In 1994 there were more than the usual number of assaults in the 1AM window.

Figure 1a. and Figure 1b. Assault rate has not changed much over 1994 to 1998.

Figure 2. There is a very small downward trend in SunWed group starting around 1996.

Figure 3b. Since 1996 there have been assaults have increased at 2AM and also at 3AM and have decreased at 11PM and held steady at 12AM. There is a slight decrease in recent assaults at 1AM.

Figure 4b From Figure 4b we see that since 1996 there has been an upward shift at 2AM and 3AM and a downward trend at 11PM for both SunWed and ThuSat. For both weekgroups, 11PM and 12PM are relatively flat. The upward shifts are larger with the ThuSat group.

Figure 5a. The seasonal component is reasonably stable. There is a slight increase in amplitude of the seasonal oscillations.

<u>Figure 5b</u>. The overall seasonal pattern is with March, July and December maxima and with January, May and September minimum. The seasonal is relatively stable. There are trends up in June and down in September and October.

<u>Figure 5c</u> Several possible outliers are noted but all statistical tests for normality are fine.. See also <u>Table 1</u>. I think that perhaps there could be an error and that some 1996 assaults in June and July were mis-recorded as occuring in 1997.

Crosstabs Analysis

Hour and wkgrp are associated

In SunWed slot, there are 30% more assaults than expected in 11PM and fewer than expected in the 1AM, 2AM and 3AM slots.

> crosstabs(counts~wkgrp+hour, data=wps.df, na.action=na.exclude, subset=wps.df\$type=="assault") Call: crosstabs(formula = counts ~ wkgrp + hour, data = wps.df, subset = wps.df\$type == "assault", na.action = na.exclude) 1612 cases in table +----+ | N 1 |N/RowTotal| |N/ColTotal| |N/Total | +----+ wkgrp |hour |11PM |12AM |1AM |2AM |3AM |RowTotl| SunWed |190 |151 |132 |112 | 56 |641 |0.3 |0.24 |0.21 |0.17 |0.087 |0.4 0.52 0.45 0.33 0.34 0.31 0.12 0.094 0.082 0.069 0.035 ThuSat |178 |185 |269 |214 |125 |971 1 0.18 0.19 0.28 0.22 0.13 0.6 |0.48 |0.55 |0.67 |0.66 |0.69 | |0.11 |0.11 |0.17 |0.13 |0.078 | ColTotl|368 |336 |401 |326 |181 |1612 | |0.23 |0.21 |0.25 |0.2 |0.11 | Test for independence of all factors Chi^2 = 43.1059 d.f.= 4 (p=9.837309e-009) Yates' correction not used > get.crosstabs.percenterror(counts~wkgrp+hour) Crosstabs - percentage error: 100*(Obs-Exp)/Exp, [,1] [,2] [,3] [,4] [,5] [1,] 30 13 -17 -14 -22 [2,] -20 -9 11 9 15 > contrib(counts~wkgrp+hour) contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [1,] **3.61** 1.50 -2.17 -1.55 -1.88 [2,] -2.93 -1.22 1.77 1.26 1.53

Hour and Year are associated

In 1994 there were more than the usual number of assaults in the 1AM window.

> crosstabs(counts~hour+year, data=wps.df, na.action=na.exclude, subset=wps.df\$type=="assault")
Call:

crosstabs(formula = counts ~ hour + year, data = wps.df, subset = wps.df\$type == "assault", na.action = na.exclude) 1612 cases in table | N |N/RowTotal| IN/ColTotal| |N/Total | +----+ hour |year |1994 |1995 |1996 |1997 |1998 |RowTotl| _____ _+____+ 11PM | 87 | 69 | 84 | 72 | 56 1368 |0.24 |0.19 |0.23 |0.2 0.15 10.23 10.27 |0.22 |0.27 0.2 |0.18 1 0.054 0.043 0.052 0.045 0.035 _____+ | 55 |0.16 12AM | 59 178 175 | 69 1336 0.23 |0.21 0.18 0.22 0.21 0.18 0.25 0.24 0.16 0.22 0.037 0.048 0.047 0.034 0.043 _____+ |**107** | 75 | 61 | 89 | 69 | 401 1 A M 1 |0.27 |0.19 |0.15 0.22 |0.17 10.25 |0.33 |0.24 |0.2 |0.25 |0.22 | |0.066 |0.047 |0.038 |0.055 |0.043 | -----+------| 43 | 55 | 55 | 87 | 86 |326 2AM |0.17 |0.17 |0.2 0.13 |0.27 0.26 |0.18 0.25 |0.18 0.27 10.13 0.027 0.034 0.034 0.054 0.053 - I _____+ | 27 | 34 | 33 | 49 3AM | 38 |181 |0.18 |0.27 0.15 0.19 0.21 0.11 |0.084 |0.11 |0.11 |0.14 0.12 1 0.017 0.021 0.02 0.03 0.024 _____+ -+ ColTotl|323 |311 |308 |352 |318 11612 1 0.19 10.19 10.22 10.2 10.2 1 _ + _ _ _ _ _+___+ Test for independence of all factors Chi^2 = 58.5987 d.f.= 16 (p=8.997489e-007) Yates' correction not used > get.crosstabs.percenterror(counts~hour+year) Crosstabs - percentage error: 100*(Obs-Exp)/Exp, [,1] [,2] [,3] [,4] [,5] -3 19 20 17 [1,] 18 -10 -23 [2,] -12 -25 4 [3,] 33 -3 -20 2 -13 34 -34 -13 -12 22 [4,]24 [5,] -26 -3 -5 6 > contrib(counts~hour+year) contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [1,] 1.54 -0.24 1.63 -0.93 -1.95 [2,] -1.01 1.64 1.35 -2.14 0.33 [3,] **2.97** -0.27 -1.78 0.15 -1.14 [4,] -2.76 -1.00 -0.92 1.87 2.70 [5,] -1.54 -0.16 -0.27 1.51 0.38

LPS Assault Data

Possible outliers detected in logged series, Figure 5c, are in bold font below.

Table 1.

> assa	ault	.ts										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1994:	22	24	25	28	30	25	45	22	18	29	33	22
1995:	27	23	19	21	19	34	44	41	24	18	18	23
1996:	16	35	33	18	25	14	16	29	25	25	31	41
1997:	26	18	33	35	16	49	50	31	15	20	29	30
1998:	30	33	19	30	27	28	40	34	18	19	26	14

> SeasonalMannKendall(assault.ts)
tau = -0.00844, sl =94.32%

Figure 1.

Assault rate has not changed much over 1994 to 1998.

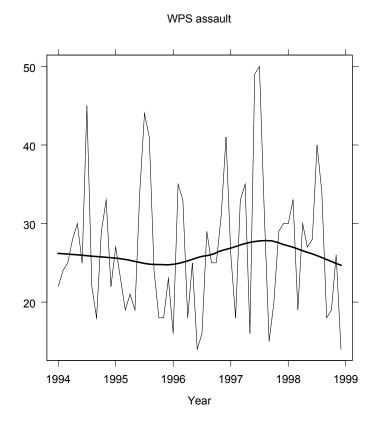


Figure 2. By wkgrp.

There is a very small downward trend in SunWed group starting around 1996.

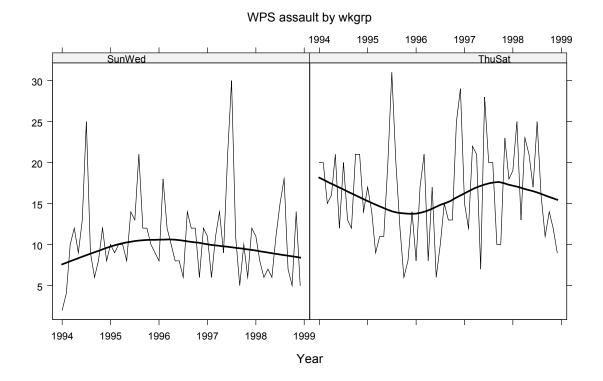


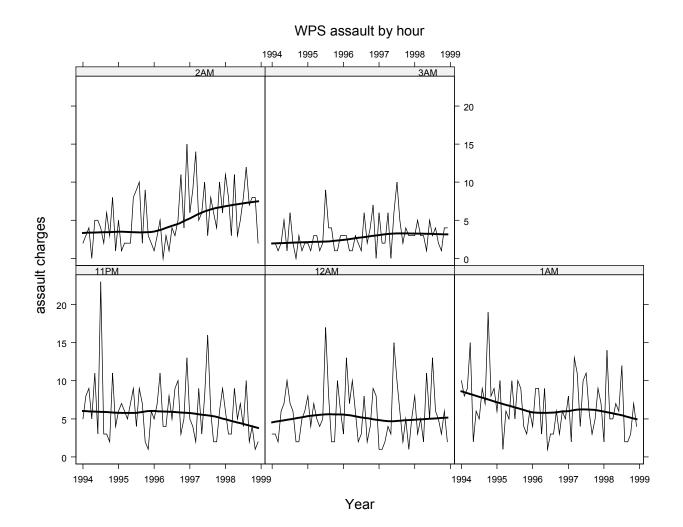
Table 2.

> assault.SunWed.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1994: 2 4 10 12 9 13 25 9 6 8 12 8 1995: 10 9 10 10 8 14 13 21 12 12 10 9 1996: 8 18 12 10 8 8 6 14 12 12 6 12 1997: 11 6 11 14 9 21 30 11 5 10 6 12 1998: 11 8 6 7 6 11 15 18 7 5 14 5 > SeasonalMannKendall(assault.SunWed.ts) tau = -0.00867, sl = 94.23%

> assa	ault.	.Thus	Sat.t	CS								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1994:	20	20	15	16	21	12	20	13	12	21	21	14
1995:	17	14	9	11	11	20	31	20	12	6	8	14
1996:	8	17	21	8	17	6	10	15	13	13	25	29
1997:	15	12	22	21	7	28	20	20	10	10	23	18
1998:	19	25	13	23	21	17	25	16	11	14	12	9
> Seas	sonal	lManı	nKend	dall	(assa	ault	.Thust	Sat.t	cs)			
tau =	0.04	126,	s	L =72	2.038	5						

Figure 3a. By hour. Common Scaling

Since 1996 there have been assaults have increased at 2AM and also at 3AM and have decreased at 11PM and held steady at 12AM. There is a slight decrease in recent assaults at 1AM.



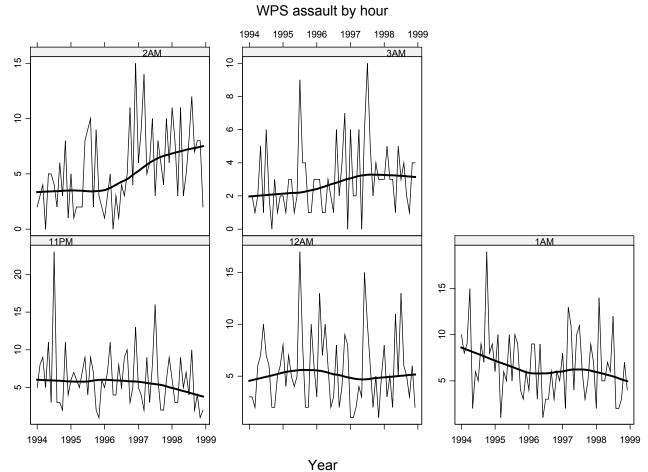


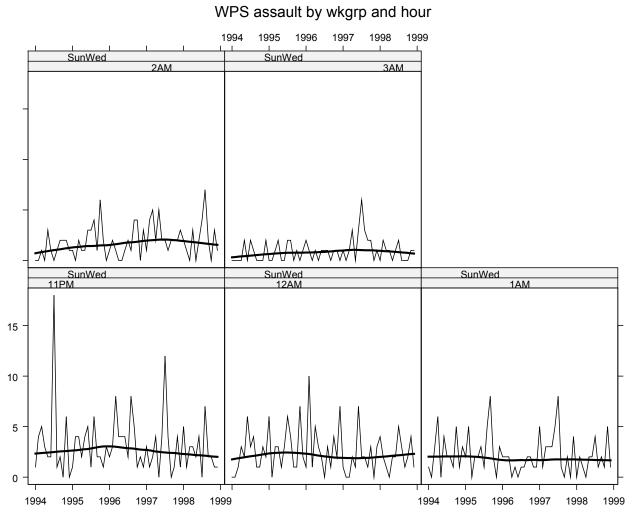
Figure 3b. By hour. Independent vertical scaling.

free vertical scaling

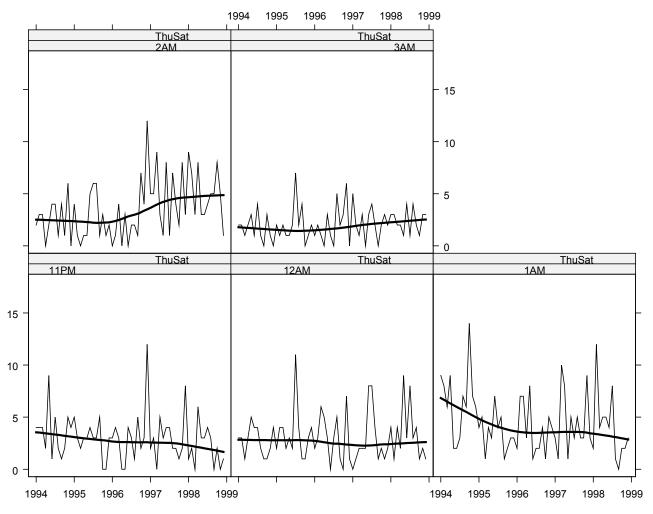
Tabulation and Mann-Kendall Tests

					18	abula	ation	anc	1 Ma	inn-	Kenc	lal
> assault	.11PM	M.ts										
Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1994: 5	8	9	5	11	3	23	3	3	2	11	4	
1995: 6	7	6	5	7	9	4	9	7	2	1	6	
1996: 5	7	11	4	4	8	5	9	10	3	5	13	
1997: 5	4	2	9	3	8	16	6	2	2	6	9	
1998: 6	3	3	9	5	7	4	10	2	4	1	2	
> Seasonal	lManı	nKend	dall	(assa	ault	.11PM	M.ts)					
tau = -0.1	116,	s	1 =33	3.74	00							
> assault	.12A1	M.ts										
Jan	Feb	Mar	Apr	May	Jun	Jul	Auq	Sep	Oct	Nov	Dec	
1994: 3	3	2	6	7	10	7	6	2	2	5	6	
1995: 8	4	7	5	4	5	17	8	2	2	10	6	
1996: 3		7	10	6	2	3	8	2	4	9	8	
1997: 1		2	4	3	15		6	2	5	1	5	
1998: 8	3	5	2	11	5	13	6	5		6	2	
> Seasonal												
tau = -0.0						•						
ouu o.	0010	, .										
> assault	.1AM	.ts										
		Mar	Apr	Mav	Jun	.Tu]	Αιια	Sep	Oct	Nov	Dec	
1994: 10	8	9	15	2	6	5	9	7	19	8	9	
1995: 6	10	1	- 6	5	10	5	10	9	4	3	6	
1996: 4		9	3	9	1	3	- 3	6		6	5	
1997: 8		13		4	10	11	6	3		9	7	
1998: 2		5	5	7	6	12	2	2		7	4	
> Seasonal								-	0	1	-	
tau = -0.1						•	• • • • •					
cuu 0.	_ , , ,	0.		.20	0							
> assault	2ам	ts										
		.es Mar	Anr	Mav	.Tun	.Tu 1	Διια	Sen	Oct	Nov	Dec	
1994: 2	3	4	0	5	5	4	2	6 000	3	8	1	
1995: 5	1	2	2	2	8	9	10	2	9	3	2	
1996: 1	3	5	0	3	1	4	- 3	5		4	15	
1997: 6	9	14		6	10	3	8	6		10	6	
1998: 11	8	3		3	5	8	12	7		8	2	
> Seasonal								/	0	0	2	
tau = 0.3						• 21111	• 007					
cau - 0.5	10,		-0		0							
> assault	ЗДМ	+ s										
		.cs Mar	Anr	Matz	Tun	.Tu 1	Aur	Son	Oct	Nov	Dec	
1994: 2	1 e.D 2	Mai 1	дрі 2	5	1	6 001	Aug 2	0 0	3	1	Dec 2	
1995: 2	1	3	2	1	2	9	4	4	1	1	2	
1995. 2	1 3	1	1	1 3	2	9 1	4	4		1 7	0	
1998: 3 1997: 6	2	2		0	6		5	2		3	3	
1997: 6 1998: 3	∠ 5	2	ю З	1	ь 5	10	5	2	4	3 4	3 4	
> Seasonal								2	Ŧ	4	4	
					au⊥t	.JAM	. LS)					
tau = 0.3	, co	ST	-1.,	2040								

Figure 4a. By hour and wkgrp. Common Scaling.



Year

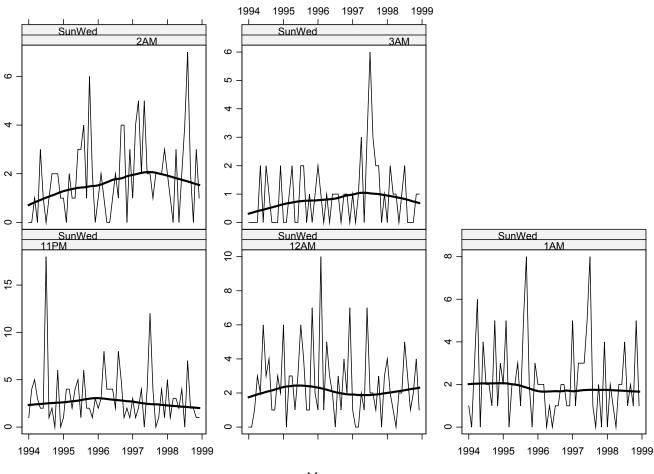


WPS assault by wkgrp and hour

Year

Figure 4b. By hour and wkgrp. Independent Scaling.

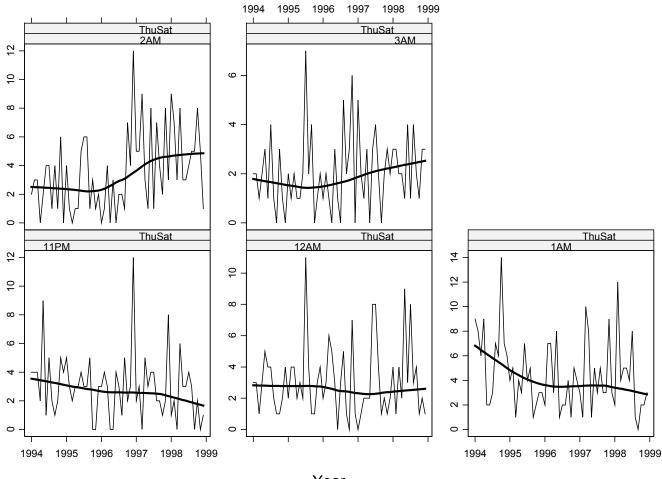
From Figure 4b we see that since 1996 there has been an upward shift at 2AM and 3AM and a downward trend at 11PM for both SunWed and ThuSat. For both weekgroups, 11PM and 12PM are relatively flat. The upward shifts are larger with the ThuSat group.



WPS assault by wkgrp and hour



free vertical scale



WPS assault by wkgrp and hour

Year

free vertical scale

Tabulations and Seasonal Mann-Kendall Tests

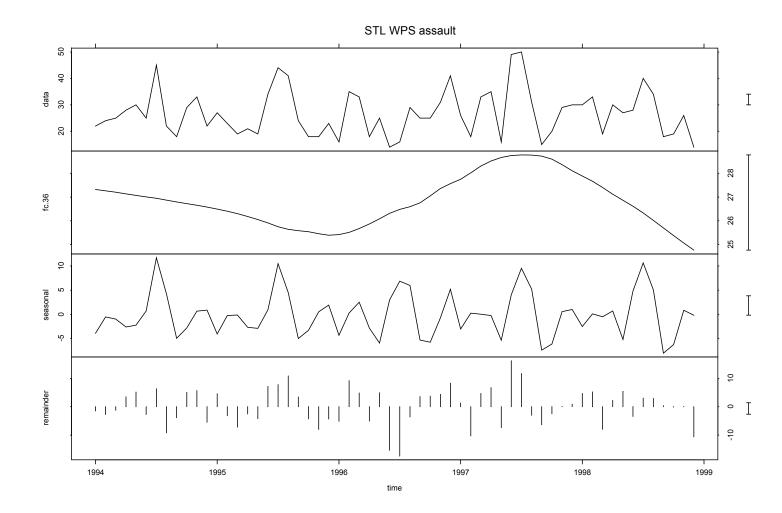
> assault.11PMSunWed.ts										
		Mar Apr		Jun	Jul	Auq	Sep	Oct	Nov	Dec
1994: 1	4	5 3	-	2	18	1	2	0	6	0
1995: 1	4	4 2		5	1	6	2	2	1	3
1996: 2	3	8 4		4	2		5	1	2	1
1997: 3	1	2 4		4	12	4	0	1	4	1
1998: 5	1	3 3		4	0	7	2	2	1	1
> Seasonal									Ŧ	Ŧ
tau = -0.0			-		•	.15 u 11 v	veu.			
tau - 0.0	JU JI,	51 -	94.0-	1 0						
> assault	127M	SunMod	+ 0							
		Mar Apr		Turn	T., 1	7110	Son	Oat	Nou	Dog
1994: 0	0			6	3		1	1	3	2
1995: 6	0	3 3		3	6	4	1	1	7	2
1996: 1	10	1 5		2	0		1	4	2	7
1997: 1	0	0 2		7	2		1	3	0	3
1998: 4	2	1 0		2	5		1	2	4	1
> Seasona					.12AI	MSun≬	Ved.t	cs)		
tau = -0.2	108,	sl =3	9.47	010						
> assault										
Jan	Feb 1	Mar Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1994: 1	0	36	0	4	2	2	1	5	1	3
1995: 2	5	0 2	2	3	1	5	8	2	0	3
1996: 2	2	2 0	1	0	1	1	2	2	1	1
1997: 5	1	3 3	3	5	8	1	0	2	0	4
1998: 0	2	1 0	2	2	4	1	2	1	5	1
> Seasonal	lMannH	Kendall	(assa	ault	.1AM	SunWe	ed.ts	5)		
tau = -0.0										
> assault	.2AMSı	unWed.t	s							
Jan	Feb 1	Mar Apr	May	Jun	Jul	Auq	Sep	Oct	Nov	Dec
1994: 0	0	1 0	_	1	0	1	2	2	2	1
1995: 1	0	2 1	1	3	3	4	1	6	2	0
1996: 1		1 0		1			4	4		3
1997: 1	4	5 2	5	2	2		2	2	2	3
1998: 2	1	0 3		2	4		2	0	3	1
> Seasonal								-	0	-
tau = 0.24			-			0 41111		.,		
0.12	-, -									
> assault	3AMS1	unWed +	S							
		Mar Apr		מנוד	.T11]	Aur	Sen	$0c^+$	Nov	Dec
1994: 0	0	0 0		0 0 0	2	Aug 1	0 0	0000	0	2
1994. 0 1995: 0	0	1 2		0	2		0	1	0	1
1995: 0 1996: 2	1			1	2		0	1 1	1	1
		1 3		⊥ 3			2	1 2		1
	0 2		-	-	6 2				0 1	1
1998: 0		11		1		0	0	0	T	T
> Seasonal				au⊥t	. JAM	SUNWE	εα.τ	5)		
tau = 0.20	, כנ	sl =10	. 228							

tau = 0.211, sl = 8.443%

Figure 5a. STL Analysis

The seasonal component is reasonably stable. There is a slight increase in amplitude of the seasonal oscillations.

R-sq = 43%



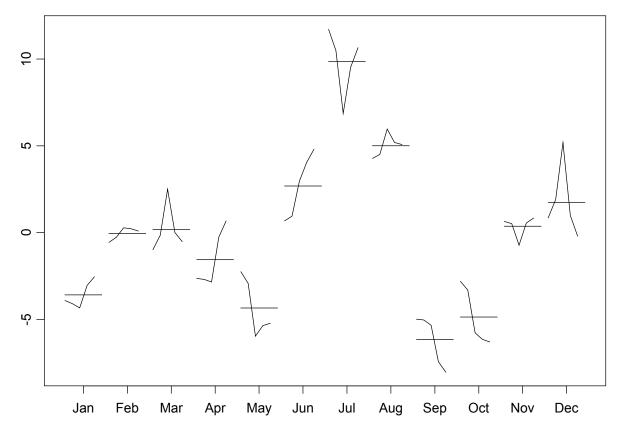
ss.window = 5 ,ss.robust = TRUE , fc.window = 36 , fc.degree = 1

WPS ASSAULTS

Figure 5b. Monthplot of Seasonal Component

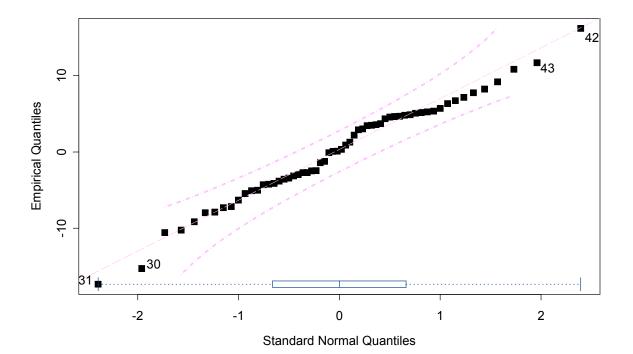
The overall seasonal pattern is with March, July and December maxima and with January, May and September minimum. The seasonal is relatively stable. There are trends up in June and down in September and October.

WPS assault, seasonal



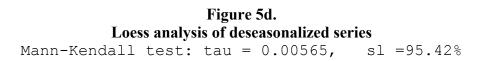
WPS ASSAULTS

Several possible outliers are noted but all statistical tests for normality are fine.

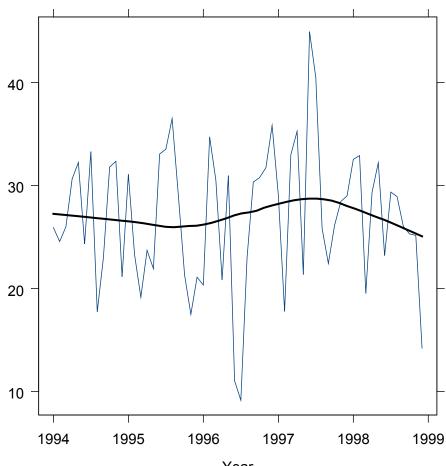


Skewness Coefficient: $g_1 = -0.2457219$, s.l. = 0.1990065 Michael's Statistic: $D_sp = 0.06000243$, s.l. = 0.7382427 Wilk-Shapiro Statistic: W = 0.9830784, s.l. = 0.7981366

> cour	nter	.ts										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1994:	1	2	3	4	5	6	7	8	9	10	11	12
1995:	13	14	15	16	17	18	19	20	21	22	23	24
1996:	25	26	27	28	29	30	31	32	33	34	35	36
1997:	37	38	39	40	41	42	43	44	45	46	47	48
1998:	49	50	51	52	53	54	55	56	57	58	59	60
See Ta	able	1.										



WPS assaults, deseasonalilzed



Year

20

H. WPS Dataset: Impaired

Summary

There were 1573 impaired driving charges in Windsor from January 1994 to December 1998.. <u>Table 1</u>.

<u>Crosstabs</u>. The ThuSat slot accounts for 56% of the impaired charges. About 33% of charges were in the 1AM slot and only 10% in the 3AM. Almost the same as with LPS impaired.

<u>HOUR-WKGP</u>: Unlike the LPS impaired the association between hour and wkgrp is statistically significant at less than 1%. SunWed at 1AM window has more than expected and SunWed at 2AM has fewer than expected if these factors were independent.

<u>HOUR-YEAR</u>: As would be expected from our loess analysis in <u>Figure 3b</u> and <u>Figure 3c</u> there has been a swing to later hours post 1996.

Figure 1a. and Figure 1b. In order to facilitate comparison with the LPS IMPAIRED series we will again use a log transformaton. The slope of the least-squares line is about -0.113619 which implies that impaired charges have been declining at the rate of about 7.6% per year over 1994 to 1998. There is a slight lack of fit of the least-squares line post 1996.

<u>Figure 2</u>. Both exhibit a nearly linear downward slope on log scale. The downward slope for SunWed is slightly greater. The table below shows the slope of the linear least squares fit to each group. Seasonal Mann-Kendall test is significant at <<1% for both cases.

Wkgrp	Slope of fit to log(data,2)	Annual rate of decrease %
SunWed	-0.119	7.9%
ThuSat	-0.109	7.3%

<u>Figure 3a</u>. and <u>Figure 3b</u>. The curves are all different. It is interesting that 2AM and 3AM actually increases after 1996 and the rate of decline for 1AM seems to have increased after 1996. I intend to model this data further with an intervention analysis.

Figure 4a and Figure 4b For SunWed 2AM and 3AM windows show an initial increase post-1996 but then a decline in recent years. SunWed 12AM the decline starting in 1994 flattens out in post-1996. SunWed-1AM there is a drop after 1996 but then it levels off. ThuSat shows strong monotonic decrease at 1AM and the rate seemed to increase slightly post-1996. ThuSat also shows a strong increase at 2AM post-1996 and a small increase post-1996 for the ThuSat-3AM window. The linear decrease for ThuSat at 12AM seems to have flattened out post-1996 for both 11PM and 12AM.

Figure 5a. There is a linear downward trend on the log scale. The seasonal component has lots of local maxima and minima. R-sq = 63.4%

Figure 5b. Local maximum in January, March, April, August-September and November and local minimum in February, April, July and December. But the seasonal component is variable and shows a lot of changes over time. There is are increasing trends for February and March and a decreasing trend for October.

Figure 5c. The normality tests indicate no significant departure from normality in the remainder component.

Figure 5d. Deseasonalized time series trend analysis. Mann-Kendall test: tau = -0.424, sl = 1.779e-4% The thicker line shows the linear least-squares fit and the thinner line shows the loess curve. The least-squares line has a slope of -0.1129019 almost the same as in Figure 1b. A slight lack of fit of the linear least-squares line is evident. The residuals from the least-squares line are approximately normally distributed with no evident outliers and spread-level diagnostics do not reveal any monotone spread.

Crosstabs Analysis

Hour and wkgrp are associated

The ThuSat slot accounts for 56% of the impaired charges. About 33% of charges were in the 1AM slot and only 10% in the 3AM. Almost the same as with LPS impaired.

Unlike the LPS impaired the association between hour and wkgrp is statistically significant at less than 1%. SunWed at 1AM window has more than expected and SunWed at 2AM has fewer than expected if these factors were independent.

```
> crosstabs(counts~wkgrp+hour, data=wps.df, subset=wps.df$type=="impaired",
na.action=na.omit)
Call:
crosstabs(formula = counts ~ wkgrp + hour, data = wps.df, subset = wps.df$type ==
"impaired", na.action = na.omit)
2855 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
wkgrp |hour
   |11PM |12AM |1AM |2AM |3AM |RowTotl|
SunWed |163 |209 |463 |296 |119 |1250 |
    |0.13 |0.17 |0.37 |0.24 |0.095 |0.44
     |0.48 |0.41 |0.49 |0.39 |0.41
                                     1
    |0.057 |0.073 |0.16 |0.1 |0.042 |
_____+
ThuSat |178 |301 |486 |468 |172 |16
|0.11 |0.19 |0.3 |0.29 |0.11 |0.
|0.52 |0.59 |0.51 |0.61 |0.59 |
                                     |1605
                                           1
                                     0.56
     0.062 0.11 0.17 0.16 0.06
_____+
ColTot1|341 |510 |949 |764 |291 |2855 |
    0.12 0.18 0.33 0.27 0.1
                                    Test for independence of all factors
     Chi^2 = 22.39313 d.f.= 4 (p=0.0001673517)
     Yates' correction not used
> get.crosstabs.percenterror(counts~wkgrp+hour)
Crosstabs - percentage error: 100*(Obs-Exp)/Exp,
    [,1] [,2] [,3] [,4] [,5]
[1,] 9 -6 11 -12 -7
[2,] -7 5 -9 9 5
> contrib(counts~wkgrp+hour)
contribution to chi-sq:
    [,1] [,2] [,3] [,4] [,5]
[1,] 1.12 -0.96 2.33 -2.11 -0.74
[2,] -0.99 0.84 -2.06 1.86 0.66
```

Hour and Year are associated: As would be expected from our loess analysis in <u>Figure 3b</u> and <u>Figure 3c</u> there has been a swing to later hours post 1996.

> > crosstabs(counts~hour+year, data=wps.df, subset=wps.df\$type=="impaired", na.action=na.omit) Call: crosstabs(formula = counts ~ hour + year, data = wps.df, subset = wps.df\$type == "impaired", na.action = na.omit) 2855 cases in table ΙN |N/RowTotal| IN/ColTotal| |N/Total | +----+ hour |year |1994 |1995 |1996 |1997 |1998 |RowTotl| _____ _+____+ _+___+ 11PM | 81 | 85 | 60 | 48 | 67 |341 |0.24 |0.25 |0.18 |0.14 0.2 10.12 0.12 0.14 0.1 0.092 |0.14 1 0.021 0.017 0.023 |0.028 |0.03 | 88 |0.17 12AM 1138 1108 186 1 90 1510 0.27 0.21 |0.17 |0.18 0.18 0.21 0.18 0.15 0.17 0.18 0.048 0.038 0.03 0.031 0.032 _____+ |241 |189 |127 |116 1276 1949 1 A M 1 |0.29 |0.25 0.2 |0.13 |0.12 10.33 0.41 0.41 0.33 0.24 0.24 0.097 0.084 0.066 0.044 0.041 ------+----+----+------+----|124 |106 |185 |187 |162 |764 2AM 0.14 0.24 |0.16 0.24 0.21 10.27 0.32 0.18 10.36 0.33 10.19 |0.043 |0.037 |0.065 |0.065 |0.057 | _____+ | 53 | 59 | 74 | 55 3AM | 50 291 0.17 0.18 0.2 0.25 0.19 0.1 |0.075 |0.089 |0.1 |0.14 0.11 1 0.018 0.019 0.021 0.026 0.019 _____+ -+ ColTot1|669 |593 |579 |524 |490 |2855 | |0.23 |0.21 10.2 0.18 10.17 1 ___+ ____ -+---Test for independence of all factors Chi^2 = 141.2984 d.f.= 16 (p=0) Yates' correction not used > get.crosstabs.percenterror(counts~hour+year) Crosstabs - percentage error: 100*(Obs-Exp)/Exp, [,1] [,2] [,3] [,4] [,5] [1,] 20 -13 -23 14 1 2 -17 [2,] 15 -6 3 [3,] 24 22 -2 -27 -29 19 33 0 39 -31 -33 24 [4,]10 [5,] -27 -12 > contrib(counts~hour+year) contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [1,] 0.12 1.68 -1.10 -1.84 1.11 [2,] 1.69 0.20 -1.71 -0.58 0.26 [3,] 3.60 3.13 -0.25 -3.57 -3.67 [4,] -4.11 -4.18 2.41 3.95 2.70 [5,] -2.20 -0.96 0.00 2.82 0.72

•

WPS Impaired Data

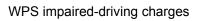
The normality tests on the remainder component indicate the data may be assumed normal with no outliers. However the observations corresponding to the largest and smallest two residuals are identified below in bold red. Test for monotonic trend is significant at $\ll 1\%$.

Table 1.

> imp	.ts											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1994:	66	47	47	81	47	47	66	51	63	46	56	52
1995:	66	45	46	51	55	40	37	46	57	65	49	36
1996:	50	43	51	38	50	53	37	49	45	54	69	40
1997:	63	49	55	27	49	38	28	46	30	50	56	33
1998:	47	42	48	51	33	22	43	39	33	37	51	44

> SeasonalMannKendall(imp.ts)
tau = -0.4, sl =0.07634%

Figure 1.



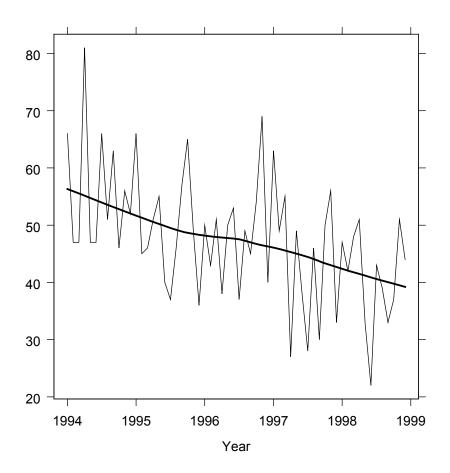
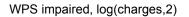


Figure 1b. Logged series

In order to facilitate comparison with the LPS IMPAIRED series we will again use a log transformaton. Figure 1b shows the time series after a log transformation. Specifically a log to the base 2 transformation. This particular version of the log transformation means that we can interpret a one unit change as a doubling and an increase by 0.5 units on the vertical scale corresponds to an increase of about 41% in the original data.

The transformation reveals that the decline in impaired charges is almost linear. Figure 1b shows a thicker line that corresponds to the least squares linear fit. The loess curve is the thinner line which stops at the last observed data value. There is a slight lack of fit of the least-squares line post 1996. The slope of the least-squares line is about -0.113619 which implies that impaired charges have been declining at the rate of about 7.6% per year over 1994 to 1998. Note: $1-2^{-0.113619} \approx 0.07573337$



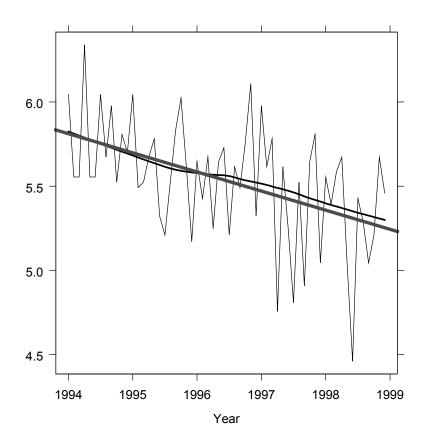
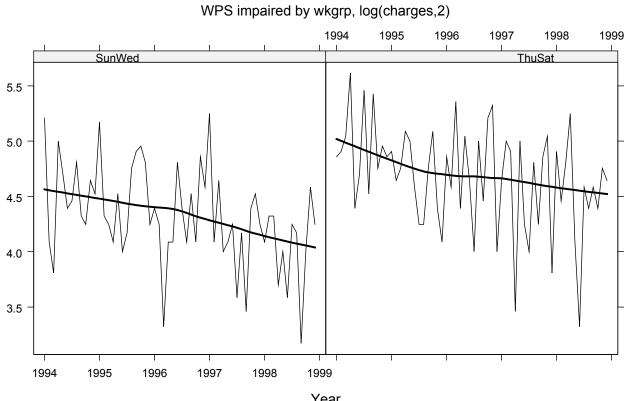


Figure 2. By wkgrp.

Both exhibit a nearly linear downward slope on log scale. The downward slope for SunWed is slightly greater. The table below shows the slope of the linear least squares fit to each group. Seasonal Mann-Kendall test is significant at <<1% for both cases.

Wkgrp	Slope of fit to log(data,2)	Annual rate of decrease %
SunWed	-0.119	7.9%
ThuSat	-0.109	7.3%



Year

Table 2.

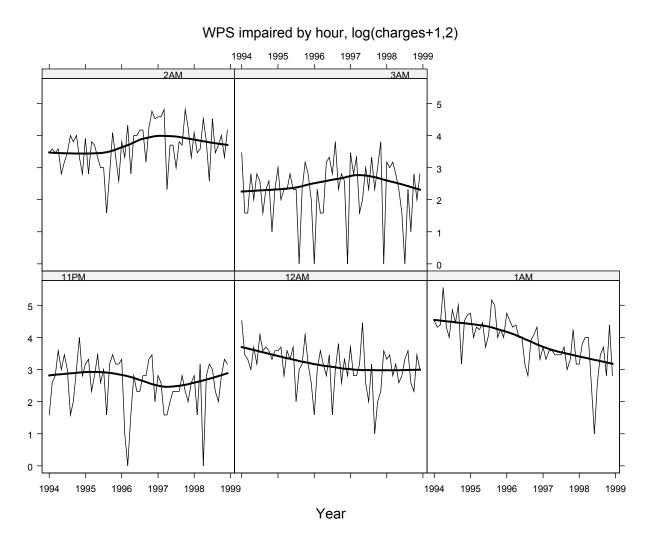
> imp.SunWed.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1994: 37 17 14 32 26 21 22 28 20 19 25 23 16 18 27 1995: 36 20 19 17 23 30 31 28 19 19 10 17 17 28 21 17 23 17 1996: 21 29 24 1997: 38 17 25 16 17 19 12 18 11 21 23 19 1998: 17 20 20 13 16 12 19 18 9 16 24 19 > SeasonalMannKendall(imp.SunWed.ts) tau = -0.38, sl = 0.1468% > imp.ThuSat.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1994: 29 30 33 49 21 26 44 23 43 27 31 29 1995: 30 25 27 34 32 24 19 19 27 34 21 17 1996: 29 24 41 21 33 25 16 32 22 37 40 16 1997: 25 32 30 11 32 19 16 28 19 29 33 14 25 1998: 30 22 28 38 17 10 24 21 24 21 27

> SeasonalMannKendall(imp.ThuSat.ts)

tau = -0.288, sl = 1.516%

Figure 3a. By hour. Common Scaling

The curves are all different. It is interesting that 2AM and 3AM actually increases after 1996 and the rate of decline for 1AM seems to have increased after 1996. I intend to model this data further with an intervention analysis.



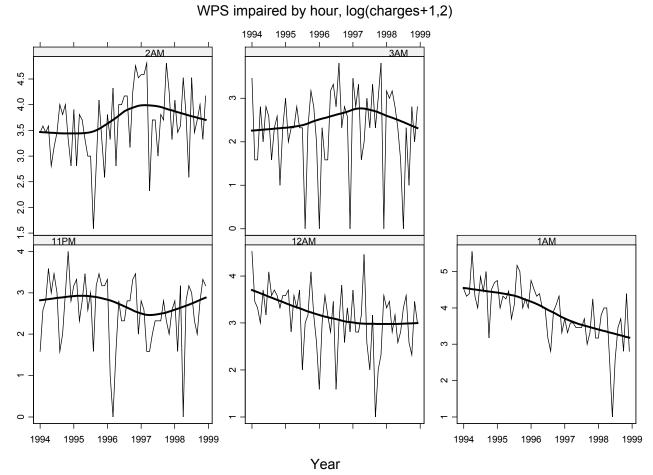
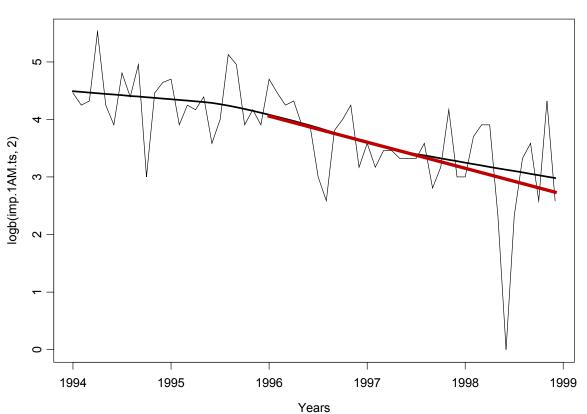


Figure 3b. By hour. Independent vertical scaling.

free vertical scaling

Figure 3c.

Least square line fit (thick red line) to data from 1996 to present. The slope of the least squares line to post-1996 is -0.453 which corresponds to a annual rate of decrease of about 27.0%. Since 1998, the actual impaired driving charges are decreasing slightly slower than this.

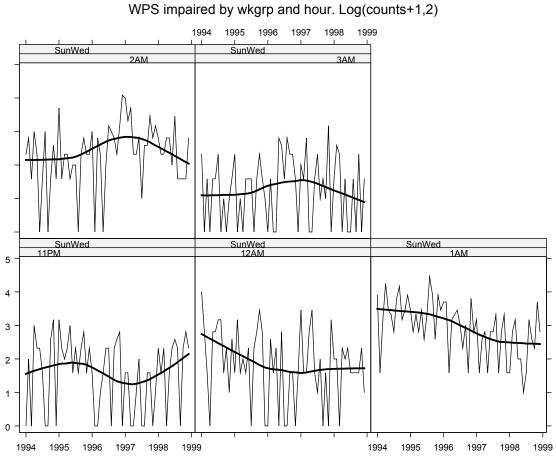


WPS impaired at 1AM, log(charges,2)

Tabulation and Mann-Kendall Tests

Figure 4a. By hour and wkgrp. Common Scaling.

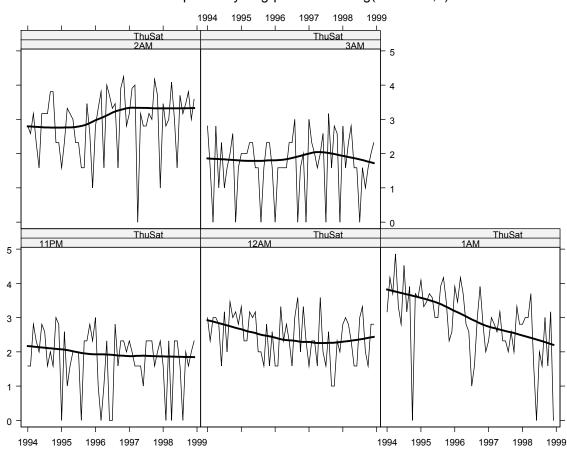
For SunWed 2AM and 3AM windows show an initial increase post-1996 but then a decline in recent years. SunWed 12AM the decline starting in 1994 flattens out in post-1996. SunWed-1AM there is a drop after 1996 but then it levels off.



Year

WPS IMPAIRED

ThuSat shows strong monotonic decrease at 1AM and the rate seemed to increase slightly post-1996. ThuSat also shows a strong increase at 2AM post-1996 and a small increase post-1996 for the ThuSat-3AM window. The linear decrease for ThuSat at 12AM seems to have flattened out post-1996 for both 11PM and 12AM.

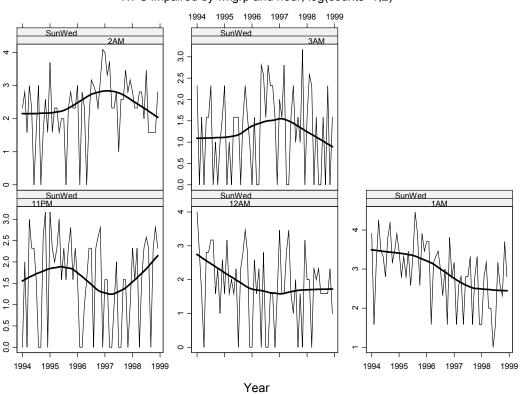


WPS impaired by wkgrp and hour. Log(counts+1,2)

Year

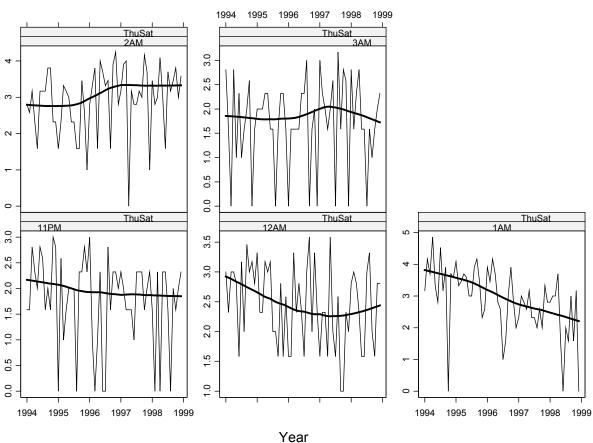
Figure 4b. By hour and wkgrp. Independent Scaling.

Cf. Figure 1b



WPS impaired by wkgrp and hour, log(counts+1,2)

free vertical scale



WPS impaired by wkgrp and hour, log(counts+1,2)

free vertical scale

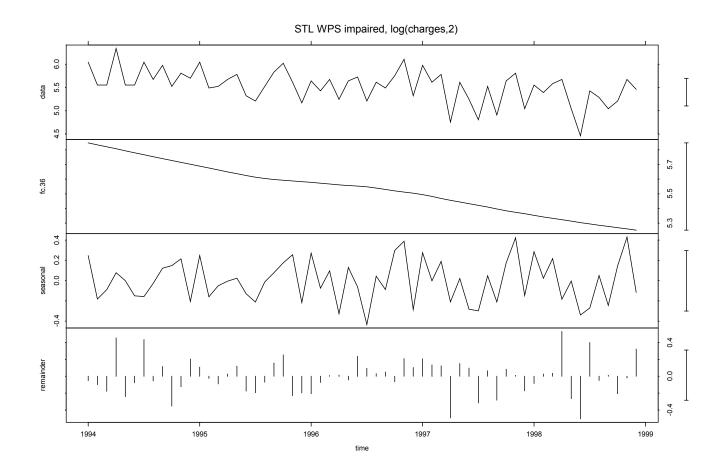
Tabulations and Seasonal Mann-Kendall Tests

WPS IMPAIRED

Figure 5a. STL Analysis

There is a linear downward trend on the log scale. The seasonal component has lots of local maxima and minima.

R-sq = 63.4%



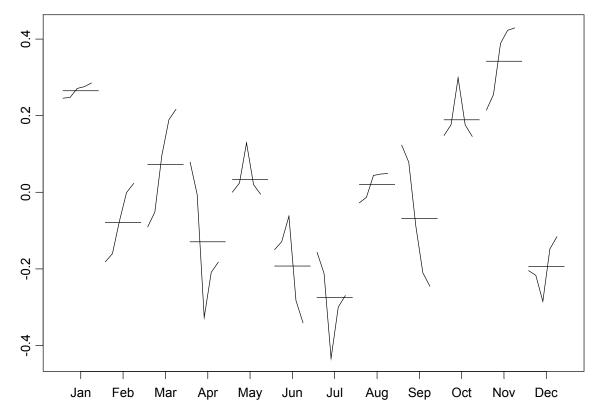
ss.window = 5 ,ss.robust = TRUE , fc.window = 36 , fc.degree = 1

WPS IMPAIRED

Figure 5b. Monthplot of Seasonal Component

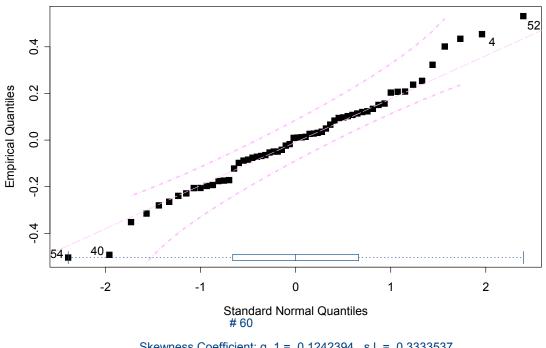
Local maximum in January, March, April, August-September and November and local minimum in February, April, July and December. But the seasonal component is variable and shows a lot of changes over time. There is are increasing trends for February and March and a decreasing trend for October.

WPS impaired, seasonal, logged



WPS IMPAIRED

Figure 5c. Normal Probability Plot The normality tests indicate no significant departure from normality in the remainder component.



Skewness Coefficient: g 1 =	0.1242394 , s.l. = 0.3333537
Michael's Statistic: D sp =	0.0486953, s.l. = 0.580161
Wilk-Shapiro Statistic: W =	0.982025 , s.l. = 0.7582085

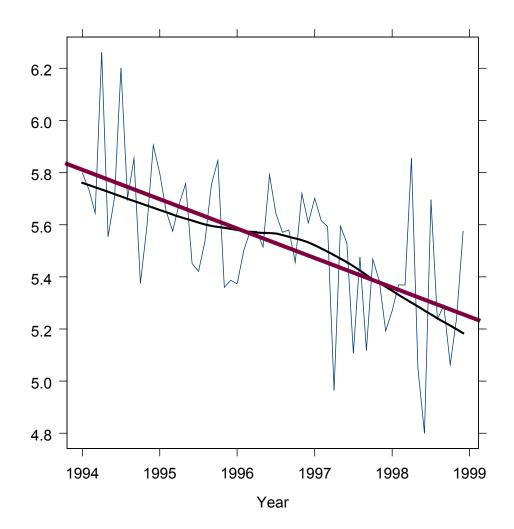
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1994:	1	2	3	4	5	6	7	8	9	10	11	12
1995:	13	14	15	16	17	18	19	20	21	22	23	24
1996:	25	26	27	28	29	30	31	32	33	34	35	36
1997:	37	38	39	40	41	42	43	44	45	46	47	48
1998:	49	50	51	52	53	54	55	56	57	58	59	60
See Ta	able	1.										

Figure 5d. Loess analysis of deseasonalized series

Mann-Kendall test: tau = -0.424, sl =1.779e-4%

The thicker line shows the linear least-squares fit and the thinner line shows the loess curve. The least-squares line has a slope of -0.1129019 almost the same as in Figure 1b. A slight lack of fit of the linear least-squares line is evident. The residuals from the least-squares line are approximately normally distributed with no evident outliers and spread-level diagnostics do not reveal any monotone spread.

WPS impaired, deseasonalilzed, logged



I. FARS Dataset

Introduction

Fatality Accident Reporting System (FARS) Michigan and New York States May 1992 to December 1999, 3956 records Each record in the database corresponds to an accident in which one or more fatalities occurred between the hours of 11PM and 4AM. The number of fatalities is in the field occ#. This field was used to compute the number of deaths in an aggregated dataset. This aggregated datasets gives the number of deaths for each combination of the following factors:

Factors/Covariates of Interest:

year - year of accident (ordered factor)
month - month of accident (ordered factor)
wkgrp - ordered factor variable with labels "SunWed" and "ThuSat"
hour - ordered factor with levels "11PM", "12AM", "1AM", "2AM", "3AM"
bacfactor - levels: "absent", "present"

Response Variables: FARS Deaths (Totals) FARS Deaths with bacfactor present FARS Deaths with bacfactor absent It should be noted that Totals is not just the sum of the other two. There are 45% more data in the totals because so many BAC measurements are missing.

I. FARS Dataset

Summary

ThuSat accounts for 86% of deaths. Among the hour groups 2AM slot (ie. 2AM to 3AM) accounted for most deaths (24%) while 3AM to 4AM (16%) the least and deaths where uniformly distributed over 11PM to 2AM. Of course the traffic intensities might be expected to be less in some cases. The larger number of deaths in the 2AM slot could be due to the patrons of bars leaving at closing time.

BAC is missing in 38% of the records which corresponds to 45% of the deaths. For the remaining 3205 deaths, alcohol was involved about 70% of the time. Crosstabs analysis shows that this percentage, 70%, has not changed over the years. This suggests that the no-drinking & driving campaigns have not been very effective and improvements are due to other safety factors such as day-time running lights, increased use of seatbealts and airbags and improved car design.

Many of the crosstab analyses that were done are statistically significant.

Crosstab Analysis

Wkgrp and hour not independent (significant at 0.09%).

The SunWed group has the strongest interactions with hour. There are 20% more deaths at in the 12AM slot in SunWed than would be expected under the hypothesis of independence. Recall this actually means deaths on Monday not Sunday.

Call: crosstabs(deaths ~ wkqrp + hour, data = fars.df, na.action = na.exclude) 5794 cases in table +----+ | N |N/RowTotal| |N/ColTotal| |N/Total | +----+ wkgrp |hour |11PM |12AM |1AM |2AM |3AM |RowTotl| SunWed | 156 | 195 | 180 | 173 | 109 |813 |
 |0.192
 |0.240
 |0.221
 |0.134
 |0.14
 |

 |0.131
 |0.169
 |0.159
 |0.122
 |0.121
 |

 |0.027
 |0.034
 |0.031
 |0.030
 |
 0.019
 |
 ThuSat |1038 | 959 | 952 |1242 | 790 |4981 | |0.208 |0.193 |0.191 |0.249 |0.159 |0.86 | |0.869 |0.831 |0.841 |0.878 |0.879 | | |0.179 |0.166 |0.164 |0.214 |0.136 | | ColTotl|1194 |1154 |1132 |1415 |899 |5794 | |0.21 |0.20 |0.20 |0.24 |0.16 | | Test for independence of all factors Chi^2 = 18.59653 d.f.= 4 (p=0.0009431474) Yates' correction not used >Contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [1,] -0.89 2.60 1.68 -1.81 -1.53 [2,] 0.36 -1.05 -0.68 0.73 0.62

(Observed-Expected)/Expected Times 100 rounded to nearest integer:

> get.crosstabs.percenterror(deaths~wkgrp+hour)
 [,1] [,2] [,3] [,4] [,5]
[1,] -7 20 13 -13 -14
[2,] 1 -3 -2 2 2

present|2236 | |0.7 | -----+

Bacfactor is missing in about 38% of the cases but in the remaining 3205 deaths where bacfactor is available, 70% are with bacfactor at a positive level.

```
> sum(is.na(fars.df$bacfactor))/nrow(fars.df)
[1] 0.3789181
```

alcohol is present in 70% of the deaths

Call: crosstabs(deaths ~ bacfactor, data = fars.df, na.action = na.exclude) 3205 cases in table +-----+ |N | |N/Total| +-----+ bacfactor| -----+ absent | 969 | |0.3 | -----+

Bacfactor and hour are not independent

The most marked effect is occurs with bacfactor interaction with 11PM. There are 40% more deaths with bacfactor absent at 11PM than would be expected under independence of the classifications.

```
Call:
crosstabs(deaths ~ bacfactor + hour, data = fars.df, na.action = na.exclude)
3205 cases in table
+----
IN
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
bacfactor|hour
  |11PM |12AM |1AM |2AM |3AM |RowTotl|
absent |266 |175 |179 |213 |136 |969 |
|0.275 |0.181 |0.185 |0.220 |0.140 |0.3 |
     |0.423 |0.282 |0.275 |0.268 |0.266 |
|0.083 |0.055 |0.056 |0.066 |0.042 |
present|363 |445 |471 |582 |375 |2236 |
|0.162 |0.199 |0.211 |0.260 |0.168 |0.7 |
|0.577 |0.718 |0.725 |0.732 |0.734 |
|0.113 |0.139 |0.147 |0.182 |0.117 |
ColTotl|629 |620 |650 |795 |511 |3205 |
|0.20 |0.19 |0.20 |0.25 |0.16 | |
Test for independence of all factors
      Chi^2 = 54.40057 d.f.= 4 (p=4.338307e-011)
      Yates' correction not used
```

Contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [1,] 5.50 -0.91 -1.25 -1.76 -1.49 [2,] -3.62 0.60 0.82 1.16 0.98

(Observed-Expected)/Expected Times 100 rounded to nearest integer:

> get.crosstabs.percenterror(deaths~bacfactor+hour)

	[,1]	[,2]	[,3]	[,4]	[,5]	
[1,]	40	-7	-9	-11	-12	
[2,]	-17	3	4	5	5	

Bacfactor and wkgrp are nearly independent.

Although, bacfactor and wkgrp are statistically not independent (at 5% just), the degree of dependence is very small. There are 14% fewer deaths in the SunWed group with bacfactor absent than would be expected.

Call: crosstabs(deaths ~ bacfactor + wkgrp, data = fars.df, na.action = na.exclude) 3205 cases in table +----+ IN |N/RowTotal| |N/ColTotal| |N/Total | +----+ bacfactor | wkgrp |SunWed |ThuSat |RowTotl| ----+ absent | 111 | 858 | 969 | |0.115 |0.885 |0.3 | |0.261 |0.309 | | |0.035 |0.268 | | -----+ present| 314 |1922 |2236 | |0.140 |0.860 |0.7 | |0.739 |0.691 | |0.098 |0.600 | ----+ ColTotl|425 |2780 |3205 | |0.13 |0.87 | | ----+ Test for independence of all factors Chi^2 = 3.936043 d.f.= 1 (p=0.04726182) Yates' correction not used

> contrib(deaths~bacfactor+wkgrp)
Contribution to chi-sq:
 [,1] [,2]
[1,] -1.54 0.6
[2,] 1.02 -0.4

(Observed-Expected)/Expected Times 100 rounded to nearest integer:

> get.crosstabs.percenterror(deaths~bacfactor+wkgrp)
 [,1] [,2]
[1,] -14 2

[2,] 6 -1

Month and bacfactor are independent. The last table on this page shows that there are 18% more deaths in BAC absent in Jun than expected. But the overall test is not close to being significant.

+ N N/RowT N/ColT N/Tota +	 otal												
bacfact	or mont Jan	h Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	RowTot]
absent	+ 50 0.052 0.286 0.016	-+ 49 0.051 0.262 0.015	-+ 58 0.060 0.283 0.018	-+ 78 0.080 0.328 0.024	-+ 91 0.094 0.302 0.028	-+ 106 0.109 0.358 0.033	-+ 108 0.111 0.314 0.034	+ 115 0.119 0.330 0.036	-+ 94 0.097 0.311 0.029	-+ 94 0.097 0.276 0.029	-+ 67 0.069 0.264 0.021	-+ 59 0.061 0.276 0.018	-+ 969 0.3
present	125 0.056 0.714 0.039	138 0.062 0.738 0.043	147 0.066 0.717 0.046	160 0.072 0.672 0.050	210 0.094 0.698 0.066	+ 190 0.085 0.642 0.059	236 0.106 0.686 0.074	234 0.105 0.670 0.073	208 0.093 0.689 0.065	246 0.110 0.724 0.077	187 0.084 0.736 0.058	155 0.069 0.724 0.048	2236 0.7
ColTotl	+ 175 0.055	-+ 187 0.058	-+ 205 0.064	-+ 238 0.074	-+ 301 0.094	-+ 296 0.092	-+ 344 0.107	-+ 349 0.109	-+ 302 0.094	-+ 340 0.106	-+ 254 0.079	-+ 214 0.067	-+ 3205

Test for independence of all factors Chi^2 = 12.27257 d.f.= 11 (p=0.3435074) Yates' correction not used

> contrib(deaths~bacfactor+month)

Contribution to chi-sq:

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]	[,9]	[,10]	[,11]	[,12]
[1,]	-0.40	-1.00	-0.51	0.71	0	1.74	0.39	0.92	0.28	-0.87	-1.12	-0.71
[2,]	0.26	0.66	0.33	-0.47	0	-1.15	-0.26	-0.61	-0.19	0.57	0.74	0.47

(Observed-Expected)/Expected Times 100 rounded to nearest integer:

> ge	t.cros	sstabs	s.perd	center	rror (d	deaths	s~baci	facto	r+mont	ch)		
	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]	[,9]	[,10]	[,11]	[,12]
[1,]	-5	-13	-6	8	0	18	4	9	3	-9	-13	-9
[2,]	2	6	3	-4	0	-8	-2	-4	-1	4	6	4

Month and hour are not independent. In the 11PM slot there are more deaths in Jan, Mar & Dec and fewer in October. The largest difference is the excess deaths in Jun in the 3AM to 4AM slot (38% difference between observed and expected). See table at bottom for more comparisons.

```
> crosstabs( deaths~hour+month, data = fars.df, na.action = na.exclude)
Call:
crosstabs (deaths ~ hour + month, data = fars.df, na.action = na.exclude)
5794 cases in table
+----
| N
IN/RowTotall
|N/ColTotal|
|N/Total |
+----+
hour |month
    |Jan |Feb |Mar |Apr |May |Jun |Jul |Aug |Sep |Oct |Nov |Dec |RowTotl|
11PM | 86 | 52 | 98 | 81 | 117 | 98 | 135 | 124 | 106 | 89 | 97 | 111 | 1194 |
    0.0720 |0.0436 |0.0821 |0.0678 |0.0980 |0.0821 |0.1131 |0.1039 |0.0888 |0.0745 |0.0812 |0.0930 |0.21
    0.2552 |0.1775 |0.2841 |0.2177 |0.2143 |0.1728 |0.2119 |0.1959 |0.1910 |0.1471 |0.2055 |0.2569 |
    0.0148 |0.0090 |0.0169 |0.0140 |0.0202 |0.0169 |0.0233 |0.0214 |0.0183 |0.0154 |0.0167 |0.0192 |
12AM | 72 | 71 | 80 | 75 | 131 | 125 | 100 | 127 | 108 | 120 | 78 | 67 | 1154 |
    0.0624 |0.0615 |0.0693 |0.0650 |0.1135 |0.1083 |0.0867 |0.1101 |0.0936 |0.1040 |0.0676 |0.0581 |0.20
    0.2136 |0.2423 |0.2319 |0.2016 |0.2399 |0.2205 |0.1570 |0.2006 |0.1946 |0.1983 |0.1653 |0.1551 |
    |0.0124 |0.0123 |0.0138 |0.0129 |0.0226 |0.0216 |0.0173 |0.0219 |0.0186 |0.0207 |0.0135 |0.0116 |
1AM | 47 | 64 | 47 | 91 | 106 | 85 | 146 | 103 | 147 | 131 | 86 | 79 | 1132 |
    0.0415 0.0565 0.0415 0.0804 0.0936 0.0751 0.1290 0.0910 0.1299 0.1157 0.0760 0.0698 0.20
    0.1395 |0.2184 |0.1362 |0.2446 |0.1941 |0.1499 |0.2292 |0.1627 |0.2649 |0.2165 |0.1822 |0.1829 |
    0.0081 |0.0110 |0.0081 |0.0157 |0.0183 |0.0147 |0.0252 |0.0178 |0.0254 |0.0226 |0.0148 |0.0136 |
_____+
2AM | 85 | 63 | 72 | 60 | 123 | 134 | 163 | 184 | 120 | 186 | 133 | 92 | 1415 |
    0.0601 |0.0445 |0.0509 |0.0424 |0.0869 |0.0947 |0.1152 |0.1300 |0.0848 |0.1314 |0.0940 |0.0650 |0.24
    0.2522 |0.2150 |0.2087 |0.1613 |0.2253 |0.2363 |0.2559 |0.2907 |0.2162 |0.3074 |0.2818 |0.2130 |
    0.0147 |0.0109 |0.0124 |0.0104 |0.0212 |0.0231 |0.0281 |0.0318 |0.0207 |0.0321 |0.0230 |0.0159 |
_____+
3AM | 47 | 43 | 48 | 65 | 69 | 125 | 93 | 95 | 74 | 79 | 78 | 83 | 899
    0.0523 |0.0478 |0.0534 |0.0723 |0.0768 |0.1390 |0.1034 |0.1057 |0.0823 |0.0879 |0.0868 |0.0923 |0.16
    0.1395 |0.1468 |0.1391 |0.1747 |0.1264 |0.2205 |0.1460 |0.1501 |0.1333 |0.1306 |0.1653 |0.1921 |
    0.0081 |0.0074 |0.0083 |0.0112 |0.0119 |0.0216 |0.0161 |0.0164 |0.0128 |0.0136 |0.0135 |0.0143 |
ColTotl|337 |293 |345 |372 |546 |567 |637 |633 |555 |605 |472 |432 |5794 |
    0.058 0.051 0.060 0.064 0.094 0.098 0.110 0.109 0.096 0.104 0.081 0.075
```

Test for independence of all factors

Chi^2 = 170.65 d.f.= 44 (p=1.110223e-016) Yates' correction not used

Contribution to chi-sq:

(Observed-Expected)/Expected Times 100 rounded to nearest integer:

	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
11PM	22	-16	34	4	2	-18	1	-7	-9	-30	-2	22
12AM	7	20	15	1	20	10	-22	1	-3	0	-17	-23
1AM	-30	8	-32	23	-3	-25	15	-18	32	9	-8	-9
2AM	5	-11	-14	-33	-6	-2	7	21	-10	29	18	-12
ЗАМ	-13	-9	-14	10	-21	38	-9	-6	-17	-18	4	19

Wkgrp and month are not independent. The last table shows the percentage differences between observed and expected. There are larger fluctuations in the SunWed group, for example there are 34% more deaths in March than would be expected if the classifications SunWed and Month were independent.

5794 ca	ses in t + otal 0tal 1	2	rp + mont	h, data	= fars.c	lf, na.ac	tion = n	a.exclud	le)				
21	month Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	lOct	Nov	Dec	RowTotl
	35 0.0431 0.1039 0.0060	43 0.0529 0.1468 0.0074		+ 54 0.0664 0.1452 0.0093	99 0.1218 0.1813 0.0171	64 0.0787 0.1129 0.0110	82 0.1009 0.1287 0.0142		80 0.0984 0.1441 0.0138		57 0.0701 0.1208 0.0098	44 0.0541 0.1019 0.0076	813 0.14
	302 0.0606 0.8961 0.0521	250 0.0502 0.8532 0.0431		318 0.0638 0.8548 0.0549	+ 447 0.0897 0.8187 0.0771	503 0.1010 0.8871 0.0868	+ 555 0.1114 0.8713 0.0958	551 0.1106 0.8705 0.0951	475 0.0954 0.8559 0.0820	497 0.0998 0.8215 0.0858	415 0.0833 0.8792 0.0716	388 0.0779 0.8981 0.0670	++ 4981 0.86
ColTotl	+ 337 0.058 +	293 0.051	+ 345 0.060	+ 372 0.064 +	+ 546 0.094 +	+ 567 0.098	+ 637 0.110 +	633 0.109	+ 555 0.096	605 0.104	472 0.081	+ 432 0.075	++ 5794 ++

```
Test for independence of all factors
Chi^2 = 37.14727 d.f.= 11 (p=0.0001088637)
Yates' correction not used
```

> contrib(deaths~wkgrp+month)

Contribution to chi-sq:

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]	[,9]	[,10]	[,11]	[,12]
[1,]	-1.79	0.29	2.38	0.25	2.56	-1.74	-0.78	-0.72	0.24	2.51	-1.13	-2.13
[2,]	0.72	-0.12	-0.96	-0.10	-1.03	0.70	0.32	0.29	-0.10	-1.01	0.46	0.86

(Observed-Expected)/Expected times 100 rounded to nearest integer:

> get	c.cros	sstabs	s.perd	centei	cror(d	deaths	s~wkgı	rp+mor	nth)			
	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]	[,9]	[,10]	[,11]	[,12]
[1,]	-26	5	34	3	29	-20	-8	-8	3	27	-14	-27
[2,]	4	-1	-6	-1	-5	3	1	1	0	-4	2	4

Bacfactor and year. There are no statistically significant changes over time.

Call: crosstabs(deaths ~ bacfactor + year, data = fars.df, na.action = na.exclude) 3205 cases in table +----+ | N |N/RowTotal| |N/ColTotal| |N/Total | +----+ bacfactor|year |1992 |1993 |1994 |1995 |1996 |1997 |1998 |1999 |RowTotl| +----+----+ absent | 102 | 154 | 149 | 148 | 126 | 98 | 100 | 92 | 969

 |0.105
 |0.159
 |0.154
 |0.153
 |0.130
 |0.101
 |0.103
 |0.095
 |0.3

 |0.266
 |0.301
 |0.333
 |0.304
 |0.316
 |0.288
 |0.316
 |0.287
 |

 |0.032
 |0.048
 |0.046
 |0.039
 |0.031
 |0.031
 |0.029
 |

 present|281 |357 |299 |339 |273 |242 |216 |229 |2236 | |0.126 |0.160 |0.134 |0.152 |0.122 |0.108 |0.097 |0.102 |0.7 | |0.734 |0.699 |0.667 |0.696 |0.684 |0.712 |0.684 |0.713 | |0.088 |0.111 |0.093 |0.106 |0.085 |0.076 |0.067 |0.071 | _____+ ColTotl|383 |511 |448 |487 |399 |340 |316 |321 |3205 | 0.120 |0.159 |0.140 |0.152 |0.124 |0.106 |0.099 |0.100 | Test for independence of all factors Chi^2 = 5.645549 d.f.= 7 (p=0.5816909) Yates' correction not used > contrib(deaths~bacfactor+year) Contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [1,] -1.28 -0.04 1.16 0.06 0.49 -0.47 0.46 -0.51

[2,] 0.84 0.03 -0.77 -0.04 -0.32 0.31 -0.30 0.34

(Observed-Expected)/Expected times 100 rounded to nearest integer:

> get.crosstabs.percenterror(deaths~bacfactor+year) [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [1,] -12 0 10 1 4 -5 5 -5 [2,] 5 0 -4 0 -2 2 -2 2

Wkgrp and Year.

There are 24% more deaths in the SunWed group in 1996 than expected but overall the dependence between wkgrp and year is not statistically significant at 5%.

```
Call:
crosstabs (deaths ~ wkqrp + year, data = fars.df, na.action = na.exclude)
5794 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
wkgrp |year
  |1992 |1993 |1994 |1995 |1996 |1997 |1998 |1999 |RowTotl|
_____+
SunWed | 79 |113 | 93 |100 |128 |101 | 85 |114 |813
                                                                             1

      |0.097
      |0.139
      |0.114
      |0.123
      |0.157
      |0.124
      |0.105
      |0.140
      |0.14

      |0.134
      |0.131
      |0.131
      |0.132
      |0.174
      |0.133
      |0.126
      |0.162
      |

      |0.014
      |0.020
      |0.016
      |0.017
      |0.022
      |0.017
      |0.015
      |0.020
      |

ThuSat |509 |752 |615 |656 |608 |660 |591 |590 |4981 |
|0.102 |0.151 |0.123 |0.132 |0.122 |0.133 |0.119 |0.118 |0.86 |
|0.866 |0.869 |0.869 |0.868 |0.826 |0.867 |0.874 |0.838 |
      0.088 |0.130 |0.106 |0.113 |0.105 |0.114 |0.102 |0.102 |
ColTot1|588 |865 |708 |756 |736 |761 |676 |704 |5794 |
   0.10 0.15 0.12 0.13 0.13 0.13 0.12 0.12
Test for independence of all factors
      Chi^2 = 12.89018 d.f.= 7 (p=0.07482989)
       Yates' correction not used
> contrib(deaths~wkgrp+year)
Contribution to chi-sq:
       [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8]
[1,] -0.39 -0.76 -0.64 -0.59 2.43 -0.56 -1.01 1.53
[2,] 0.16 0.31 0.26 0.24 -0.98 0.23 0.41 -0.62
```

(Observed-Expected)/Expected Times 100 rounded to nearest integer:

> get.crosstabs.percenterror(deaths~wkgrp+year)
 [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8]
[1,] -4 -7 -6 -6 24 -5 -10 15
[2,] 1 1 1 -4 1 2 -3

Hour and year are not independent.

But I don't see any interesting trends. The changes seem to be strongest in the 11PM and 12AM slots. The largest effect occurred in 1999 when there was a 32% excess of deaths in the 12AM slot.

5794 ca: + N N/RowT0 N/ColT0 N/Tota	ses in t + otal otal		+ year,	data =	fars.df,	na.acti	on = na.	exclude)	
hour	year	1993	1994	1995	1996	1997	1998	1999	RowTotl ++
	0.121 0.247	167 0.140 0.193 0.029	0.135 0.227	0.113 0.179	184 0.154 0.250 0.032	0.179	150 0.126 0.222 0.026	116 0.097 0.165 0.020	1194 0.21
	0.084 0.165	0.140	0.105 0.171	0.146 0.224	123 0.107 0.167 0.021	0.156 0.237	117 0.101 0.173 0.020	185 0.160 0.263 0.032	1154 0.20
	0.106		0.133 0.213	0.138	+ 141 0.125 0.192 0.024		146 0.129 0.216 0.025	135 0.119 0.192 0.023	1132 0.20
	0.098 0.235		0.127 0.253	0.126 0.235	162 0.114 0.220 0.028		148 0.105 0.219 0.026	170 0.120 0.241 0.029	++ 1415 0.24
	0.098		0.107	0.131	+ 126 0.140 0.171 0.022		115 0.128 0.170 0.020	+ 98 0.109 0.139 0.017	++ 899 0.16
	0.10 +	0.15 +		0.13 +	736 0.13 +	761 0.13 +	676 0.12 +	704 0.12 +	++ 5794 ++

Test for independence of all factors Chi^2 = 86.38715 d.f.= 28 (p=7.19933e-008) Yates' correction not used

> contrib(deaths~hour+year)

Contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [1,] 2.16 -0.84 1.25 -1.67 2.63 -1.66 0.91 -2.41 [2,] -1.86 -0.78 -1.69 1.50 -1.95 2.31 -1.52 3.78 [3,] 0.48 -1.92 1.08 0.68 -0.23 -0.79 1.21 -0.22 [4,] -0.47 2.05 0.46 -0.49 -1.32 0.96 -1.33 -0.15 [5,] -0.34 1.45 -1.32 0.06 1.10 -1.02 0.99 -1.07

(Observed-Expected)/Expected Times 100 rounded to nearest integer:

> get	t.cros	sstabs	s.perd	centei	eror(d	leaths	s~hour	r+year)
	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]
[1,]	20	-6	10	-13	21	-13	8	-20
[2,]	-17	-6	-14	12	-16	19	-13	32
[3,]	4	-15	9	6	-2	-7	11	-2
[4,]	-4	14	4	-4	-10	7	-10	-1
[5,]	-4	13	-13	1	10	-9	10	-10

I. FARS Dataset – Time Series

Summary

There are downward trends in FARS deaths, death with BAC positive and deaths with BAC negative.

<u>Table 1.</u> FARS deaths per month for (a) totals, (b) BAC positive and (c) BAC negative. Note that (b) + (c) \neq (a) since BAC is n/a for about 38% of the records.

Figure 1. Deaths per Month Time Series Plot. The lowess trend line decreases from 73.4 deaths per month to 57.7 deaths per month which corresponds to an average annual rate of decrease of 3.1%.

Figure 2. Fars Deaths with BAC Present. The lowess trend line decreases from 32.9 deaths per month to 17.7 deaths per month which corresponds to an average annual rate of about 7.8%.

Figure 3. Fars Deaths with BAC Absent. The lowess trend line decreases from 13.0 deaths per month to 6.0 deaths per month which corresponds to about 9.6% per year.

Figure 3Bi and 3Bii. The rate of decrease is higher in both cases with BAC present and BAC absent. There must be some sort of selection effect.

<u>Table 2</u>. Mean monthly death rates, percentage change and Mann-Kendall trend tests. Remark: the mean monthly Total is not the sum of the mean monthly values for BAC present and absent since BAC is missing in about 38% of the cases. The magnitude of the downward trend is the same in both BAC present and absent but is somewhat less than in Total.

	1993	1998	%change	tau	sl
Total	72.1	58.7	18.6	-0.14	5%
BAC Present	29.8	19.1	35.8	-0.38	<10 ⁻⁶
BAC Absent	12.8	7.7	40.2	-0.27	<10 ⁻⁴

Table 4a Annual deaths and Mann-Kendall Trend Tests, 2AM and ThuSat have significant downward trends at 7% and 3% significance levels on a two-sided Mann-Kendall trend test.

Figure <u>4a</u>, <u>4b</u> show trellis time series displays of the annual deaths.

Figure 5 shows a drop in the deaths starting in 1996 in both cases (BAC present and BAC absent). The decline in BAC present deaths is very large but is actually less than the decline in BAC absent. But these declines are in reality perhaps not so important because the percentage decline in FARS deaths totals is much less (recall that BAC is missing in 45% of the deaths). This summarized:

Year	1993	1999	% change
Total	865	704	19%
Bac present	357	229	35%
Bac absent	154	92	40%

<u>Multipanel trellised time series plots</u> showing the monthly time series of deaths totals, deaths with BAC present and deaths BAC absent were constructed for each combination of the factors wkgrp and hour. Robust loess locally-linear 60% smooth trend lines are drawn on the plots. Each trellis plot decomposes the aggregated time series according to the factor levels. Two types of trellis plots are used:

- (i) the scales are kept the same on each trellis plot to make comparisons between the different panels easy
- (ii) the deaths are log transformed and separate scaling is used on the vertical axis to allow better visualization of the trend

Figure 6a-ii, FARS deaths exhibit a downward trend for 11PM both SunWed and ThuSat, an upward trend for SunWed 3AM and an upward trend for ThuSat-2AM. There are many more deaths in the ThuSat weekgroup.

Comparison of panels for 3AM-ThuSat and 2AM-ThuSat in <u>Figure 6a-i</u> suggest that some spikes occur at common times. Such an temporal association could be caused by poor driving conditions on particular nights. However closer investigation shows that there is really no apparent association, see <u>Figure 6a-iiia</u> and <u>Figure 6a-iiiβ</u> which shows comparisons when the aspect-ratio is reduced to improve visualization on the horizontal scale and a log transformation is made to reduce the effect of outliers on the resolution.

Figure 6b-ii, FARS deaths with BAC present, exhibits a downward trend for ThuSat at 2AM and 12AM and possibly also SunWed at 2AM.

Figure 6c-ii, FARS deaths with BAC absent shows a downward trend for all levels of hour for ThuSat but little change for SunWed.

Seasonal-Trend-Loess Decompositions were examined. For all series, the downward trend is evident as well as a change in the seasonal pattern occurring in 1996 (Figures 7, 9, 11).

<u>Figures 7a,b</u>. STL FARS deaths. Peak in July and reach minimum in February. These deaths are relatively high from May through to October. $R^2 = 53.4\%$.

Figure 7c. The deseasonalized series showns an apparent change in variability around 1997. The trend is highly significant with Mann-Kendall test, tau = -0.214, sl = 1.261%. The loess trend declines from about 70.5 deaths/month at the beginning of 1992 to 56.7 deaths/month by the end of 1999. This corresponds to a percentage annual rate of decline of about 2.8%.

<u>Figures 8a,b</u> STL FARS deaths with BAC present. Peak in October and are a minimum in Jan. These deaths are also relatively high in July and August. There are strong increasing trends in February and September and a decreasing trend in October and to a lesser extend in March and April too. $R^2 = 90.5\%$.

Figure 8c. Trend in the deseasonalized component is obviously significant. Mann-Kendall trend test: tau = -0.461, sl =8.001e-6%. The loess trend declines from about 70.5 deaths/month at the beginning of 1992 to 56.7 deaths/month by the end of 1999 which corresponds to a percentage annual rate of decline of about 2.8%.

<u>Figures 9a,b</u>. STL FARS deaths with BAC absent. Peak in August but are relatively high from April through to October and are relatively low in December, January, February and March. There has been a sharp downward trend in deaths in Nov. $R^2 = 93.2\%$.

Figure 9c. There is a significant downward trend. The Mann-Kendall test yields, tau = -0.338, sl = 0.008465%. The deseasonalized FARS BAC absent deaths decline from about 12.7 at the beginning of 1992 to about 7.2% by the end of 1999. The average annual percentage decrease is about 7.2%.

Figure 9d. There are some outliers evident. Observations 31 and 85 corresond to Nov. 1994 and May 1999 respectively. See Table 1c.

Time Series: Tables and Plots

Monthly time series are plotted and a 60% locally linear robust smoothed trend line is shown.

Table 1a. FARS, deaths, total number of deaths per month

> deaths.fars.ts												
	Jan	Feb	Mar	Apr	Мау	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					89	94	73	77	72	73	48	62
1993:	66	52	67	46	82	60	99	95	72	107	62	57
1994:	41	16	31	73	54	61	92	49	57	95	84	55
1995:	42	36	85	77	66	48	86	93	78	57	35	53
1996:	48	46	33	38	74	102	59	74	67	74	68	53
1997:	60	47	50	53	71	66	85	71	70	79	55	54
1998:	39	46	46	39	47	65	83	73	77	48	70	43
1999:	41	50	33	46	63	71	60	101	62	72	50	55

Table 1b. FARS, deaths BAC present, deaths per month > deaths.farsbacp.ts

> deaths.farsbacp.ts														
		Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
	1992:					47	39	42	41	29	38	24	21	
	1993:	23	25	34	22	27	22	31	36	26	53	34	24	
	1994:	18	10	20	25	32	31	30	25	20	45	23	20	
	1995:	27	16	32	46	30	14	37	33	28	27	23	26	
	1996:	19	26	16	9	21	26	23	29	31	20	29	24	
	1997:	10	18	20	17	20	19	31	22	37	22	19	7	
	1998:	14	17	19	23	15	19	25	14	18	19	21	12	
	1999:	14	26	6	18	18	20	17	34	19	22	14	21	

Table 1c. FARS, deaths BAC absent, deaths per month

> deaths.farsbaca.ts

	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					12	22	9	11	20	7	10	11
1993:	9	11	17	15	15	12	11	19	12	17	8	8
1994:	8	2	8	16	6	13	22	9	7	19	32	7
1995:	7	11	11	18	7	10	12	26	23	7	5	11
1996:	9	5	4	14	18	25	12	10	8	12	2	7
1997:	8	11	6	8	16	8	7	12	3	9	4	6
1998:	5	3	5	6	14	12	23	8	9	4	4	7
1999:	4	6	7	1	3	4	12	20	12	19	2	2

Nov. 1994 (observation 31) and May 1999 (observation 85) seem to be outliers in Table 1c – see STL Analysis.

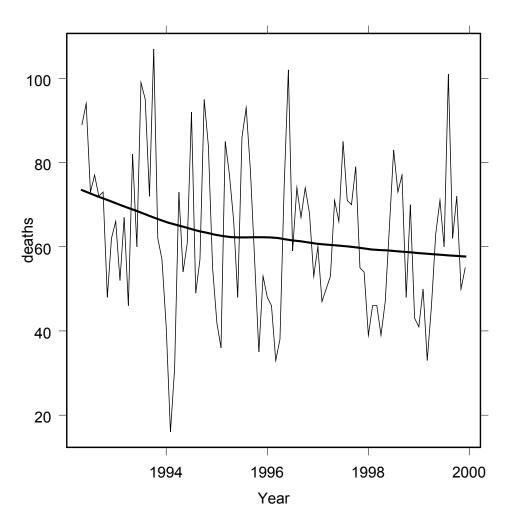
Table 2. Mean monthly death rates, percentage change and Seasonal Mann-Kendall trend tests. Remark: the mean monthly Total is not the sum of the mean monthly values for BAC present and absent since BAC is missing in about 38% of the cases. The magnitude of the downward trend is the same in both BAC present and absent but is somewhat less than in Total.

	1993	1998	%change	tau	sl
Total	72.1	58.7	18.6	-0.18	48
BAC Present	29.8	19.1	35.8	-0.42	<10 ⁻⁶
BAC Absent	12.8	7.7	40.2	-0.30	<10 ⁻⁴

```
Seasonal Mann-Kendall Tests
> SeasonalMannKendall(deaths.fars.ts)
tau = -0.181, sl =3.668%
> SeasonalMannKendall(deaths.farsbacp.ts)
tau = -0.425, sl =9.23e-5%
> SeasonalMannKendall(deaths.farsbaca.ts)
tau = -0.295, sl =0.07391%
Annual Series
> aggregate(deaths.fars.ts,1,mean)
1993: 72.08333 59.00000 63.00000 61.33333 63.41667 56.33333 58.66667
start deltat frequency
 1993 1 1
> pc.diff(aggregate(deaths.fars.ts,1,mean))
[1] 0.1861272
> aggregate(deaths.farsbacp.ts,1,mean)
1993: 29.75000 24.91667 28.25000 22.75000 20.16667 18.00000 19.08333
start deltat frequency
 1993
        1 1
> pc.diff(aggregate(deaths.farsbacp.ts,1,mean))
[1] 0.3585434
> aggregate(deaths.farsbaca.ts,1,mean)
1993: 12.833333 12.416667 12.333333 10.500000 8.166667 8.333333 7.666667
start deltat frequency
 1993
        1 1
> pc.diff(aggregate(deaths.farsbaca.ts,1,mean))
[1] 0.4025974
```

Figure 1: Deaths per Month Time Series Plot

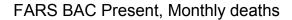
The lowess trend line decreases from 73.4 deaths per month to 57.7 deaths per month which corresponds to an average annual rate of decrease of 3.1%. > 1-(57.7/73.4)^(1/(92/12)) [1] 0.03090372



FARS, Monthly deaths

Figure 2: Fars Deaths with BAC Present

The lowess trend line decreases from 32.9 deaths per month to 17.7 deaths per month which corresponds to an average annual rate of about 7.8%. > $1-(17.7/32.9)^{(1/(92/12))}$ [1] 0.07767495



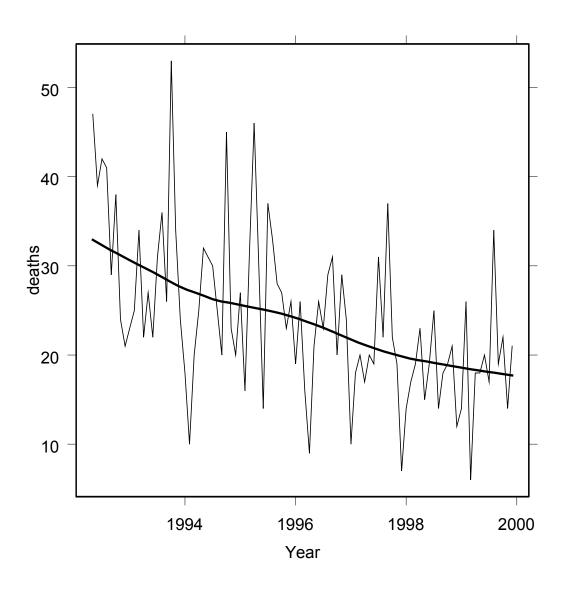
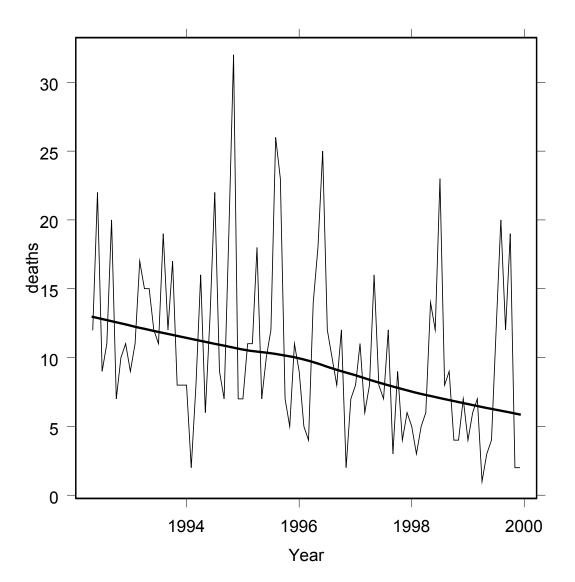


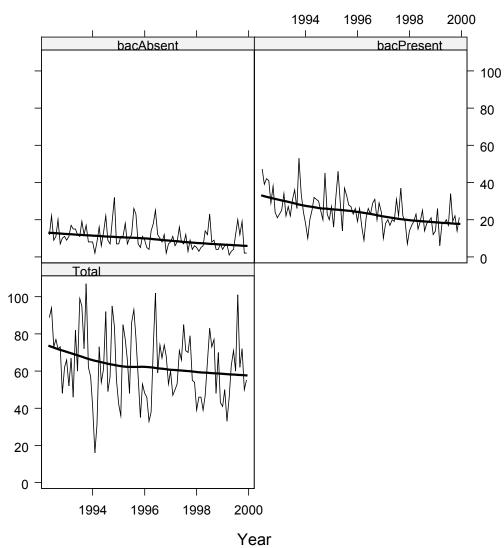
Figure 3: Fars Deaths with BAC Absent

The lowess trend line decreases from 13.0 deaths per month to 6.0 deaths per month which corresponds to about 9.6% per year. > 1-(6.0/13.0)^(1/(92/12)) [1] 0.09593214



FARS BAC Absent, Monthly deaths

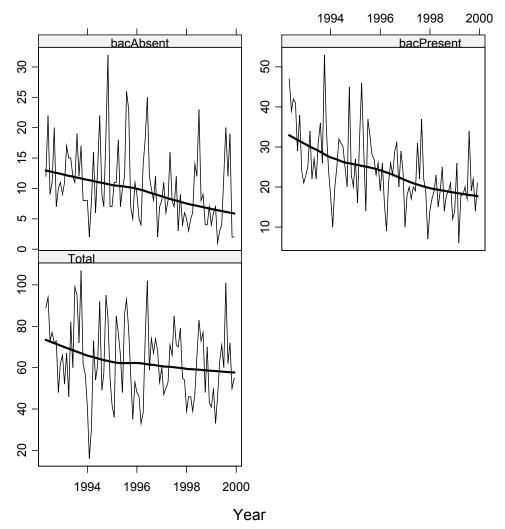
Figure 3Bi. Comparison of FARS deaths for total, BAC Present and BAC Absent. Same scale.



FARS, deaths, same scale

Figure 3Bii. Comparison of FARS deaths for total, BAC Present and BAC Absent. Sliced scale.

With the sliced scale the slopes are comparable since the number of data units per inch is the same on each graph. We clearly see that the slopes are greater in bacAbsent and bacPresent .



FARS, deaths, sliced scale

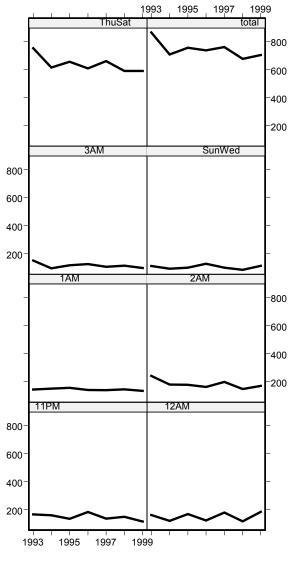
Annual Totals

Table 4a. Annual deaths and Mann-Kendall Trend Tests, 2AM and ThuSat have significant downward trends at 7% and 3% significance levels on a two-sided Mann-Kendall trend test.

Aggregate annual number of deaths

<pre>> (fars.death.annual[,-1])</pre>										
	Total	PM11	AM12	AM1	AM2	AM3	SunWed	ThuSat		
1993	865	167	162	144	241	151	113	752		
1994	708	161	121	151	179	96	93	615		
1995	756	135	169	156	178	118	100	656		
1996	736	184	123	141	162	126	128	608		
1997	761	136	180	139	199	107	101	660		
1998	676	150	117	146	148	115	85	591		
1999	704	116	185	135	170	98	114	590	>	
MannKei	ndallAnnua	l(fars.c	leath.ar	nual[,	-1])					
	Total	PM1	.1	AM12		AM1	AM2	AM3	SunWed ThuSat	
tau -0	.4285714 -	0.428571	4 0.238	80952 -	-0.4285	714 -0	.52380949	-0.3333333	0.04761904 -0.61904758	
sl O	.1361111	0.136111	1 0.561	9048	0.1361	.111 0	.06904757	0.2388889	1.00000000 0.03015876	

Figure 4a: Annual Deaths



Year

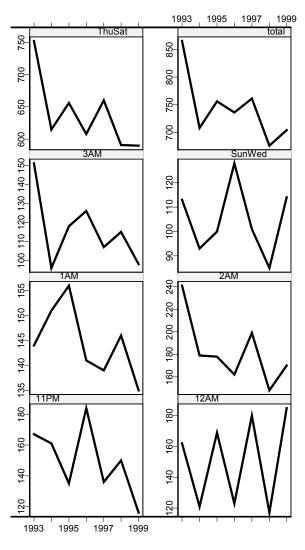


Figure 4b: Annual Deaths (Panels Scaled Independently)

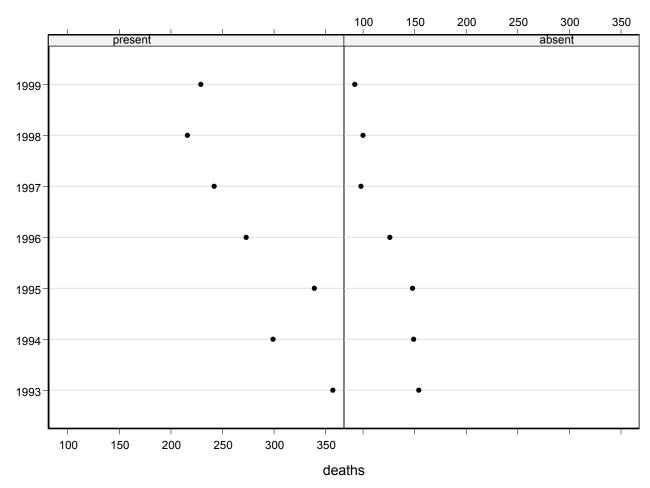
Year

Figure 5: Annual FARS Deaths, Comparison BAC Absent and Present

Number of Deaths, FARS, BAC Absent, 1993 to 1999 > aggregate(deaths.farsbaca.ts, ndeltat=1) 1993: 154 149 148 126 98 100 92

Number of Deaths, FARS, BAC Present, 1993 to 1999 > aggregate(deaths.farsbacp.ts, ndeltat=1) 1993: 357 299 339 273 242 216 229

Figure 6 shows a drop in the deaths starting in 1996 in both cases (BAC present and BAC absent).



Annual FARS Death by BAC Level

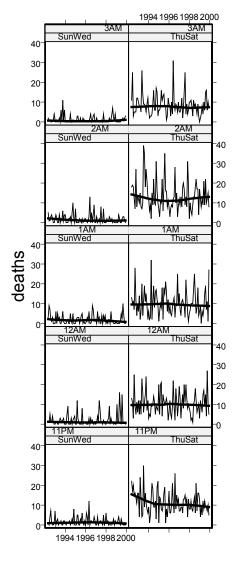
Trellis Time Series Plots by wkgrp and hour

There are many more deaths in the ThuSat weekgroup. There may be some temporal relationships between the series which could be due to weather and bad driving conditions. But on further investigation there was found to be no significant cross-correlation between the series and Figure 6a-iii α and Figure 6a-iii β shows comparisons when the aspect-ratio is reduced to improve visualization on the horizontal scale and a log transformation is made to reduce the effect of outliers on the resolution.

Figure 6a-i. FARS deaths. Common Scale.

Comparison of panels for 3AM-ThuSat and 2AM-ThuSat in Figure 6a-i suggest that some spikes occur at common times. Such an temporal association could be caused by poor driving conditions on particular nights. However closer investigation shows that there is really no apparent association, see Figure 6a-iii α and Figure 6a-iii β .

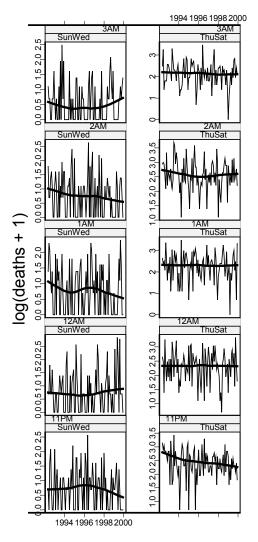
FARS, Monthly deaths



Year

Figure 6a-ii. FARS deaths. Logged and Scaled Independently

FARS deaths exhibit a downward trend for 11PM both SunWed and ThuSat, an upward trend for SunWed 3AM and an upward trend for ThuSat-2AM.

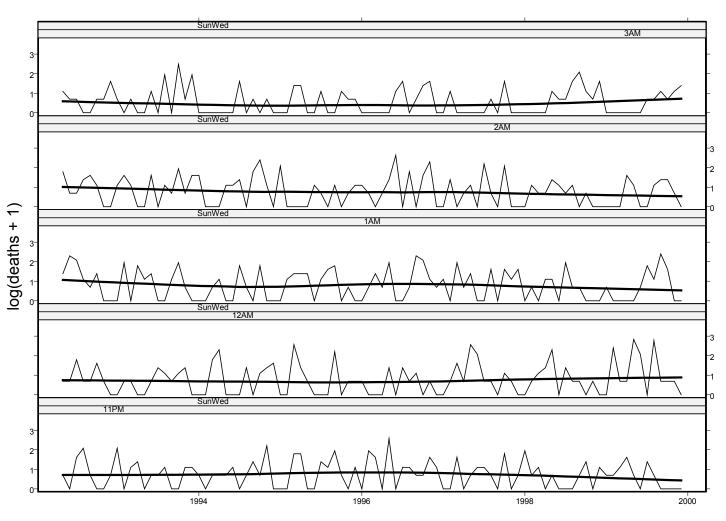


FARS log(deaths+1). Panels scaled independently.

Year

FARS: Time Series

Figure 6a-iiia FARS deaths, SunWed, small aspect-ratio and log transformation

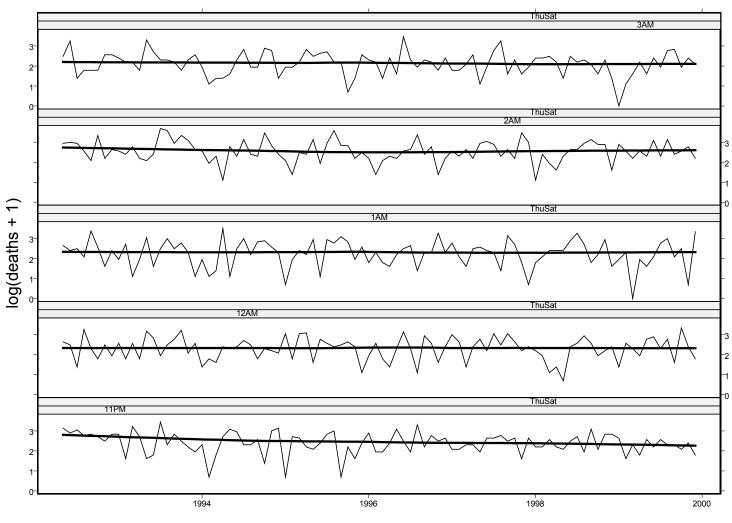


FARS log(deaths+1)

Year

FARS: Time Series

Figure 6a-iii β FARS deaths, ThuSat, small aspect-ratio and log transformation

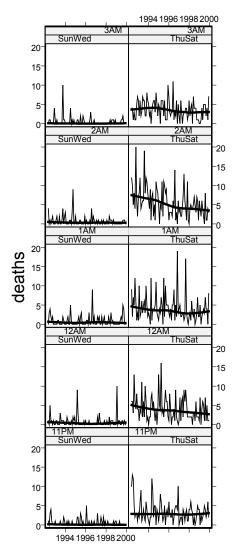


FARS log(deaths+1)

Year

Figure 6b-i. Deaths with BAC present. Common Scaling

FARS BAC Present, Monthly deaths



Year

Figure 6b-ii. Deaths with BAC present. Logged and Scaled Independently

Figure 6b-ii, FARS deaths with BAC present, exhibits a downward trend for ThuSat at 2AM and 12AM and possibly also SunWed at 2AM.

FARS log(deaths+1). Panels scaled independently.

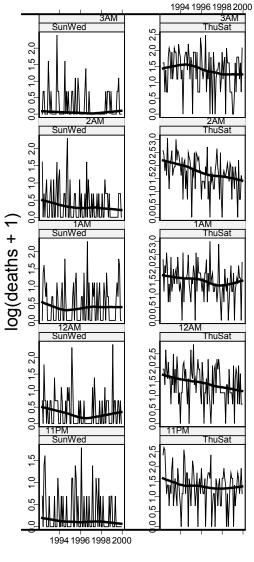
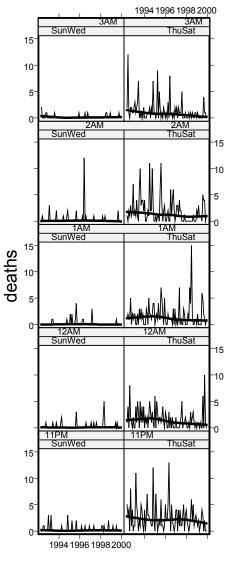




Figure 6c-i. Deaths with BAC absent.

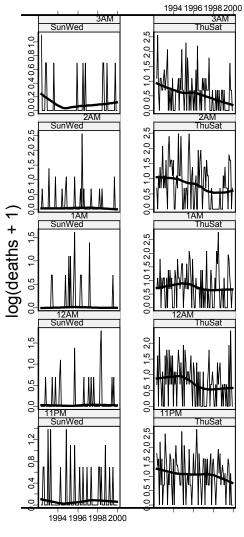
FARS BAC Absent, Monthly deaths



Year

Figure 6c-ii. Deaths with BAC absent. Logged and Scaled Independently.

Figure 6c-ii, FARS deaths with BAC absent shows a downward trend for all levels of hour for ThuSat but little change for SunWed.



FARS log(deaths+1). Panels scaled independently.

Year

STL Analysis

The changes in seasonal patterns may be of interest. An additive robust decomposition of the time series into trend, seasonal and remainder was done for each of the monthly time series. A robust locally quadratic 60% smoother was used for trend estimation.

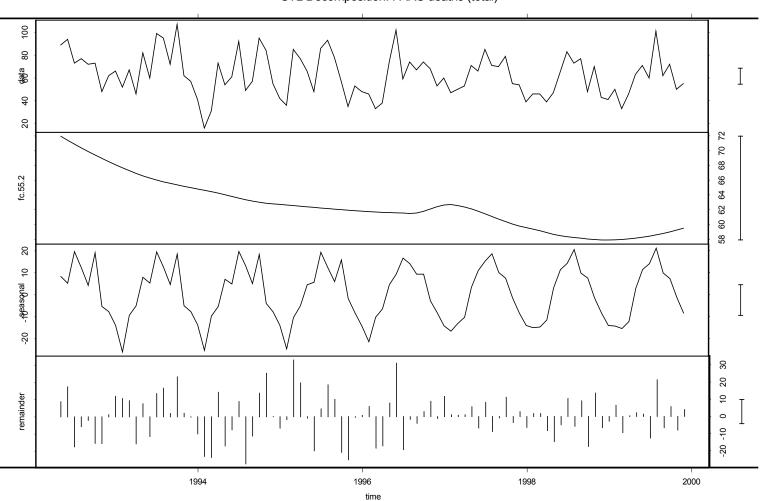
For all series, a change in the seasonal pattern occurs in 1996.

Notice for FARS Deaths BAC Present the seasonal component has shifted so there have been large increases in February and September while July and August have consistently remained high. This information may be useful for planning traffic safety campaigns.

Figure 7a: STL - All Deaths

 $R^2 = 53.4\%$

There was a slight hump in 1997 followed by decrease. The seasonality and variability of the data changed in about August 1996.

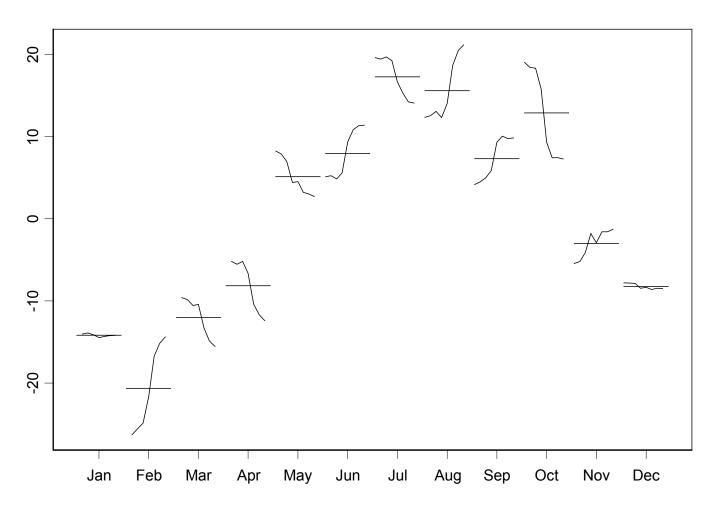


STL Decomposition: FARS deaths (total)

ss.window = 7 , ss.robust = TRUE , fc.window = 55.2 , fc.degree = 2

Figure 7b: Monthplot, FARS, totals

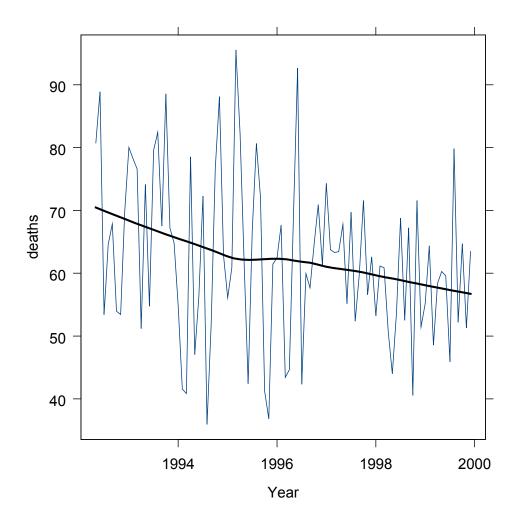
FARS deaths totals peak in Jul and reach minimum in Feb. These deaths are relatively high from May through to October (Figure 8).



FARS deaths (total), Monthplot of Seasonal

Figure 7c. The deseasonalized series showns an apparent change in variability around 1997. The trend is highly significant with Mann-Kendall test, tau = -0.214, sl = 1.261%. The loess trend declines from about 70.5 deaths/month to 56.7 deaths/month which corresponds to a percentage annual rate of decline of about 2.8%.

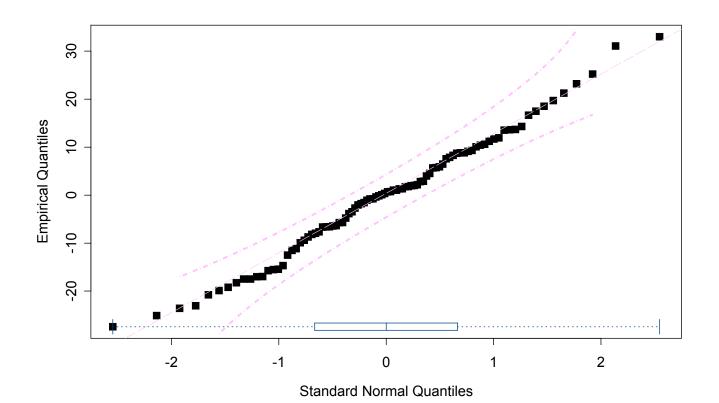
```
> pc.change(deaths.fars.deseasonalized.ts)
    start end change% annual.rate%
70.46393 56.71454 19.51266 2.791644
```



FARS, Deseasonalilzed monthly deaths

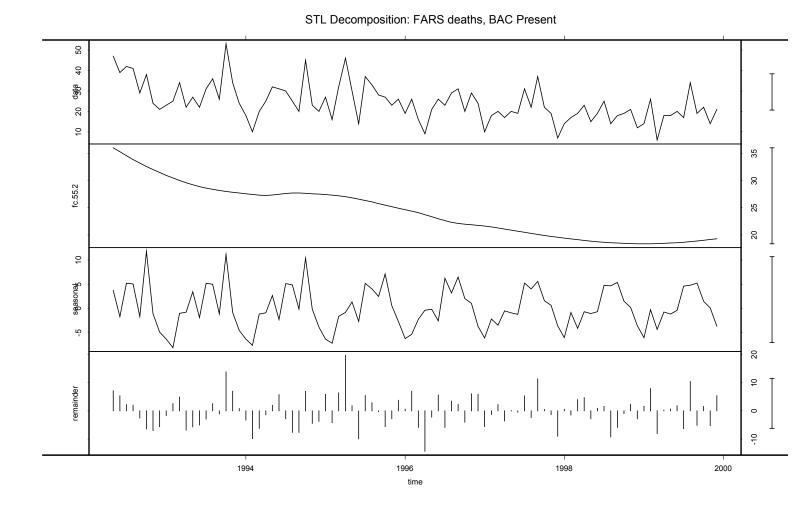
Figure 7d.

Remainder term is normally distributed.



Skewness Coefficient: $g_1 = 0.113532$, s.l. = 0.3177048 Michael's Statistic: $D_sp = 0.0422705$, s.l. = 0.609863 Wilk-Shapiro Statistic: W = 0.9802573, s.l. = 0.5645543

Figure 8a: STL, FARS Deaths, BAC Present R-sq = 60 %



ss.window = 7 , fc.window = 55.2 , fc.degree = 2

Figure 8b: Monthplot, FARS, BAC Present

FARS deaths with BAC Present (Fig 10) peak in October and are a minimum in Jan. These deaths are also relatively high in Jul and Aug. There are strong increasing trends in Feb and Sep and a decreasing trend in Oct and to a lesser extend in Mar and Apr too.

FARS deaths BAC Present, Monthplot of Seasonal

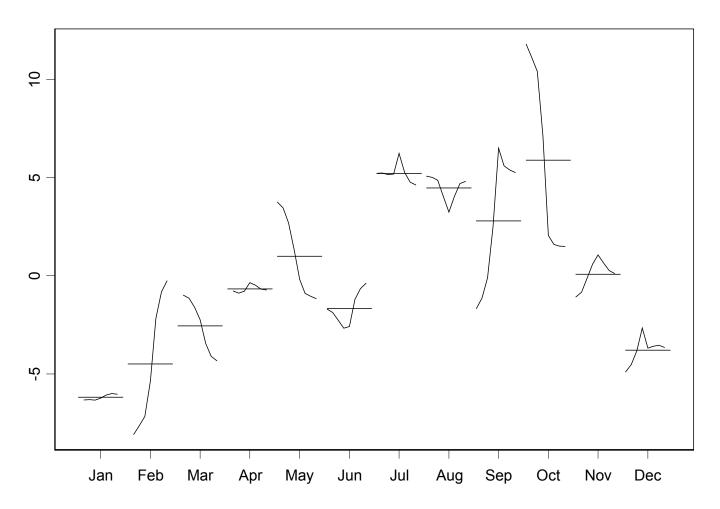
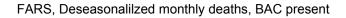
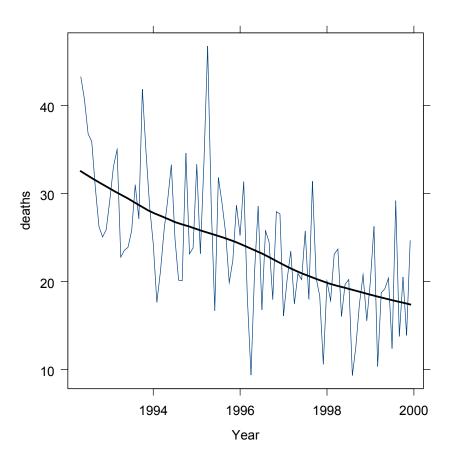


Figure 8c.
Trend in the deseasonalized component is obviously significant.
Mann-Kendall trend test: tau = -0.461, sl =8.001e-6%
> pc.change(deaths.fars.deseasonalized.ts)
 start end change% annual.rate%
32.52797 17.38038 46.56791 7.849873



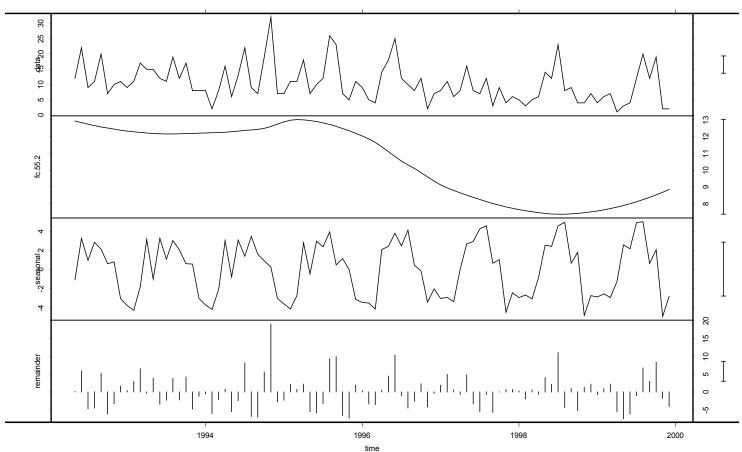


FARS: Time Series

Figure 9a: STL FARS Deaths: BAC Absent R²= 38.3%

Mann-Kendall trend test: tau = -0.338, sl = 0.008465%

STL Decomposition: FARS deaths, BAC Absent



ss.window = 7 , fc.window = 55.2 , fc.degree = 2

Figure 9b: Monthplot, FARS, BAC absent

FARS deaths with BAC absent peak in Aug but are relatively high from Apr through to Oct and are relatively low in Dec, Jan, Feb and Mar. There has been a sharp downward trend in deaths in Nov (Figure 12).



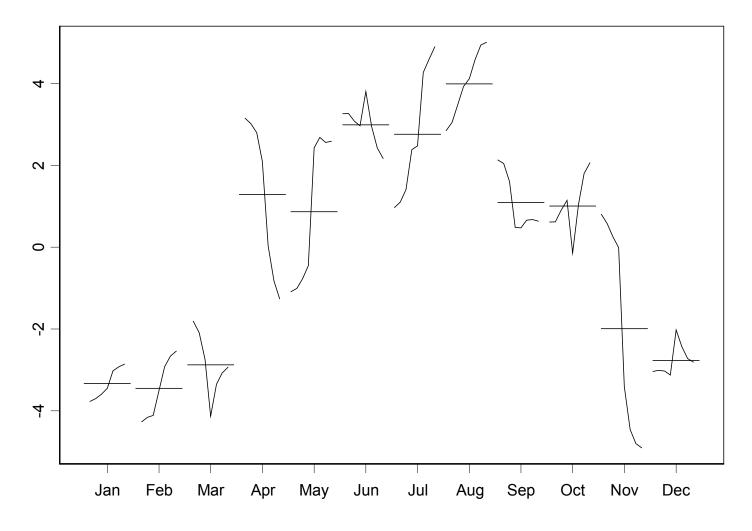
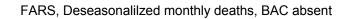
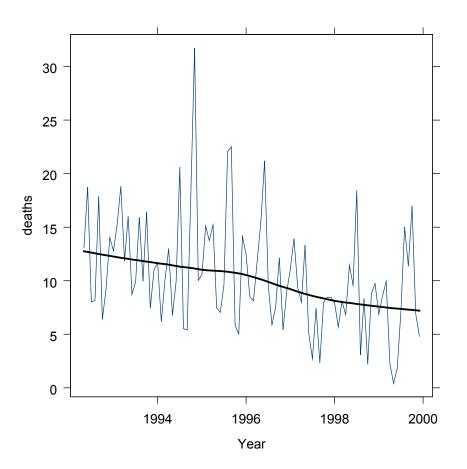


Figure 9c.

There is a significant downward trend. The Mann-Kendall test yields, tau = -0.338, sl = 0.008465%. The deseasonalized FARS BAC absent deaths decline from about 12.7 at the beginning of 1992 to about 7.2% by the end of 1999. The average annual percentage decrease is about 7.2%.

```
> pc.change(deaths.fars.deseasonalized.ts)
    start end change% annual.rate%
12.74778 7.199321 43.5249 7.181718
```

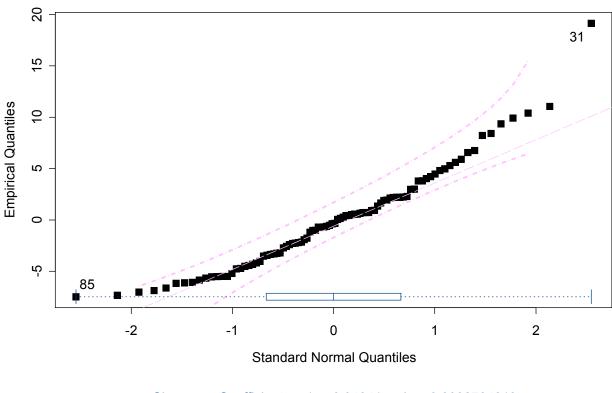




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Figure 9d

There are some outliers evident. Observations 31 and 85 corresond to Nov. 1994 and May 1999 respectively. <u>See Table 1c</u>.



Skewness Coefficient: $g_1 = 0.94341$, s.l. = 0.0002704813 Michael's Statistic: $D_sp = 0.111165$, s.l. = 0.0004617088 Wilk-Shapiro Statistic: W = 0.9450933, s.l. = 0.001567573

J. MISTPOL Dataset

Introduction

Data on 164,931 car accidents from three counties in Michigan for the period January 1992 to December 1999 are available, broken down by county below,

County	Frequency	Percent	
Macomb	23725	14.4	
Oakland	31936	19.4	
Wayne	109270	66.3	
Total	164931	100.0	

The day, month and year were available for all but 45 of these accidents. In addition the following variables of possible interest were observed for many of these accidents.

Factor	Levels	% missing
hour	11PM, 12AM, 1AM, 2AM	0
wkgrp	SunWed, ThuSat	0.02
sex	male, female	28.1
drink	yes, no	47.6
injury	fatal, incapacitating, non-incapacitating, none	29.6

The data after May 1999 appear to be incomplete so our analysis was based on the data from May 1992 to April 1999. After removing records from May 1999 onwards there remained 155,207 records.

Three separate datasets were analyzed:

- (a) All Accidents (155207 records)
- (b) Fatal Accidents (465 records)
- (c) Incapacitating Accidents (4516 records)

J. MISTPOL Dataset All Accidents

Summary

Figure 1. Montly number of accidents, in thousands. Seasonal Mann-Kendall test not significant. The trend line suggests a 4% drop from 1.78 to 1.70 thousands of accidents per month from May 1992 to April 1999. It is clear that the total number accidents increased and has now begun to decrease. There are some outliers in 1992 in May, Jul, Nov and Dec.

There is a mild upward trend to about 1997 and then a steeper downward trend (Figure 2, 3). Seasonal maximums occur in Jun-Jul and also in Dec-Jan (Fig. 4).

Drink-Hour are associated: there are fewer deaths in 11PM-YES and more in 2AM-YES than expected under independence.

<u>Drink-Weekgroup</u> are associated: there are slightly more (5%) in the ThuSat-YES group than expected under independence.

Drink and Injury are associated: there is great increase in the frequency of drinking with fatal and more serious accidents.

Drink-Gender are not independent since drinking males are more frequently in accidents than drinking females.

<u>Hour-Injury</u> are associated: more serious accidents (fatalities, incapacitating and noncapacitating) tend to happen later in the 2AM window.

Weekgroup-Injury are independent.

<u>Gender-Injury</u> are associated since males tend to be more involved in fatal accidents.

<u>Figure 4</u> shows that ThuSat group has more accidents. There is an evident downward trend starting around 1996 in SunWed and shortly after in ThuSat.

No significant trends are present in the trellis plots for hour (Figure 6).

Figure 7 shows there has been a downward trend in incapacitating and non-incapacitating accidents but little other trends in the trellis plots for injury.

Figure 8 shows accidents by males more frequent and there is a downward trend in both groups but more pronounced for males.

The trellis plot for drink shows that both the yes and no groups show a trend downward but the yes group has a more pronounced trend (Figure 9). This suggests that perhaps there has been a downward trend in the number of accidents where alcohol testing has been done as well as perhaps some decrease in drinking-and-driving.

An STL analysis of accidents with drinking indicates a downward trend which increases after 1996 (Figure 10). The seasonal variation in accidents with drinking involved is complicated (Figure 10, 11): there are two peaks one in Jul and the other in Dec. The July peak shows little variation in its amplitude over the years. The troughs are in Feb and Sep.

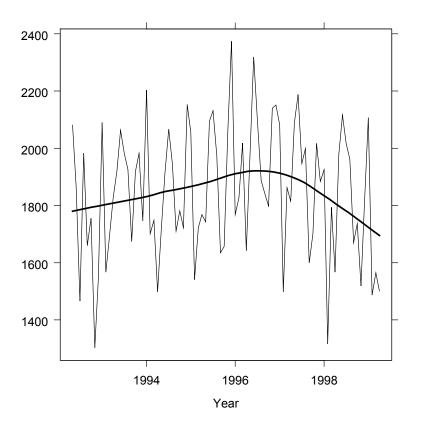
Time Series Plots

Montly number of accidents, in thousands. Seasonal Mann-Kendall test not significant. The trend line suggests a 4% drop from 1.78 to 1.70 thousands of accidents per month from May 1992 to April 1999. It is clear that the total number accidents increased and has now begun to decrease. There are some outliers in 1992 in May, Jul, Nov and Dec.

```
> accidents.mistpol.total.ts
        Jan Feb Mar Apr
                                May
                                      Jun
                                            Jul
                                                  Aug
                                                        Sep
                                                              Oct
                                                                    Nov
                                                                          Dec
1992:
                             2.081 1.865 1.466 1.983 1.660 1.755 1.302 1.549
1993: 2.090 1.567 1.693 1.827 1.922 2.063 1.981 1.926 1.674 1.919 1.982 1.746
1994: 2.203 1.701 1.748 1.499 1.716 1.913 2.067 1.945 1.711 1.781 1.722 2.153
1995: 2.056 1.541 1.722 1.768 1.743 2.095 2.131 1.974 1.635 1.657 2.061 2.374
1996: 1.768 1.827 2.018 1.642 1.984 2.319 2.095 1.886 1.842 1.798 2.140 2.151
1997: 2.082 1.498 1.862 1.815 2.089 2.188 1.947 1.999 1.599 1.703 2.017 1.884
1998: 1.925 1.316 1.794 1.567 1.978 2.116 2.017 1.961 1.670 1.737 1.519 1.827
1999: 2.107 1.487 1.565 1.501
> print.Kendall(SeasonalMannKendall(accidents.mistpol.total.ts))
tau = 0.0159, sl =86.23%
> pc.change(accidents.mistpol.total.ts)
[1] 1.779563 1.695388 4.730113
```

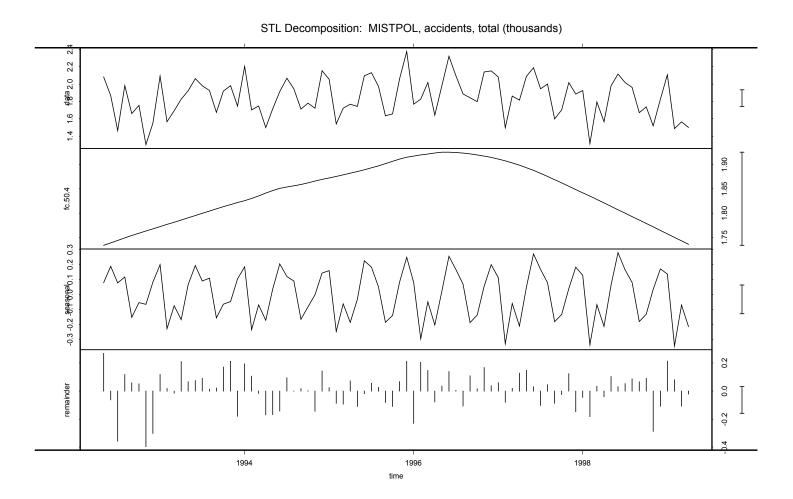
Figure 1.

All Accidents, MISTPOL



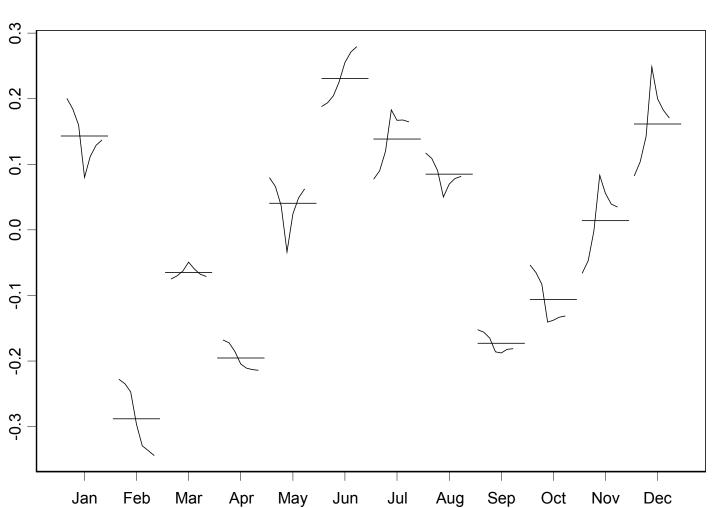
STL Analysis

There is a upward trend to 1996 followed by a downward trend. R-sq = 65.9%Figure 2a.



ss.window = 7 , fc.window = 50.4 , fc.degree = 1

Figure 2b. Seasonal maxima in Jun-Jul and also in Dec-Jan (Fig. 4).



Monthplot, MISTPOL, accidents, total

Figure 2c. Trend Analysis of the Deseasonalized Series

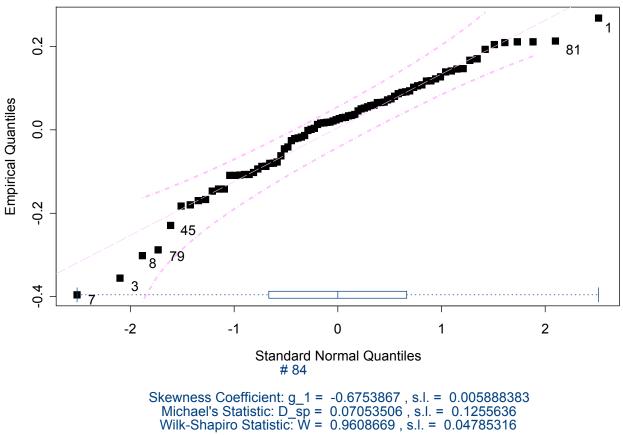
Mann-Kendall: tau = 0.0321, sl =66.9% However there is clearly a downward trend evident in the last 2 years. Outliers at the beginning are also evident.

2.0 - 0 2.0 - 0 1.8 - 0 1.6 - 0 1.4 - 0 1.994 1994 1996 1998Year

MISTPOL, all accidents, deseasonalilzed

Figure 2d. Normal Probability Plot of Remainder

There are some outliers in 1992 in May, Jul, Nov and Dec.												
> counter.ts												
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					1	2	3	4	5	6	7	8
1993:	9	10	11	12	13	14	15	16	17	18	19	20
1994:	21	22	23	24	25	26	27	28	29	30	31	32
1995:	33	34	35	36	37	38	39	40	41	42	43	44
1996:	45	46	47	48	49	50	51	52	53	54	55	56
1997:	57	58	59	60	61	62	63	64	65	66	67	68
1998:	69	70	71	72	73	74	75	76	77	78	79	80
1999:	81	82	83	84								



Drink and Hour

There are fewer accidents in 11PM-YES and more in 2AM-YES than expected under independence.

	bs(~ dr ases in		ur, data	= mistp	ol.df, na	a.action = na.exclude)
+ N N/RowT N/ColT N/Tota +	otal otal 1					
	11PM	12AM +			RowTotl	
yes	4696 0.228 0.154	4546 0.221 0.217 0.053	4786 0.233 0.282	6536 0.318 0.365	20564	•
	0.393 0.846 0.299	16386 0.249 0.783 0.190	12208 0.185 0.718 0.141	0.173 0.635 0.132		•
	30536 0.35	+ 20932 0.24 +	16994 0.20	17914	Ì	+
contri [1,] -	r indepe Chi^2 = Yates' c .bution [,1] -29.49	ndence o 3011.41 correctic to chi [,2] -6.08 3.41	f all fa d.f.= 3 on not us i-sq: [,3] 11.55	(p=0) sed [,4] 33.73	+	Ŧ

Drink and Weekgroup

There are more in the ThuSat-YES group than expected under independence. So we see there tends to be a higher proportion of drunk driving accidents in the ThuSat wkgrp.

Call: crosstabs(~ drink + wkgrp, data = mistpol.df, na.action = na.exclude) 86340 cases in table +----+ | N |N/RowTotal| |N/ColTotal| |N/Total | +----+ drink |wkgrp |SunWed |ThuSat |RowTotl| yes | 7007 |13551 |20558 | |0.341 |0.659 |0.24 | |0.218 |0.250 | |0.081 |0.157 | ----+----+-----+-------+ no |25111 |40671 |65782 | |0.382 |0.618 |0.76 | 0.782 |0.750 | | |0.291 |0.471 | | ColTotl|32118 |54222 |86340 | |0.37 |0.63 | | Test for independence of all factors Chi^2 = 112.1013 d.f.= 1 (p=0) Yates' correction not used contribution to chi-sq: [,1] [,2] [1,] -7.20 5.51 [2,] 4.04 -3.09

Drink and Injury

An great increase in the frequency of drinking with fatal and more serious accidents.

```
Call:
crosstabs( ~ drink + injury, data = mistpol.df, na.action = na.exclude)
84421 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
drink |injury
|fatal |incpctt|noncap |possibl|none |RowTotl|
           yes | 273 | 1672 | 2942 | 2372 |12902 |20161 |
    |0.0135 |0.0829 |0.1459 |0.1177 |0.6399 |0.24
     0.7479 0.4963 0.4867 0.2114 0.2034 |
     0.0032 0.0198 0.0348 0.0281 0.1528
                                         1
no | 92 | 1697 | 3103 | 8847 |50521 |64260 |
|0.0014 |0.0264 |0.0483 |0.1377 |0.7862 |0.76 |
     0.2521 0.5037 0.5133 0.7886 0.7966
     |0.0011 |0.0201 |0.0368 |0.1048 |0.5984 |
ColTotl|365 |3369 |6045 |11219 |63423 |84421 |
    |0.0043 |0.0399 |0.0716 |0.1329 |0.7513 |
-----+
Test for independence of all factors
     Chi^2 = 4275.379 d.f.= 4 (p=0)
     Yates' correction not used
contribution to chi-sq:
     [,1] [,2] [,3] [,4] [,5]
[1,] 19.24 29.98 38.58 -6.04 -17.74
[2,] -10.82 -16.85 -21.69 3.40 9.97
```

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Drink and Gender

Drinking and gender are not independent. Drinking males are more frequently in accidents than drinking females.

```
Call:
crosstabs( ~ drink + sex, data = mistpol.df, na.action = na.exclude)
83792 cases in table
+----+
| N
|N/RowTotal|
|N/ColTotal|
|N/Total |
+----+
drink |sex
      |male |female |RowTotl|
yes |16033 | 4129 |20162 |
|0.795 |0.205 |0.24 |
|0.276 |0.161 |
     |0.191 |0.049 |
                           -----+
no |42157 |21473 |63630 |
    |0.663 |0.337 |0.76 |
|0.724 |0.839 |
|0.503 |0.256 |
                            1
                            ----+
ColTotl|58190 |25602 |83792 |
|0.69 |0.31 | |
-----+
Test for independence of all factors
      Chi^2 = 1270.155 d.f.= 1 (p=0)
Yates' correction not used
contribution to chi-sq:
      [,1] [,2]
[1,] 16.66 -25.16
[2,] -9.41 14.21
```

Hour and Injury

The 2AM window has a higher than expected share of the three most serious injury classes and correspondingly the 11PM windows has fewer of the three most serious injury classes and more of the none-injury accidents.

	cases in + Cotal Cotal N		ury, dat	a = mist	pol.dI,	na.actior	n = na.exclud	1e)
hour	injury fatal	· •	· -	· •		RowTotl		
11PM	129 0.0032 0.2575 0.0011	1285 0.0321 0.2689 0.0111	2122 0.0529 0.2782 0.0183	5010 0.1250 0.3385 0.0432	31546 0.7868 0.3569 0.2717	0.35 	 	
12AM	117 0.0042 0.2335 0.0010	1170	1747 0.0632 0.2290 0.0150	3512 0.1271 0.2373 0.0303	21091 0.7631 0.2386 0.1817		l	
1AM	120 0.0051 0.2395 0.0010	983 0.0417 0.2057 0.0085	1702 0.0722 0.2231 0.0147	3070 0.1303 0.2074 0.0264	17686 0.7506 0.2001 0.1523	23561 0.20 	 	
2AM	135 0.0054 0.2695 0.0012	1340 0.0540 0.2805 0.0115	2057 0.0829 0.2697 0.0177	3208 0.1293 0.2168 0.0276	18069 0.7283 0.2044 0.1556	Ì	 	
ColTotl	501 0.0043	4778 0.0412	7628 0.0657	14800 0.1275	88392 0.7614			
Test fo contri [1,] - [2,] - [3,]	or indepe Chi ² = Yates' o bution 5 [,1] [, 3.36 -8 0.24 0 1.82 0	+	f all fad d.f.= 1 in not us iq: 3] [,4 85 -1.2 33 -0.0 86 1.2	ctors .2 (p=0) sed [5 5.64 9 0.25 26 -1.88		+	÷	

Weekgroup and Injury Injury and weekgroup are not associated.

Call: crosstabs(~ wkgrp + injury, data = mistpol.df, na.action = na.exclude) 116063 cases in table +----+ | N |N/RowTotal| |N/ColTotal| |N/Total | +----+ wkgrp |injury |fatal |incpctt|noncap |possibl|none |RowTotl| -----SunWed | 188 | 1764 | 2863 | 5642 |33072 |43529 | |0.0043 |0.0405 |0.0658 |0.1296 |0.7598 |0.38 0.3752 0.3693 0.3756 0.3813 0.3743 | |0.0016 |0.0152 |0.0247 |0.0486 |0.2849 | 1 ThuSat | 313 | 3012 | 4760 | 9156 |55293 |72534 | |0.0043 |0.0415 |0.0656 |0.1262 |0.7623 |0.62 | 0.6248 0.6307 0.6244 0.6187 0.6257 |0.0027 |0.0260 |0.0410 |0.0789 |0.4764 | ColTotl|501 |4776 |7623 |14798 |88365 |116063 | |0.0043 |0.0412 |0.0657 |0.1275 |0.7614 | -----+ Test for independence of all factors Chi^2 = 3.344447 d.f.= 4 (p=0.501921) Yates' correction not used contribution to chi-sq: [,1] [,2] [,3] [,4] [,5] [1,] -0.02 -0.94 -0.28 1.31 -0.24 [2,] 0.02 0.72 0.21 -1.01 0.18

October 19, 2001

Gender and Injury

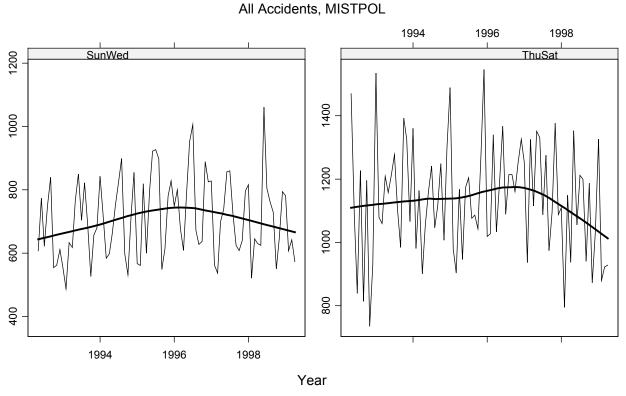
Males tend to be more involved in fatal accidents. The biggest contribution to the gender difference though comes from the possibly-injury accidents where males account for only 60.2% whereas for all accidents they account for 69%. So there contribution to minor accidents is much less than females.

```
Call:
crosstabs( ~ sex + injury, data = mistpol.df, na.action = na.exclude)
109948 cases in table
+----+
IN
|N/RowTotal|
|N/ColTotal|
|N/Total
       +----+
sex |injury
    |fatal |incpctt|noncap |possibl|none |RowTotl|
male | 386 | 3303 | 5342 | 8706 |58346 |76083 |
   0.005070.043410.070210.114430.766870.69
    |0.79261|0.70865|0.71772|0.60182|0.70389|
    0.003510.030040.048590.079180.53067
female | 101 | 1358 | 2101 | 5760 |24545 |33865
    0.002980.040100.062040.170090.724790.31
    0.207390.291350.282280.398180.29611
    0.000920.012350.019110.052390.22324
---+-
ColTotl|487
          |4661 |7443 |14466 |82891 |109948 |
   |0.0044 |0.0424 |0.0677 |0.1316 |0.7539 |
Test for independence of all factors
     Chi^2 = 659.1499 d.f.= 4 (p=0)
     Yates' correction not used
contribution to chi-sq:
    [,1] [,2] [,3] [,4] [,5]
[1,] 2.47 1.10 2.66 -12.80 4.13
[2,] -3.72 -1.65 -3.99 19.23 -6.20
```

Time Series Trellis Plots of Weekgroup

Figure 4 shows that ThuSat group has more accidents. There is an evident downward trend starting around 1996 in SunWed and shortly after in ThuSat.

Figure 4. Sliced vertical scale.



Sliced Scaling

Wkgrp time series and seasonal Mann-Kendall trends

> mistpol.total.SunWed.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 774 622 750 1992: 607 840 554 562 611 1993: 555 488 633 618 765 850 704 822 690 526 657 680 1994: 843 720 584 598 663 754 826 899 597 532 715 855 1995: 567 561 819 600 797 922 927 898 549 614 780 828 749 800 678 609 793 952 1006 672 628 637 889 825 1996: 1997: 828 562 537 700 738 857 860 724 625 608 640 796 1998: 816 521 645 630 625 1061 805 762 730 550 647 794 1999: 781 607 642 573 > SeasonalMannKendall(mistpol.total.SunWed.ts) tau = 0.143, sl =11.86%

> mistpol.total.ThuSat.ts

Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1470 1085 839 1227 814 1196 735 930 1993: 1535 1079 1060 1209 1157 1213 1277 1104 984 1393 1325 1066 1994: 1360 981 1164 901 1053 1159 1241 1046 1114 1249 1007 1298 1995: 1489 980 903 1168 946 1173 1204 1076 1086 1043 1281 1546 1996: 1019 1027 1340 1033 1191 1367 1089 1214 1214 1161 1251 1326 1997: 1254 936 1325 1115 1351 1331 1087 1275 974 1095 1377 1088 1998: 1109 795 1149 937 1353 1055 1212 1199 940 1187 872 1033 1999: 1326 880 923 928 > SeasonalMannKendall(mistpol.total.ThuSat.ts) tau = -0.111, sl =22.48%

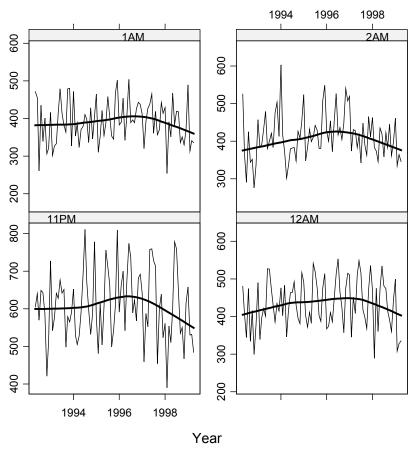
Time Series Trellis Plots of Hour

Here is a breakdown by hour from the original SPSS dataset. The 1474 accidents recorded at Midnight were put in the "12AM" group.

	Frequency	Percent
12 mid1:00 a.m.	37279	22.6
1:00-2:00 a.m.	34966	21.2
2:00-3:00 a.m.	36834	22.3
11:00 p.m 12 mid	54378	33.0
Midnight	1474	.9

No significant trends are present in the trellis plots for hour.

Figure 6.



All Accidents, MISTPOL

Sliced Scaling

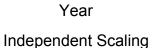
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Time Series Trellis Plots of Injury Class

From the crosstab analysis below, there were 109,948 records that contained information of injury. Since there were 164,931 records in total, this means that the injury was reported in only about 2/3 of the accidents. Figure 7 shows there has been a downward trend in incapacitating and non-incapacitating accidents but little other trends. Note that each panel was scaled independently.

Figure 7. Independent vertical scale.

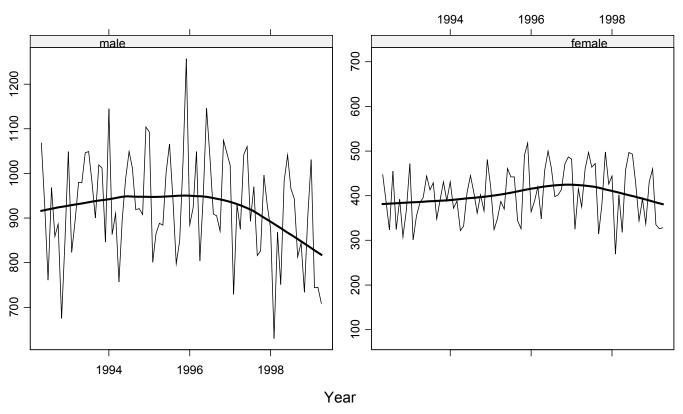
All Accidents, MISTPOL, injury



Time Series Trellis Plots of Gender

Figure 8 shows accidents by males more frequent and there is a downward trend in both groups but more pronounced for males.

Figure 8. Sliced vertical scale.



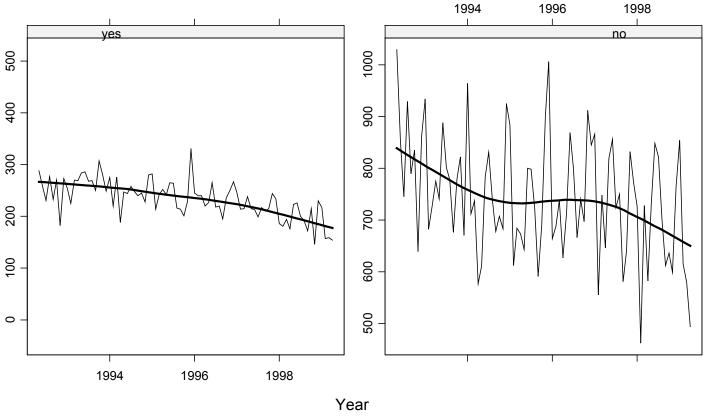
Sliced Scaling

All Accidents, MISTPOL

Time Series Trellis Plots of Drink

In 52.4% of the accidents, drinking information was obtained. Our analysis indicates a significant downward trend in the number of accidents with drinking and without drinking. The downward trend is more pronounced among the drinking group.

Figure 9. Sliced vertical scale.



All Accidents, MISTPOL, drink

Sliced Scaling

Tables and Seasonal Mann-Kendall test for total accidents by drinking group.

> accidents.mistpol.sober.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1029 854 745 929 789 835 861 639 1993: 934 682 727 775 741 779 888 801 778 676 822 670 1994: 964 712 737 577 611 783 831 740 678 707 683 925 800 798 722 1995: 884 612 684 673 643 591 691 900 1006 664 688 738 627 714 869 798 666 740 697 912 845 1996: 1997: 866 555 748 646 819 856 724 750 581 641 832 771 1998: 727 462 728 582 731 848 822 703 612 636 599 764 1999: 854 615 580 494 564 > SeasonalMannKendall(accidents.mistpol.sober.ts) tau = -0.325, sl = 0.03499%

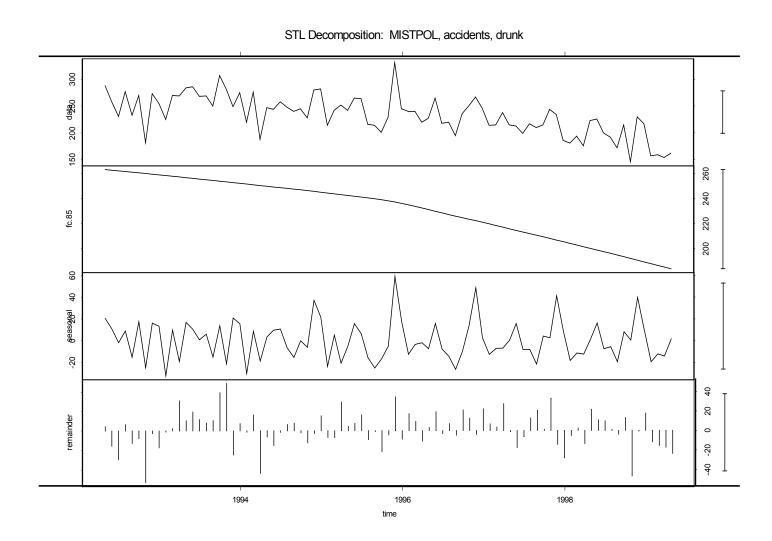
```
> accidents.mistpol.drunk.ts
```

Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 288 258 231 277 233 270 182 273 1993: 255 225 270 269 284 286 268 269 250 307 281 249 1994: 275 220 276 188 247 244 258 248 240 245 228 280 1995: 282 214 242 252 242 265 264 216 214 201 229 331 1996: 245 240 240 220 227 265 218 220 195 236 250 267 1997: 246 214 215 238 215 213 199 217 210 215 244 234 1998: 186 181 194 176 223 226 200 192 172 215 146 230 1999: 217 157 159 154 162 > SeasonalMannKendall(accidents.mistpol.drunk.ts) tau = -0.606, sl =2.894e-9%

STL Analysis of Accidents with Drinking

There is a downward trend over 1992 to 1999 which is more pronounced after 1996. The seasonal cycle is complicated and shows a lot of variability. R-sq = 74.9 % so pretty good fit.

Figure 10a



ss.window = 5 , fc.window = 85 , fc.degree = 1

The seasonal variation is not very stable. There are two peaks one in Jul and the other in Dec. The July peak shows little variation in its amplitude over the years. The troughs are in Feb and Sep.

Figure 10b.

Monthplot, MISTPOL, drunk

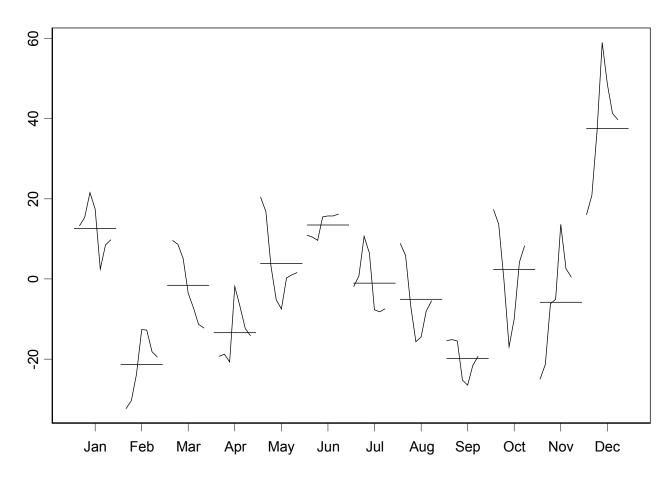


Figure 10c. Loess trend analysis of deseasonalized series.

tau = -0.58, sl = 2.22e-14%

MISTPOL, all accidents, deseasonalilzed, drink=yes

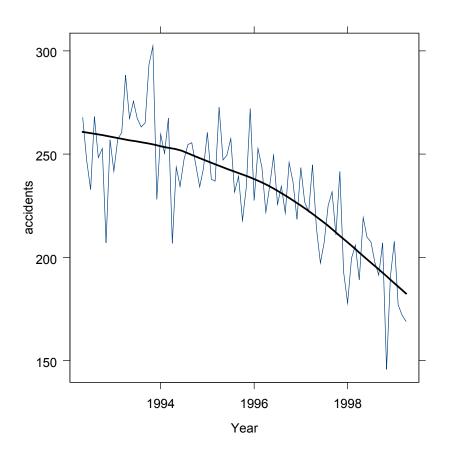
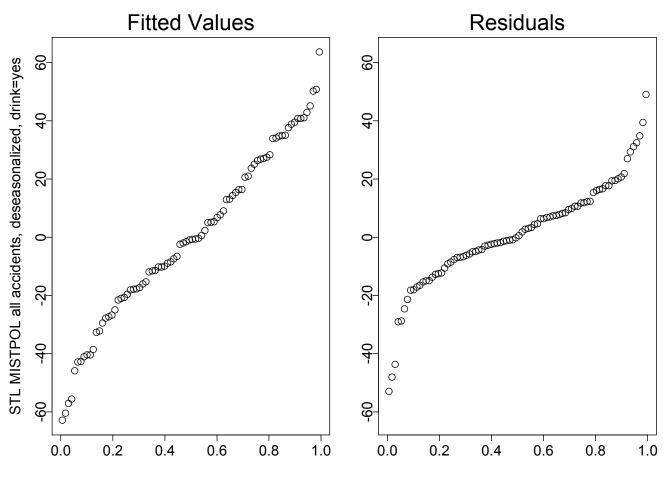


Figure 10d. RF-spread,

R-sq = 74.9 %



f-value

J. MISTPOL Dataset *Fatal Accidents* Summary

No much sign of any trend in the number of fatal accidents per month (Fig 1).

STL analysis confirms no trend and unstable seasonal component is seen (Figure 2a and Figure 2b). There is a peak in Jul but Mar though May is also quite high and there is a minimum in Jan-Feb (Figure 2b). Some outliers were detected Figure 2d and Table 1. Low value, R-sq = 37.4 %, means a lot of unexplained variation.

Drink-Hour are not associated.

<u>Drink-Weekgroup</u>: there are slightly more (17%) in the ThuSat-YES group than expected under independence but this excess is not statistically significant.

Drink-Gender: males are strongly associated with more fatal accidents involving drinking.

Figure 4 shows that ThuSat group has more accidents and there is a small downward trend starting around 1996 in ThuSat.

No monotonic trend but loess shows a slight increasing trend at 1AM and 11PM windows in recent years (Figure 5).

Males have more fatal accidents. No trend in fatalities with males or females (Figure 6).

There has been a striking increase in fatalities involving drinking since 1996 (Figure 7a and Figure 7b). Also confirmed by Seasonal Mann-Kendall test - <u>see Table</u>.

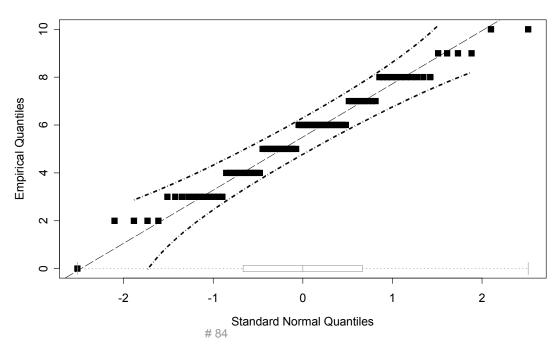
STL analysis confirms that fatal accidents with drinking have risen since 1994 (Figure 8a). Also see loess trend analysis and Mann-Kendall test for deseasonalized fatal accident with drinking – Figure 8e. The R-sq = 50.8 % is better than in Figure 2a.

Time Series Plots

No much sign of any trend (Fig 1). There is no significant lag one autocorrelation but seasonality is evident in the lag 12 autocorrelation of 0.92. The data are approximately Poisson distributed with mean of 5.5.

The discreteness of the data is evident from the normal probability plot but otherwise the normal distribution is not too bad.

Figure 1.



Skewness Coefficient: g_1 = -0.03869677 , s.l. = 0.4381817 Michael's Statistic: D_sp = 0.07047588 , s.l. = 0.126429 Wilk-Shapiro Statistic: W = 0.9644368 , s.l. = 0.08429486

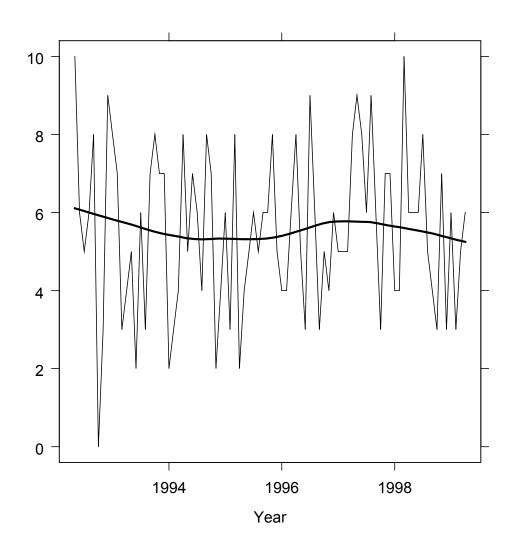
Table 1.Time series tabulation & outliers.

Outliers discovered in Figure 2d and 2f are shown.

> accidents.mistpol.death.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1993: 8 1994: **2** 1995: 6 3 8 **2** 4 5 6 5 6 6 8 5 1996: 4 4 6 8 5 3 9 6 3 5 4 1997: 5 5 5 8 1998: 4 4 **10** 6 1999: 6 > SeasonalMannKendall(accidents.mistpol.death.ts) tau = 0.0169, sl = 85.76% > counter.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1993: 1994: **21** 22 1995: 33 34 37 38 40 41 42 1996: 45 46 47 48 49 50 51 52 53 54 **:** 57 58 60 61 62 64 65 66 1998: 69 70 72 73 74 75 76 77 78 1999: 81 82 83

Figure 1. No monotonic trend.



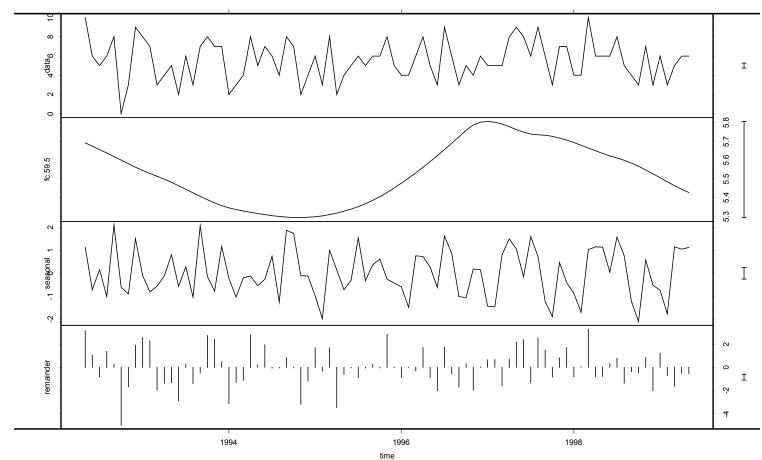


STL Analysis

STL analysis indicates no monotonic trend. Unstable seasonal component. R-sq = 37.4 %

Figure 2a.

STL Decomposition: MISTPOL, fatal accidents



ss.window = 5 , fc.window = 59.5 , fc.degree = 1

There is a lot of variability in the seasonal component. There is a peak in Jul but Mar though May is also quite high and there is a minimum in Jan-Feb (Figure 2b).

Figure 2b.

Monthplot, MISTPOL, fatalities

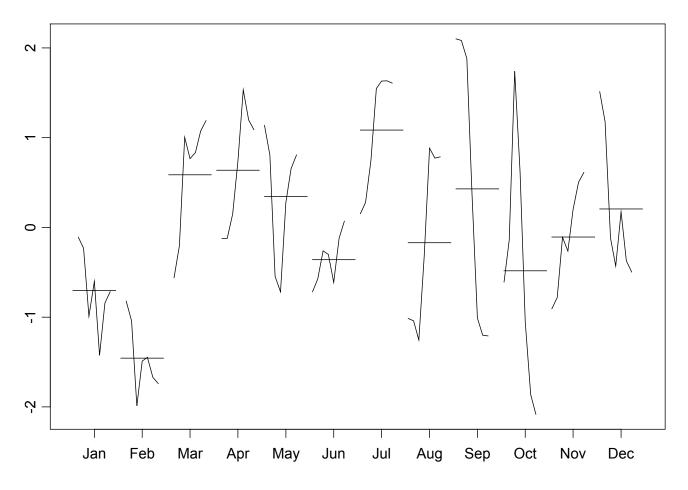
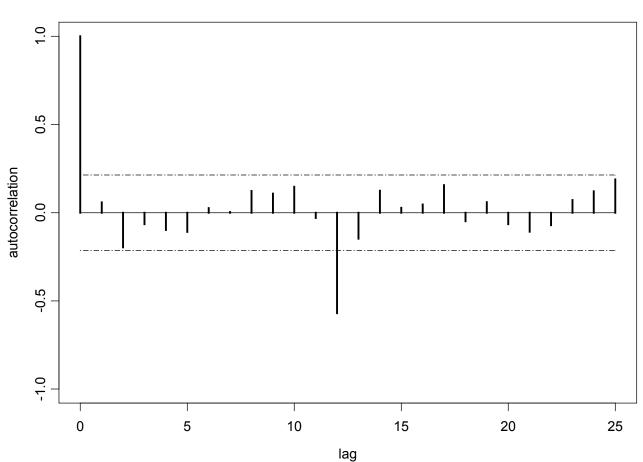


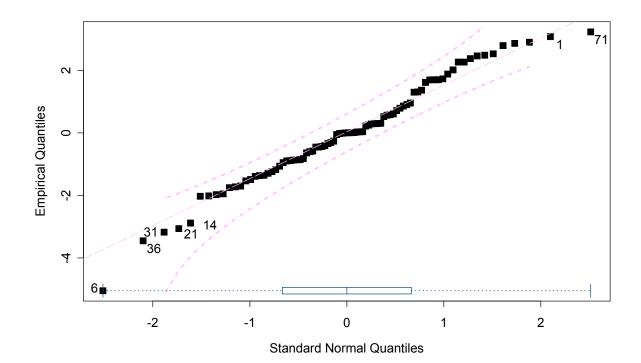
Figure 2c.

No unexpected autocorrelation in the remainder term.



STL Remainder, MISTPOL, fatalities

Figure 2d Normal probability plot of remainder indicates some outliers.



Skewness Coefficient: $g_1 = -0.1824367$, s.l. = 0.2329958 Michael's Statistic: D_sp = 0.05067573, s.l. = 0.8947265 Wilk-Shapiro Statistic: W = 0.9766002, s.l. = 0.4186626

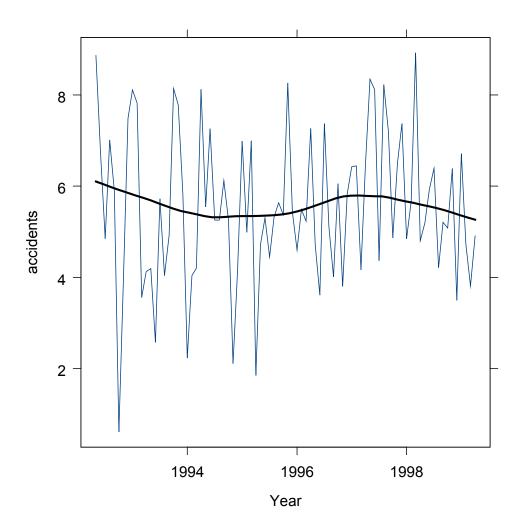
> cour	nter	.ts										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					1	2	3	4	5	6	7	8
1993:	9	10	11	12	13	14	15	16	17	18	19	20
1994:	21	22	23	24	25	26	27	28	29	30	31	32
1995:	33	34	35	36	37	38	39	40	41	42	43	44
1996:	45	46	47	48	49	50	51	52	53	54	55	56
1997:	57	58	59	60	61	62	63	64	65	66	67	68
1998:	69	70	71	72	73	74	75	76	77	78	79	80
1999:	81	82	83	84								

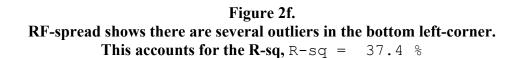
Figure 2e.

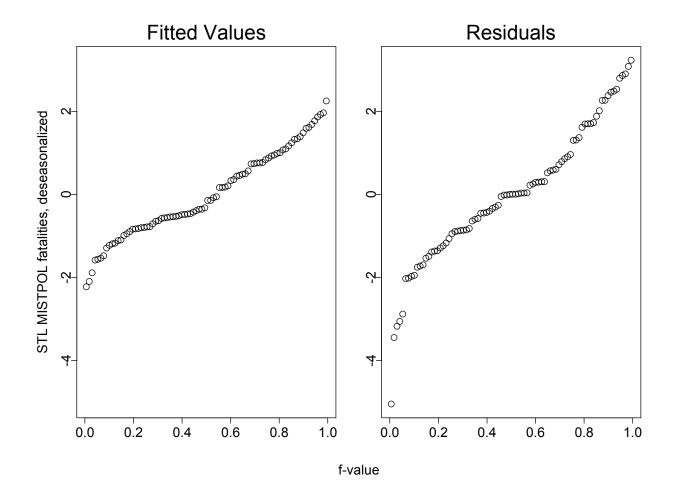
The loess trend analysis of the deseasonalized series, MISTPOL, fatalities. There is no montonic trend.

tau = 0.0212, sl =77.86%

MISTPOL, fatalities, deseasonalilzed







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Crosstab Analysis

Drink and Hour

Drink-Hour are not associated.

> crosstabs(~ drink + hour, data = mistpol.fatal.df, na.action = na.exclude) Call: crosstabs(formula = ~ drink + hour, data = mistpol.fatal.df, na.action = na.exclude) 333 cases in table +----+ | N |N/RowTotal| |N/ColTotal| |N/Total | +----+ drink |hour |11PM |12AM |1AM |2AM |RowTotl| yes |58 |58 |60 |76 |252 | |0.23 |0.23 |0.24 |0.3 |0.76 | |0.69 |0.77 |0.75 |0.81 | |0.17 |0.17 |0.18 |0.23 | no |26 |17 |20 |18 |81 | |0.32 |0.21 |0.25 |0.22 |0.24 | |0.31 |0.23 |0.25 |0.19 | |0.078 |0.051 |0.06 |0.054 | ColTotl|84 |75 |80 |94 |333 |0.25 |0.23 |0.24 |0.28 | 1 -----+ Test for independence of all factors Chi^2 = 3.504304 d.f.= 3 (p=0.3202043) Yates' correction not used > get.crosstabs.percenterror(~ drink + hour) [,1] [,2] [,3] [,4] [1,] -9 2 -1 7 [2,] 27 -7 3 -21 > contrib(~ drink + hour) contribution to chi-sq: [,1] [,2] [,3] [,4] [1,] -0.70 0.17 -0.07 0.58 [2,] 1.23 -0.29 0.12 -1.02

Drink and Weekgroup

There are slightly more (17%) in the ThuSat-YES group than expected under independence but this excess is not statistically significant.

```
> crosstabs( ~ drink + wkgrp, data = mistpol.fatal.df, na.action = na.exclude)
Call:
crosstabs(formula = ~ drink + wkgrp, data = mistpol.fatal.df, na.action = na.exclude)
333 cases in table
+----+
IN
|N/RowTotal|
|N/ColTotal|
|N/Total
         +----+
drink |wkgrp
     |SunWed |ThuSat |RowTotl|
-----+
yes | 89 |163 |252 |
   |0.35 |0.65 |0.76 |
|0.71 |0.79 |
|0.27 |0.49 |
-----+
no | 37 | 44 | 81 | |
|0.46 |0.54 |0.24 |
|0.29 |0.21 | |
|0.11 |0.13 | |
----+
ColTotl|126 |207 |333 |
|0.38 |0.62 | |
----+
Test for independence of all factors
      Chi^2 = 2.797941 d.f.= 1 (p=0.09438545)
      Yates' correction not used
> get.crosstabs.percenterror( ~ drink + wkgrp)
   [,1] [,2]
[1,] -7 4
[2,] 21 -13
> contrib( ~ drink + wkgrp)
contribution to chi-sq:
[,1] [,2]
[1,] -0.65 0.51
[2,] 1.15 -0.90
```

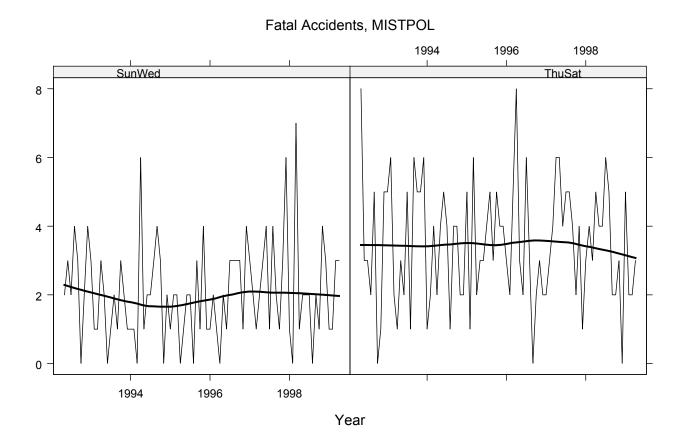
Drink and Gender Males are strongly associated with more fatal accidents involving drinking.

Call: crossta 326 cas + N N/Row! N/Row! N/Col! N/Tota +	abs(form ses in t fotal Total al +	ula = ~		a = mistpol.fatal.df, na.action = na.exclude) sex, data = mistpol.fatal.df, na.action = na.exclude)
drink	male	female		
yes	200 0.81 0.8 0.61		246 0.75 	
no	49 0.61 0.2 0.15	31 0.39 0.4 0.095	80 0.25 	
ColTot	L 249	-+ 77 0.24	326	
> get.c [, [1,]	Chi^2 = Yates'	corrections.percent	2 d.f.= 1 on not us	(p=0.000244625)
contrib [1,] (59	x)	

Time Series Trellis Plots of Weekgroup

Figure 4 shows that ThuSat group has more accidents. There is a very small downward trend starting around 1996 in ThuSat.

Figure 4.



Tables and Seasonal Mann-Kendall test for fatal accidents by wkgrp

> accid	lent	s.fa	atal	.Sun	Ved							
J	「an	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					2	3	2	4	3	0	2	4
1993:	3	1	1	3	2	0	1	2	1	3	2	1
1994:	1	1	0	6	1	2	2	3	4	3	0	2
1995:	1	2	2	0	1	2	2	0	3	1	4	1
1996:	1	2	1	0	2	1	3	3	3	3	1	4
1997:	3	2	1	2	3	4	1	4	2	1	3	6
1998:	1	0	7	1	2	2	2	0	2	1	4	3
1999:	1	1	3	3	5							
> Seasc	nal	Manr	nKend	dall	(acc:	Ldent	cs.fa	atal	.Sun	Ved)		
tau = C				L =33						,		
		,										
> accid	_											
/ accit	lent	s.fa	atal.	.Thust	Sat							
				-		Jun	Jul	Aug	Sep	Oct	Nov	Dec
				-		Jun 3	Jul 3	Aug 2	Sep 5	Oct 0	Nov 1	Dec 5
J				-	May			-	-			
J 1992:	Jan	Feb	Mar	Apr	May 8 3	3	3	2	5	0	1	5
J 1992: 1993:	Jan 5	Feb 6	Mar 2	Apr 1	May 8 3 4	3 2	3 5	2 1	5 6	0 5	1 5	5 6
J 1992: 1993: 1994:	Jan 5 1	Feb 6 2	Mar 2 4	Apr 1 2	May 8 3 4 3	3 2 5	3 5 4	2 1 1	5 6 4	0 5 4	1 5 2	5 6 2
1992: 1993: 1994: 1995:	Jan 5 1 5	Feb 6 2 1	Mar 2 4 6	Apr 1 2 2	May 8 3 4 3	3 2 5 3	3 5 4 4	2 1 1 5 3	5 6 4 3	0 5 4 5	1 5 2 4	5 6 2 4
1992: 1993: 1994: 1995: 1996:	Jan 5 1 5 3	Feb 6 2 1 2	Mar 2 4 6 5	Apr 1 2 2 8	May 8 3 4 3 3	3 2 5 3 2	3 5 4 4 6	2 1 1 5 3	5 6 4 3 0	0 5 4 5 2	1 5 2 4 3	5 6 2 4 2
1992: 1993: 1994: 1995: 1996: 1997:	Jan 5 1 5 3 2	Feb 6 2 1 2 3	Mar 2 4 6 5 4	Apr 1 2 2 8 6	May 8 3 4 3 3 6	3 2 5 3 2 4	3 5 4 6 5	2 1 5 3 5	5 6 4 3 0 4	0 5 4 5 2 2	1 5 2 4 3 4	5 6 2 4 2 1
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999:	Jan 5 1 5 3 2 3 5	Feb 6 2 1 2 3 4 2	Mar 2 4 6 5 4 3 2	Apr 1 2 2 8 6 5 3	May 8 3 4 3 3 6 4 1	3 2 5 3 2 4 4	3 5 4 6 5 6	2 1 5 3 5 5	5 6 4 3 0 4 2	0 5 4 5 2 2 2	1 5 2 4 3 4	5 6 2 4 2 1
1992: 1993: 1994: 1995: 1996: 1997: 1998:	Jan 5 1 5 3 2 3 5 0nal	Feb 6 2 1 2 3 4 2 .Manr	Mar 2 4 6 5 4 3 2 0Kenc	Apr 1 2 2 8 6 5 3 3 1 all	May 8 3 4 3 3 6 4 1	3 2 5 3 2 4 4 4	3 5 4 6 5 6	2 1 5 3 5 5	5 6 4 3 0 4 2	0 5 4 5 2 2 2	1 5 2 4 3 4	5 6 2 4 2 1

Time Series Trellis Plots of Hour

Increasing trend at 1AM and 11PM in recent years (Figure 5)

Figure 5.

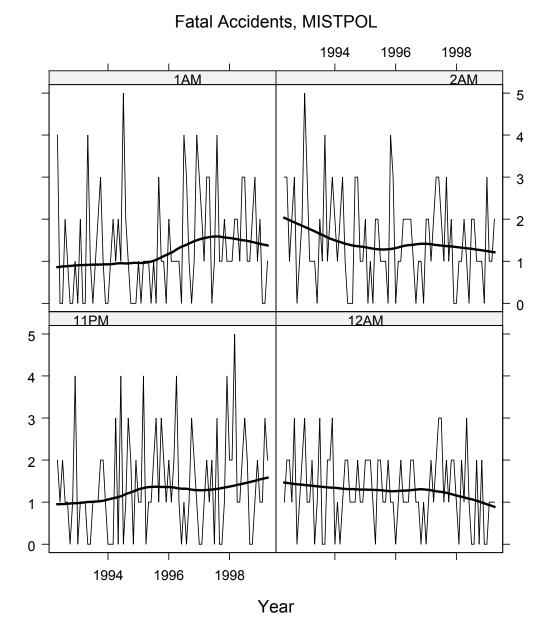


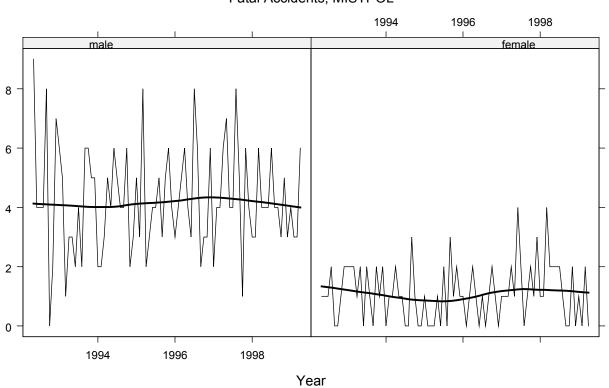
Table and Mann-Kendall test, fatal accidents by hour

> acc	ident	cs.fa	atal	.11PI	4								
		Feb				Jun	Jul	Auq	Sep	Oct	Nov	Dec	
1992:			-	I	2	1	2	1	1	0	1	4	
1993:	0	1	1	1	0	0	1	1	1	2	2	1	
1994:	0	0	0	3	0	4	0	1	3	2	0	2	
1995:	1	1	4	0	1	1	2	3	1	3	2	1	
1996:	2	1	2	4	1	0	1	0	1	3	2	1	
1997:	0	0	1	2	1	2	0	3	0	0	1	4	
1998:	2	2	5	1	1	2	3	2	0	0	1	2	
1999:	1	1	3	2	3	2	5	2	0	0	1	2	
> Sea:		_	-	_	-	idant	-c f	-+1	11 אס	Л)			
				=20		Luein	-9•T0	atar		.1)			
tau =	0.12	,	SI	-20	.40								
> acc:	idant	- a f	+ - 1	1 א ד	л								
> acc.						Turn	т., 1	⊅ ∝	Com	Oat	Nor	Dee	
1002.	Jan	Feb	Mar	Apr									
1992:	2	1	1	0	1	2	2	1	3	0	1	2	
1993:	3	1	1	2	0	1	3	0	0	2	2	3	
1994:	0	1	0	1	2	2	1	1	1	2	1	1	
1995:	2	2	2	1	0	2	2		1	2	1	1	
1996:	0	1	2	1	1	1	2		1	1	-	1	
1997:	0	1	2	1	2	3	3	1	2	1		2	
1998:	1	0	2	1	3	1	0	0	2	0	2	0	
1999:	0	1	1	1	2			_					
> Sea							cs.fa	atal	.12AM	4)			
tau =	-0.0)56,	S	L =50	5.739	0							
> acc													
		ts.fa Feb			Мау	Jun	Jul		Sep	Oct		Dec	
1992:	Jan	Feb	Mar	Apr	4	0	0	2	1	0	0	1	
1992: 1993:	Jan 0	Feb 2	Mar O	Apr 0	4 4	0 1	0 0	2 1	1 2	0 3	0 1		
1992: 1993: 1994:	Jan 0 0	Feb 2 1	Mar 0 2	Apr 0 1	4 4 2	0	0 0 5	2 1 2	1 2 1	0 3 0	0 1 0	1	
1992: 1993: 1994: 1995:	Jan 0 0 1	Feb 2 1 0	Mar 0 2 1	Apr 0	4 4 2 1	0 1	0 0	2 1 2 0	1 2 1 3	0 3	0 1 0 1	1 0 0 0	
1992: 1993: 1994: 1995: 1996:	Jan 0 0 1 2	Feb 2 1 0 1	Mar 0 2 1 1	Apr 0 1 1 1	4 4 2 1 1	0 1 1	0 0 5	2 1 2 0	1 2 1	0 3 0	0 1 0 1 1	1 0 0 4	
1992: 1993: 1994: 1995: 1996: 1997:	Jan 0 0 1	Feb 2 1 0	Mar 0 2 1 1 1	Apr 0 1 1	4 4 2 1	0 1 1 0 0	0 0 5 1	2 1 2 0 3 4	1 2 1 3 1 1	0 3 0 1 0 1	0 1 0 1 1	1 0 0 4 1	
1992: 1993: 1994: 1995: 1996: 1997: 1998:	Jan 0 1 2 3 1	Feb 2 1 0 1	Mar 0 2 1 1	Apr 0 1 1 3 2	4 4 2 1 1	0 1 1 0 0	0 0 5 1 4	2 1 2 0 3	1 2 1 3 1	0 3 0 1 0	0 1 0 1 1	1 0 0 4	
1992: 1993: 1994: 1995: 1996: 1997:	Jan 0 0 1 2 3	Feb 2 1 0 1 2	Mar 0 2 1 1 1	Apr 0 1 1 3	4 4 2 1 1 3	0 1 1 0 0	0 0 5 1 4 1	2 1 2 0 3 4	1 2 1 3 1 1	0 3 0 1 0 1	0 1 0 1 1 2	1 0 0 4 1	
1992: 1993: 1994: 1995: 1996: 1997: 1998:	Jan 0 1 2 3 1 2	Feb 2 1 0 1 2 1 0	Mar 0 2 1 1 1 2 0	Apr 0 1 1 1 3 2 1	4 2 1 3 1	0 1 1 0 0 0 3	0 5 1 4 3	2 1 2 0 3 4 1	1 2 1 3 1 1	0 3 0 1 0 1 2	0 1 0 1 1 2	1 0 0 4 1	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999:	Jan 0 1 2 3 1 2 sonal	Feb 2 1 0 1 2 1 0 1Manr	Mar 0 2 1 1 2 0 0 Keno	Apr 0 1 1 1 3 2 1	4 2 1 3 1 1 (acci	0 1 1 0 0 0 3	0 5 1 4 3	2 1 2 0 3 4 1	1 2 1 3 1 1	0 3 0 1 0 1 2	0 1 0 1 1 2	1 0 0 4 1	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea:	Jan 0 1 2 3 1 2 sonal	Feb 2 1 0 1 2 1 0 1Manr	Mar 0 2 1 1 2 0 0 Keno	Apr 0 1 1 3 2 1 dall	4 2 1 3 1 1 (acci	0 1 1 0 0 0 3	0 5 1 4 3	2 1 2 0 3 4 1	1 2 1 3 1 1	0 3 0 1 0 1 2	0 1 0 1 1 2	1 0 0 4 1	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea:	Jan 0 1 2 3 1 2 sonal 0.15	Feb 2 1 0 1 2 1 0 1 Manr 5,	Mar 0 2 1 1 2 0 NKeno sl =	Apr 0 1 1 3 2 1 1 dall =11.8	4 2 1 3 1 1 (acci	0 1 1 0 0 0 3	0 5 1 4 3	2 1 2 0 3 4 1	1 2 1 3 1 1	0 3 0 1 0 1 2	0 1 0 1 1 2	1 0 0 4 1	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea: tau =	Jan 0 1 2 3 1 2 sonal 0.15	Feb 2 1 0 1 2 1 0 1 Manr 5,	Mar 0 2 1 1 2 0 NKeno sl =	Apr 0 1 1 3 2 1 dall =11.8	4 2 1 3 1 1 3 3 1 8	0 1 0 0 3 ident	0 5 1 4 3 cs.fa	2 1 2 0 3 4 1 atal	1 2 1 3 1 1 1 AM)	0 3 0 1 0 1 2	0 1 0 1 1 2 3	1 0 0 4 1 1	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea: tau =	Jan 0 1 2 3 1 2 sonal 0.15	Feb 2 1 0 1 2 1 0 1 Manr 5, 5	Mar 0 2 1 1 2 0 NKeno sl =	Apr 0 1 1 3 2 1 dall =11.8	4 2 1 3 1 1 3 3 1 8	0 1 0 0 3 ident	0 5 1 4 3 cs.fa	2 1 2 0 3 4 1 atal	1 2 1 3 1 1 1 AM)	0 3 0 1 0 1 2	0 1 0 1 1 2 3	1 0 0 4 1 1	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea: tau = > acc:	Jan 0 1 2 3 1 2 sonal 0.15 ident Jan	Feb 2 1 0 1 2 1 0 1 Manr 5, 5	Mar 0 2 1 1 2 0 NKeno sl = Atal Mar	Apr 0 1 1 3 2 1 dall =11.8	4 2 1 3 1 (acc) 31% May	0 1 0 0 3 ident	0 5 1 4 3 cs.fa Jul 1	2 1 2 0 3 4 1 atal Aug 2	1 2 1 3 1 1 .1AM) Sep	0 3 0 1 0 1 2 0 0 ct 0	0 1 0 1 2 3 Nov	1 0 0 4 1 1 2	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > acc: 1992:	Jan 0 1 2 3 1 2 sonal 0.15	Feb 2 1 0 1 2 1 0 1 Manr 5, 5 5 5	Mar 0 2 1 1 2 0 NKeno sl =	Apr 0 1 1 3 2 1 dall =11.8	4 4 2 1 3 1 31% May 3	0 1 0 0 3 ident Jun 3	0 5 1 4 3 cs.fa	2 1 2 0 3 4 1 atal	1 2 1 3 1 1 1 .1AM) Sep 3	0 3 0 1 0 1 2 0 0 0	0 1 0 1 2 3 Nov	1 0 0 4 1 1 2 Dec	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > acc: 1992: 1993:	Jan 0 1 2 3 1 2 sonal 0.15 ident Jan 5	Feb 2 1 0 1 2 1 0 1 Manr 5, 5 5 5 5 3	Mar 0 2 1 1 2 0 NKeno sl = Atal Mar 1	Apr 0 1 1 3 2 1 dall =11.8 .2AM Apr 1	4 2 1 3 1 (acc) 31% May 3 1	0 1 0 0 3 ident Jun 3 0	0 5 1 4 1 3 cs.fa Jul 1 2	2 1 2 0 3 4 1 atal Aug 2 1	1 2 1 1 1 1 .1AM) Sep 3 4	0 3 0 1 2 0 0 1 2	0 1 0 1 2 3 Nov 1 2	1 0 0 4 1 1 2 3	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea: tau = > acc: 1992: 1993: 1994:	Jan 0 1 2 3 1 2 sonal 0.15 ident Jan 5 2	Feb 2 1 0 1 2 1 0 1 Manr 5, 5 5 5 5 1	Mar 0 2 1 1 2 0 Keno sl = atal Mar 1 2 1	Apr 0 1 1 2 1 dall =11.8 .2AM Apr 1 3	4 4 2 1 3 1 (acc) 3 1% May 3 1 1	0 1 0 0 3 ident Jun 3 0 2	0 5 1 4 1 3 cs.fa Jul 1 2 0 1	2 1 2 0 3 4 1 atal Aug 2 1 0	1 2 1 1 1 1 .1AM) Sep 3 4 3	0 3 0 1 2 0 0 1 2 0 0 1 3	0 1 1 2 3 Nov 1 2 1	1 0 4 1 1 2 2 3 1 3	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea: tau = > acc: 1992: 1992: 1993: 1994: 1995: 1996:	Jan 0 1 2 3 1 2 sonal 0.15 ident Jan 5 2 2	Feb 2 1 0 1 2 1 0 1 Manr 5, 5 5 5 5 5 1 0	Mar 0 2 1 1 2 0 NKeno sl = atal Mar 1 2	Apr 0 1 1 3 2 1 dall =11.8 .2AM Apr 1 3 0	4 4 2 1 3 1 (acc) 3 1% May 3 1 1 2	0 1 0 0 3 ident Jun 3 0 0	0 5 1 4 1 3 cs.fa Jul 2 0 1 2	2 1 2 0 3 4 1 atal Aug 2 1 0 1 1	1 2 1 1 1 1 .1AM) Sep 3 4 3 1	0 3 0 1 2 0 0 0 0 1 3 0	0 1 1 2 3 Nov 1 2 1 4 1	1 0 0 4 1 1 2 3 1	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea: tau = > acc: 1992: 1992: 1993: 1994: 1995: 1996: 1997:	Jan 0 1 2 3 1 2 sonal 0.15 ident Jan 5 2 2 0	Feb 2 1 0 1 2 1 0 1 Manr 5, 5 5 5 5 1 0 1 0 1	Mar 0 2 1 1 2 0 NKeno sl = atal Mar 1 2 1 1	Apr 0 1 1 2 1 dall =11.8 .2AM Apr 1 3 0 2	4 4 2 1 3 1 (acc) 3 1% May 3 1 1 2 2	0 1 0 0 3 ident Jun 3 0 2 2	0 5 1 4 1 3 cs.fa Jul 1 2 0 1	2 1 2 0 3 4 1 atal Aug 2 1 0 1	1 2 1 3 1 1 1 .1AM) Sep 3 4 3 1 0	0 3 0 1 2 0 1 2 0 1 3 0 1 3 0	0 1 1 2 3 Nov 1 2 1 4	1 0 0 4 1 1 1 Dec 2 3 1 3 0	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea: tau = > acc: 1992: 1993: 1994: 1995: 1996: 1997: 1998:	Jan 0 1 2 3 1 2 5 0.15 1 0.15 1 1 0.15 2 2 2 0 2 0 2 0 2 0	Feb 2 1 0 1 2 1 0 1 Manr 5, 5 5 5 5 6 7 7 1 0 1 2 1 0 1 2 1 0 1 2 1 0 1 1 2 1 0 1 2 1 0 1 2 1 0 1 2 1 0 1 2 1 0 1 2 1 0 1 2 1 0 1 0	Mar 0 2 1 1 2 0 NKeno sl = atal. Mar 1 2 1 1 1 1	Apr 0 1 1 2 1 dall =11.8 .2AM Apr 1 3 0 2 2 2 2	4 4 2 1 3 1 (acc 3 1% May 3 1 1 2 2 3	0 1 0 0 3 ident Jun 3 0 2 2 3	0 5 1 4 1 3 5.fa Jul 1 2 0 1 2 2	2 1 2 0 3 4 1 atal Aug 2 1 0 1 1 1	1 2 1 3 1 1 1 .1AM) Sep 3 4 3 1 0 3	0 3 0 1 2 0 0 1 2 0 1 3 0 1 1	0 1 2 3 Nov 1 2 1 4 1 2	1 0 4 1 1 2 3 1 3 0 0	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea: tau = > acc: 1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999:	Jan 0 1 2 3 1 2 5 0.15 1 0.15 1 dent Jan 5 2 2 0 2 0 3	Feb 2 1 0 1 2 1 0 1 Mann 5, 5 5 5 5 1 0 1 2 1 1 2 1 1	Mar 0 2 1 1 2 0 MKeno sl = atal Mar 1 2 1 1 1 1 1	Apr 0 1 1 2 1 dall =11.8 .2AM Apr 1 3 0 2 2 2 2 2	4 4 2 1 3 1 (acc 3 1 % May 3 1 2 2 3 1 0	0 1 0 0 3 ident 3 0 2 2 3 0	0 5 1 4 1 3 5.fc Jul 1 2 0 1 2 2 2	2 1 2 0 3 4 1 atal 2 0 1 1 1 2	1 2 1 3 1 1 1 .1AM) Sep 3 4 3 1 0 3 1	0 3 0 1 2 0 1 2 0 1 3 0 1 1 1 1	0 1 2 3 Nov 1 2 1 4 1 2	1 0 4 1 1 2 3 1 3 0 0	
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Sea: tau = > acc: 1992: 1993: 1994: 1995: 1996: 1997: 1998:	Jan 0 1 2 3 1 2 5 5 0.15 1 dent Jan 5 2 2 0 2 0 2 0 3 3 50na]	Feb 2 1 0 1 2 1 0 1 Manr 5, 5 5 5 5 6 7 1 0 1 2 1 1 1 1 1 1 1 1 1	Mar 0 2 1 1 2 0 NKeno sl = atal 1 2 1 1 1 1 1 1 1 1 1 1 1 1	Apr 0 1 1 2 1 1 3 2 1 1 3 2 1 1 3 0 2 2 2 2 2 2 2 1 1	4 4 2 1 3 1 (acc: 31% May 3 1 2 2 3 1 0 (acc: 3 1 0 0	0 1 0 0 3 ident 3 0 2 2 3 0 0 ident	0 5 1 4 1 3 5.fc Jul 1 2 0 1 2 2 2	2 1 2 0 3 4 1 atal 2 0 1 1 1 2	1 2 1 3 1 1 1 .1AM) Sep 3 4 3 1 0 3 1	0 3 0 1 2 0 1 2 0 1 3 0 1 1 1 1	0 1 2 3 Nov 1 2 1 4 1 2	1 0 4 1 1 2 3 1 3 0 0	

Time Series Trellis Plots of Gender

Males have more fatal accidents. No trend in fatalities with males or females (Figure 6a and Figure 6b).

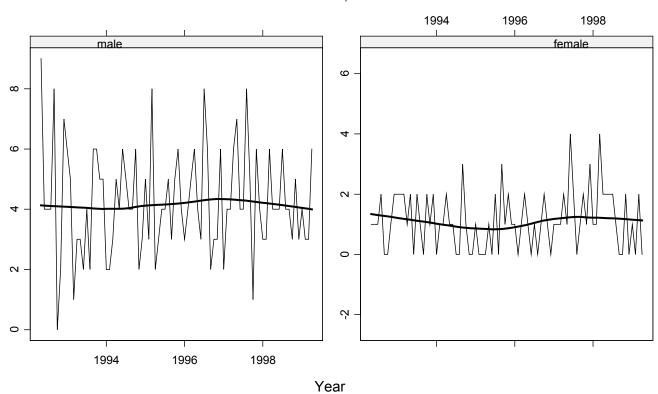
Figure 6a. Common scale.



Fatal Accidents, MISTPOL

Figure 6b. Sliced vertical scale.

The plot is redone allowing sliced scaling on the vertical axis between the panels.



Fatal Accidents, MISTPOL

Tables and Seasonal Mann-Kendall tests

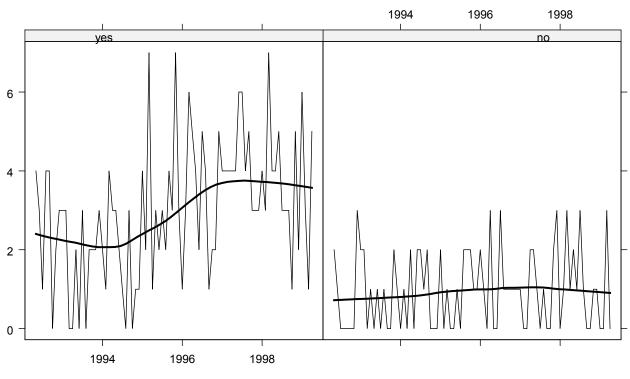
accidents.fatal.male													
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec	
1992:					9	4	4	4	8	0	2	7	
1993:	6	5	1	3	3	2	4	2	6	6	5	5	
1994:	2	2	3	5	4	6	5	4	4	6	2	3	
1995:	5	3	8	2	3	4	4	5	3	5	6	4	
1996:	3	4	5	6	4	3	8	6	2	3	3	6	
1997:	2	4	4	6	7	4	4	8	5	1	6	4	
1998:	3	3	6	4	4	4	6	4	4	3	5	3	
1999:	4	3	3	6	6								
> Seas	sonal	LManı	nKend	dall	(acci	ident	cs.fa	atal	.male	∋)			
tau =	0.05	501,	s	L = 59	9.44%	0							

> accidents.fatal.female												
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					1	1	1	2	0	0	1	2
1993:	2	2	2	1	2	0	2	1	0	2	1	2
1994:	0	1	1	2	1	1	0	0	3	1	0	0
1995:	1	0	0	0	1	0	2	0	3	1	2	1
1996:	1	0	1	2	1	0	1	0	1	2	1	0
1997:	1	1	1	2	1	4	2	0	1	2	1	3
1998:	1	1	4	2	2	2	2	1	0	0	2	0
1999:	1	0	2	0	0							
> Seas	sonal	LManı	nKend	dall	(acc:	ident	cs.fa	atal.	.fema	ale)		
tau =	Ο,	sl	=100) 응								

Time Series Trellis Plots of Drink

There has been a striking increase in fatalities involving drinking since 1996 (Figure 7a and 7b).

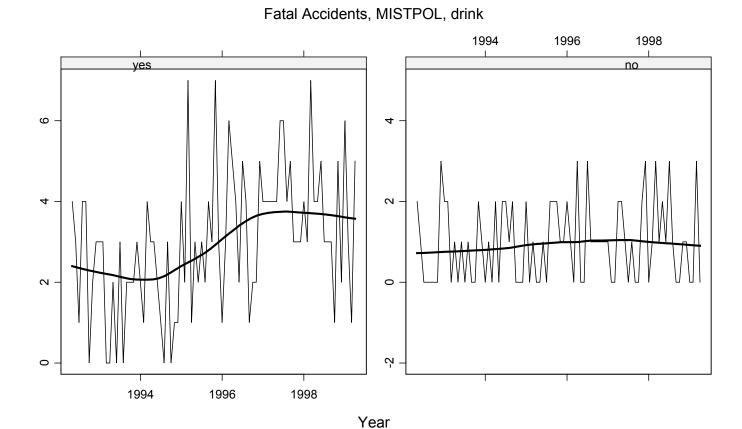
Figure 7a. Fatal accidents by drinking. Common scale.



Fatal Accidents, MISTPOL, drink

Year







Tables and Trend Tests: fatal accidents

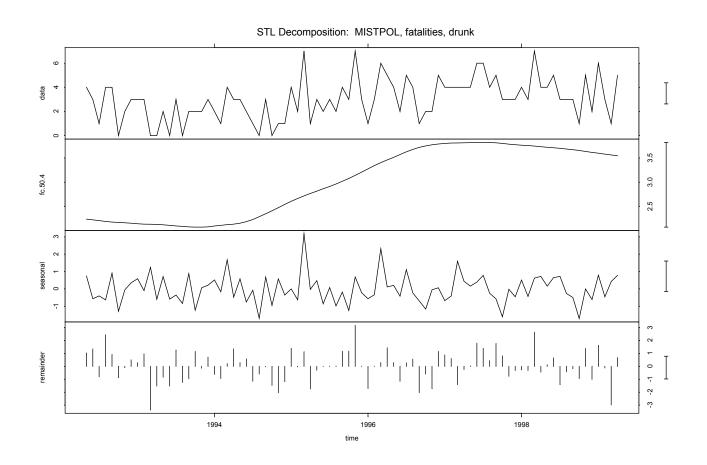
There is an alarming upward trend in fatal accidents associated with drinking. No trend in other.

```
> accidents.fatal.yes
    Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec
1992:
                  4
                     3
                        1
                           4
                              4
                                 0
                                    2
                                       3
                  2 0
                        3
                              2
                                 2
                                    2
                                       3
1993:
     3
         3
            0
                           0
               0
1994: 2 1 4
               3
                3 2
                            3
                                    1
                                       1
                       1
                           0
                               0
1995: 4 2 7 1 3 2
                        3 2
                            4 3
                                    7
                                       3
1996: 1 3 6 5 4 2 5 4 1 2 2 5
1997: 4 4 4 4 6
                                      3
                        6
                          4
                              5 3
                                   3
                                1
1998: 4 3 7 4 4 5
                        3
                           3
                              3
                                   5
                                       2
1999: 6
        3 1
               5
                  4
> SeasonalMannKendall(accidents.fatal.yes)
tau = 0.352, sl = 0.02077%
> accidents.fatal.no
    Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec
1992:
                  2
                     1
                        0
                           0
                             0
                                 0
                                    0
                                       3
1993: 2 2
                  0
                                    2
           0
                     1
                        0
                           1
                              0
                                 0
                                       1
               1
1994: 0 1
               2
                 0 2
                        2
                              2
           0
                           1
                                0
                                   0
                                       0
1995: 2 0 1 0 0 1 0 2 2 2
                                  1
                                       1
1996: 2 1 0 3 0 0 3
                           1 1 1
                                    1
                                       1
1997: 1 0 0 2 2 1 0 1 0 0 2 3
1998: 0 1 3 1 2 1 3
                           1 0 0 1
                                       1
1999: 0 0 3 0
                  2
> SeasonalMannKendall(accidents.fatal.no)
tau = 0.0527, sl = 59.31%
```

STL Analysis of Fatal Accidents with Drinking

Fatal accidents with drinking have risen since 1994 (Figure 8). This analysis is based on very small counts. There is no significant lag one autocorrelation but seasonality is evident in the lag 12 autocorrelation of 0.92. The data are approximately Poisson distributed with mean of 3.5. R-sq = 50.8 %. No monotone spread.

Figure 8a.



ss.window = 5 , fc.window = 50.4 , fc.degree = 1

ო

2

 $\overline{}$

0

7

L

Feb

Mar

Jan

Apr

May

Jun

The seasonal variation is not very stable (Figure 8b). Figure 8b.



Т

Jul

Aug

Τ

Sep

Oct

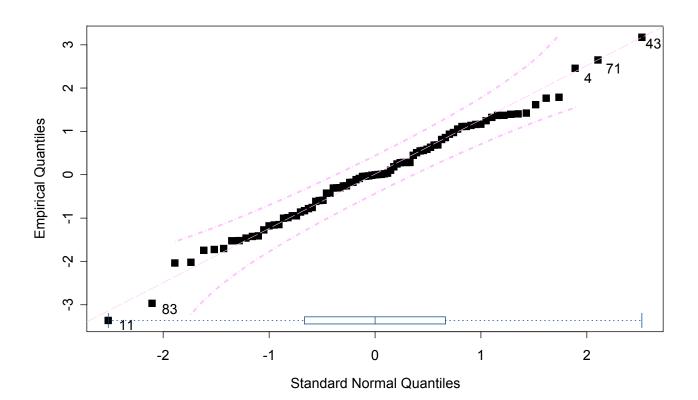
Monthplot, MISTPOL, fatalities, drunk

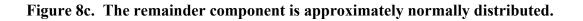
Τ

Nov

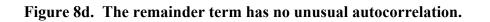
Т

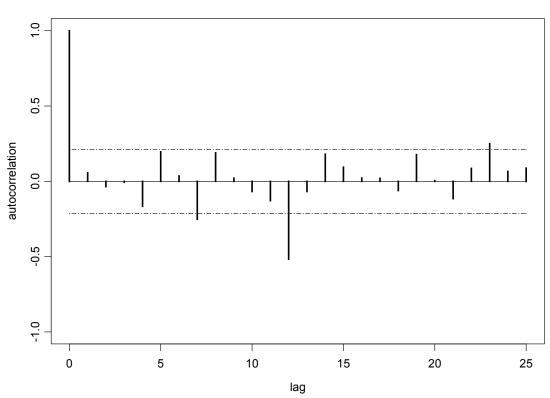
Dec





Skewness Coefficient: $g_1 = -0.1005119$, s.l. = 0.3425395 Michael's Statistic: $D_sp = 0.04366982$, s.l. = 0.6218337 Wilk-Shapiro Statistic: W = 0.9896425, s.l. = 0.9538782

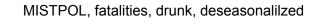




STL Remainder, MISTPOL, fatalities, drunk

Figure 8e. Loess trend analysis of the deseasonalized series. Mann-Kendall test: tau = 0.329, sl =6.569e-

```
sl =6.569e-4%
```



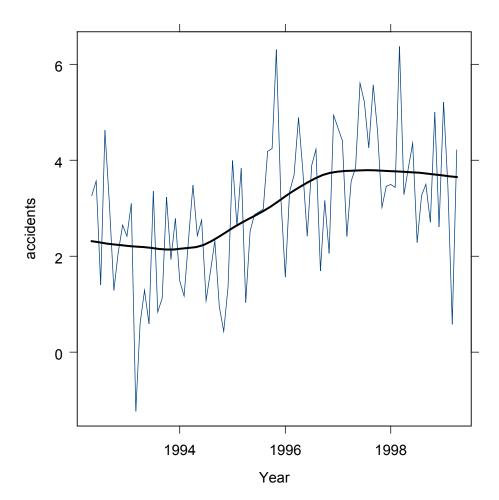
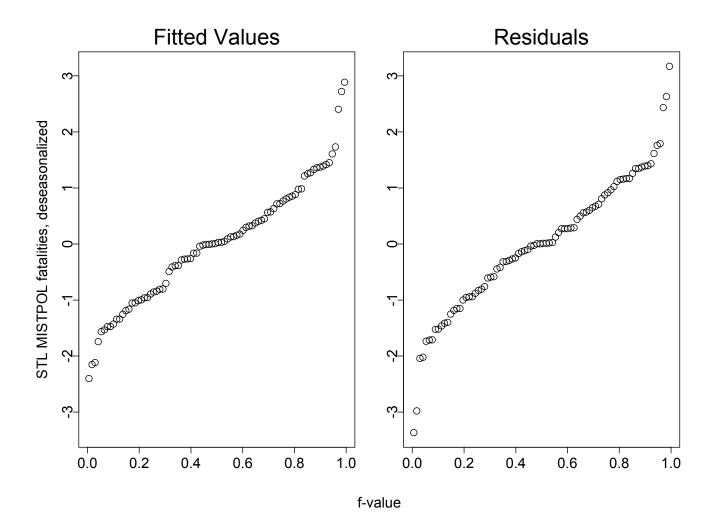


Figure 8f. RF spread plot. Agrees with R-sq = 50.8 %.



J. MISTPOL Dataset Incapacitating Accidents

Summary

Figure 1. There is a clear downward trend in the number of incapacitating accidents (Fig. 1). Since May 1992 the trend line has declined from 60.6 to 37.4 by April 1999. The average annual rate of decline is about 6.7%. The decline is much steeper during the second half of the series (ie. from 1996). Seasonal Mann-Kendall indicates strong trend. Possible outliers noted (Figure 3c and Figure 3e).

The data are approximately Gaussian distributed with mean of 53.7 (Figure 2).

STL analysis confirms the downward trend and shows a change in the seasonal pattern occurred after 1996 (Figure 3a). There is a lot of variability in the seasonal component. There is a peak in May-Jun, Jun-Aug are also high and there is a minimum in Jan-Feb and in Sep (Figure 3b).

Drink-Hour are dependent: there are fewer YES-11PM and more YES-2AM.

Drink-Weekgroup are associated – there are less YES-SunWed as might be supposed.

Drink-Gender are associated: males are strongly associated with more incapacitating accidents involving drinking.

Figure 4 shows that ThuSat group has more incapacitating accidents. There is an evident downward trend in both groups.

By hour. Downward trend in all panels but the slope seems to decrease as it gets later (Figure 5) and the downward trend for 2AM does not start until after 1996. Monotonic trend not significant for 2AM, Table 5.

Males have more incapacitating accidents. Much greater downward trend for males (Figure 6). Monotonic trend for males but not females – Table 6. Test for monotonic trend significant for males bu not significant for females. The loess trend line for incapacitating accidents by males has declined from about 43.0 per month in May 1992 to about 25.3 per month by April 1999. The average annual rate of decline is about 7.3%. It is clear that there is no change for females until after 1996 and then only a very small decline.

Both drinking and non-drinking have downward strong trends (Figure 7).

Strong downward trend in incapacitating accidents involving drinking. Seasonal component has changed since about 1995 and is very changeable (Figure 8a). R-sq = 65.4 %Possible outliers May 1993 (+), Nov 1993 (+) and Feb 1994 (-) – see Figures 8c and 8e. Figure 8b, Apr-Aug and Oct are relatively high and Dec is also relatively low.

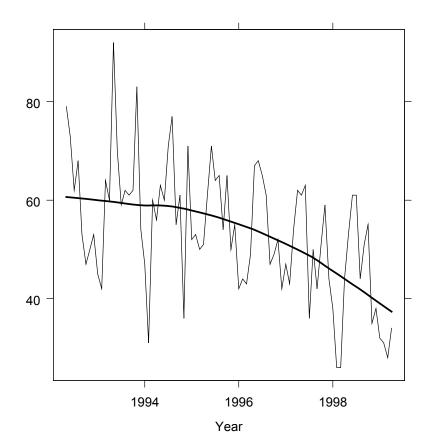
Time Series Tabulation and Loess Analysis

There is a clear downward trend in the number of incapacitating accidents (Fig. 1). Since May 1992 the trend line has declined from 60.6 to 37.4 by April 1999. The average annual rate of decline is about 6.7%. The decline is much steeper during the second half of the series (ie. from 1996). Seasonal Mann-Kendall indicates strong trend. Possible outliers noted (Figure 3c and Figure 3e).

> acc	> accidents.mistpol.incap.ts											
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					79	73	62	68	53	47	50	53
1993:	45	42	64	60	92	70	59	62	61	62	83	54
1994:	47	31	60	56	63	60	71	77	55	61	36	71
1995:	52	53	50	51	61	71	64	65	54	65	50	55
1996:	42	44	43	49	67	68	65	61	47	49	52	42
1997:	47	43	54	62	61	63	36	50	42	51	59	44
1998:	38	26	26	44	53	61	61	44	51	55	35	38
1999:	32	31	28	34								
> Seas	sonal	LManı	nKend	dall	(acc	ident	s.m	istpo	ol.in	ncap	.ts)	
tau =	-0.4	132,	s	L =2	.6e-4	18						
> pc.	char	nge (acci	lden	ts.n	nist	pol.	inca	ap.t	s)		
s	tart	5	e	end	cha	ange	% ar	inua	l.ra	te%		
60.6	5727	7 37	.365	547	38.3	3990	3		6.68	723		

Figure 1.

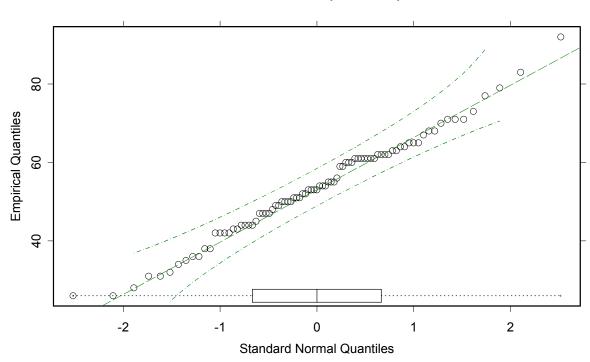
Incapacitating Accidents, MISTPOL



Normal Probability Plot

There is no significant lag one autocorrelation but seasonality is evident in the lag 12 autocorrelation of 0.92. The data are approximately Gaussian distributed with mean of 53.7 (Figure 2).





accidents.mistpol.incap.ts

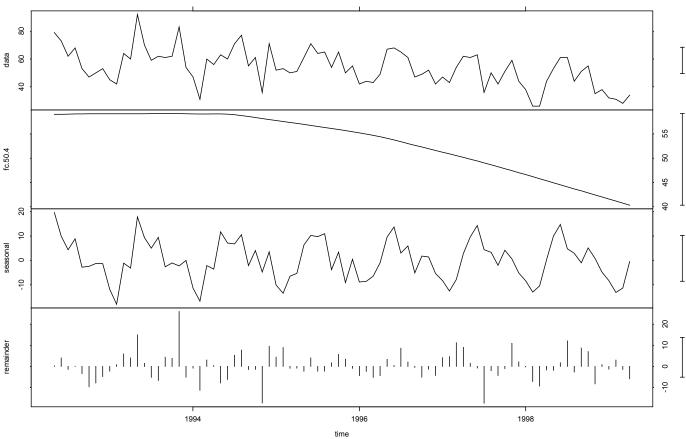
Skewness Coefficient: $g_1 = 0.1004565$, s.l. = 0.3426213 Michael's Statistic: $D_sp = 0.04588675$, s.l. = 0.7811733 Wilk-Shapiro Statistic: W = 0.9823758, s.l. = 0.6912928

STL Analysis

STL analysis confirms the downward trend and shows a change in the seasonal pattern occurred after 1996 (Figure 3a). R-sq = 74.5 %

Figure 3a.

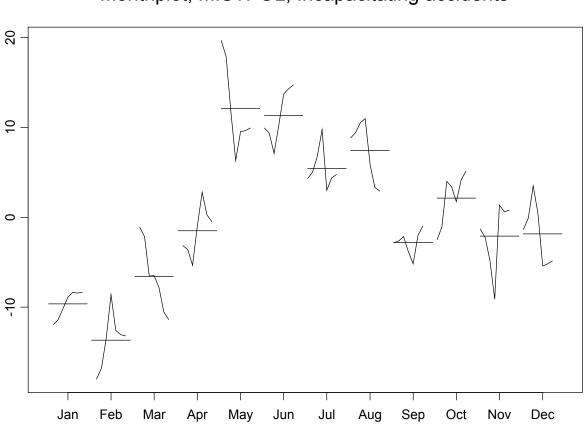
STL Decomposition: MISTPOL, Incapacitating accidents



ss.window = 5 , fc.window = 50.4 , fc.degree = 1

There is a lot of variability in the seasonal component. There is a peak in May-Jun, Jun-Aug are also high and there is a minimum in Jan-Feb and in Sep (Figure 3b).

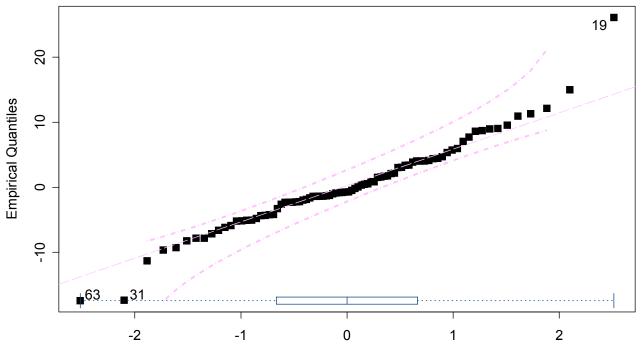
Figure 3b.



Monthplot, MISTPOL, Incapacitating accidents

Figure 3c.

Remainder term is reasonably normal. Some possible outliers noted.



Standard Normal Quantiles

Skewness Coefficient: g_1 = 0.4815951 , s.l. = 0.03155919
Michael's Statistic: $D_{sp} = 0.04984248$, s.l. = 0.9513645
Wilk-Shapiro Statistic: W = 0.9720545 , s.l. = 0.2469895

> cour	nter	.ts										
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					1	2	3	4	5	6	7	8
1993:	9	10	11	12	13	14	15	16	17	18	19	20
1994:	21	22	23	24	25	26	27	28	29	30	31	32
1995:	33	34	35	36	37	38	39	40	41	42	43	44
1996:	45	46	47	48	49	50	51	52	53	54	55	56
1997:	57	58	59	60	61	62	63	64	65	66	67	68
1998:	69	70	71	72	73	74	75	76	77	78	79	80
1999:	81	82	83	84								

Figure 3d.

Strong downward trend in deseasonalized series.

tau = -0.468, sl = 0%

MISTPOL, Incapacitating accidents, deseasonalilzed

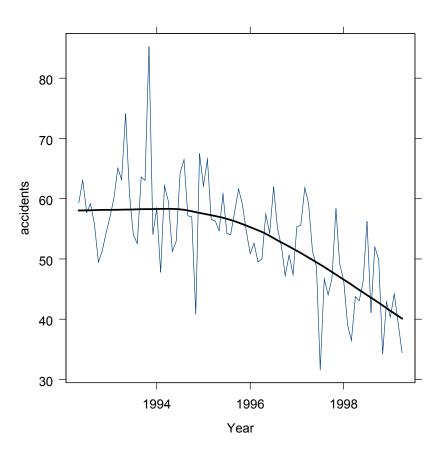
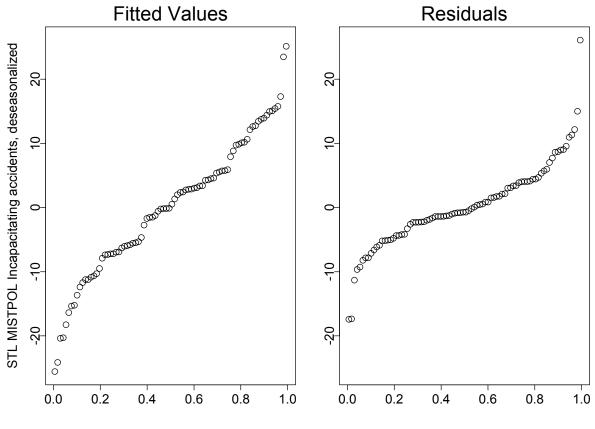


Figure 3e. RF-spread plot

Outliers are evident otherwise agrees more or less with R-sq = 74.5 %



f-value

Crosstab Analysis

Drink and Hour

Drink-Hour are dependent: there are fewer YES-11PM and more YES-2AM.

> crosstabs(~ drink + hour, data = mistpol.incap.df, na.action = na.exclude) Call: crosstabs(formula = ~ drink + hour, data = mistpol.incap.df, na.action = na.exclude) 3210 cases in table +----+ | N |N/RowTotal| |N/ColTotal| |N/Total | +----+ drink |hour |11PM |12AM |1AM |2AM |RowTotl| yes |322 |345 |373 |563 |1603 | |0.2 |0.22 |0.23 |0.35 |0.5 | |0.37 |0.45 |0.56 |0.62 | | |0.1 |0.11 |0.12 |0.18 | | -----+ no |540 |423 |294 |350 |1607 | |0.34 |0.26 |0.18 |0.22 |0.5 | |0.63 |0.55 |0.44 |0.38 | |0.17 |0.13 |0.092 |0.11 | ColTotl|862 |768 |667 |913 |3210 | |0.27 |0.24 |0.21 |0.28 | | Test for independence of all factors Chi^2 = 122.0984 d.f.= 3 (p=0) Yates' correction not used > get.crosstabs.percenterror(~ drink + hour) [,1] [,2] [,3] [,4] [1,] -25 -10 12 23 [2,] 25 10 -12 -23 > contrib(~ drink + hour) contribution to chi-sq: [,1] [,2] [,3] [,4] [1,] -5.23 -1.97 2.19 5.01 [2,] 5.22 1.96 -2.18 -5.01

Drink and Weekgroup

Drink-Weekgroup are associated – there are less YES-SunWed as might be supposed.

```
> crosstabs( ~ drink + wkgrp, data = mistpol.incap.df, na.action = na.exclude)
Call:
crosstabs(formula = ~ drink + wkgrp, data = mistpol.incap.df, na.action = na.exclude)
3208 cases in table
+----+
IN
|N/RowTotal|
|N/ColTotal|
|N/Total
         +----+
drink |wkgrp
     |SunWed |ThuSat |RowTotl|
-----+
yes | 547 |1055 |1602 |
   |0.34 |0.66 |0.5 |
|0.46 |0.52 |
|0.17 |0.33 |
-----+
no | 639 | 967 | 1606 |
|0.4 |0.6 |0.5 |
|0.54 |0.48 | |
|0.2 |0.3 | |
----+
ColTotl|1186 |2022 |3208 |
|0.37 |0.63 | |
----+
Test for independence of all factors
      Chi^2 = 10.96149 d.f.= 1 (p=0.0009302476)
      Yates' correction not used
> get.crosstabs.percenterror( ~ drink + wkgrp)
[,1] [,2]
[1,] -8 4
[2,] 8 -4
> contrib(~ drink + wkgrp)
contribution to chi-sq:
[,1] [,2]
[1,] -1.86 1.42
[2,] 1.86 -1.42
```

October 19, 2001

Drink and Gender

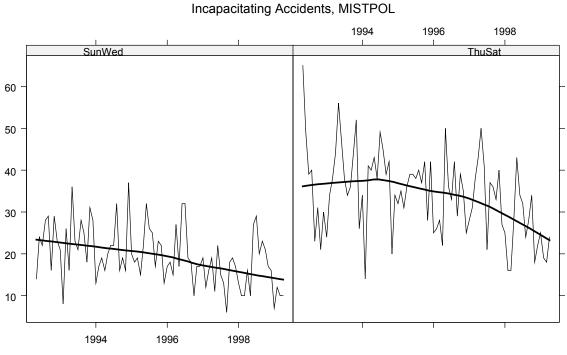
Drink-Gender are associated: males are strongly associated with more incapacitating accidents involving drinking.

```
> crosstabs( ~ drink + sex, data = mistpol.incap.df, na.action = na.exclude)
Call:
crosstabs(formula = ~ drink + sex, data = mistpol.incap.df, na.action = na.exclude)
3169 cases in table
+----+
IN
|N/RowTotal|
|N/ColTotal|
|N/Total
         +----+
drink |sex
     |male |female |RowTotl|
----+
yes |1241 | 340 |1581 |
   0.78 | 0.22 | 0.5 |
0.56 | 0.36 |
0.39 | 0.11 |
-----+
no | 986 | 602 | 1588 |
|0.62 | 0.38 | 0.5 |
|0.44 | 0.64 | |
|0.31 | 0.19 |
----+
ColTotl|2227 |942 |3169 |
|0.7 |0.3 | |
----+
Test for independence of all factors
      Chi^2 = 102.054 d.f.= 1 (p=0)
      Yates' correction not used
> get.crosstabs.percenterror( ~ drink + sex)
[,1] [,2]
[1,] 12 -28
[2,] -12 28
> contrib(~ drink + sex)
contribution to chi-sq:
[,1] [,2]
[1,] 3.90 -5.99
[2,] -3.89 5.98
```

Time Series Trellis Plots of Weekgroup

Figure 4 shows that ThuSat group has more incapacitating accidents. There is an evident downward trend in both groups.

Figure 4.



Year

Table 4

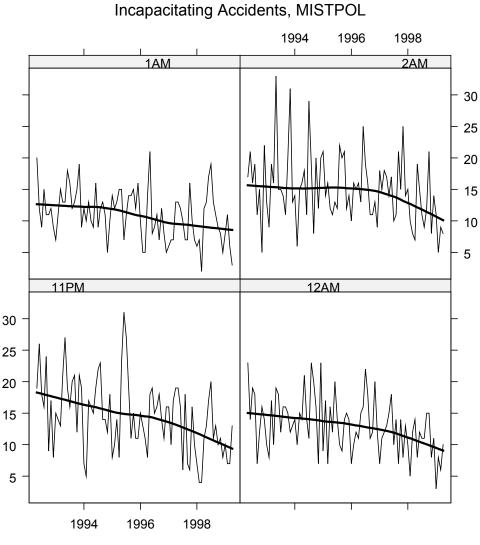
Data and Seasonal Mann-Kendall Tests for Incapacitating Accidents by WKGRP

> acci	ldent	cs.in	ncap	.Sun	Ved							
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					14	24	22	28	29	16	29	23
1993:	21	8	26	16	36	23	21	28	25	18	31	28
1994:	13	17	19	16	20	22	22	32	16	19	16	37
1995:	20	18	19	15	22	32	26	25	17	23	22	13
1996:	17	18	15	27	17	32	32	19	18	10	17	17
1997:	19	12	16	19	11	22	15	13	6	18	19	17
1998:	13	10	10	16	10	27	29	20	23	21	17	16
1999:	7	12	10	10								
> Seas	sonal	lManı	nKend	dall	(acci	Ldent	cs.ir	ncap.	.SunV	Ved)		
tau =	-0.2	274,	s	L =0	.3193	38						
> acci	dent	-s in	lcan	Thus	Sat							
> acci			-			านา	.T11]	Αιια	Sep	Oct	Nov	Dec
			-		Мау			-	-		Nov 21	
1992:	Jan	Feb	Mar	Apr	May 65	49	39	40	23	31	21	30
1992: 1993:	Jan 24	Feb 34	Mar 38	Apr 44	May 65 56	49 47	39 38	40 34	23 36	31 44	21 52	30 26
1992: 1993: 1994:	Jan 24 34	Feb 34 14	Mar 38 41	Apr 44 40	May 65 56 43	49 47 38	39 38 49	40 34 45	23 36 39	31 44 42	21 52 20	30 26 34
1992: 1993: 1994: 1995:	Jan 24 34 32	Feb 34 14 35	Mar 38 41 31	Apr 44 40 36	May 65 56 43 39	49 47 38 39	39 38 49 38	40 34 45 40	23 36 39 37	31 44 42 42	21 52 20 28	30 26 34 42
1992: 1993: 1994: 1995: 1996:	Jan 24 34 32 25	Feb 34 14 35 26	Mar 38 41 31 28	Apr 44 40 36 22	May 65 56 43 39 50	49 47 38 39 36	39 38 49 38 33	40 34 45 40 42	23 36 39 37 29	31 44 42 42 39	21 52 20 28 35	30 26 34 42 25
1992: 1993: 1994: 1995: 1996: 1997:	Jan 24 34 32 25 28	Feb 34 14 35 26 31	Mar 38 41 31 28 38	Apr 44 40 36 22 43	May 65 56 43 39 50 50	49 47 38 39 36 41	39 38 49 38 33 21	40 34 45 40 42 37	23 36 39 37 29 36	31 44 42 42 39 33	21 52 20 28 35 40	30 26 34 42 25 27
1992: 1993: 1994: 1995: 1996: 1997: 1998:	Jan 24 34 32 25 28 25	Feb 34 14 35 26 31 16	Mar 38 41 31 28 38 16	Apr 44 40 36 22 43 28	May 65 56 43 39 50	49 47 38 39 36	39 38 49 38 33	40 34 45 40 42	23 36 39 37 29	31 44 42 42 39	21 52 20 28 35	30 26 34 42 25
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999:	Jan 24 34 25 28 25 25	Feb 34 14 35 26 31 16 19	Mar 38 41 31 28 38 16 18	Apr 44 40 36 22 43 28 24	May 65 56 43 39 50 50 43	49 47 38 39 36 41 34	39 38 49 38 33 21 32	40 34 45 40 42 37 24	23 36 39 37 29 36 28	31 44 42 42 39 33 34	21 52 20 28 35 40	30 26 34 42 25 27
1992: 1993: 1994: 1995: 1996: 1997: 1998:	Jan 24 34 25 28 25 25 25 50na	Feb 34 14 35 26 31 16 19 IMan	Mar 38 41 31 28 38 16 18 0Keno	Apr 44 40 36 22 43 28 24 24 dall	May 65 56 43 39 50 50 43 (acci	49 47 38 39 36 41 34	39 38 49 38 33 21 32	40 34 45 40 42 37 24	23 36 39 37 29 36 28	31 44 42 42 39 33 34	21 52 20 28 35 40	30 26 34 42 25 27

Time Series Trellis Plots of Hour

Downward trend in all panels but the slope seems to decrease as it gets later (Figure 5) and the downward trend for 2AM does not start until after 1996.





Year

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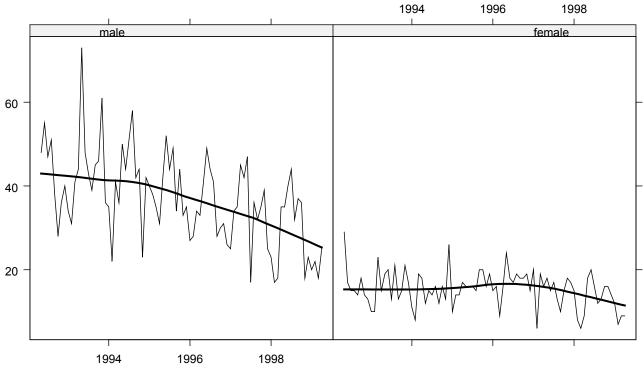
Table 5. Incapacitating Accidents by Hour.Significant downward trend except for 2AM.

Time Series Trellis Plots of Gender

Males have more incapacitating accidents. Much greater downward trend for males (Figure 6a and 6b). Downward trend for females starts after 1996.



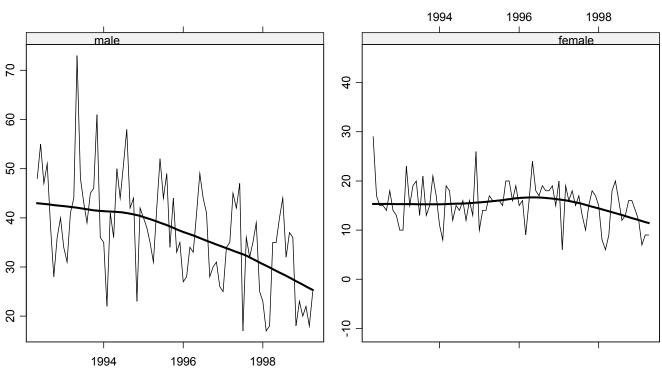
Incapacitating Accidents, MISTPOL



Year

Figure 6b. Sliced Vertical Scale.

It is clear that there is no change for females until after 1996 and then only a very small decline.



Incapacitating Accidents, MISTPOL

Year

Table 6. Incapacitating Accidents by Gender.

Test for monotonic trend significant for males bu not significant for females. The loess trend line for incapacitating accidents by males has declined from about 43.0 per month in May 1992 to about 25.3 per month by April 1999. The average annual rate of decline is about 7.3%.

```
> accidents.incap.male
```

Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1993: 34 1994: 1995: 40 38 31 42 1996: 27 28 34 33 41 1997: 25 45 42 1998: 23 35 35 44 32 37 36 1999: 20 22 18 > SeasonalMannKendall(accidents.incap.male) tau = -0.497, sl =7.373e-6% > pc.change(accidents.incap.male) end change% annual.rate% start 42.98232 25.33472 41.05781 7.273524

> accidents.incap.female

	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					29	17	15	15	14	18	14	13
1993:	10	10	23	15	19	20	13	21	13	15	21	17
1994:	11	8	19	18	12	15	14	16	12	16	13	26
1995:	10	14	14	17	16	16	16	15	20	20	16	19
1996:	15	16	9	16	24	18	17	19	18	18	19	15
1997:	20	6	19	16	18	15	17	13	10	15	18	17
1998:	15	8	6	9	18	20	16	12	13	16	16	14
1999:	12	7	9	9								
> Seas	sonal	lManı	nKend	dall	(acci	Ldent	cs.ir	ncap.	.fema	ale)		
tau =	-0.1	L2,	sl	=20	.04%							

Time Series Trellis Plots of Drink

Both drinking and non-drinking have downward strong trends (Figure 7).

Figure 7.

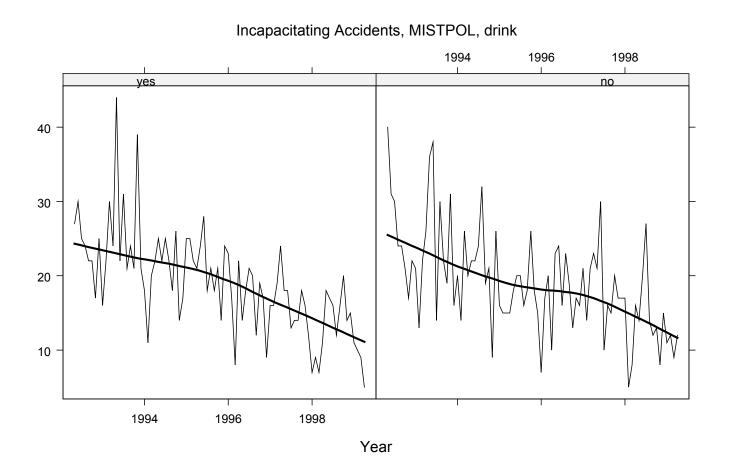


Table 7.Incapacitating accidents by drinking status.Highly significant downward trends in both groups.

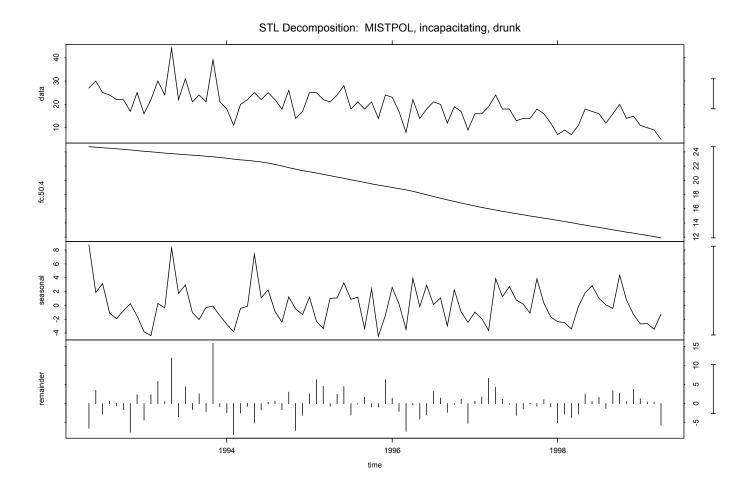
> acc	ident	cs.in	ncap	.yes								
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					27	30	25	24	22	22	17	25
1993:	16	22	30	24	44	22	31	21	24	21	39	21
1994:	18	11	20	22	25	22	25	22	18	26	14	17
1995:	25	25	22	21	24	28	18	21	18	21	14	24
1996:	23	17	8	22	14	18	21	20	12	19	17	9
1997:	16	16	19	24	18	18	13	14	14	18	16	12
1998:	7	9	7	11	18	17	16	12	16	20	14	15
1999:	11	10	9	5								
> Seas	sonal	lManı	nKend	dall	(acc:	ident	ts.in	ncap	.yes))		
tau =	-0.5	596,	S	l =1	.372€	∋-8%						
> pc.0	chang	ge (a	ccide	ents	.inca	ap.ye	es)					
st	tart		end	d cl	hange	e% ar	nnua	l.rat	ce⊗			
24.33	3004	11.1	10533	1 54	.3555	55	1(0.599	926			
> acc												
											Nov	
1992:	Jan	Feb	Mar	Apr	40	31	30	24	24	21	17	22
1992: 1993:	Jan 21	Feb 13	Mar 22	Apr 26	40 36	31 38	30 14	24 30	24 22	21 19	17 31	22 16
1992: 1993: 1994:	Jan 21 20	Feb 13 14	Mar 22 26	Apr 26 20	40 36 22	31 38 22	30 14 24	24 30 32	24 22 19	21 19 21	17 31 9	22 16 26
1992: 1993: 1994: 1995:	Jan 21 20 16	Feb 13 14 15	Mar 22 26 15	Apr 26 20 15	40 36 22 18	31 38 22 20	30 14 24 20	24 30 32 16	24 22 19 18	21 19 21 26	17 31 9 18	22 16 26 15
1992: 1993: 1994: 1995: 1996:	Jan 21 20 16 7	Feb 13 14 15 17	Mar 22 26 15 20	Apr 26 20 15 10	40 36 22 18 23	31 38 22 20 24	30 14 24 20 16	24 30 32 16 23	24 22 19 18 19	21 19 21 26 13	17 31 9 18 17	22 16 26 15 16
1992: 1993: 1994: 1995: 1996: 1997:	Jan 21 20 16 7 21	Feb 13 14 15 17 14	Mar 22 26 15 20 21	Apr 26 20 15 10 23	40 36 22 18 23 21	31 38 22 20 24 30	30 14 24 20 16 10	24 30 32 16 23 16	24 22 19 18 19 15	21 19 21 26 13 20	17 31 9 18 17 17	22 16 26 15 16 17
1992: 1993: 1994: 1995: 1996: 1997: 1998:	Jan 21 20 16 7 21 17	Feb 13 14 15 17 14 5	Mar 22 26 15 20 21 8	Apr 26 20 15 10 23 16	40 36 22 18 23	31 38 22 20 24	30 14 24 20 16	24 30 32 16 23	24 22 19 18 19	21 19 21 26 13	17 31 9 18 17	22 16 26 15 16
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999:	Jan 21 20 16 7 21 17 11	Feb 13 14 15 17 14 5 12	Mar 22 26 15 20 21 8 9	Apr 26 20 15 10 23 16 12	40 36 22 18 23 21 14	31 38 22 20 24 30 20	30 14 24 20 16 10 27	24 30 32 16 23 16 14	24 22 19 18 19 15 12	21 19 21 26 13 20	17 31 9 18 17 17	22 16 26 15 16 17
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas	Jan 21 20 16 7 21 17 11 sona	Feb 13 14 15 17 14 5 12 IMan	Mar 22 26 15 20 21 8 9 0Keno	Apr 26 20 15 10 23 16 12 dall	40 36 22 18 23 21 14 (acc:	31 38 22 20 24 30 20 ident	30 14 24 20 16 10 27	24 30 32 16 23 16 14	24 22 19 18 19 15 12	21 19 21 26 13 20	17 31 9 18 17 17	22 16 26 15 16 17
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau =	Jan 21 20 16 7 21 17 11 50na -0.4	Feb 13 14 15 17 14 5 12 IMann 456,	Mar 22 26 15 20 21 8 9 mKeno	Apr 26 20 15 10 23 16 12 dall 1 =8	40 36 22 18 23 21 14 (acc: .8936	31 38 22 20 24 30 20 ident e-5%	30 14 24 20 16 10 27	24 30 32 16 23 16 14	24 22 19 18 19 15 12	21 19 21 26 13 20	17 31 9 18 17 17	22 16 26 15 16 17
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > pc.c	Jan 21 20 16 7 21 17 11 sona -0.4	Feb 13 14 15 17 14 5 12 IMann 456,	Mar 22 26 15 20 21 8 9 mKeno si	Apr 26 20 15 10 23 16 12 dall 1 =8 ents	40 36 22 18 23 21 14 (acc: .8936 .inca	31 38 22 20 24 30 20 ident e-5% ap.no	30 14 24 20 16 10 27 cs.in	24 30 32 16 23 16 14	24 22 19 18 19 15 12 .no)	21 19 21 26 13 20	17 31 9 18 17 17	22 16 26 15 16 17
1992: 1993: 1994: 1995: 1996: 1997: 1998: 1999: > Seas tau = > pc.c	Jan 21 20 16 7 21 17 11 50na -0.4 chang	Feb 13 14 15 17 14 5 12 IMann 456, ge (ad	Mar 22 26 15 20 21 8 9 nKeno si ccide eno	Apr 26 20 15 10 23 16 12 dall 1 =8 ents d cha	40 36 22 18 23 21 14 (acc: .893e .inca	31 38 22 20 24 30 20 ident e-5% ap.no % ann	30 14 24 20 16 10 27 cs.in	24 30 32 16 23 16 14	24 22 19 18 19 15 12 .no)	21 19 21 26 13 20	17 31 9 18 17 17	22 16 26 15 16 17

STL Analysis of Incapacitating Accidents with Drinking

Strong downward trend in incapacitating accidents involving drinking. Seasonal component has changed since about 1995 and is very changeable (Figure 8). R-sq = 65.4 % Possible outliers May 1993 (+), Nov 1993 (+) and Feb 1994 (-) – see Figures <u>8c</u> and <u>8e</u>.

> acci	ldent	.m	istpo	ol.in	ncap.	.drum	nk.ts	5				
	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					27	30	25	24	22	22	17	25
1993:	16	22	30	24	44	22	31	21	24	21	39	21
1994:	18	11	20	22	25	22	25	22	18	26	14	17
1995:	25	25	22	21	24	28	18	21	18	21	14	24
1996:	23	17	8	22	14	18	21	20	12	19	17	9
1997:	16	16	19	24	18	18	13	14	14	18	16	12
1998:	7	9	7	11	18	17	16	12	16	20	14	15
1999:	11	10	9	5	9							
> Seas	sona	lManı	nKend	dall	(acc	ident	.m	istpo	ol.in	ncap	.dru	nk.ts)
tau =	-0.0	508,	s	L =4	.2740	e-9%						
> pc.c	chang	ge (ad	ccide	ents	.mist	pol.	.inca	ap.dı	runk	.ts)		
st	tart		end	d cl	nange	e% ar	nnua	l.rat	ce∛			
24.33	3004	11.1	10531	L 54	.3555	55	1(0.599	926			
											F	auro Q

Figure 8a.

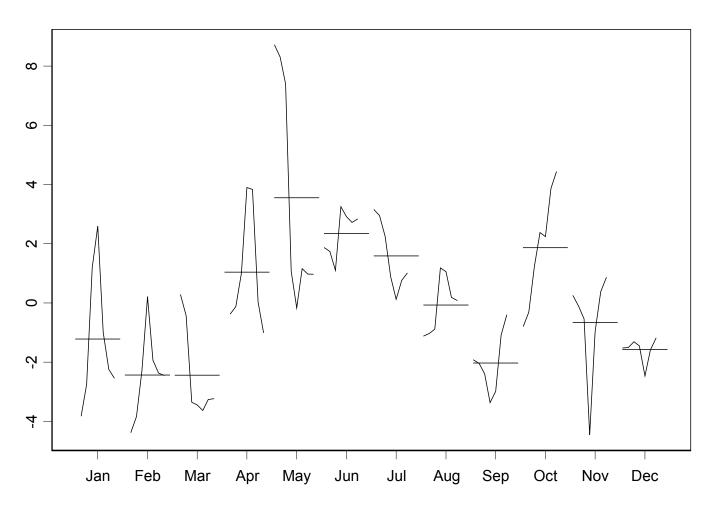


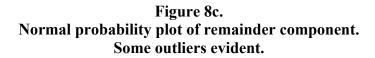
ss.window = 5 , fc.window = 50.4 , fc.degree = 1

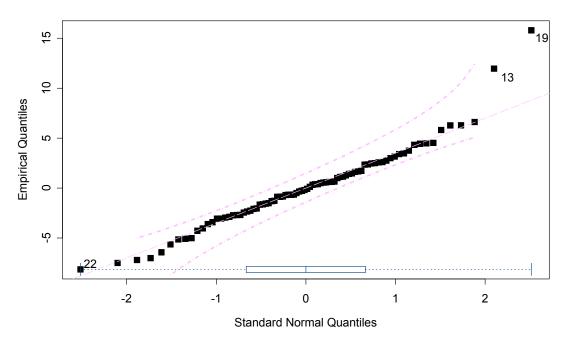
The seasonal variation is not very stable (Figure 9). Apr-Aug and Oct are relatively high and Dec is also relatively low.

Figure 8b

Monthplot, MISTPOL, incapacitating, drunk





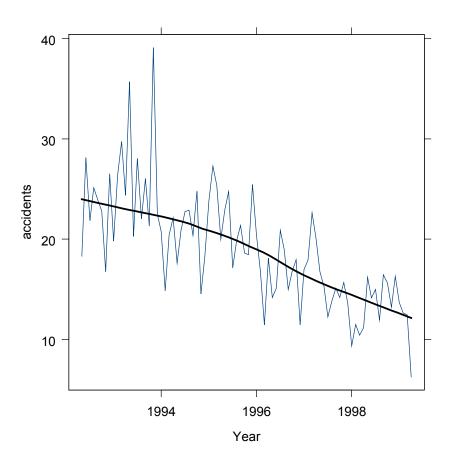


Skewness Coefficient: g_1 = 0.8364291, s.l. = 0.001263527Michael's Statistic: D_sp = 0.06378794, s.l. = 0.2666077Wilk-Shapiro Statistic: W = 0.9563681, s.l. = 0.02245304

> cour	nter	.ts										
	Jan	Feb	Mar	Apr	Мау	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1992:					1	2	3	4	5	6	7	8
1993:	9	10	11	12	13	14	15	16	17	18	19	20
1994:	21	22	23	24	25	26	27	28	29	30	31	32
1995:	33	34	35	36	37	38	39	40	41	42	43	44
1996:	45	46	47	48	49	50	51	52	53	54	55	56
1997:	57	58	59	60	61	62	63	64	65	66	67	68
1998:	69	70	71	72	73	74	75	76	77	78	79	80
1999:	81	82	83	84								

Figure 8d. Loess trend analysis of deasonalized component. tau = -0.557, sl =0%

```
> pc.change(deseasonalized.ts)
    start end change% annual.rate%
23.99675 12.13592 49.42682 9.28003
```



MISTPOL, incapacitating, drunk, deseasonalilzed

Wkgrp time series and seasonal Mann-Kendall trends

> mistpol.total.SunWed.ts Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 774 622 750 1992: 607 840 554 562 611 1993: 555 488 633 618 765 850 704 822 690 526 657 680 1994: 843 720 584 598 663 754 826 899 597 532 715 855 1995: 567 561 819 600 797 922 927 898 549 614 780 828 749 800 678 609 793 952 1006 672 628 637 889 825 1996: 1997: 828 562 537 700 738 857 860 724 625 608 640 796 1998: 816 521 645 630 625 1061 805 762 730 550 647 794 1999: 781 607 642 573 > SeasonalMannKendall(mistpol.total.SunWed.ts) tau = 0.143, sl =11.86%

> mistpol.total.ThuSat.ts

Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec 1992: 1470 1085 839 1227 814 1196 735 930 1993: 1535 1079 1060 1209 1157 1213 1277 1104 984 1393 1325 1066 1994: 1360 981 1164 901 1053 1159 1241 1046 1114 1249 1007 1298 1995: 1489 980 903 1168 946 1173 1204 1076 1086 1043 1281 1546 1996: 1019 1027 1340 1033 1191 1367 1089 1214 1214 1161 1251 1326 1997: 1254 936 1325 1115 1351 1331 1087 1275 974 1095 1377 1088 1998: 1109 795 1149 937 1353 1055 1212 1199 940 1187 872 1033 1999: 1326 880 923 928 > SeasonalMannKendall(mistpol.total.ThuSat.ts) tau = -0.111, sl =22.48%

Aspect Ratio

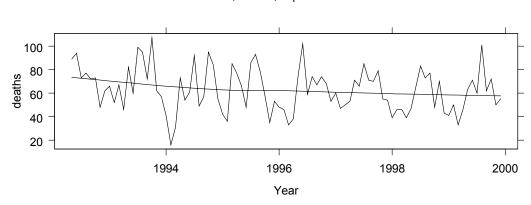
The aspect-ratio is defined as

(physical length of vertical axis)÷(physical length of horizontal axis)

The aspect-ratio for some displays is shown in the table below.

	Horizontal	Vertical	Aspect-ratio
My TV	19.5 inches	25.5 inches	0.46
Video formatted for movie theatre on my TV	11 5/8 inches	25.5 inches	0.76
My Computer Display	1200 pixels	1600 pixels	0.75

As an illustrative example the aspect-ratio in Figure 1 is chosen to be 0.25. Figure 1. FARS deaths, Aspect-ratio is 0.25



FARS, deaths, aspect-ratio=0.25

October 6, 2001 Golden Ratio

Mathematica uses by default an aspect-ratio which is the reciprocal of the Golden Ratio, approximately 0.61803. The Golden Ratio may be defined as the proportion of the division of a line so that the larger is to the smaller as the the smaller is to the whole. Let the smaller line segment have length one and the larger line segment have length Y then the golden ratio is the positive solution of the equation,

$$1/Y = Y/(Y+1)$$
 (1)

this may also be written

Y =	- 1	+	1/Y	(2	2)
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or

1/Y = Y-1 (3)

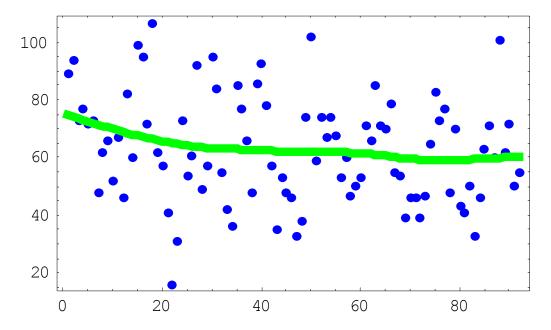
An equivalent definition is for rectangles. In this case the golden ratio is the ratio of the difference between the sides of the rectangle and the smaller side is equal to the ratio of the smaller side to the larger side. Let Y denote the larger side and letting the smaller side have unit length then we see that this is expressed by equation (3).

The golden ratio is denoted by $\varphi = (1 + \sqrt{5})/2$, with numerical value approximately 1.61803 which has reciprocal value approximately equal to 61803.

The aspect ratio is supposed in classical asthetic theory to be uniquely pleasing to the eye.

By default, *Mathematica* uses the reciprocal of the Golden Ratio as the default aspect ratio. Below is a plot of the FARS deaths with a locally-linear robust loess 60% smoother drawn in Mathematica.

Figure 2.

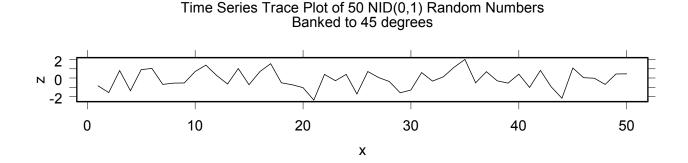


The Banking to 45° Principle

Cleveland (Elements of Graphing Data, 1994; Visualizing Data, 1993) developed a more scientific approach to the choice of aspect-ratio for data visualization. Cleveland's suggestion, which he calls banking to is to 45°, is to choose the aspect-ratio so the average absolute weighted of what is of interest is approximately 45°. Banking to 45° improves the accuracy of our visualization of data and trends. This has been demonstrated experimentally and also using Cleveland's theory of graphical perception (Cleveland, 1994, Ch. 4).

If you have a stationary time series (with essentially zero slope) and you bank to 45° all the successive line segments joining the individual values then the aspect-ratio will be a very small number, typically << 0.25. On the other hand if you bank to 45° the estimated trend which itself has slope close to zero then the aspect-ratio will be a large number >>4. Sometimes this automatic principle gives aspect-ratios which are too small or too large for the resolution or size of the paper or display. In this case we just regard the optimal aspect-ratio, obtained by banking to 45° , as a general indication for what direction we should move in choose an aspect-ratio for our plot. Another problem is that banking to a loess curve may sometimes be at the expense of good data resolution.

Figure 3. Banking a stationary time series trace plot to 45°.



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Appendix 1: Aspect-Ratios

Illustrative FARS Time Series Examples

Figure 4 below shows the FARS deaths time series in which the successive line segments joining the data values are banked to 45°. This is very different from the case of a stationary time series due to the presence of seasonality.

Figure 4. The data is banked to 45° to choose the aspect-ratio.

100 - 100

FARS, deaths aspect-ratio determined by banking data to 45deg

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Appendix 1: Aspect-Ratios

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Figure 5 below shows the FARS deaths with an aspect-ratio chosen to bank the loess trend line to 45°. This aspect ratio optimizes our visualization of the trend line but since it compresses the data area into a small rectangle, data visualization is impaired.

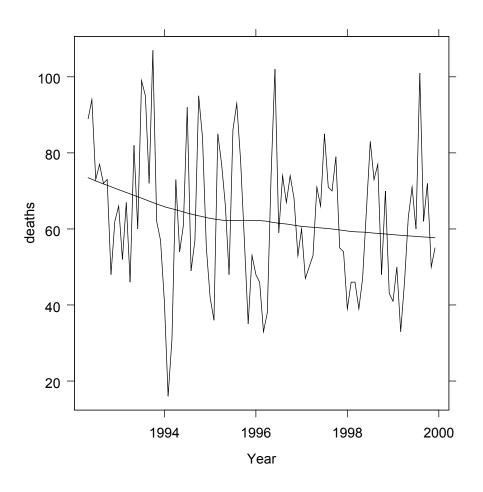
Figure 5. The aspect ratio is choose by banking to 45° the locally-linear loess trend curve.

Fars, deaths

Conclusion

A square aspect ratio (aspect-ratio of 1) provides good data visualization as well as a reasonable visualization of the trend. It is very important when when interpreting these graphs to note the tick mark labels on the vertical axes. So for example, in Figure 6, we see that initially the loess trend starts in 1992 at around 74 deaths per month and decreases to about 58 deaths per month. There is a lot of variability about the trend line. In general we will use an aspect-ratio of 1 as our default for this project.

Figure 6. Aspect-ratio is 1.0



FARS, deaths, aspect-ratio=1.0

Figure 7 below illustrates an exception to the 45° banking rule. We have delibrated chosen a reduced aspect ratio of to improve visualization on the horizontal scale so that we can look for temporal associations between the panels. If a square aspect-ratio were used we would not be able to see so well that there does not appear to be any association between the panels. We have also used a log transformation to improve the data visualiation. Without a log transformation the bulk of the data would be squished into a relatively small area of the panel and the resolution of individual data points would not be satisfactory.

Figure 7. Common scales used in panels.

Year

FARS log(deaths+1)

Appendix 2: Loess

Gaussian Case

Given bivariate observations (x_i, y_i) , i = 1, 2, ..., n, the basic model that can be fit may be written,

 $y_i = g(x_i) + \epsilon_i, i = 1, \dots, n.$

where $e_i \sim \text{NID}(0, \sigma^2)$ and g(x) is a local linear polynomial of degree $\lambda \ge 0$. The local linear polynomial may be written,

$$g(x) = \beta_0^{(x)} + \beta_1^{(x)} x + \dots + \beta_\lambda^{(x)} x^\lambda$$

The parameters $\beta_0^{(x)}$, $\beta_1^{(x)}$, ..., $\beta_{\lambda}^{(x)}$ are estimated by weighted least squares for each value of x. The weight function weights the data, (x_i, y_i) , so that data values near to x have greater weight than those farther away from x. Following Cleveland (1996) we use the tricube weight function,

$$T(z) = \frac{(1 - |z|^3)^3}{0} |z| \le 1$$

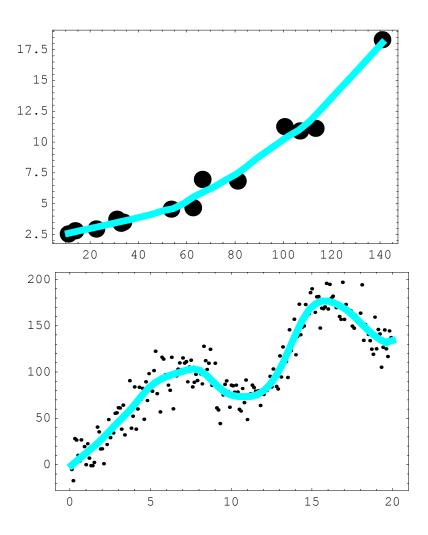
to define the local neighborhood weights for the data at the point x,

$$w_i(x) = T(\Delta_i(x) / \Delta(x, \alpha)),$$

where $\Delta_i(x) = |x - x_i|$ and $\Delta(x, \alpha)$ controls the amount smoothing. Larger values of $\Delta(x, \alpha)$ result in more smoothing. As $\Delta(x, \alpha) \to \infty$, $w_i(x) \to 1$ for each i = 1, 2, ..., n and the local linear model reduces to the standard parametric polynomial regression. For $0 < \alpha \le 1$, $\Delta(x, \alpha)$ is the distance to the *q*-th nearest neighbor where $q = [\alpha n]$, where $[\bullet]$ is the integer part. Hence, $\Delta(x, \alpha) = \Delta_{(q)}(x)$, where $\Delta_{(q)}(x)$ denotes the *q*-th largest value of $\Delta_i(x)$, i = 1, ..., n. For $\alpha > 1$, $\Delta(x, \alpha) = \alpha \Delta_{(n)}(x)$. It follows that as $\alpha \to \infty$, the local linear model reduces to a parametric polynomial regression of degree λ .

Approximate statistical inference methods are available for Gaussian loess models (Cleveland and Devlin, 1988).

Here are two examples of loess fit first to a small dataset of 13 points and then second to a dataset with 200 points.



Robust or Symmetric Errors Case

In the robust case it is only assumed that the errors are symmetric. Hence for data (x_i, y_i) , i = 1, 2, ..., n, the basic model that can be fit may be written,

$$y_i = g(x_i) + \epsilon_i, i = 1, \dots, n.$$

where $e_i \sim \text{IID}(0, \sigma^2)$ with a symmetric distribution and g(x) is again a local linear polynomial of degree $\lambda \ge 0$.

The model is fit to each of the *n* observed data values using a Iteratively Reweighted Least Squares (IRWLS) algorithm. Each iteration fits the model using the current weights which are the product of the robustness weights times the local neighborhood weights. So for fitting at the point x_j the weights are given by $w_i(x_j)r_j$, j = 1, ..., n.

An additional set of weights are used to provide robustness. Given the residuals from the last fit, $\hat{\epsilon}_i$, i = 1, ..., n, the robustness weights are then defined by

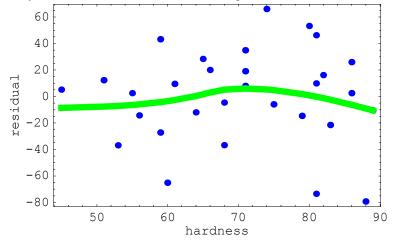
$$r_i = B\left(\frac{\hat{\epsilon}_i}{6\,s}\right)$$

where $s = \text{median} \{ | \hat{\epsilon}_1 |, ..., | \hat{\epsilon}_n | \}$ and B(z) is the bisquare function,

$$B(z) = \frac{(1-z^2)^2}{0} |z| \le 1$$

Robust loess fits are often used for preliminary investigations of a dataset to guard against outliers and for the purpose of detecting outliers. In various types of residual analysis plots with many types of other models such as in linear models it is useful to use a robust loess curve. For example in a residual dependency plot the residuals are plotted against some covariate. A trend in the loess curve in this plot would indicate that the model could be improved.

Here is an an example of a residual dependency plot. Since the trend line is curved a little in the middle of the plot there is a small amount of dependency on the variable hardness remaining in the residuals.



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Appendix 3: Mann-Kendall Trend Tests

Introduction

Hipel & McLeod (1994) and McLeod et al. (1990) have used the Mann-Kendall trend test in the analysis of various types of environmental data.

Kendall Rank Correlation

Let $(X_1, Y_1), \dots, (X_n, Y_n)$ be a bivariate random sample of size n. The Pearson correlation coefficient provides an optimal measure of the degree of association between the X's and the Y's when the sample is drawn from a bivariate normal distribution. The Pearson correlation coefficient is reasonably robust for many other distributions as well.

The Kendall correlation coefficient, denoted by τ , provides a more general non-parametric measure of monotonic association. It is said to be monotonic since making a monotonic transformation on either the X's or the Y's does not change the numerical value of τ .

Kendall's rank correlation coefficient (Kendall, 1970, equation 1.5) may be written,

$$\tau = \frac{S}{D}.$$
 (2)

where S, is the Kendall score given by

$$S = \sum_{i>j} \operatorname{sign}(X_j - X_i) \operatorname{sign}(Y_j - Y_i),$$
(1)

where $sign(\bullet)$ denotes the sign function and D is the maximum possible value of S. In the case where there are no ties among either the X's or the Y's,

$$D = \binom{n}{2}.$$

More generally, if there are n_x distinct ties of extent $t_i, i = 1, ..., n_x$ among the X's and n_y distinct ties of extent $u_i, i = 1, ..., n_y$ among the Y's then

$$D = \sqrt{\left(\binom{n}{2} - T\right)\left(\binom{n}{2} - U\right)},$$

where

$$T = \frac{1}{2} \sum_{i=1}^{n_x} t_i(t_i - 1),$$

$$U = \frac{1}{2} \sum_{i=1}^{n_y} u_i (u_i - 1).$$

In the case where there are no ties in either ranking, it is known (Kendall, 1975, p.51) that under the null hypothesis, the distribution of S may be well approximated by a normal distribution with mean zero and variance,

Var
$$(S) = \frac{1}{18}n(n-1)(2n+5),$$

provided that $n \ge 10$. Valz and McLeod (1990) have given a simplified derivation of this formula for Var (S).

In the case of ties, the variance of S is more complicated,

$$\operatorname{Var}(S) = \left\{\frac{1}{18}n(n-1)(2n+5) - \sum t_i(t_i-1)(2t_i+5) - \sum u_i(u_i-1)(2u_i+5)\right\} \\ + \frac{1}{9n(n-1)(n-2)} \left\{\sum t_i(t_i-1)(t_i-2)\right\} \left\{\sum u_i(u_i-1)(u_i-2)\right\} \\ + \frac{1}{2n(n-1)} \left\{\sum t_i(t_i-1)\right\} \left\{\sum u_i(u_i-1)\right\}.$$

Valz, McLeod and Thompson (1994) have examined the adequacy of the normal approximation in this general case.

The test of the null hypothesis $\mathcal{H}_0: \tau = 0$ is equivalent to testing $\mathcal{H}_0: S = 0$. If there are no ties and if $n \geq 10$ the normal approximation based on Var(S) is adequate. When $n \leq 10$ and there are ties present in only of the variables then the efficient exact algorithm of Panneton & Robillard (1972a, 1972b) may be used. Otherwise if ties are present in both variables then the exact enumeration algorithm given by Valz (1990) may be used or alternatively bootstrapping (Efron and Tibshirani, 1993). Our S-Plus function Kendall implements these algorithms for computing τ and its significance level under a two-sided test.

Mann-Kendall Trend Test

Given *n* consecutive observations of a time series $z_t, t = 1, \dots, n$, Mann (1945) suggested using the Kendall rank correlation of z_t with $t, t = 1, \dots, n$ to test for monotonic trend. The null hypothesis of no trend assumes that the $z_t, t = 1, \dots, n$ are independently distributed. Our S-Plus function, MannKendall(z) implements the Mann-Kendall test using Kendall(x, y) to compute τ and its significance level under the null hypothesis. The Mann-Kendall trend test has some desirable features. In the simple linear trend model with independent Gaussian errors, $z_t = \alpha + \beta t + e_t$, where e_t is Gaussian white noise, it is known that the Mann-Kendall trend test has 98% efficiency relative to the usual least squares method of testing $\beta = 0$. Also, an empirical simulation study of Hipel, McLeod and Fosu (1986) showed that the Mann-Kendall test outperformed the lag one autocorrelation test for detecting a variety of deterministic trends such as a step-intervention or a linear trend.

and

In the case of no ties in the values of $z_t, t = 1, \dots, n$ the Mann-Kendall rank correlation coefficient τ has an interesting interpretation. In this case, the Mann-Kendall rank correlation for a trend test can be written

$$\tau = \frac{S}{\binom{n}{2}},$$

where

$$S = 2P - \binom{n}{2},$$

where P is the number of times that $z_{t_2} > z_{t_1}$ for all $t_1, t_2 = 1, \ldots, n$ such that $t_2 > t_1$. Thus $\tau = 2\pi_c - 1$, where π_c is the relative frequency of positive concordance, i.e., the proportion of time for which $z_{t_2} > z_{t_1}$ when $t_2 > t_1$. Equivalently, the relative frequency of positive concordance is given by $\pi_c = 0.5(\tau + 1)$.

The Mann-Kendall test is essentially limited to testing the null hypothesis that the data are independent and identically distributed. Our time series data may diverge from this assumption in two ways. First there may be autocorrelation and second may be a seasonal component. To eliminate these factors we can use annual data but this has the effect of reducing the power. For strong positive autocorrelation in the series, the effect of using annual totals will reduce the effect of this autocorrelation substantially and the loss of power is, perhaps, not expected to be too much — this is something we will investigate further in a methodological study.

The method of Brillinger (1989) deals with both the problems of seasonality and autocorrelation but it also requires an estimate of the spectral density at zero. However the test of Brillinger (1989) is not suitable for testing for long-term trend with monthly data with a strong seasonal component since the running-average smoother used will not be useful in this case. Another model-building approach to trend analysis is intervention analysis (Box & Tiao, 1975; Hipel & McLeod, 1994) which can also handle both seasonality and autocorrelation. This assumes a known intervention time and the development of a suitable time series model.

Seasonal Mann-Kendall Trend Test

The Seasonal-Mann-Kendall trend test is a test for monotonic trend in a time series with seasonal variation. Hirsch et al. (1982) developed such a test by computing the Kendall score separately for each month. The separate monthly scores are then summed to obtain the test statistic. The variance of the test statistic is obtained by summing the variances of the Kendall score statistic for each month. The normal approximation may then be used to evaluate significance level. In this test, the null hypothesis is that the time series is of the form $z_t = \mu_m + e_t$ where e_t is white noise error and μ_m represents the mean for period m. The τ coefficient is defined by

$$\tau = \frac{\sum_{i=1}^{s} S_i}{\sum_{i=1}^{s} D_i},$$

where $S_i, D_i, i = 1, \dots, s$ denote the Kendall scores and denominators for the *i*-th season and *s* is the seasonal period. We implemented this procedure in S-Plus in our function SeasonalMannKendall(z)

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