Monte-Carlo Testing for Preservation of Historical Statistics in Synthetic Time Series

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Please see:

http://www.stats.uwo.ca/faculty/aim/TIES2012/ for pdf copies of this presentation and papers referred to therein.

- Simulated or synthetic weather and hydrological time series used for planning
- Does the model adequately capture important characteristics of the series: seasonal means and variances, long-range dependence, critical drought, etc.
- informal and formal testing

Weather Generator, Downscaling, Climate Change¹

- Empirical nearest neighbour model for weather simulation.
- Daily time series for 22 weather stations in southwestern Ontario (Blyth,Brantford,Dorchester,Embro,Exeter,Foldens,Glen Allan,Ilderton,London A, etc.)
- 7 weather variables (precipitation, max temperature, min temperature, humidity, wind speed (eastward, northward), pressure.
- Long time series from January 1, 1979 to December 31, 2005 (length 9862).
- Are temporal and spatial correlations preserved as well as seasonal means, variances, extremes, etc.

¹with Leanna King and Dr. Simonovic

Spatial Correlation with London A



Ian McLeod

Synthetic Time Series

- given data, D
- statistic of interest, $K_{obs} = K(\mathcal{D})$
- **model fitted to data**, $\hat{\mathcal{M}}$
- H₀ : K_{obs} ~ K(Â), where K(Â) denotes the distribution of the statistic K if the fitted model is correct.

It is standard practice to check the adequacy of the a fitted statistical model using a significance test.

These tests can be tricky!

For example if we fit $z_t = \phi z_{t-1} + a_t$, t = 1, ..., n, where $a_t \sim \text{NID}(0, 1)$, we may use the lag one residual autocorrelation, $\hat{r}_{\hat{a}}(1)$ is the test statistic.

But,

$$\hat{r}_{\hat{a}}(1) = \phi^2/n.$$

Not as might be expected since

$$\hat{r}_{a}(1) = 1/n.$$

If we knew the model then we could obtain $a_t, t = 1, ..., n$. In this case $\hat{r}_a(1)$ is normal with mean zero and variance 1/n but it turns out that $\hat{r}_{\hat{a}}(1)$ is normal mean zero and variance ϕ^2/n . If wrong variance is used, test is much too conservative.

This occurs because $\hat{\phi}$ and $\hat{r}_{\hat{a}}(1)$ may be highly correlated asymptotically!

This phenomenon is very general and occurs for other statistics such as Hurst's K and the critical period statistics.

- Fit model, $\hat{\mathcal{M}}$. Compute observed test statistic, K_{obs} .
- Set N ← 1000, number of MC iterations. Repeat next step N times.
- Simulate $\hat{\mathcal{M}}$ and compute value of *K* with new data (fitting model again if needed).

$$\rho = (N+1)^{-1}(1 + \sum I(|K| > |K_{obs}|)).$$

Just using a few simulations and comparing the null distribution of the test statistic with the observed value is often adequate as in the spatial correlation with the weather generator. Dufour (2006, Proposition 2.6) establishes that under regularity conditions the usual Monte-Carlo test procedure provides a valid test in the presence of a nuisance parameter provides that consistent estimators are used.

Dufour, J.-M. (2006). Monte Carlo Tests with Nuisance Parameters: A General Approach to Finite-Sample Inference and Non-Standard Asymptotics. *Journal of Econometrics*, 113/2, 443-477.

Monte-Carlo Test Theorem: Applications in Diagnostic Checking

Mahdi & McLeod (2011) and references therein discuss many applications of this theorem to portmanteau residual diagnostic checks in time series.

Mahdi, E. and McLeod, A.I. (Online, 05/09/2011). Improved Multivariate Portmanteau Diagnostic Test. Journal of Time Series Analysis. DOI: 10.1111/j.1467-9892.2011.00752.x.

Hurst Phenomenon & Long Memory

- Hurst (1951,1957) fit the equation $R_n = (n/2)^K$, where R_n is the rescaled cumulative range to 690 geophysical time series.
- Obtained K = 0.73 with $\sigma_K = 0.09$ based on these 690 series.
- W. Feller & others noted *K* ≈ 0.5 for all ARMA-like time series.
- Long-memory time series can be used. More complicated and computationally difficult models.

Hurst, H.E. (1957). A Suggested Statistical Model of some Time Series which occur in Nature. *Nature* 180, 494-494 doi:10.1038/180494a0

- Simulated 25,000 replications of an AR(1) with $\phi = 0.5$ and n = 100.
- **•** Obtained joint distribution for $\hat{\phi}$ and *K*
- Mean and standard deviation of K are 0.73 and 0.06 respectively.

Empirical Distribution $\hat{\phi}$ and *K*



Simple ARMA models were fit to 22 annual geophysical time series and the MC test for preservation of Hurst's K suggested no evidence of lack of fit.

Similarly we showed that PAR models preserve the critical period statistics in monthly riverflows.

Hipel, K.W. & McLeod, A.I. (1994), *Time Series Modelling of Water Resources and Environmental Systems*, Elsevier. http://www.stats.uwo.ca/faculty/aim/1994Book/

Espanola Mud Varves, Annual, K = 0.84



Ian McLeod Synthetic Time Series

Nile River Minima, 622 to 1284

× //



	Loglikelihood	AIC	BIC
ARMA(2,1)	237.61	-469.22	-455.73
FGN(mle)	236.52	-471.04	-466.54

Table: AIC/BIC for best models

The series length is n = 663 so the BIC penalty for parameters is 6.5k as opposed to 2k for the AIC.

Distribution of Hurst K in Fitted ARMA(2,1)



Available using *Mathematica's* cdf format and in the Wolfram Demonstrations.

- Visual Comparison of FGN and ARMA(2,1) for Nile Minima Series
- Visualizing Fractional Gaussian Noise

- Long time series from January 1, 1979 to December 31, 2005 (length 9862).
- Detrended by fitting harmonic regression with period 365.25.
- New PACF plot presented next.

New PACF Plot



McLeod, A. I., Ying Zhang & Changjiang Xu (2011), FitAR: Subset AR Model Fitting. R package version 1.92. http://CRAN.R-project.org/package=FitAR

McLeod, A.I. & Zhang, Ying (2008). Improved Subset Autoregression: With R Package. *Journal of Statistical Software* 28/2, 1-28. http://www.jstatsoft.org/v28/i02.

McLeod, A.I. & Zhang, Y. (2006). Partial Autocorrelation Parameterization for Subset Autoregression. *Journal of Time Series Analysis*, 27, 599-612. Couillard, M. and Davison, M. (2005). A comment on measuring the Hurst exponent of financial time series. *Physica A* 3 48, 404-418.